



## Abstract

We report global long-term trends in surface ocean pH using a new pH data set computed by combining  $f\text{CO}_2$  observations from the Surface Ocean  $\text{CO}_2$  Atlas (SOCAT) version 2 with surface alkalinity estimates based on temperature and salinity.

Trends were determined over the periods 1981–2011 and 1991–2011 for a set of 17 biomes using a weighted linear least squares method. We observe significant decreases in surface ocean pH in  $\sim 70\%$  of all biomes and a global mean rate of decrease of  $-0.0018 \pm 0.0004 \text{ yr}^{-1}$  for 1991–2011. We are not able to calculate a global trend for 1981–2011 because too few biomes have enough data for this. In two-thirds of the biomes, the rate of change is commensurate with the trends expected based on the assumption that the surface ocean pH change is only driven by the surface ocean carbon chemistry remaining in a transient equilibrium with the increase in atmospheric  $\text{CO}_2$ . In the remaining biomes deviations from such equilibrium may reflect changes in the trend of surface ocean  $f\text{CO}_2$ , most notably in the equatorial Pacific Ocean, or changes in the oceanic buffer (Revelle) factor. We conclude that well-planned and long-term sustained observational networks are key to reliably document the ongoing and future changes in ocean carbon chemistry due to anthropogenic forcing.

## 1 Introduction

The concentration of atmospheric carbon dioxide ( $\text{CO}_2$ ) is rapidly increasing due to the burning of fossil fuels, cement production, and land use changes (Le Quéré et al., 2014). This drives a net flux of  $\text{CO}_2$  into the ocean, causing its concentration of  $\text{CO}_2$  to increase, which drives a decrease in pH and in the concentration of the carbonate ion ( $\text{CO}_3^{2-}$ , Doney et al., 2009b; Zeebe and Wolf-Gladrow, 2001). These changes in the chemistry of the oceanic carbonate system, collectively referred to as ocean acidification (Gattuso and Hansson, 2011), are a source of concern due to their potential impact on organisms, ecosystems and biogeochemical cycles (Doney et al.,

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observations have only a limited effect on the uncertainties and error of the computed pH (Lauvset and Gruber, 2014).

Even though the use of pH computed from  $f\text{CO}_2$  generates a global data set containing millions of pH observations, the resulting data are still sparse in time and space on a global scale, making the determination of global long-term trends challenging. For surface  $p\text{CO}_2$  this challenge has historically been overcome by binning the data into a very coarse grid (order of  $5\text{--}10^\circ$  in latitude and longitude) by e.g. Lenton et al. (2012) and Takahashi et al. (2002, 2009b), but more recently Fay and McKinley (2013) proposed to aggregate the data into biomes. This type of aggregation is more likely to capture the correct long-term dynamics of a region, as one expects a biome to respond in a more coherent manner to perturbations than a region defined by a latitude/longitude range.

Given the absence of a global observation-based analysis of pH trends, models have so far been the only source of information. The Norwegian Earth System Model (NorESM1-ME), as part of the Coupled Model Intercomparison Project phase 5 (CMIP5, Taylor et al., 2012), simulates a global average pH decrease of  $-0.0017\text{ yr}^{-1}$  (1981–2011), which is largely commensurate with observations reported from the time series stations. A recent study using ten different CMIP5 models, including NorESM1-ME, showed that all models give similar global average pH trends – both in the historical and future scenarios (Bopp et al., 2013).

This secular pH trend of  $-0.0017\text{ yr}^{-1}$  and the low spread between models is expected for an ocean where (i) the surface ocean  $f\text{CO}_2$  follows that in the atmosphere due to the sufficiently rapid exchange of the excess  $\text{CO}_2$  between the atmosphere and the surface ocean, and (ii) where the buffer (Revelle) factor remains spatially uniform and close to constant, as the partial derivative  $\partial[\text{H}^+]/\partial f\text{CO}_2$  is directly related to this quantity (Orr, 2011; Sarmiento and Gruber, 2006). While the latter may be a good assumption for the global average, this may not be the case regionally and locally, where most of the impacts of ocean acidification will occur. Bates et al. (2014) show, for example, not only variations between the high- and low latitude secular trends,

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but also that the Revelle factor is increasing at all time-series stations. This causes a decrease in the buffer capacity and hence a faster increase in surface ocean pH for a given increase in surface  $f\text{CO}_2$  from the addition of  $\text{CO}_2$ . This trend in the Revelle factor is expected as much of the  $\text{CO}_2$  newly added to the surface ocean from the atmosphere will be titrated away by  $\text{CO}_3^{2-}$  causing a decrease in its concentration. This decreases the ability of the surface ocean to “buffer” the pH against further uptake of  $\text{CO}_2$ , thus increasing the Revelle factor (Sarmiento and Gruber, 2006). But there are also other factors influencing the Revelle factor, mainly those processes that affect DIC and alkalinity, such as changes in ocean productivity and calcification, while changes in temperature and salinity are of minor importance (Sarmiento and Gruber, 2006).

Thus local and regional changes in the buffer factor have the potential to substantially decouple the pH trends from those of the surface ocean  $f\text{CO}_2$ , potentially causing a more variable pattern in the pH trends. The complex spatial variability, identified by Bates et al. (2014) and others (e.g. Tjiputra et al., 2014) support this hypothesis. This also shows that analyses of global pH trends are required for a comprehensive picture, including the regional distribution of changes and the dynamics of the changing ocean carbon system. Global analyses are also necessary for the validation of model results, for underpinning and interpreting response studies from organism to ecosystem level, and for optimizing the planning of continued and future observational networks.

