



*Supplement of*

## **A Bayesian sequential updating approach to predict phenology of silage maize**

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# Supplementary material

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## S1. Sensitivity analysis

The Morris or elementary effects screening method (Morris, 1991) was used to conduct a qualitative global sensitivity analysis on phenological development of maize. Sensitivity analysis was only performed for site-year 6\_2010 under the assumption that ranks of the most sensitive parameters would not change significantly due to the different weather and initial conditions in Kraichgau and the Swabian Alb. The *sensitivity* package in R (Bertrand Iooss et al., 2020) was used. The one-at-a-time (OAT) design in the *morris* function was used to define the parameter vectors. A total of 11 parameters that influence phenological development in the SPASS model were pre-selected based on expert knowledge. Uniform parameter distributions with a range equal to three standard deviations from the expected value were used. It is noted that different distributions have been used for Bayesian calibration (platykurtic prior distribution) and sensitivity analysis (uniform distribution). However, this is assumed to have a limited influence in identifying the most sensitive parameters. Settings to the *morris* function were provided: 1000 samples, 10 levels and a grid jump-size of 2 units. Phenology was simulated using the SPASS model in XN5 software for all the proposed parameter vectors. The *morris* function was then used to estimate elementary effects (Cuntz et al., 2015; Morris, 1991) of phenological development at an interval of every 5 days within the growing season. The sensitivity measures, namely, the mean ( $\mu^*$ ) of the absolute value of the elementary effects of a parameter and the standard deviation ( $\sigma$ ) were calculated on these days to evaluate parameter sensitivity over the growing season.

$$\mu_{\theta_i}^* = \frac{1}{N} \sum_{n=1}^N |ee_{n,\theta_i}| \quad \text{S1.}$$

$$\sigma_{\theta_i} = \sqrt{\frac{\sum_{n=1}^N (ee_{n,\theta_i} - \mu_{\theta_i}^*)^2}{N}} \quad \text{S2.}$$

where  $\mu_{\theta_i}^*$  and  $\sigma_{\theta_i}$  are the  $\mu^*$  and  $\sigma$  sensitivity measures for the  $i^{\text{th}}$  parameter in the parameter vector  $\theta$ ,  $ee_n$  is the elementary effects for the  $n^{\text{th}}$  parameter vector,  $N$  are the total parameter vectors and  $\mu_{\theta_i}$  is given by:

$$\mu_{\theta_i} = \frac{1}{N} \sum_{n=1}^N ee_{n,\theta_i} \quad \text{S3.}$$

Based on  $\mu^*$ , the effective sowing depth (SOWDEPTH) was the most and only sensitive parameter during emergence, which is intuitive as the other parameters influence development after emergence (Figure S1). Then the relative importance of parameters that define the cardinal temperatures (DELTMAX1, DELTOPT1 and TMINDEV1) and the physiological development days (PDD1) of the vegetative phase increased. These parameters continued to be the most influential parameters even through the generative phase of development. Even though DELTOPT2 and PDD2 are important parameters for the generative phase of development, their influence was small and over-shadowed by the influence of the vegetative phase parameters.

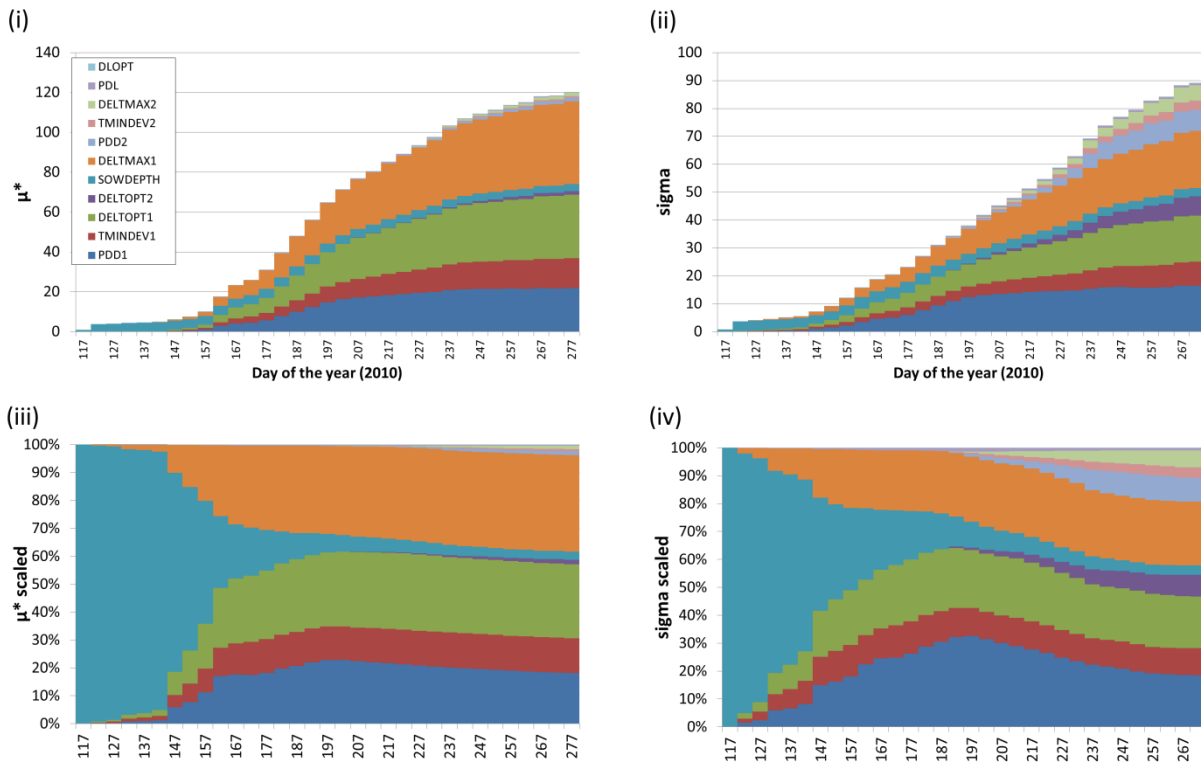


Figure S1 Plots of (i)  $\mu^*$  and (ii) sigma of elementary effects calculated for simulated phenological development at an interval of 5 days over the growing season of silage maize (between sowing day 112 and harvest day 278 of the year) at site 6 in the year 2010. The parameters that influence phenological development in the SPASS model are listed in the legend. Plots (iii) and (iv) are the normalized  $\mu^*$  and sigma values per day, respectively, expressed as a percentages.

## S2. Estimation of information entropy

Information entropy ( $H$ ) for a continuous distribution is given by:

$$H = - \int f(\theta) \ln(f(\theta)) d\theta \quad \text{S4.}$$

where  $f(\theta)$  is the probability density function of  $\theta$ .

Information entropy estimates of the posterior parameter distributions were obtained using the redistribution estimate equation (Beirlant et al., 1997):

$$H_n = - \frac{1}{n} \sum_{i=1}^n \ln f_n(\theta_i) \quad \text{S5.}$$

where  $H_n$  is the estimate of information entropy,  $f_n$  is the Kernel Density Estimate (KDE) and  $\theta_1, \dots, \theta_n$  are independent and identically distributed (i.i.d.) parameter vector samples from the posterior distribution. The KDE was obtained by using the `kde` function from the `ks` package in R (Duong, 2020). Least Squares Cross-Validation (LSCV) was used for bandwidth selection.

### **S3. Residual analysis**

Residuals were analysed for the synthetic and true sequences for simulated phenology at the maximum a posteriori probability (MAP) estimate of the model parameters. The residual plots provided in the following sections have been separated into the synthetic sequences (section S3.1), Swabian Alb true sequence (section S3.2), and Kraichgau true sequence (section S3.3).

Homoscedasticity was checked by plotting the residuals against days-after-sowing and simulated phenology (Figure S2, Figure S3, Figure S8 – Figure S13, Figure S18 – Figure S20). In general, heteroscedasticity was not observed. Normal assumption of the error model was verified by plotting histograms of the residuals and quantile-quantile plots (Figure S4, Figure S5, Figure S14 – Figure S16, Figure S21). For the first few sequential updates, the number of observations were limited making a thorough analysis difficult. For the latter few sequential updates, the residuals were found to be nearly normal.

In the synthetic sequences, the residual error distribution was nearly normal (Figure S4, Figure S5). The slight skewness is attributed to model limitations (controlled cultivar-environment sequence) and specific site-years that had a different phenological development as compared to the remaining site-years in the calibration sequence (both synthetic sequences).

The slight skewness observed in the true sequence is attributed to model limitations where the model is unable to capture the slow development during the vegetative phase that was observed at a few site-years like 6\_2013 (Figure S14, Figure S15, Figure S16) and 5\_2016 (Figure S16). Autocorrelation was estimated after padding the dataset as the observations are not at regular time-intervals. Therefore, there is no ACF estimated at some lags. Figure S17 contains the autocorrelation (ACF) plot of the residuals after the model is calibrated to data from site-years 6\_2010, 5\_2011, 5\_2012, 6\_2013, 5\_2015, and 5\_2016. Based on the limited

data with unequal lags, no autocorrelation was detected. However, it is suspected that with state variables like phenology, which are based on cumulative sums, autocorrelation of errors could theoretically exist. However, due to data limitations, error modelling would be limited in its scope for improving the results.

### S3.1 Synthetic sequences

In the ideal sequence where there is no model structural error, the skewness in the residual distribution (Figure S4) is caused by site-year 2. This site-year exhibits a different development-behaviour as compared to other site-years in the calibration sequence (Figure S6). In the controlled cultivar-environment sequence the slight skewness (Figure S5) in the distribution of the residuals are caused due to two reasons. The site-year 9 exhibits a different phenological development-behaviour as compared to other site-years in the calibration sequence (Figure S7). Additionally, the model is unable to capture the rapid growth seen in site-years 3, 4, 5, 8 and 9 between 82 and 110 days after sowing.

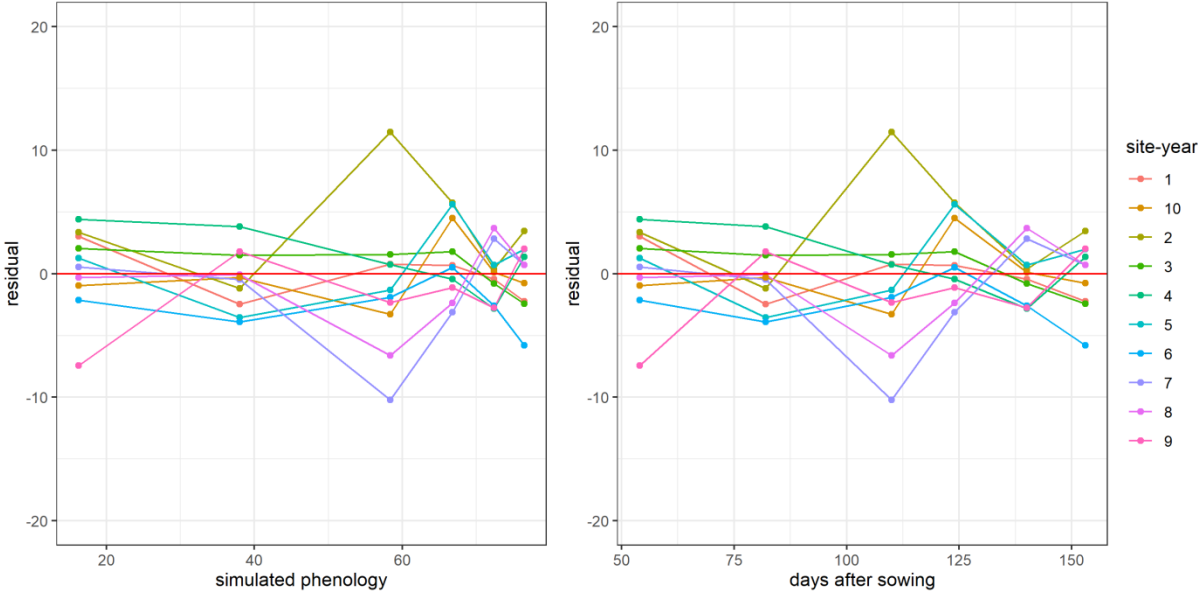


Figure S2 Residuals vs simulated phenology and days after sowing after calibration to 10 site-years in the ideal synthetic sequence