Here we take advantage of the approach of Lauvset and Gruber (2014) to determine global ocean pH trends, and their drivers, using pH data calculated from the more than 10 million observations of surface ocean  $f\text{CO}_2$  that have been made available through the Surface Ocean  $\text{CO}_2$  Atlas (SOCAT) project (Bakker et al., 2014; Pfeil et al., 2013). Although pH is the main parameter of our study  $f\text{CO}_2$  has been carried through all of our analyses in order to determine how effects of carbonate chemistry variations causes the evolution of pH to differ from that expected from  $f\text{CO}_2$  alone in many regions. Finally we use the long term pH trends derived from a global earth system model, the NorESM1-ME, in order to illustrate how important spatial variability is for the representativeness of our trend results.

## 2 Data and methods

We calculated pH in the surface ocean by a two-step calculation using observations of  $f\text{CO}_2$ , sea surface temperature (SST), and sea surface salinity (SSS) from SOCAT version 2, (Bakker et al., 2014). In the first step, alkalinity was calculated from SSS and SST using the algorithms developed by Lee et al. (2006) and Nondal et al. (2009). The Nondal et al. (2009) algorithms were developed specifically for the high-latitude (> 60° N) Atlantic Ocean, and were used only there. Whenever no measured SSS was available in the SOCATv2 data set the climatological World Ocean Atlas SSS value (Antonov et al., 2010) – which is included in the SOCATv2 data product – was used instead. The SOCAT SSS data have not been quality controlled and might therefore be biased. Lauvset and Gruber (2014) showed that this potential bias does not greatly affect the precision of the pH trends. It may affect the accuracy of the calculation but for our purpose of determining long-term trends the accuracy (i.e. the lack of bias in the data) is of less importance as long as the precision is good enough, and assuming that any bias remains constant over time. In the second step, pH on the total scale at in situ temperature was calculated from the estimated alkalinity and the observed  $f\text{CO}_2$  using CO2SYS (Lewis and Wallace, 1998). We used the  $K_1$  and  $K_2$  constants from Mehrbach et al. (1973) refit by Dickson and Millero (1987), and the borate to salinity ratio from Uppström (1974). Since we use CO2SYS this calculation also gives us dissolved inorganic carbon (DIC) and all other variables of the ocean carbon chemistry system.

Quite a few of the data fall outside the valid ranges for input data for the Lee et al. (2006) and Nondal et al. (2009) alkalinity algorithms and are lost in this step. There remains 7 381 013 data points of pH (and alkalinity) over the global ocean in the time period 1973–2011. The global calculation error (precision) for pH is  $0.0032 \pm 0.0005$ , and the calculated pH compares well to observed pH at crossover locations in the Atlantic Ocean (Lauvset and Gruber, 2014). Before analysis the pH data were bin averaged into monthly  $1^\circ \times 1^\circ$  bins, using no extrapolation or interpolation

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of the data. The global data set was divided into the 17 ocean biomes defined and named by Fay and McKinley (2014) as shown in Fig. 1. Here, we only evaluate trends in the open ocean, and data from coastal regions shallower than 250 m, based on the ETOPO2 bathymetry and those with salinity < 20 were removed.

In each biome a least squares linear regression weighted with Tukey's bisquare method was used to determine the long-term pH trend. For the long-term trend determination we required each biome to have at least three observations in each decade (1981–1990, 1991–2000, and 2001–2011). While this criterion was met in only 8 biomes for the period 1981–2011, all 15 had sufficient data for the period 1991–2011. Both ordinary and weighted least squares regressions were carried out, but we chose a weighted least squares regression over an ordinary least squares regression since this is less sensitive to outliers in the data. This makes the statistics of the regression more robust, but generally this choice does not significantly affect the results presented here. All regression results are presented with the standard error of the slope (se), which represents its 68 % confidence interval, and the root mean square error (RMSE). The RMSE is used as a measure of interannual variability.

Before the regression analysis was carried out two corrections were applied to the data: deseasonalization and removal of spatial bias. The importance of these corrections, particularly in data sparse biomes such as those in the Southern Ocean, was recently highlighted by Fay et al. (2014). The seasonal cycle in the data was removed following Takahashi et al. (2009b), using the long-term average seasonal cycle as contained in our data for each biome for this correction. However, we find that using the climatological seasonal cycle – calculated using the Takahashi  $p\text{CO}_2$  climatology (Takahashi et al., 2009a) – do not significantly affect the results. To correct for any spatial bias in the large scale biomes the difference between the climatological value in each  $1^\circ \times 1^\circ$  bin and the biome mean climatological value was subtracted from the observed value in each  $1^\circ \times 1^\circ$  bin. There is no difference between this method and simply subtracting the climatological value in each  $1^\circ \times 1^\circ$  bin, but our approach retains the biome mean values. It should be noted that the computed trends in some

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biomes are sensitive to which climatological data is used for the spatial bias correction: subtracting the climatological value vs. subtracting the long-term average in each  $1^\circ \times 1^\circ$  bin. Mostly this is because in some  $1^\circ \times 1^\circ$  bins, the long-term average is biased towards the last decade, which has significantly more data than earlier periods.

A statistical test was performed to ensure that the effects of these corrections do not corrupt our analysis: results after applying one or both corrections were compared to results after applying none using a one-way analysis of variance (ANOVA, see e.g. Vijayvargiya, 2009). The deseasonalization removes scatter in the data and leads to more robust regressions, but does not significantly ( $p$  value  $< 0.05$ ) affect the long-term trend in any biome or time period. The spatial bias correction has in most biomes no statistically significant impact on the long-term trend, but because it does affect the trend in some biomes we decided to keep it applied. The long-term pH trend is also much more sensitive to this correction than the  $f\text{CO}_2$  trend, mostly because the pH trend is very small and thus more sensitive to any data correction.

The pH change expected from a certain change in  $f\text{CO}_2$  assuming everything else (i.e. alkalinity, DIC, SST, Revelle factor) remains constant was calculated using  $\Delta\text{pH}/\Delta f\text{CO}_2 = \partial\text{pH}/\partial f\text{CO}_2$ . The partial derivative was estimated in CO2SYS using  $0.01 \mu\text{atm}$  increments in  $f\text{CO}_2$ . We used the same equation to evaluate what global average  $f\text{CO}_2$  change the global long-term trend in pH is consistent with, but then using  $-0.001$  incremental changes in pH. Using such small increments allowed us to abide by the assumption of constant DIC and Revelle factor.

In each biome the long-term trend in pH was decomposed into the effects of changes in SST, SSS, alkalinity, and DIC. First the impact of each of these drivers on the  $f\text{CO}_2$  trend was determined following Takahashi et al. (1993), Eqs. (2)–(5), we then converted our results to the impacts on  $[\text{CO}_2]$  and on  $[\text{H}^+]$  following Eq. (1.5.87) in Zeebe and Wolf-Gladrow (2001), and finally we determined the impact on pH. The DIC data and dissociation constants required for these calculations were calculated in CO2SYS from the  $f\text{CO}_2$  and alkalinity pair in the same calculation that gave us pH.

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Atlantic ice covered (NA-ICE) biomes, due to the lack of data (< 20 data points) hence they are not further discussed in the paper. Unfortunately, these are the Arctic biomes where the earliest impacts of ocean acidification are expected (Steinacher et al., 2009).

The regions with sufficient data, but without statistically significant trends, i.e., the North Pacific subpolar seasonally stratified (NP-SPSS) biome for the period 1981–2011, and the Southern Ocean subtropical seasonally stratified (SO-STSS) and ice covered (SO-ICE) biomes for the period 1991–2011, are characterized by large RMSE and a substantial amount of decadal variability, which is likely masking the long-term trends. In addition to these three biomes where the trends are statistically indistinguishable from zero, the South Pacific subtropical permanently stratified (SP-STPS) biome is likely biased by its low data density, and will not be further discussed. This decision was corroborated was by comparing the pH trend in the fully-sampled model results with the sub-sampled model results (Fig. 5): the SP-STPS biome is the only one where the difference in these trends is statistically significant at the 95 % confidence level.

The area-weighted global average pH decrease, including only statistically significant trends, is  $-0.0018 \pm 0.0004 \text{ yr}^{-1}$  for the period 1991–2011. No global trend can be computed for the 1981–2011 period, as the number of biomes with trend estimates is too small, but the Pacific Ocean trend in this period is  $-0.0019 \pm 0.0001 \text{ yr}^{-1}$ . Within the uncertainty limits the global 1991–2011 trend is comparable to the global trend in the fully-sampled NorESM1-ME model results ( $-0.0017 \text{ yr}^{-1}$ ) and to the average trend of  $-0.0018 \pm 0.0003 \text{ yr}^{-1}$  over the seven time series evaluated by Bates et al. (2014). Assuming that everything, including the Revelle factor, remains constant, this global average pH trend corresponds to a rate of increase in surface ocean  $f\text{CO}_2$  of  $1.75 \pm 0.4 \mu\text{atm yr}^{-1}$ , which is roughly the rate of increase in atmospheric  $p\text{CO}_2$ . Regionally, however, the response of the ocean carbon system to the atmospheric forcing is more variable (Fig. 1).

In the North Atlantic subpolar seasonally stratified (NA-SPSS) biome the observed pH trend is  $-0.0020 \pm 0.0004 \text{ yr}^{-1}$ . This is right in between the trend observed at the

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increase is assumed to be the only driver for the pH changes, we find that in 5 of the 12 biomes the observed pH trends significantly differ from the expected pH change. This is due either to the uncertainty in the observed trends, to associated changes in carbonate chemistry, or to the surface ocean  $f\text{CO}_2$  trends being significantly different from that in the atmosphere. However, the observed pH trends significantly differ from the expected pH change calculated using the observed  $f\text{CO}_2$  trend in 5 of the 12 biomes also (Fig. 6). Only 3 of the 5 biomes are the same in both cases. Thus, the surface ocean  $f\text{CO}_2$  trend not exactly mirroring the atmospheric cannot explain the discrepancy between expected and observed pH trends in most biomes. It may be an explanation in the equatorial Pacific biomes (EP-EQU and WP-EQU) where there is no discrepancy between observed and expected pH trends when the observed  $f\text{CO}_2$  trend is used to calculate the expected pH change (Fig. 6), but a significant difference when an atmospheric rate of change is assumed.