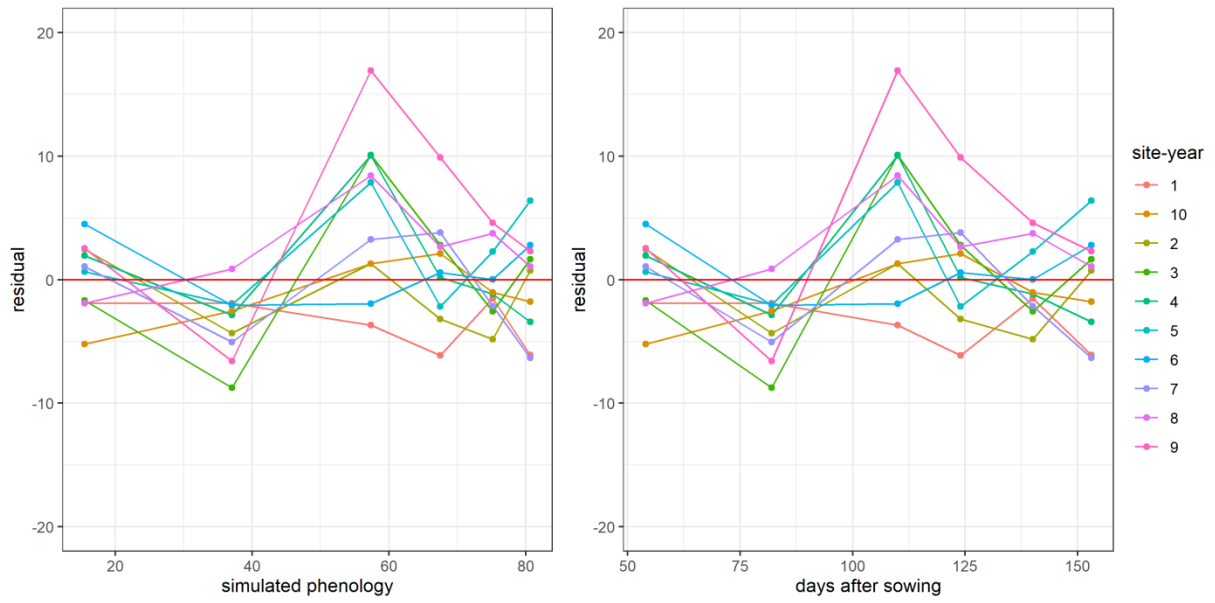


Figure S3 Residuals vs simulated phenology and days after sowing after calibration to 10 site-years in the controlled cultivar-environment synthetic sequence

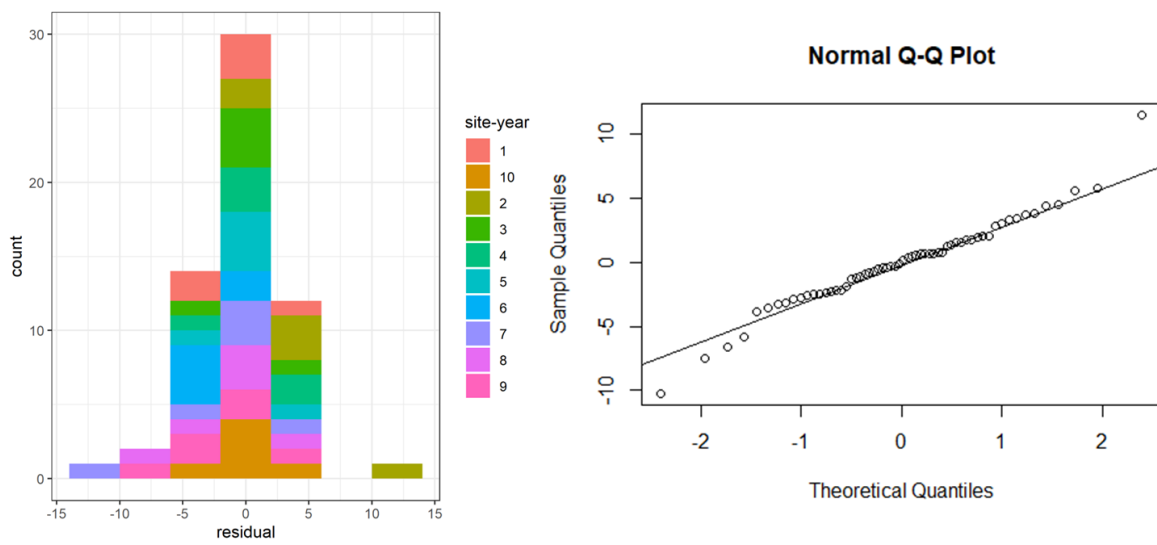


Figure S4 Histogram and quantile-quantile plots of the residuals after calibration to 10 site-years of the ideal synthetic sequence

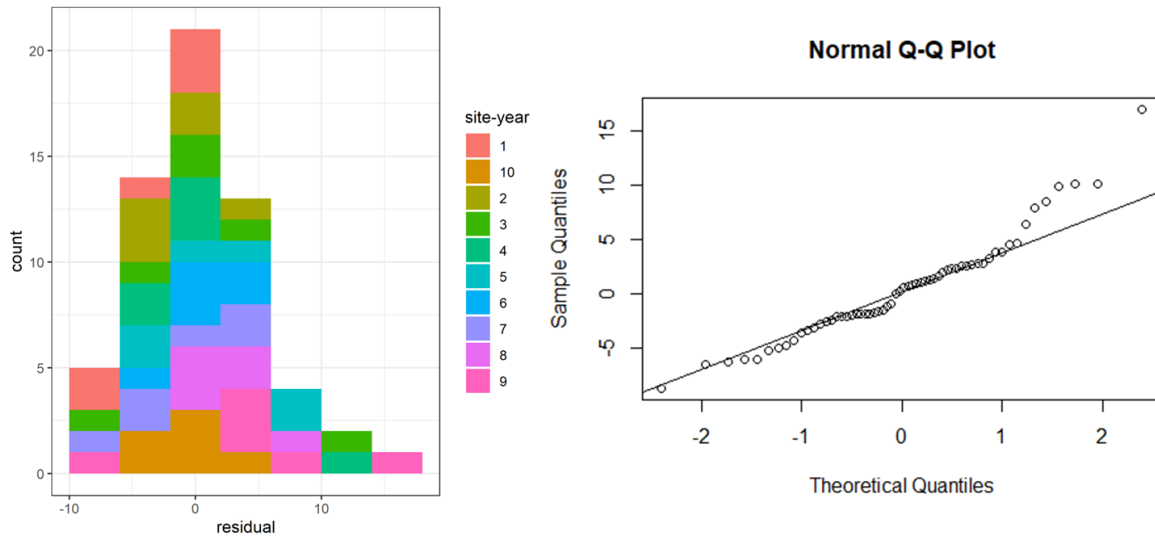


Figure S5 Histogram and quantile-quantile plots of the residuals after calibration to 10 site-years of the controlled cultivar-environment synthetic sequence

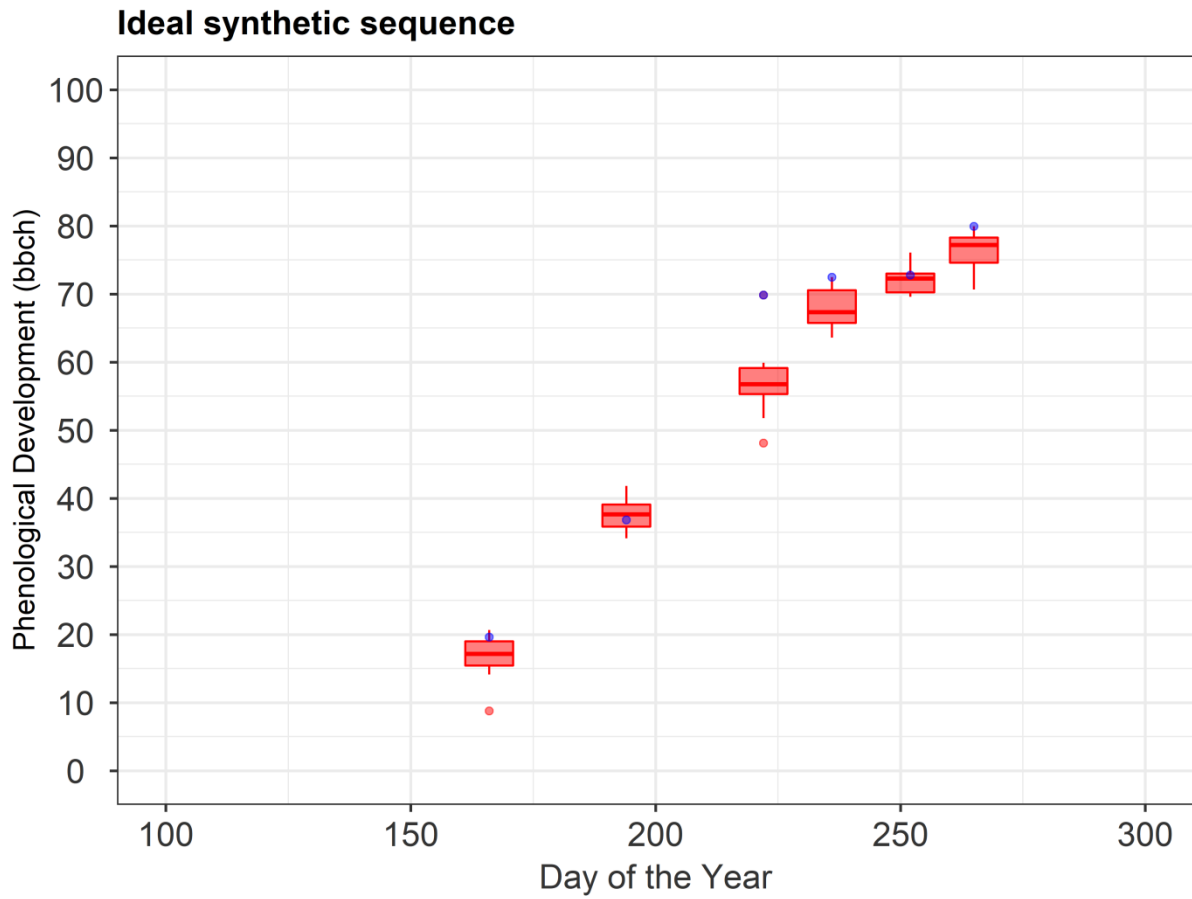


Figure S6 The boxplots show the phenological development (BBCH) of all the site-years used in calibration in the ideal synthetic sequence. The blue point corresponds to the phenological development (BBCH) for site-year 2.

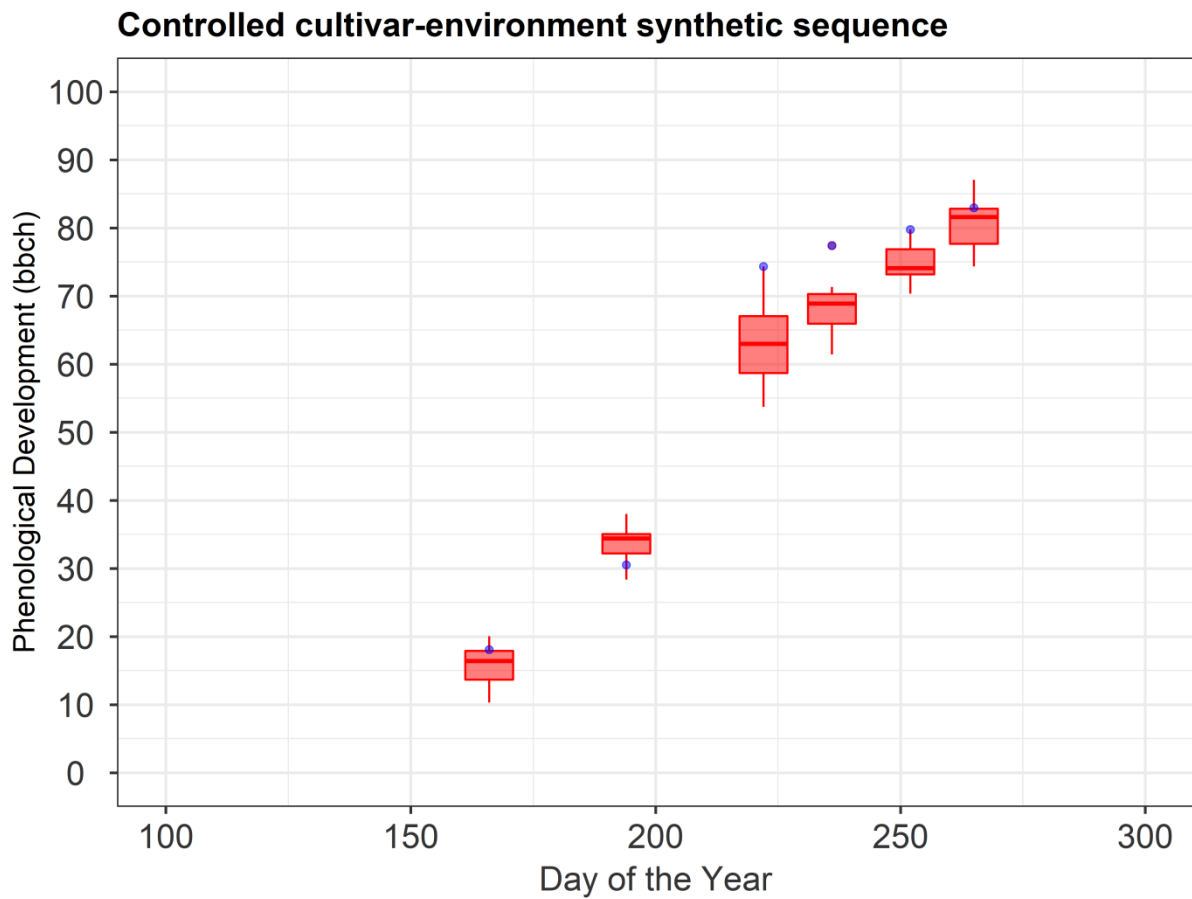


Figure S7 The boxplots show the phenological development (BBCH) of all the site-years used in calibration in the controlled cultivar-environment synthetic sequence. The blue point corresponds to the phenological development (BBCH) for site-year 9.