The observed pH trend is more often smaller than that expected for the ocean mirroring the atmospheric  $f\text{CO}_2$  change than vice versa. Only the EP-EQU biome has an observed pH change larger than the expected (Fig. 6). In the biomes where the observed trend differs from the expected trend there are indications which point to changes in carbonate chemistry, e.g. that the temporal trends in both  $\partial[\text{H}^+]/\partial p\text{CO}_2$  (Gattuso and Hansson, 2011) and Revelle factor are consistent with the differences between expected and observed pH change. However, given the combined calculation errors, generally high level of noise in our data, and relatively few data points, only some of these indications are statistically significant. Further analysis of these spatial patterns needs to be undertaken using independent pH data, preferably direct measurements in order to quantify any possible biases in the results due to our pH being a calculated variable. A combination of SOCAT data with repeat hydrography and time-series data would be ideal but this is outside the scope of this study.

### 3.3 Major driving forces behind the observed pH and trends

The decomposition of the  $f\text{CO}_2$  and pH trends confirms (Figs. 8–10) that in all biomes the long-term increase in DIC is by far the dominant driver for the long-term pH changes. Thus, knowledge about the changes in ocean DIC therefore is the most important in understanding – and predicting – changes in ocean pH (Table 2). This is not unexpected since the open ocean is in – or very close to – chemical equilibrium with the atmosphere (Lauvset and Gruber, 2014). Thus the surface ocean is taking up  $\text{CO}_2$  from the atmosphere in order to re-establish a chemical equilibrium, leading to a corresponding increase in  $p\text{CO}_2$  and DIC. It must be noted that since we do not have measurements of alkalinity this parameter is calculated from SST and SSS, and the relatively large uncertainties in these calculations may add a degree of uncertainty to the decomposition. Due to a lack of independent data this is not further evaluated in this study.

In the Atlantic Ocean biomes the second most important driver is SST (Fig. 8), which mostly has a positive change and therefore has limited the DIC increase required to maintain an  $f\text{CO}_2$  growth rate similar to that in the atmosphere. SST is the second most important driver also in the Pacific Ocean biomes (Fig. 9), but here SST decreased in many biomes leading to an enhanced increase in DIC through uptake from the atmosphere. In the Southern Ocean biomes alkalinity changes have a significant impact on the trends (Fig. 10), which also modulates the DIC changes. Decreasing alkalinity over time increases  $f\text{CO}_2$  so that the DIC change required to maintain a sea surface  $f\text{CO}_2$  growth rate similar to the atmospheric is reduced.

In most biomes there is a residual between the sum of the four components and the observed trend (Fig. 7). Lenton et al. (2012) performed a similar analysis and attributed such residuals to the use of a spatial mean Revelle factor, the approximations underlying the Takahashi et al. (1993) equations, and the assumption of linear trends in all variables. We tested whether variable data coverage is also an important contributor to this residual by subsampling the NorESM1-ME simulated pH data and comparing

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the resulting 1981–2011 decomposition with the decomposition determined using the full model output. Figure 7 illustrates that in most biomes there are similar residuals between the sum of the four components and the actual trends in the sub-sampled and fully-sampled model fields as well. We can, therefore, find no evidence to show that poor data coverage is of major importance in determining what drives the change in surface ocean pH.

### 3.4 Recent changes in the Southern Ocean biomes

While this study generally does not have statistically significant results in the Southern Ocean our results do indicate significant decadal variations in the trends (Table 1). This is consistent with the changing  $f\text{CO}_2$  trends in a recent study by Fay et al. (2014) as well as previous findings of a change in the  $\text{CO}_2$  sink in this region (e.g. Fay and McKinley, 2013; Landschützer et al., 2014). In order to investigate these recent trend changes in the Southern Ocean, we also decompose the 2001–2011 trends in the Southern Ocean biomes (Table 3). In the SO-STSS biome we find no change in what drives the pH trend for this decade compared to the longer period. However, the change in the observed pH trend (Table 3) appears to be dominated by the change in DIC as this is approximately four times larger (more negative) for the period 2001–2011 than for the period 1981–2011, and therefore likely dominates over the thermally induced reduction of the pH trends. In the SO-SPSS biome we find an increase in the DIC driven pH trend within the period 2001–2011 compared to the period 1981–2011, and since the DIC component is negative this supports the conclusion drawn by Fay and McKinley (2013) that a reduction in vertical DIC support causes the weakening of both  $f\text{CO}_2$  and pH trends in this region. In the SO-ICE biome the sign of the non-thermal drivers appears to change within the last decade, potentially driven by the recent Antarctic ice melt and ice-sheet melting driven iron fertilization (Death et al., 2014).

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### 3.5 Spatial variability

In both the observations and the sub-sampled model results we see significant regional differences in the pH trends (Fig. 5). Note that the actual simulated pH trends in each biome are not directly comparable with the observed trends since the model is a coupled climate model, which simulates its own internal climate variability. We therefore compare the fully sampled and the sub-sampled model results, and the fully sampled model results show much more uniform pH trends (Fig. 5). While these differences are mostly statistically indistinguishable within the uncertainties, it highlights the need for careful consideration of representativeness when comparing model-derived future changes and trends based on data. Figure 5 shows that in the SO-SPSS, EP-EQU, AEQU, and NP-STPS biomes the sub-sampled trend is within  $\pm 0.0001$  of the fully-sampled pH trends, and an ANOVA analysis shows that only in the SP-STPS biome are the two model trends significantly different. Thus the trends based on the existing observational coverage are overall representative of the respective biomes, and it is unlikely to be major biases in our results due to low data density. However, the uncertainties in the long-term pH trend estimates remain too large, both in observations and model (Fig. 5) and this prohibits a mechanistic understanding the observed changes in most biomes. Improved sampling strategies are necessary to reduce these uncertainties and thereby improve our understanding of surface ocean carbon chemistry changes today and in the future.

This highlights the importance of both maintaining the observational networks already in place – like the voluntary observing ship (VOS) network in the North Atlantic (Watson et al., 2009) – and instigating new ones in less well-covered ocean regions. Of particular importance is improved data coverage in the Southern Pacific Ocean (SP-STPS) where the data density as of today it too low for a robust analysis of long term pH trends.

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## 4 Conclusions

Global surface ocean pH has decreased over the past 20 years by  $-0.0018 \pm 0.0004 \text{ yr}^{-1}$ . There are however large regional variations with trends ranging from  $-0.0024 \text{ yr}^{-1}$  in the Indian Ocean (IO-STPS) biome to no significant change in the polar Southern Ocean (SO-ICE) biome. Our estimated global trend is very comparable to the trends found at time-series stations and to the global average trends in the CMIP5 models. In all biomes, the pH trend is predominantly driven by changes in DIC, implying that the surface ocean pH decline is a direct response to the increasing uptake of atmospheric  $\text{CO}_2$ . Despite this, the  $f\text{CO}_2$  and pH trends do not exactly mirror each other, which is potentially linked to trends in the surface ocean buffer (Revelle) factor over the past decades. In some biomes this leads to smaller pH changes than expected from the  $f\text{CO}_2$  change, while in others regions, the pH changes are larger than expected. Thus, knowledge of both the changing ocean DIC and the changing ocean buffer (Revelle) factor is important for understanding and accurately determining the changing ocean pH.

There are regional differences in the pH trends. It is likely that these are caused by spatial heterogeneity in the concurrent changes in buffer (Revelle) factor, while spatial heterogeneity in the surface ocean  $f\text{CO}_2$  trends seems to have only a minor effect. Our comparison between fully-sampled model and sub-sampled output from the NorESM1-ME model indicates that variable data coverage only presents a problem in the South Pacific. This nicely highlights the overall success of the scientific community in creating observational networks that reduce data coverage issues. The many scientific studies arising from this effort – among many others the recent publications by Nakaoka et al. (2013), Landschützer et al. (2013, 2014), and Schuster et al. (2013) – show that we have come a long way in understanding how ocean carbon chemistry is evolving in a world perturbed by fossil fuel emissions. The uncertainties in the trends presented here are, however, substantial and this largely prevents a more thorough understanding of current changes. Filling the remaining gaps in our surface ocean data,

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and quantifying spatial patterns in the carbon chemistry parameters, is, therefore, still of great importance in order to accurately assess the open ocean carbon chemistry changes.

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**Table 1.** Results and statistics of the regression analysis of  $f\text{CO}_2$  and  $\text{pH}_{\text{insitu}}$  vs. time. Bold text indicates biomes where the results are not statistically significant (95% confidence). No number is given if a biome does not have enough data to calculate the trend in a given time period.

Region	1981–2011				1991–2011			
	pH		$f\text{CO}_2$		pH		$f\text{CO}_2$	
	slope	rmse	slope	rmse	slope	rmse	slope	rmse
NP-SPSS	<b>-0.0003 ± 0.0005</b>	<b>0.041</b>	1.29 ± 0.17	16.2	0.0013 ± 0.0005	0.038	0.78 ± 0.22	16.1
NP-STSS	–	–	1.36 ± 0.09	10.5	-0.0010 ± 0.0005	0.031	1.54 ± 0.10	8.9
NP-STPS	-0.0016 ± 0.0002	0.020	1.45 ± 0.08	10.3	-0.0019 ± 0.0002	0.018	1.57 ± 0.11	9.9
WP-EQU	-0.0010 ± 0.0002	0.016	1.27 ± 0.15	17.8	-0.0012 ± 0.0002	0.015	1.57 ± 0.23	17.3
EP-EQU	-0.0023 ± 0.0003	0.023	2.56 ± 0.36	28.2	-0.0026 ± 0.0002	0.023	3.33 ± 0.46	27.9
SP-STPS	-0.0019 ± 0.0002	0.020	1.32 ± 0.11	12.0	-0.0022 ± 0.0003	0.020	1.15 ± 0.16	12.3
NA-SPSS	–	–	1.18 ± 0.16	15.4	-0.0020 ± 0.0004	0.028	1.11 ± 0.19	14.2
NA-STSS	–	–	1.78 ± 0.14	12.3	-0.0018 ± 0.0003	0.015	1.78 ± 0.16	12.5
NA-STPS	–	–	1.46 ± 0.09	8.5	-0.0011 ± 0.0002	0.012	1.51 ± 0.11	8.6
A-EQU	–	–	1.62 ± 0.25	16.6	-0.0016 ± 0.0003	0.014	1.76 ± 0.31	15.7
SA-STPS	–	–	0.95 ± 0.21	16.7	-0.0011 ± 0.0005	0.024	0.82 ± 0.31	17.0
IO-STPS	-0.0024 ± 0.0004	0.023	1.44 ± 0.24	13.6	-0.0027 ± 0.0005	0.025	1.56 ± 0.26	13.5
SO-STSS	<b>-0.0006 ± 0.0004</b>	<b>0.032</b>	1.77 ± 0.1	10.8	<b>-0.0004 ± 0.0004</b>	<b>0.032</b>	1.91 ± 0.12	10.8
SO-SPSS	-0.0020 ± 0.0002	0.020	1.40 ± 0.09	9.1	-0.0021 ± 0.0002	0.020	1.45 ± 0.11	9.0
SO-ICE	–	–	0.70 ± 0.28	24.4	<b>-0.0002 ± 0.0004</b>	<b>0.029</b>	0.25 ± 0.33	24.3