### S3.2 True sequence in Swabian Alb

The residual plots for the sequential updates with greater than 3 calibration site-years show high residuals in the vegetative phase (simulated phenology < 61BBCH) (Figure S11, Figure S12, Figure S13). Residuals from site-years 6\_2013 and 5\_2016 cause this skewness in the distribution of the residuals (Figure S14, Figure S15, Figure S16). This behaviour is attributed to the model's inability to capture the slow development seen in these site-years as evident from the single-site-year calibration results in Figure S22.

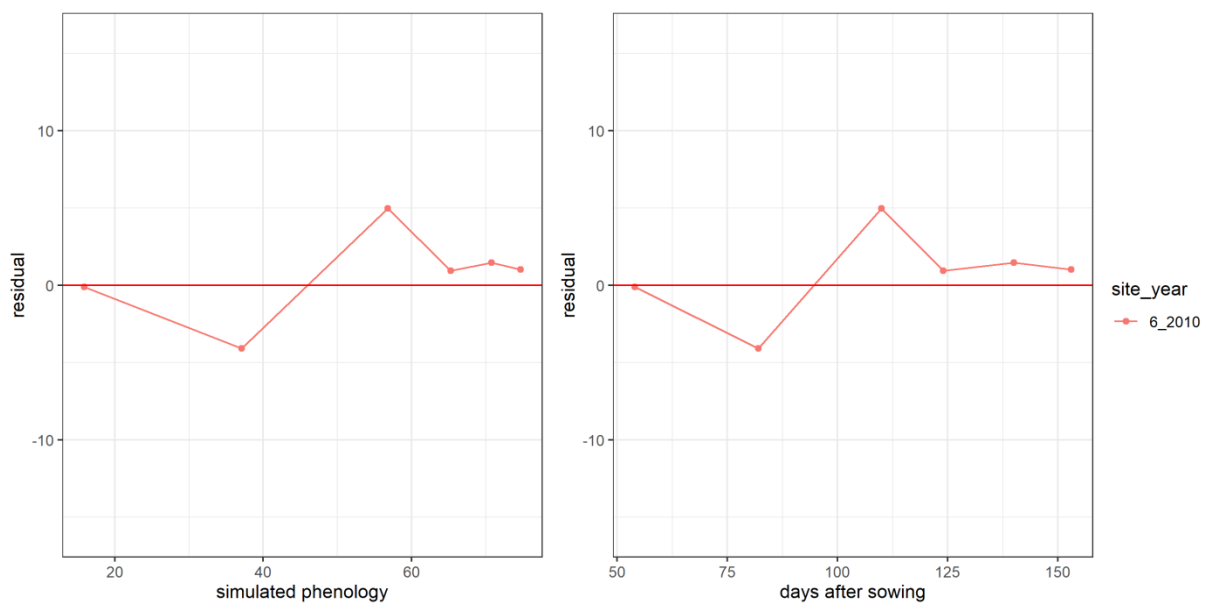


Figure S8 Residuals vs simulated phenology and days after sowing after calibration to site-year 6\_2010

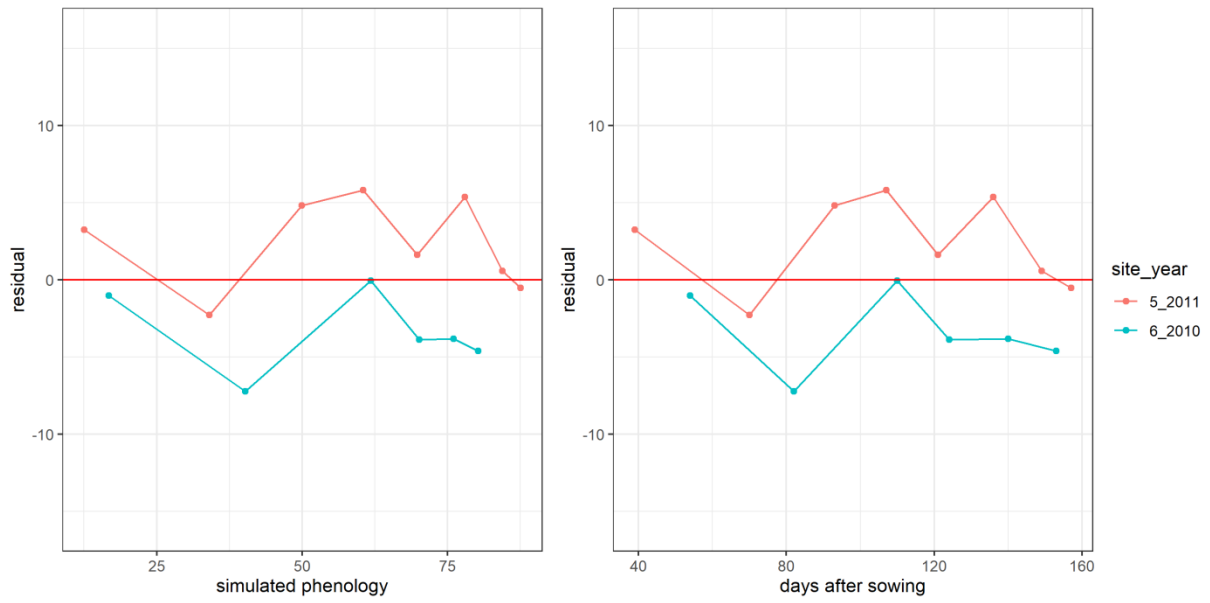


Figure S9 Residuals vs simulated phenology and days after sowing after calibration to site-years 6\_2010 and 5\_2011

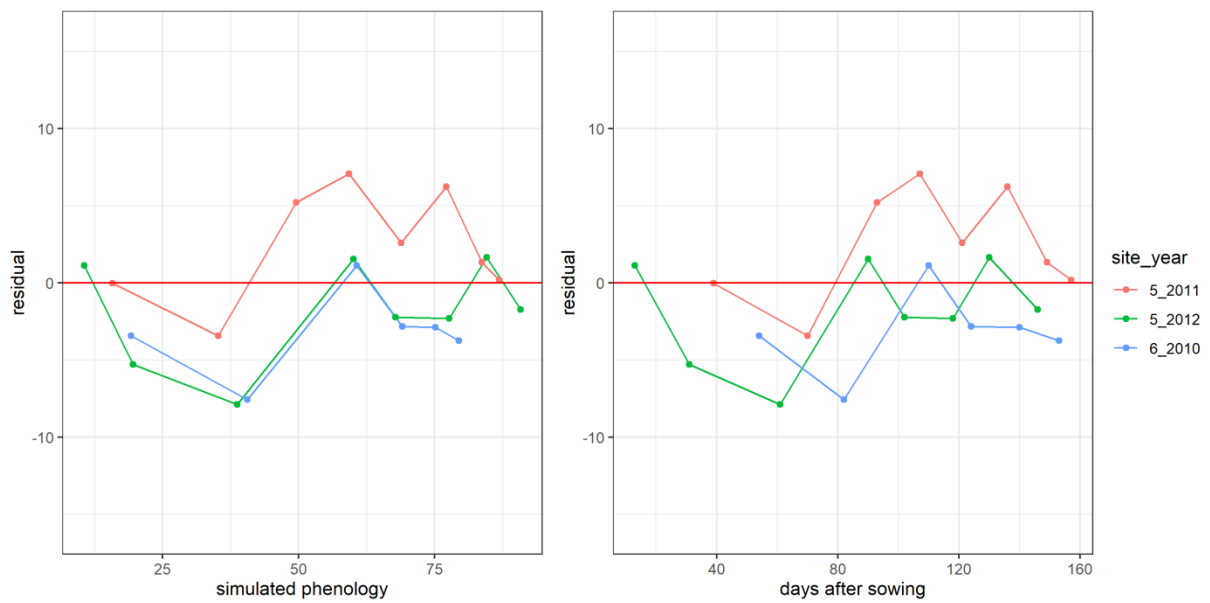


Figure S10 Residuals vs simulated phenology and days after sowing after calibration to site-years 6\_2010, 5\_2011, and 5\_2012

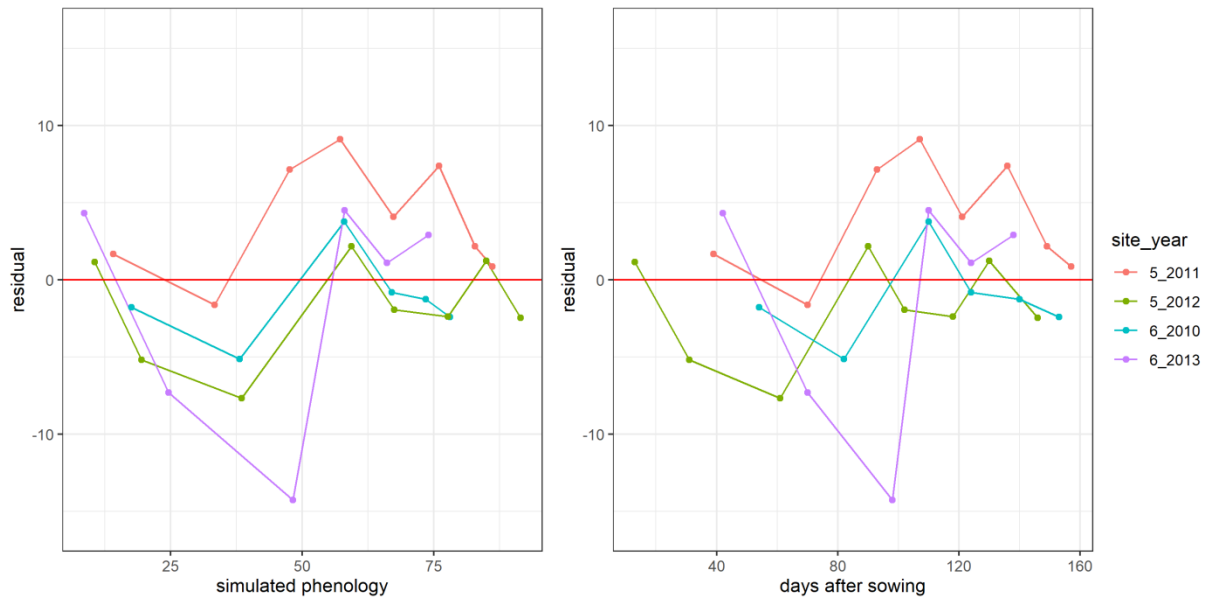


Figure S11 Residuals vs simulated phenology and days after sowing after calibration to site-years 6\_2010, 5\_2011, 5\_2012, and 6\_2013

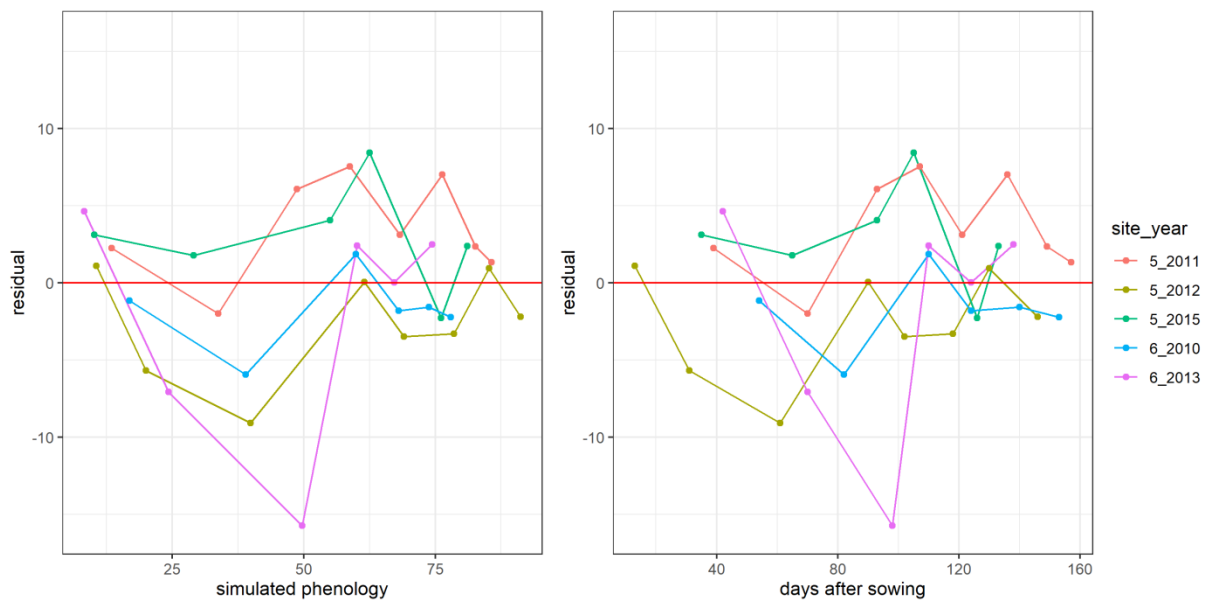


Figure S12 Residuals vs simulated phenology and days after sowing after calibration to site-years 6\_2010, 5\_2011, 5\_2012, 6\_2013, and 5\_2015

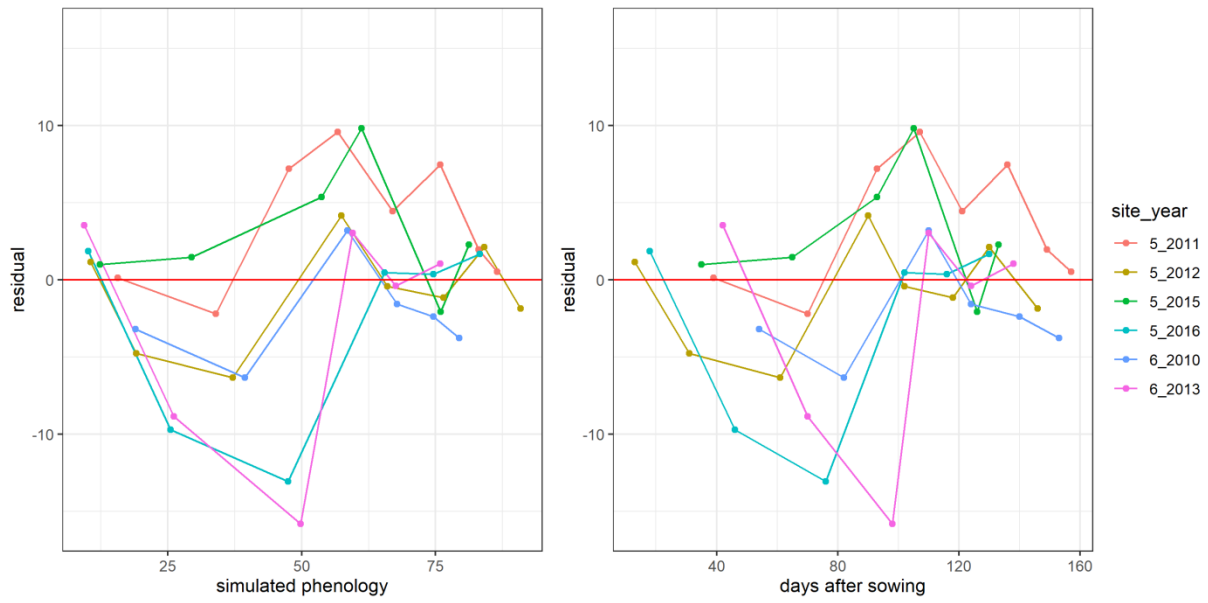


Figure S13 Residuals vs simulated phenology and days after sowing after calibration to site-years 6\_2010, 5\_2011, 5\_2012, 6\_2013, 5\_2015, and 5\_2016

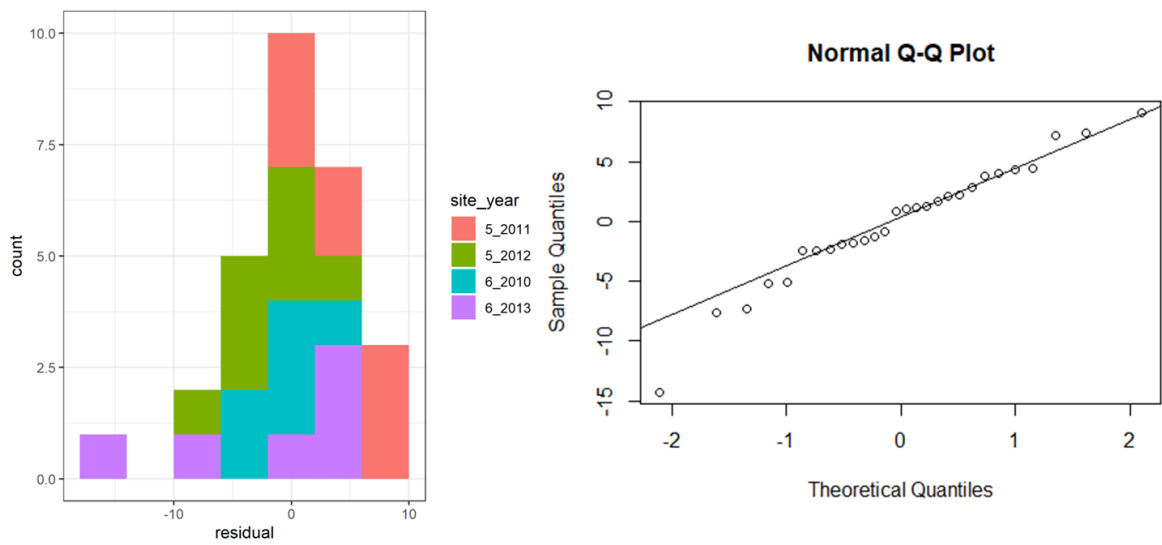


Figure S14 Histogram and quantile-quantile plots of the residuals after calibration to site-years 6\_2010, 5\_2011, 5\_2012, and 6\_2013



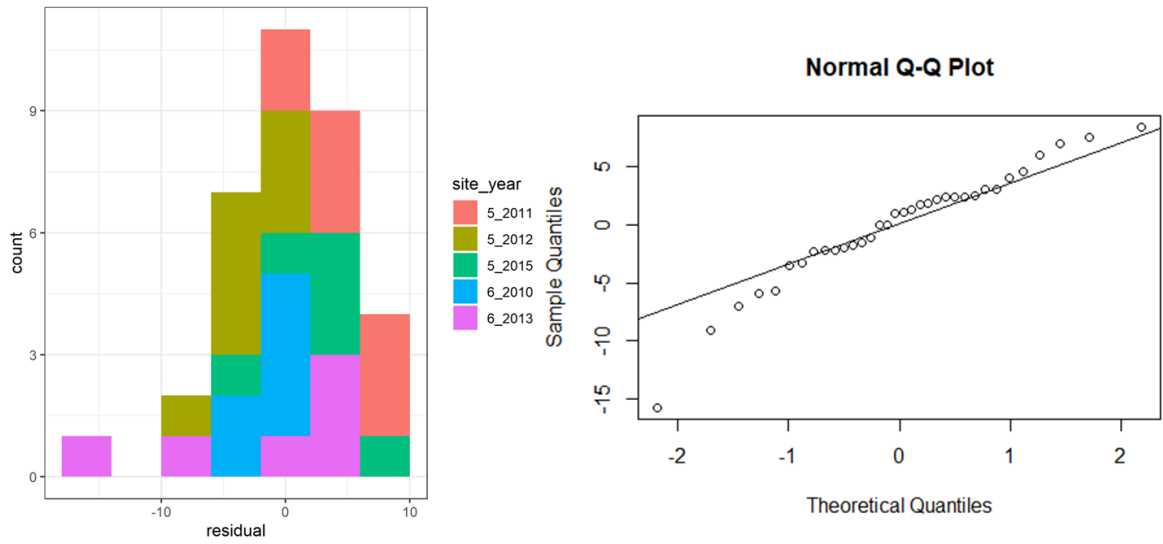


Figure S15 Histogram and quantile-quantile plots of the residuals after calibration to site-years 6\_2010, 5\_2011, 5\_2012, 6\_2013, and 5\_2015

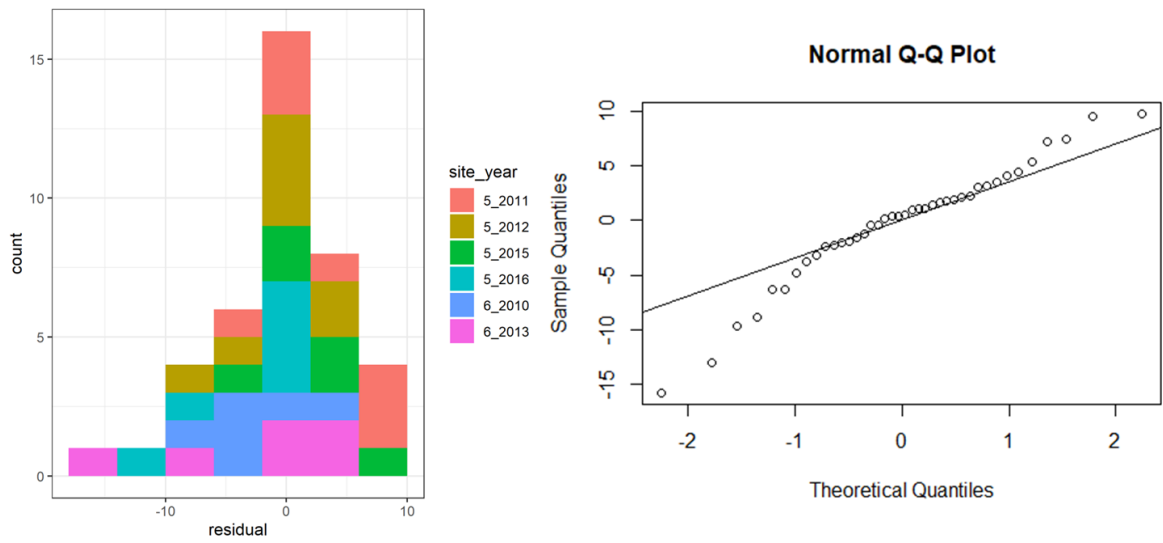


Figure S16 Histogram and quantile-quantile plots of the residuals after calibration to site-years 6\_2010, 5\_2011, 5\_2012, 6\_2013, 5\_2015, and 5\_2016

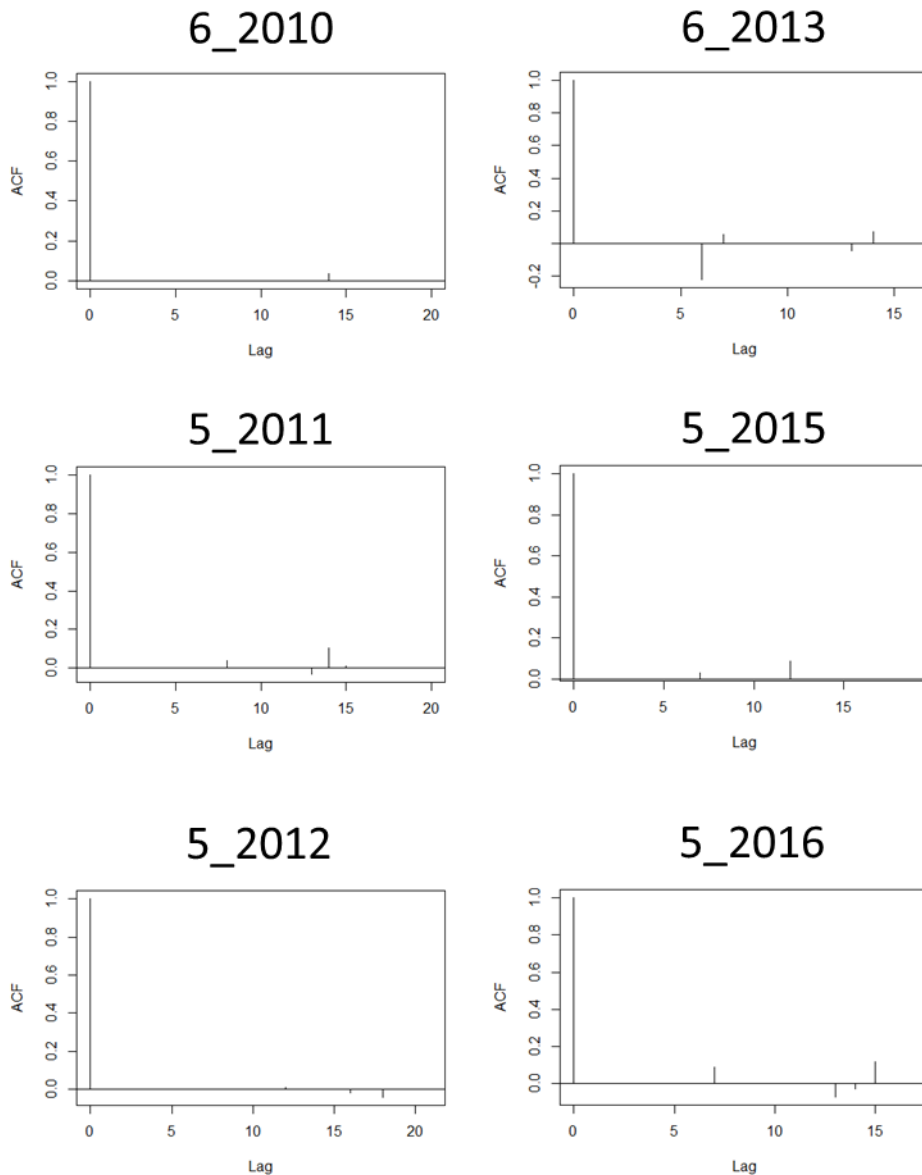


Figure S17 ACF (auto-correlation function) plots of the residuals after calibration to site-years 6\_2010, 5\_2011, 5\_2012, 6\_2013, 5\_2015, and 5\_2016

### S3.3 True sequence in Kraichgau

The residual plots for Kraichgau with limited observations show no evidence of heteroscedasticity (Figure S18, Figure S19, Figure S20) and a nearly normal distribution (Figure S21).