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**Table 3.** Decomposition of the 2001–2011  $f\text{CO}_2$  and  $\text{pH}_{\text{insitu}}$  trends in the Southern Ocean into their major drivers.

Region	pH						$f\text{CO}_2$					
	theta	salinity	DIC	alkalinity	sum	observed	theta	salinity	DIC	alkalinity	sum	observed
SO-STSS	−3.7	−0.79	8.46	−2.48	1.49	$0.0032 \pm 0.0010$	3.4	0.73	−7.78	2.28	−1.37	$1.56 \pm 0.39$
SO-SPSS	1.11	−0.07	−1.28	−0.05	−0.29	$−0.0011 \pm 0.0006$	−1.06	0.07	1.22	0.05	0.28	$0.89 \pm 0.22$
SO-ICE	−0.62	−0.05	1.00	−0.39	−0.06	$0.0006 \pm 0.0009$	0.59	0.05	−0.95	0.37	0.06	$0.21 \pm 0.74$

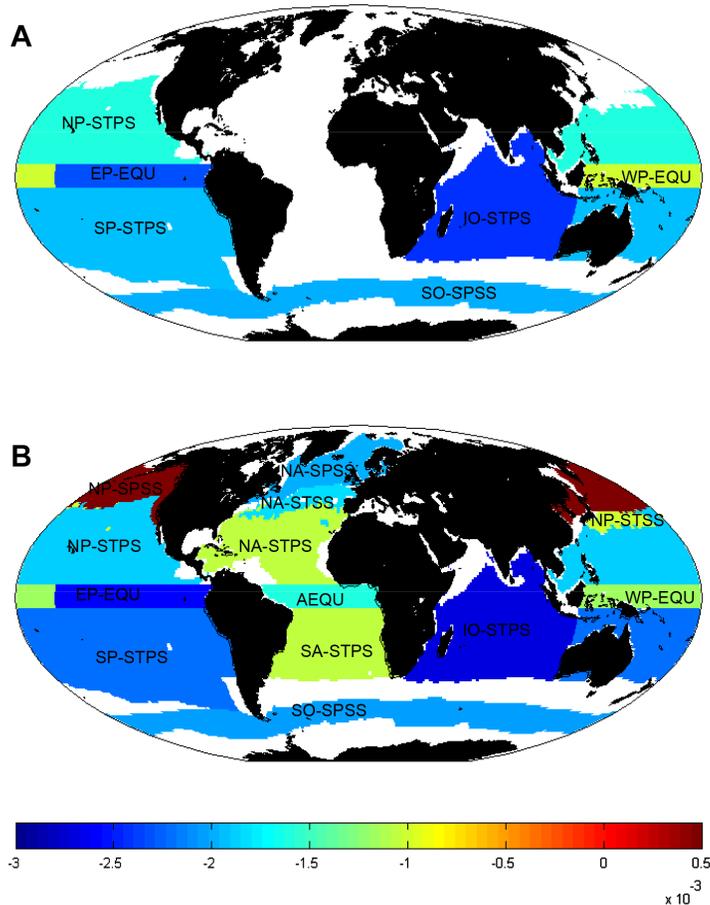
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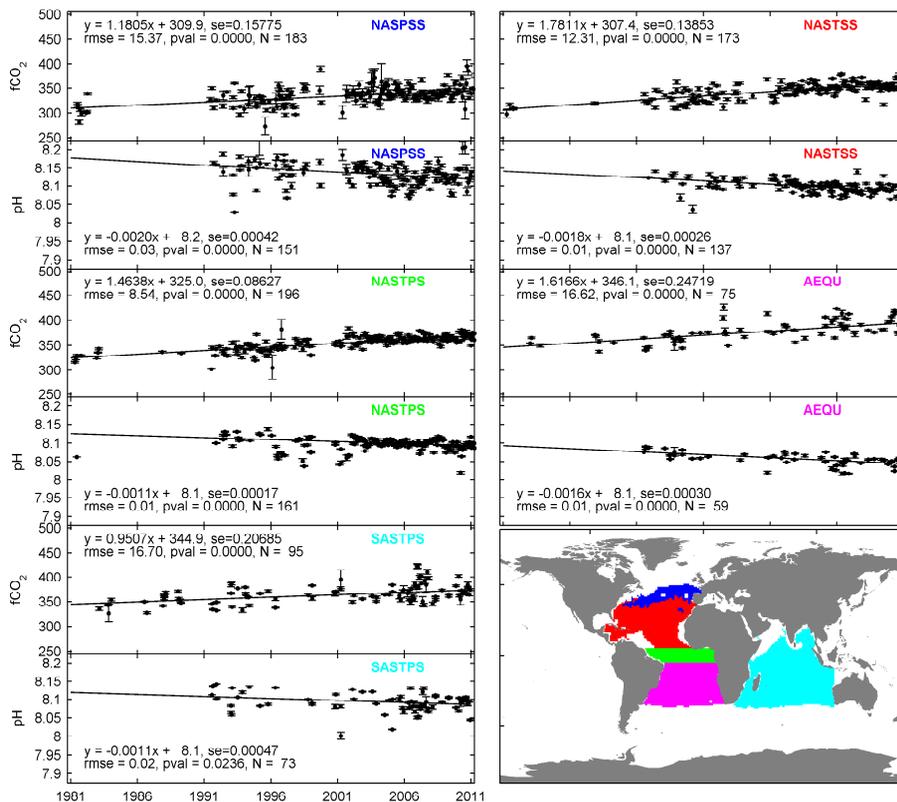
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**Figure 1.** A map of the Fay and McKinley (2014) biomes which have (a) a statistically significant pH trend in the period 1981–2011, and (b) the biomes with a statistically significant pH trend in the period 1991–2011.