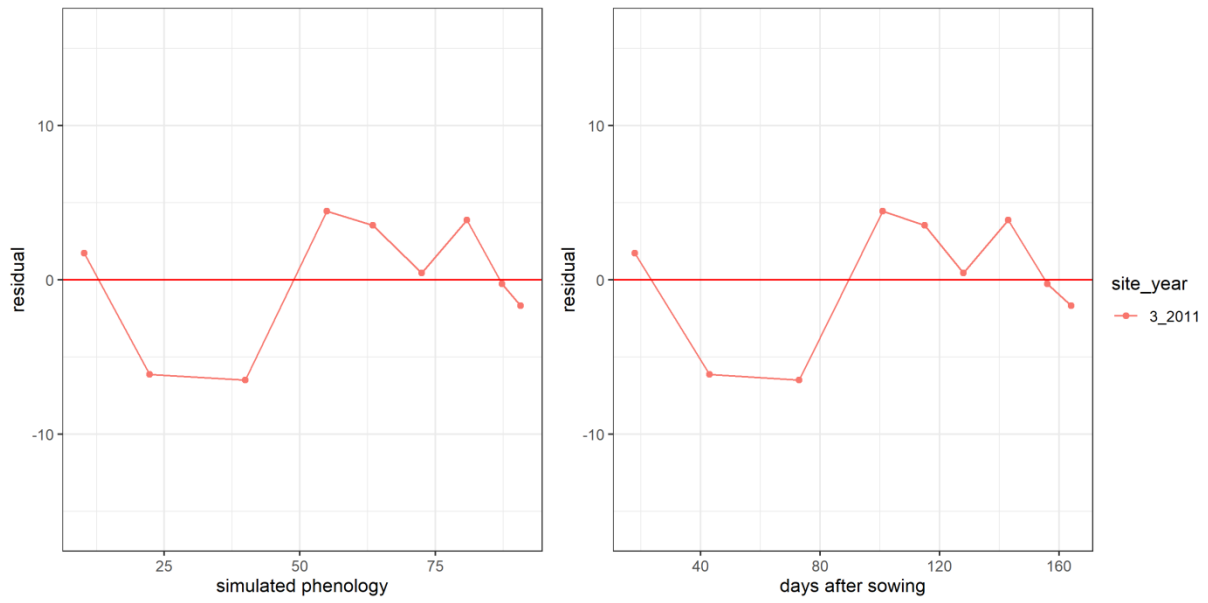


Figure S18 Residuals vs simulated phenology and days after sowing after calibration to site-years 3\_2011

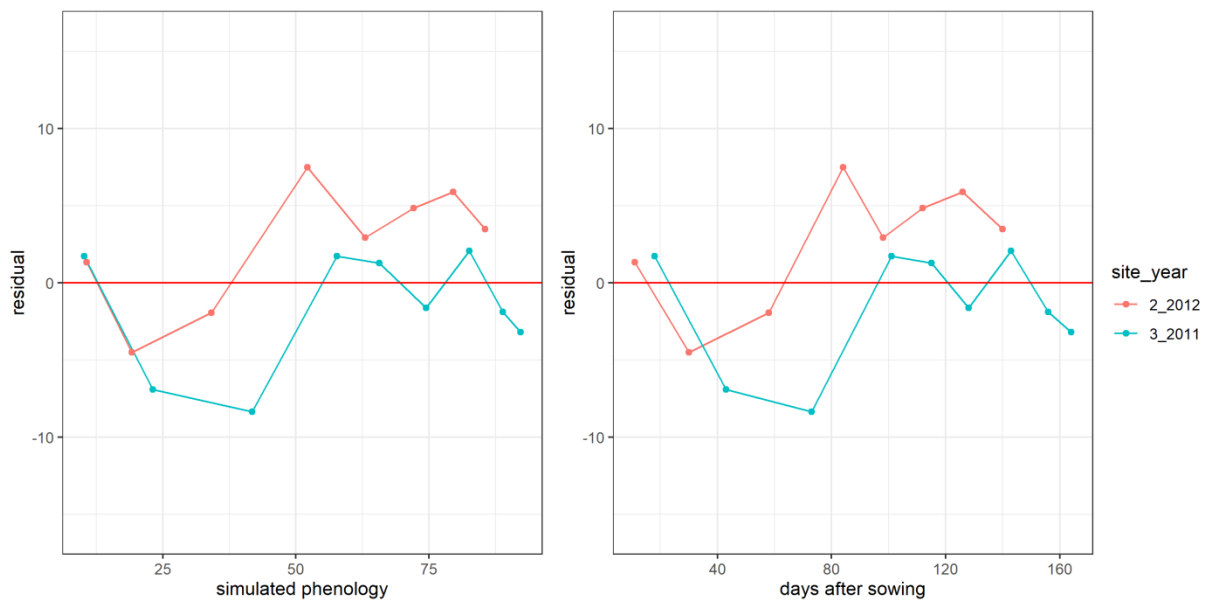


Figure S19 Residuals vs simulated phenology and days after sowing after calibration to site-years 3\_2011 and 2\_2012

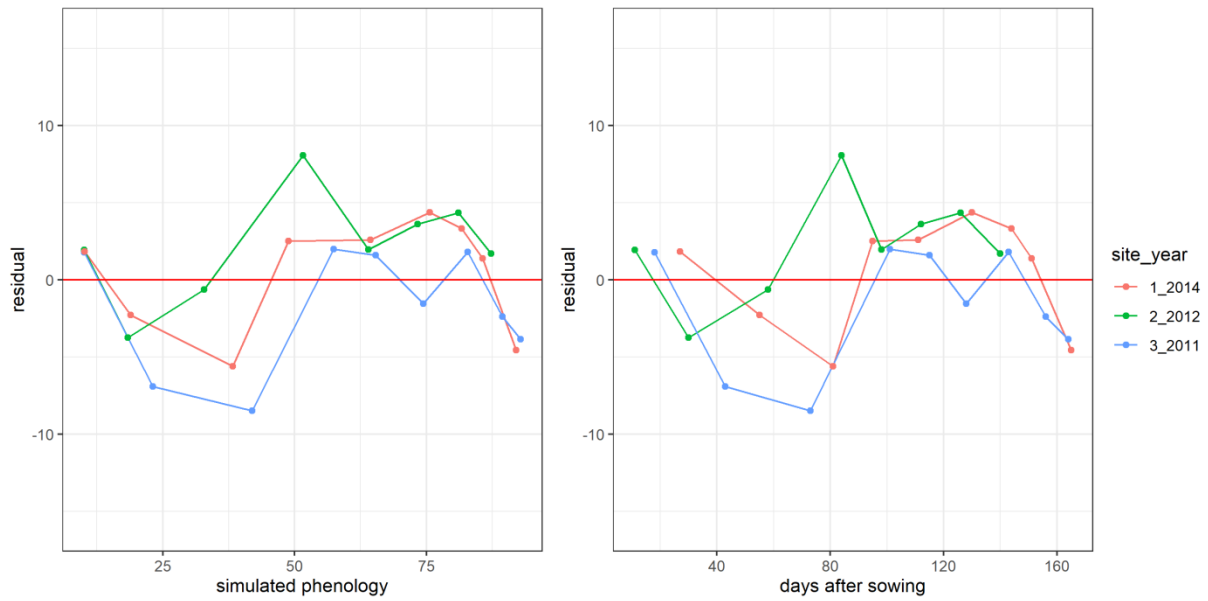


Figure S20 Residuals vs simulated phenology and days after sowing after calibration to site-years 3\_2011, 2\_2012, and 1\_2014

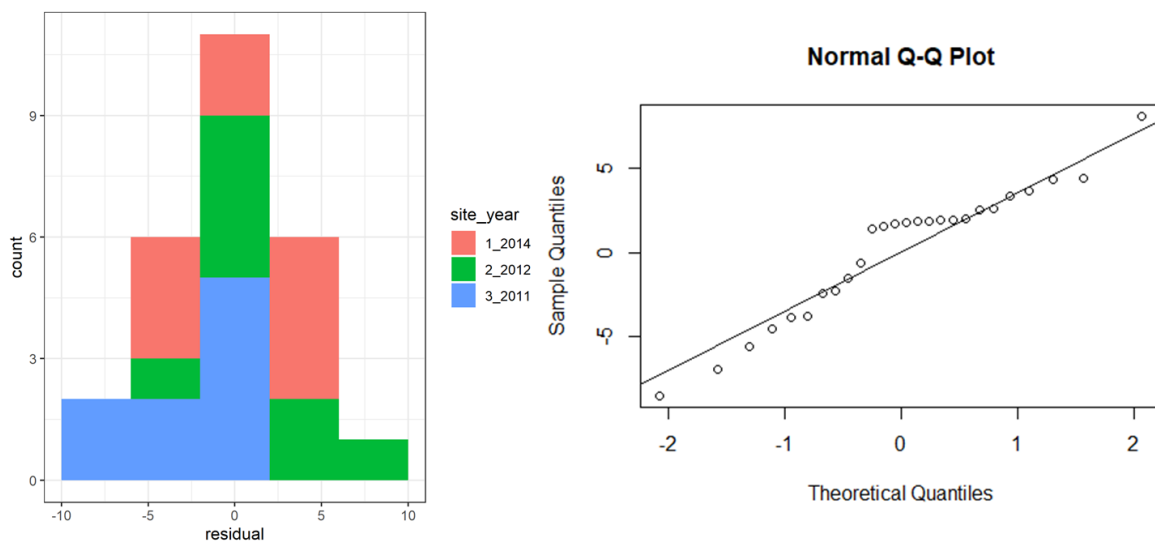


Figure S21 Histogram and quantile-quantile plots of the residuals after calibration to site-years 3\_2011, 2\_2012, and 1\_2014

## S4. Single-site-year calibration results

Observed and simulated phenology, after the SPASS model was calibrated individually to the site-years in the study, are plotted in Figure S22.

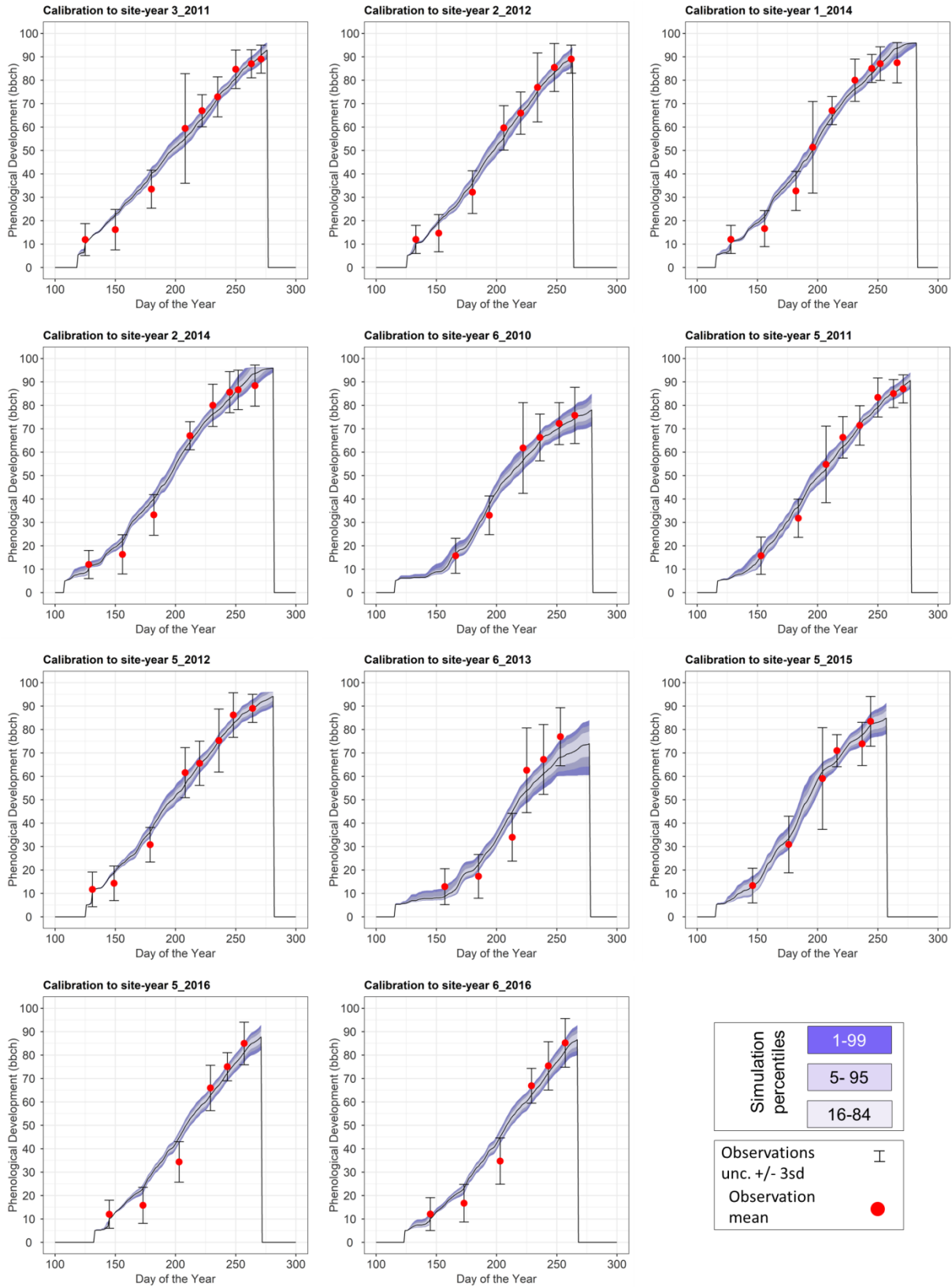


Figure S22 Observed and simulated phenological development after calibration, plotted against the day of the year. The red points are the mean observations, while the black error bars indicate  $\pm 3$  standard deviations. The mean simulation is indicated by the continuous

black line. The green bands represent the different percentiles of simulated phenology. It is noted that for some site-years, the calibrated model is unable to capture the slow development rate during the vegetative phase.

## S5. Marginal posterior parameter pdf: true sequences

Marginal posterior probability density functions (pdf) of the 6 estimated parameters for the true sequences in Kraichgau (Figure S23) and the Swabian Alb (Figure S24) are provided.

Refer to the main text (Discussion: Parameter Uncertainty) for details.

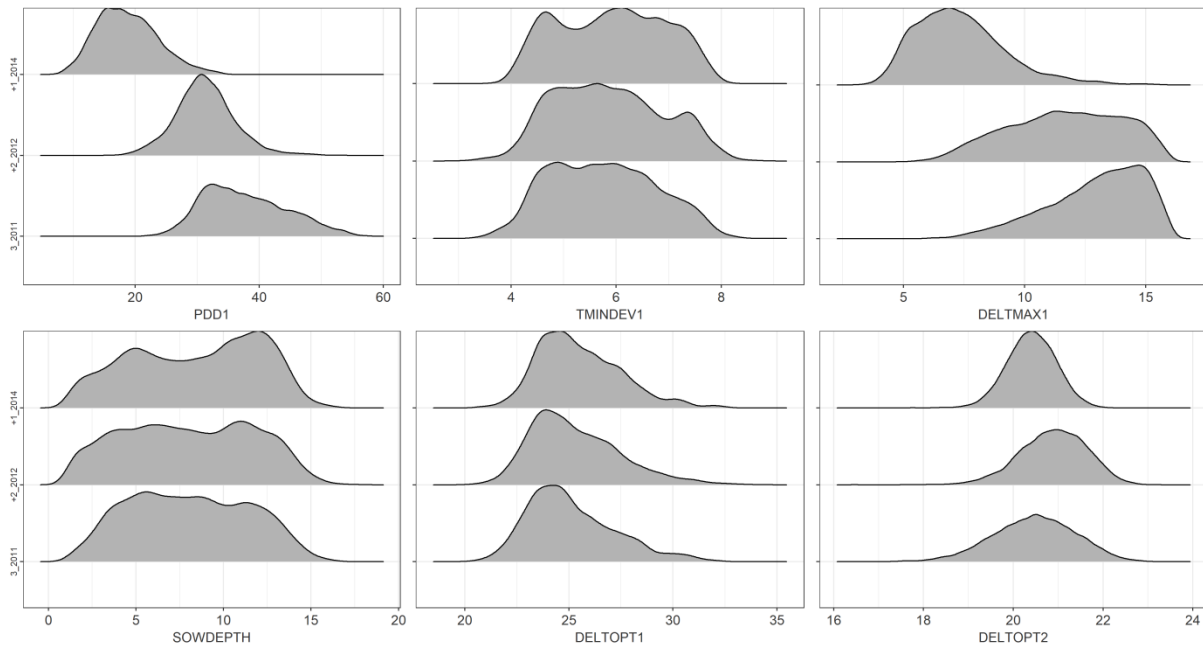


Figure S23 Marginal posterior probability density functions of the 6 estimated parameters after BSU in the true Kraichgau sequence. The y-axis read from bottom to top represent the site-year that was added in the sequential update, corresponding to each density plot. The parameter values are on the x-axis.

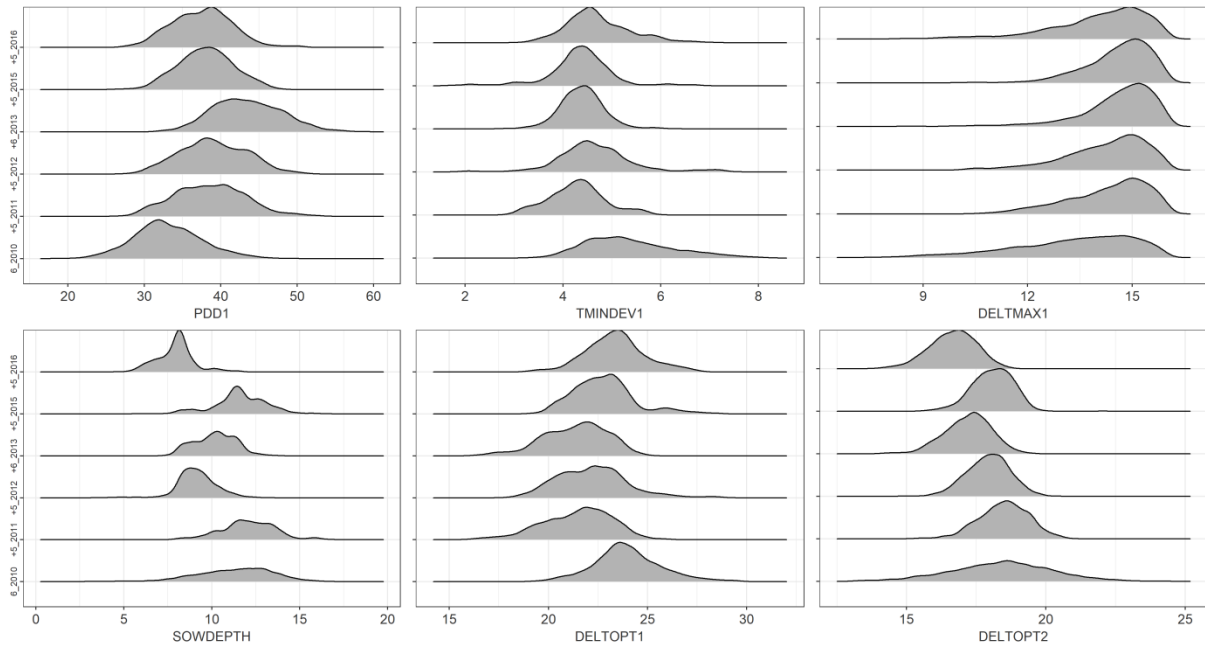


Figure S24 Marginal posterior probability density functions of the 6 estimated parameters after BSU in the true Swabian Alb sequence. The y-axis read from bottom to top represent the site-year that was added in the sequential update, corresponding to each density plot. The parameter values are on the x-axis.



## S6. Parameter distributions and entropy: synthetic sequences

Marginal prior and posterior distributions for the 6 estimated parameters of the SPASS phenology model and the entropy estimates are plotted for the ideal (Figure S25) and controlled cultivar-environment (Figure S26) synthetic sequences.

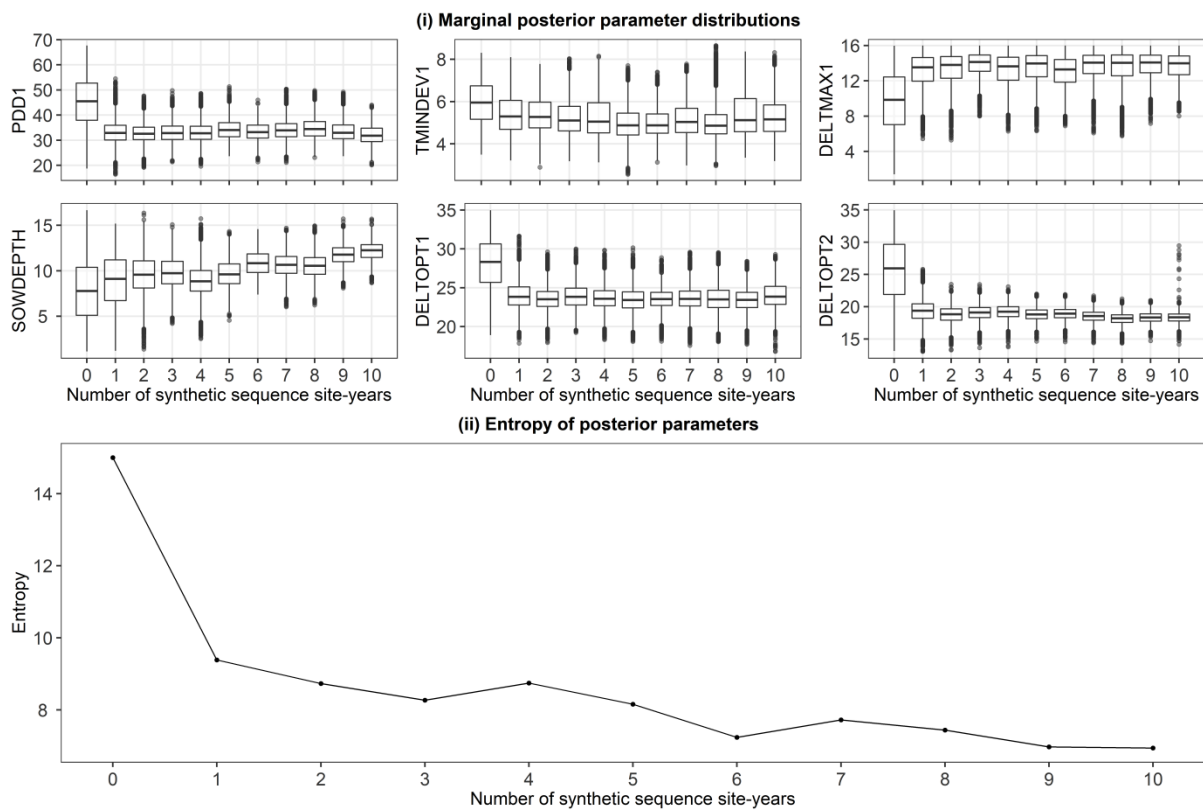


Figure S25 (i) Marginal prior and posterior parameter distributions of the 6 estimated parameters after BSU in the ideal synthetic sequence. Marginal posterior parameter values (y-axis) is plotted against the number of site-years used for calibration (x-axis), starting with the initial prior (0 on x-axis). (ii) Information entropy of the posterior parameter distributions after BSU was applied to the synthetic sequence. Length of the box represents the inter-quartile range (IQR), whiskers extend from the boxes up to  $1.5 \times \text{IQR}$  and values beyond this range are plotted as points. The ranges for parameters SOWDEPTH and DELTOPT2 narrowed through the sequential updates while the remaining parameters do not show a noticeable narrowing in range.