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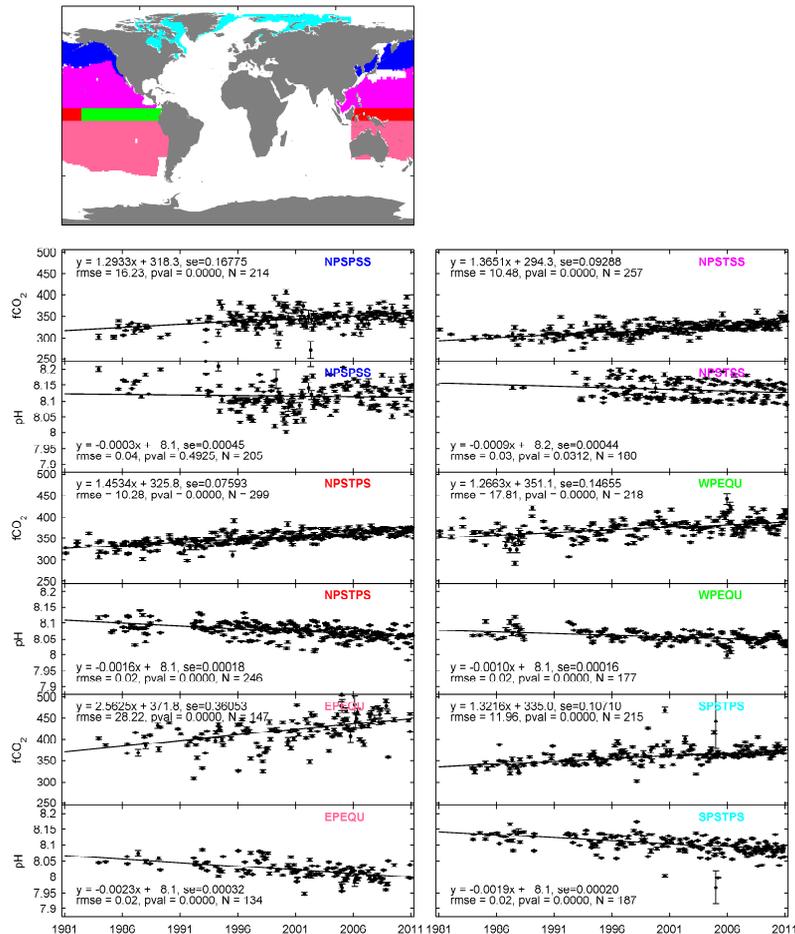
**Figure 2.** Long term  $f\text{CO}_2$  and pH trend (1981–2011) in the five Atlantic Ocean biomes.

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**Figure 3.** Long term pH trend (1981–2011) in the five Pacific Ocean biomes.