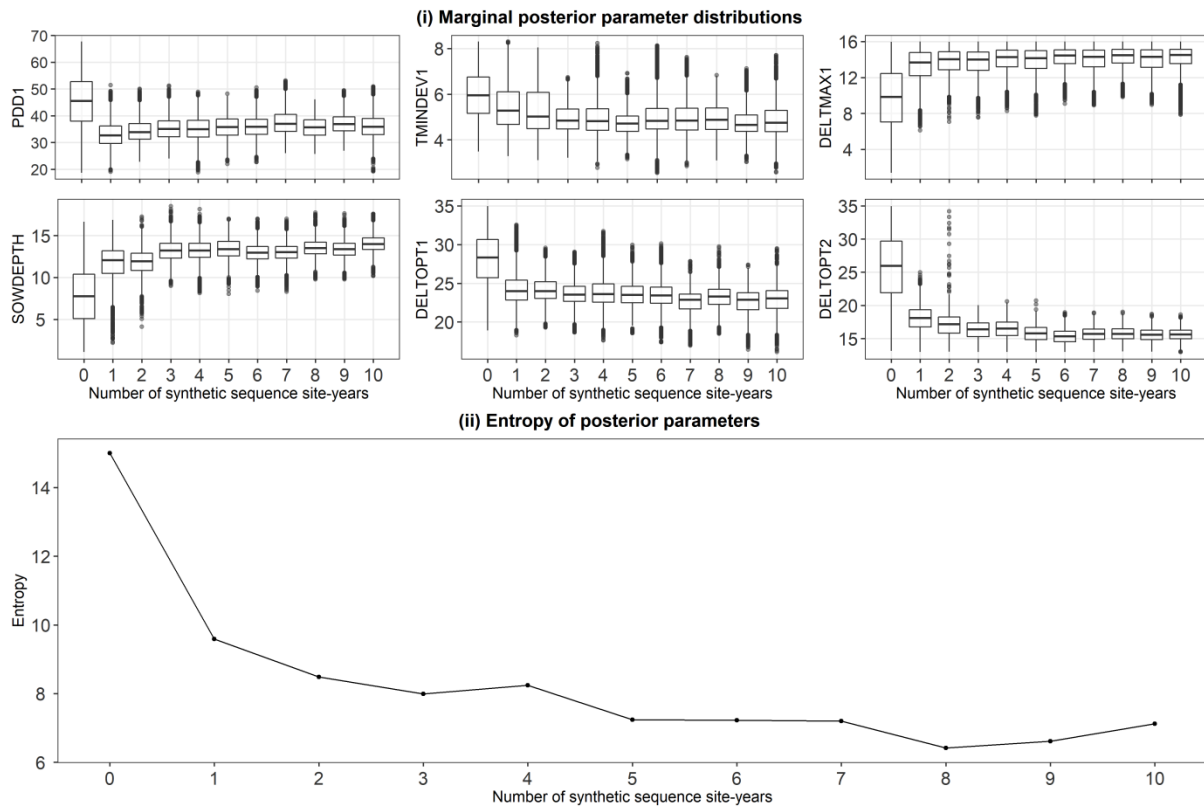


Figure S26 (i) Marginal prior and posterior parameter distributions of the 6 estimated parameters after BSU in the controlled cultivar-environment synthetic sequence. Marginal posterior parameter values (y-axis) is plotted against the number of site-years used for calibration (x-axis), starting with the initial prior (0 on x-axis). (ii) Information entropy of the posterior parameter distributions after BSU was applied to the synthetic sequence. Length of the box represents the inter-quartile range (IQR), whiskers extend from the boxes up to  $1.5 \times$  IQR and values beyond this range are plotted as points. The ranges for parameters SOWDEPTH and DELTOPT2 narrowed through the sequential updates while the remaining parameters do not show a noticeable narrowing in range.

## S7.MCMC diagnostics

The sequential update calibration cases for the true sequences in the Swabian Alb and in Kraichgau are listed in Table S7-1. The number of iterations required to adapt the jump-size (A) were variable (20-580) and dependent on the calibration case. In some cases this number was low because we set the initial pre-adaptation value for the standard deviation of the transition kernel so that the acceptance rate would be between 25% and 35%. This initial value was based on knowledge gained from preliminary calibration test runs. The jump adjustment factor (f) in Table S7-1 influences the standard deviation of the transition kernel (V) such that  $V = \frac{sd}{f}$  where sd is the standard deviation of the prior parameter distributions taken from Table 2 in the main text. With N being the total number of iterations per chain, the total number of iterations across the three chains after burn-in is given by  $T = (N - A) \times 3$ .

On adding a new site-year, the chains were re-initialized and the transition kernel was re-tuned. New data was added to the dataset and the chains were allowed to adapt. The burn-in was variable and dependent of the jump-size adaptation. We ensured that a minimum of 500 accepted samples were generated per chain, that is, a minimum of 1500 total samples across chains were drawn. However, the actual number of samples drawn (T) was higher and dependent of when the Gelman-Rubin convergence diagnostic was  $\leq 1.1$ .

To assess parameter mixing, trace-plots were analysed (examples provided in Figure S27 and Figure S28). Additionally, auto-correlation plots (Figure S29, Figure S30) are provided (coda package in R (Plummer et al., 2006)) and effective sample size (ESS in Table S7-1) were calculated (mcmcse package in R (Flegal et al., 2021), (Vats et al., 2019)). Parameter DELTOPT2 generally showed good mixing and low auto-correlation. The effective sample size between 145 and 332, together with the Gelman-Rubin convergence diagnostic ( $\leq 1.1$ ), provide sufficiently reliable posterior statistics for this study.

Table S7-1: MCMC sampling details for True sequence calibration cases in Kraichgau and the Swabian Alb

Sequence	Calibration case	Number of accepted runs per chain during jump adaptation (A)	Jump adjustment factor (f)	Total accepted samples per chain (N)	Total samples after burn-in in all chains (T) = $(N - A) \times 3$	ESS
True sequence Swabian Alb	6_2010	20	3	1480	4380	236
	6_2010, 5_2011	580	3.97	1100	1560	332
	6_2010, 5_2011, 5_2012	20	5	800	2340	145
	6_2010, 5_2011, 5_2012, 6_2013	40	4.95	820	2340	167
	6_2010, 5_2011, 5_2012, 6_2013, 5_2015	20	5	620	1800	196
	6_2010, 5_2011, 5_2012, 6_2013, 5_2015, 5_2016	240	6.7	1400	3480	159
True Sequence Kraichgau	3_2011	60	5.005	3280	9660	153
	3_2011, 2_2012	20	5	4480	13380	163
	3_2011, 2_2012, 1_2014	20	7.7	5100	15240	168

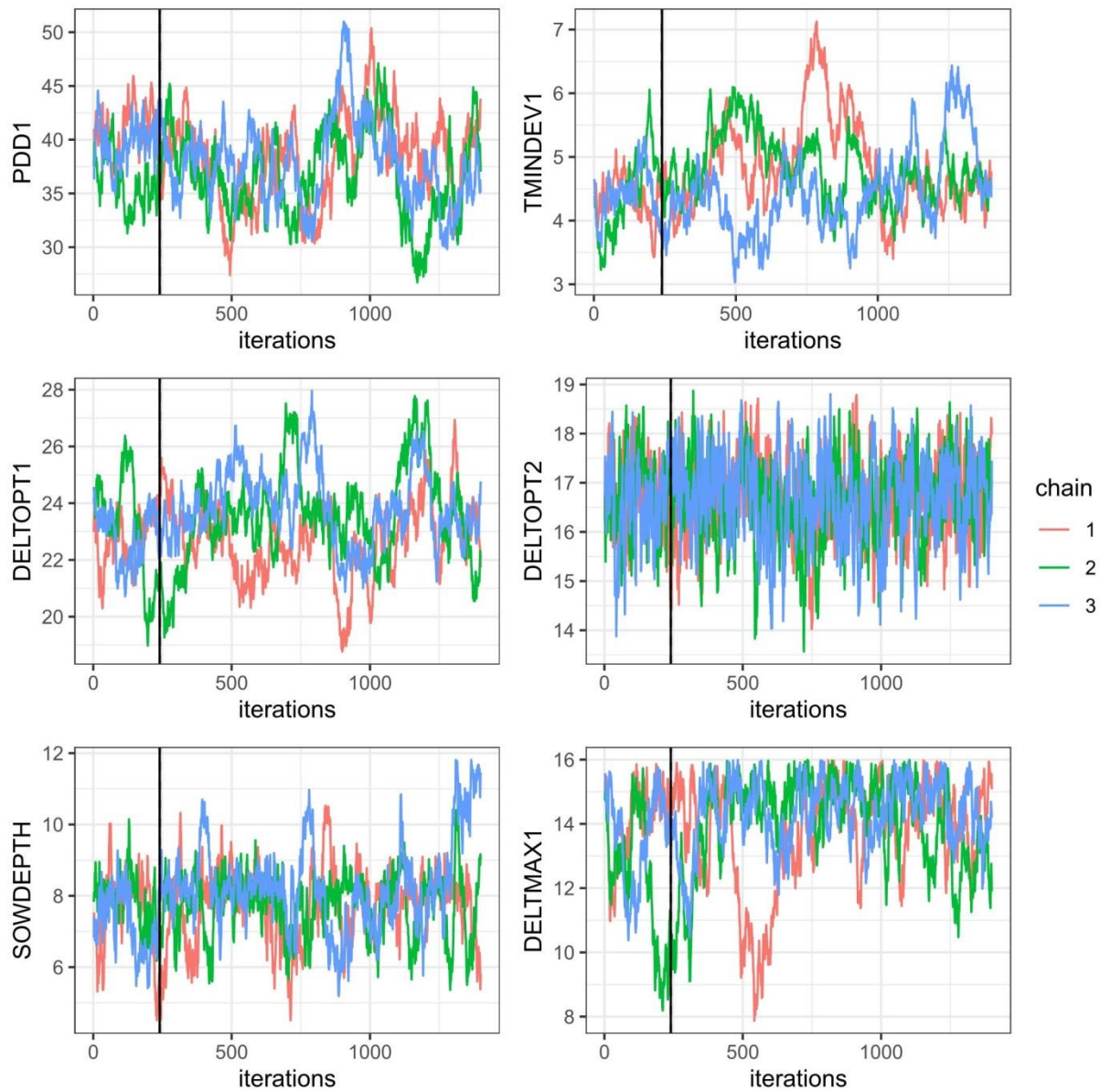


Figure S27 Trace-plots of 6 estimated parameters for the true sequence calibration of SPASS to phenology grown in the Swabian Alb at 6\_2010, 5\_2011, 5\_2012, 6\_2013, 5\_2015, and 5\_2016. The x-axis is the number of iterations and y-axis is the parameter. The colours indicate the three chains. The black solid vertical line indicates the burn-in phase during which the transition kernel was adapted.

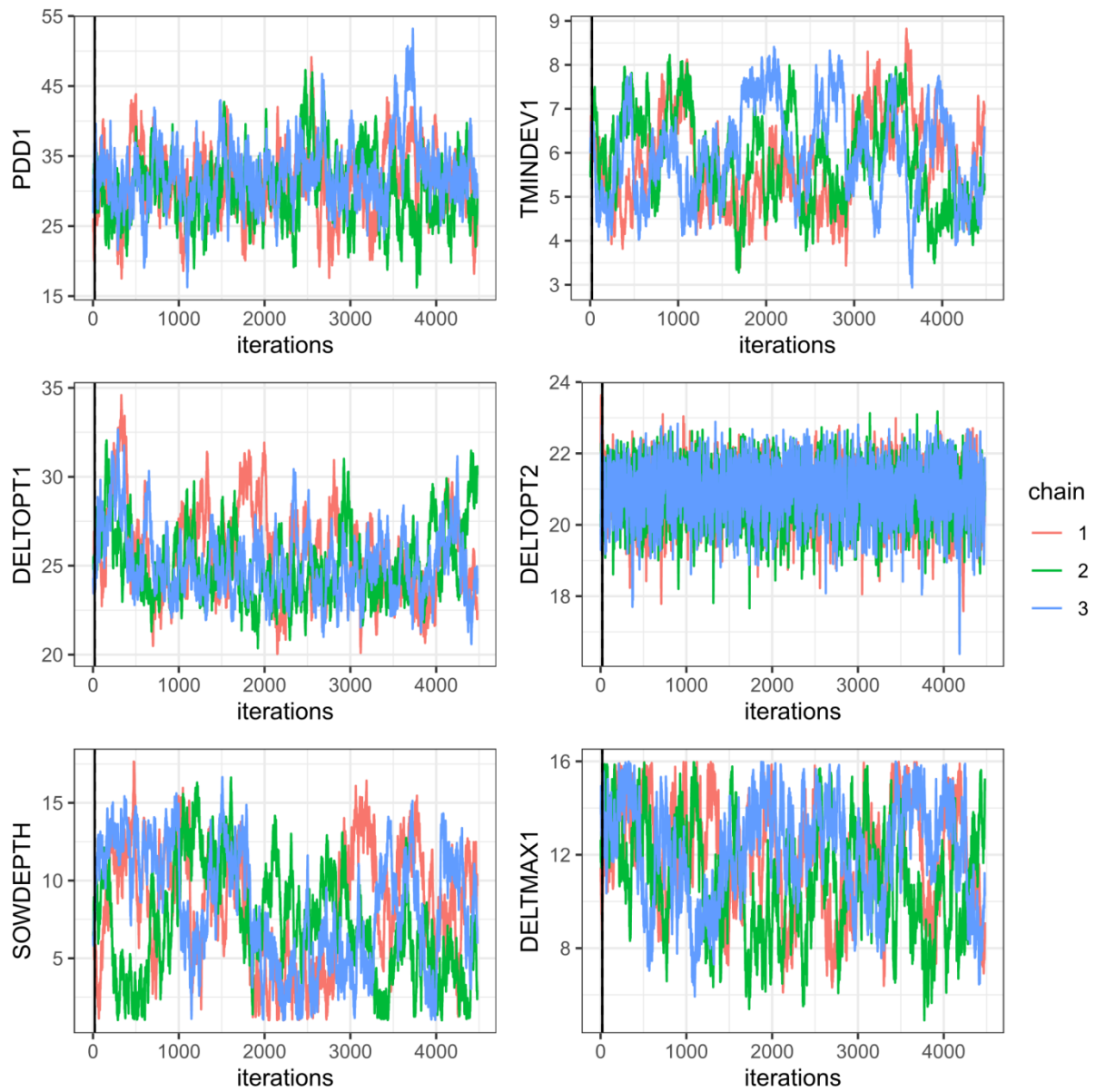


Figure S28 Trace-plots of 6 estimated parameters for the true sequence calibration of SPASS to phenology grown in Kraichgau at 3\_2011 and 2\_2012. The x-axis is the number of iterations and y-axis is the parameter. The colours indicate the three chains. The black solid vertical line indicates the burn-in phase during which the transition kernel was adapted.

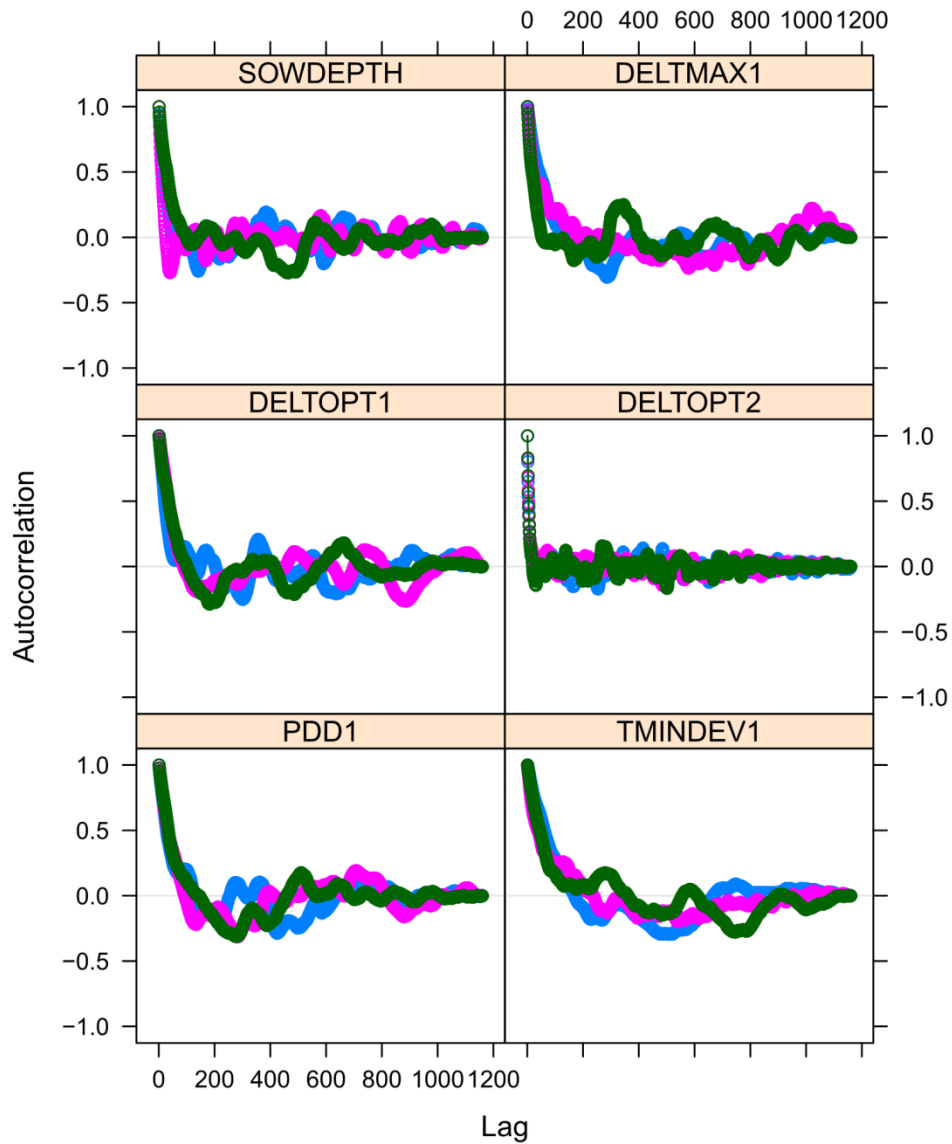


Figure S29 Auto-correlation plots of 6 estimated parameters for the true sequence calibration of SPASS to phenology grown in the Swabian Alb at 6\_2010, 5\_2011, 5\_2012, 6\_2013, 5\_2015, and 5\_2016. The x-axis is the lag distance and y-axis is the auto-correlation. The colours indicate the three chains.

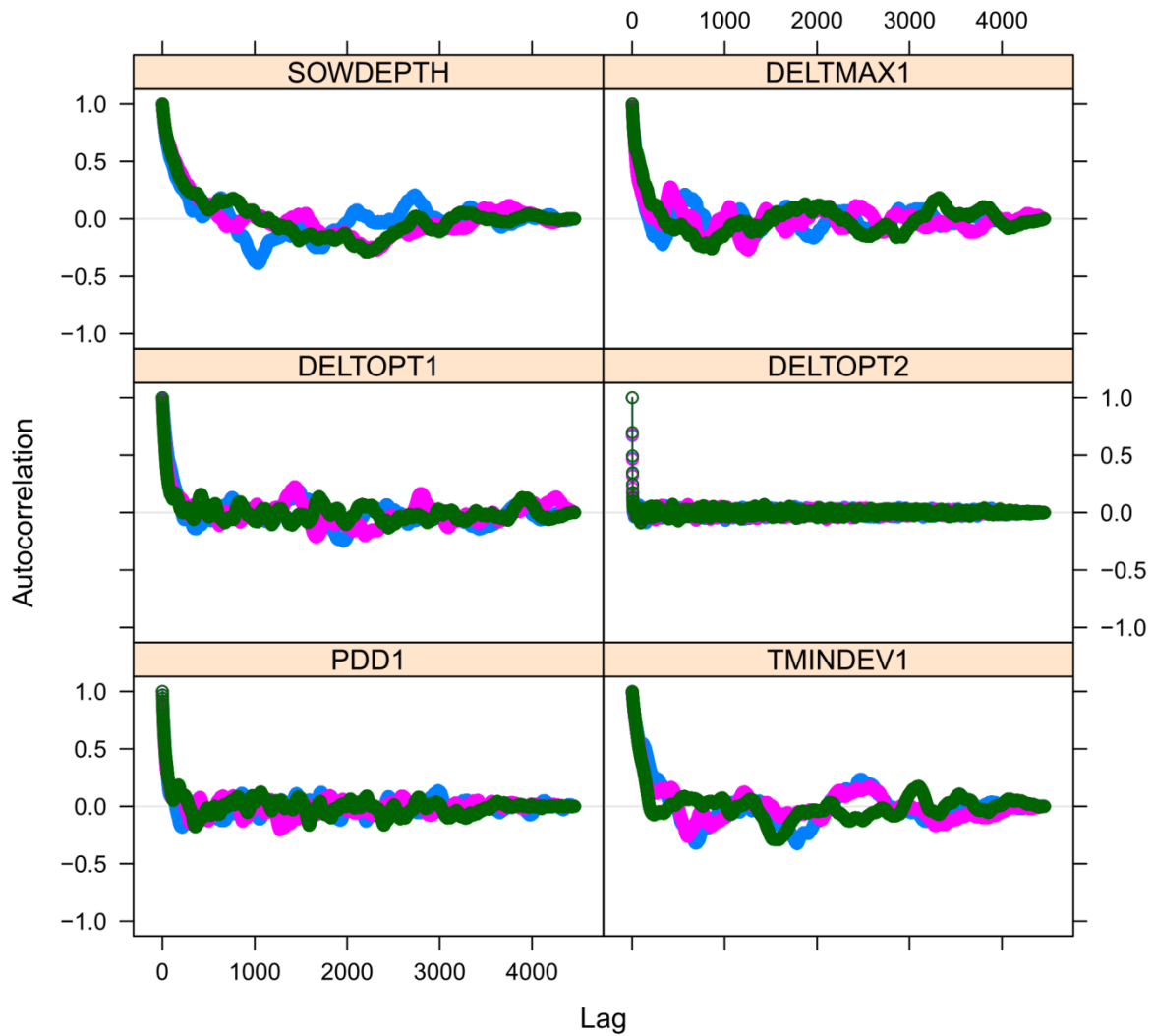


Figure S30 Auto-correlation plots of 6 estimated parameters for the true sequence calibration of SPASS to phenology grown in Kraichgau at 3\_2011 and 2\_2012. The x-axis is the lag distance and y-axis is the auto-correlation. The colours indicate the three chains.



## S8.BSU implementation

Figure S31 explains the concept of Bayesian Sequential Updating and the methodology used to implement it in this study. For the first site-year, a prior based on expert knowledge (initial prior) is used. In the next sequential update with site-year 2 data, the parameter posterior probability distribution after model calibration to site-year 1 can be used as a prior distribution. This can be repeated for n site-years. In this study, however, instead of using the previous site-year as prior for the next update, we use the initial prior and only update the likelihood function with new data.

**Site-year 1:** 
$$P(\theta|Y_{sy1}) = \frac{P(\theta) P(Y_{sy1}|\theta)}{\int_{\theta} P(\theta) P(Y_{sy1}|\theta) d\theta}$$

**Site-year 2:** 
$$P(\theta|Y_{sy2}) = \frac{P(\theta|Y_{sy1}) P(Y_{sy2}|\theta)}{\int_{\theta} P(\theta|Y_{sy1}) P(Y_{sy2}|\theta) d\theta}$$

⋮

**Site-year n:** 
$$P(\theta|Y_{syn}) = \frac{P(\theta|Y_{sy(n-1)}) P(Y_{syn}|\theta)}{\int_{\theta} P(\theta|Y_{sy(n-1)}) P(Y_{syn}|\theta) d\theta}$$

OR

under the assumption that observations from all years are independent

**Initial prior** 
$$P(\theta|Y_{syn}) = \frac{P(\theta) \prod_{x=sy1}^{syn} P(Y_x|\theta)}{\int_{\theta} P(\theta) \prod_{x=sy1}^{syn} P(Y_x|\theta) d\theta}$$

Figure S31 A schematic sketch to explain the concept of Bayesian Sequential Updating (BSU) and its implementation in this study

## S9. Prediction residuals within season

The prediction residuals are plotted against simulated phenology for the Swabian Alb true sequence at the maximum a posteriori probability (MAP) estimate of the model parameters (Figure S32). Here we analyse the prediction residuals within the growing season. The plots from left to right show the inclusion of the subsequent site-year for calibration. The model predicts poorly in the vegetative phase of development, in spite of including more site-years to the calibration sequence. The prediction residuals of the individual site-years are in agreement with the pattern observed in Figure 7 of the main text.

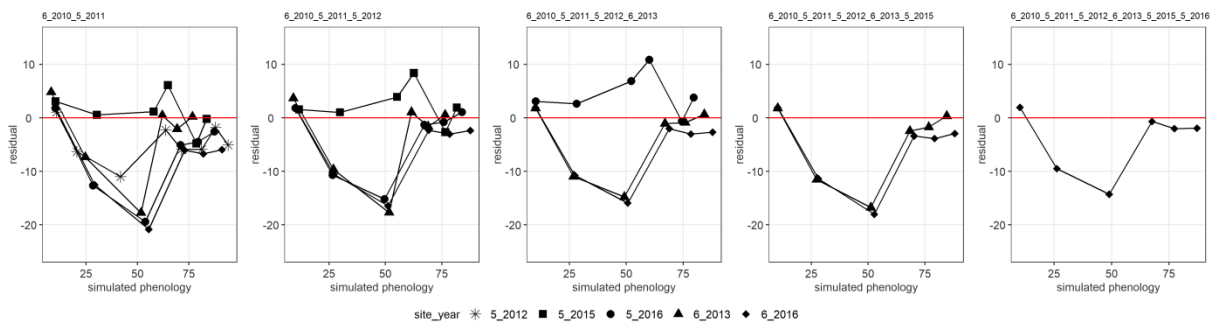


Figure S32 The plots from left to right show the inclusion of the subsequent site-year for calibration in the Swabian Alb sequence. Prediction residuals at the maximum a posteriori probability (MAP) estimate of the model parameters are plotted against simulated phenology. The points correspond to the prediction site-years. The zero residual reference is indicated by the red line.

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