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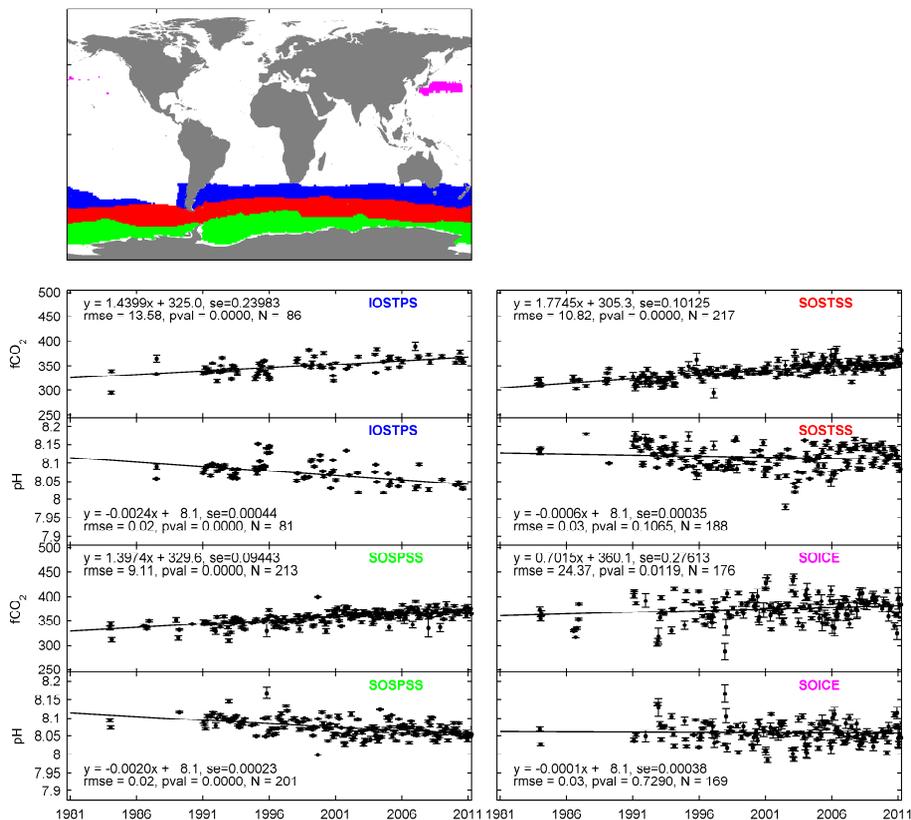
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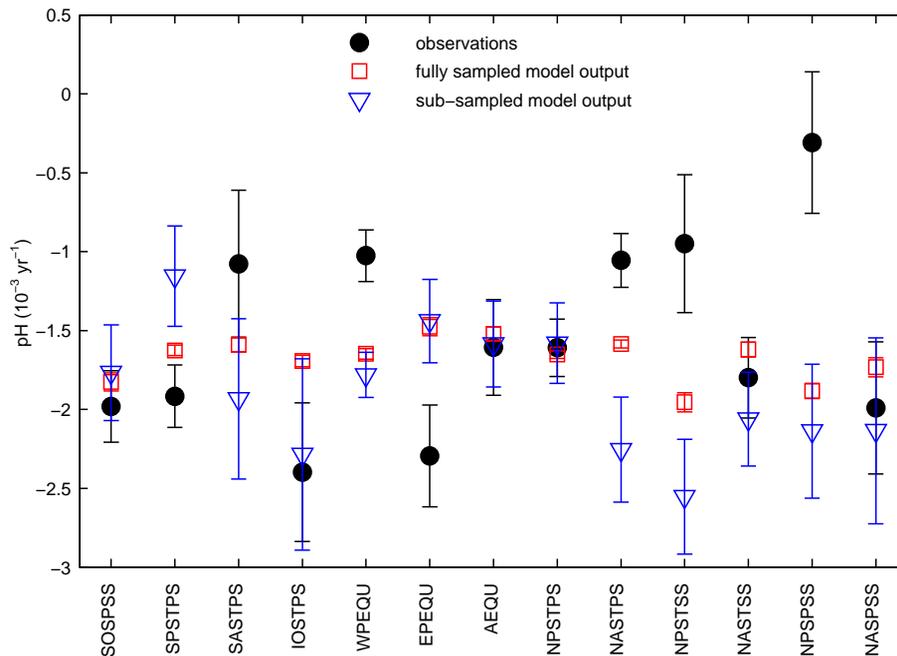
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**Figure 4.** Long term pH trend (1981–2011) in the Indian Ocean biome and the three Southern Ocean biomes.

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**Figure 5.** Summary of the pH trends in all biomes. The error bars show the  $1\sigma$  confidence interval.

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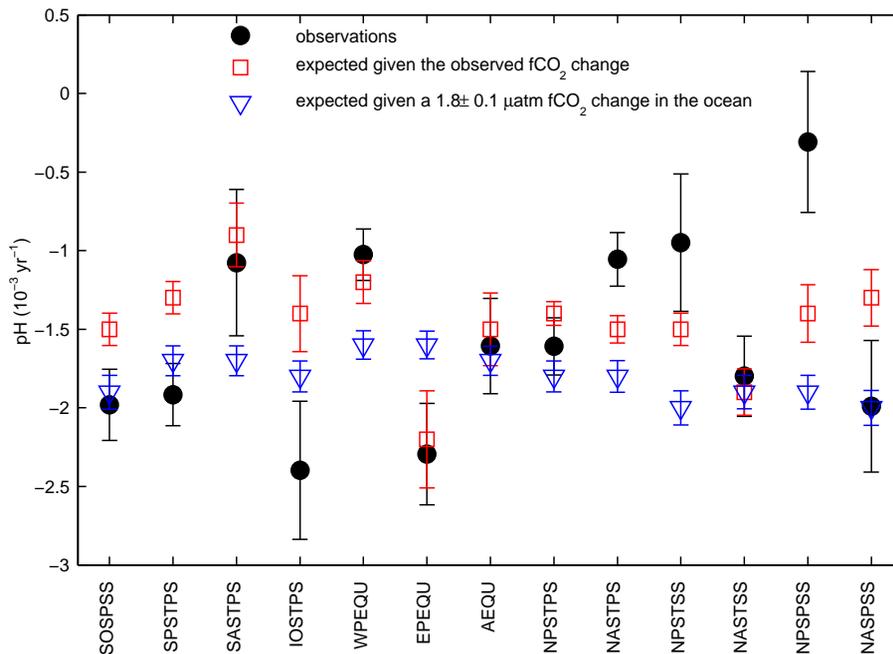
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**Figure 6.** Comparison between the observed pH trend in each biome (either 1981–2011 or 1991–2011) in black and the pH trends expected if the surface ocean  $f\text{CO}_2$  changed equal to the atmosphere (blue) and expected for the observed ocean  $f\text{CO}_2$  trends (red).

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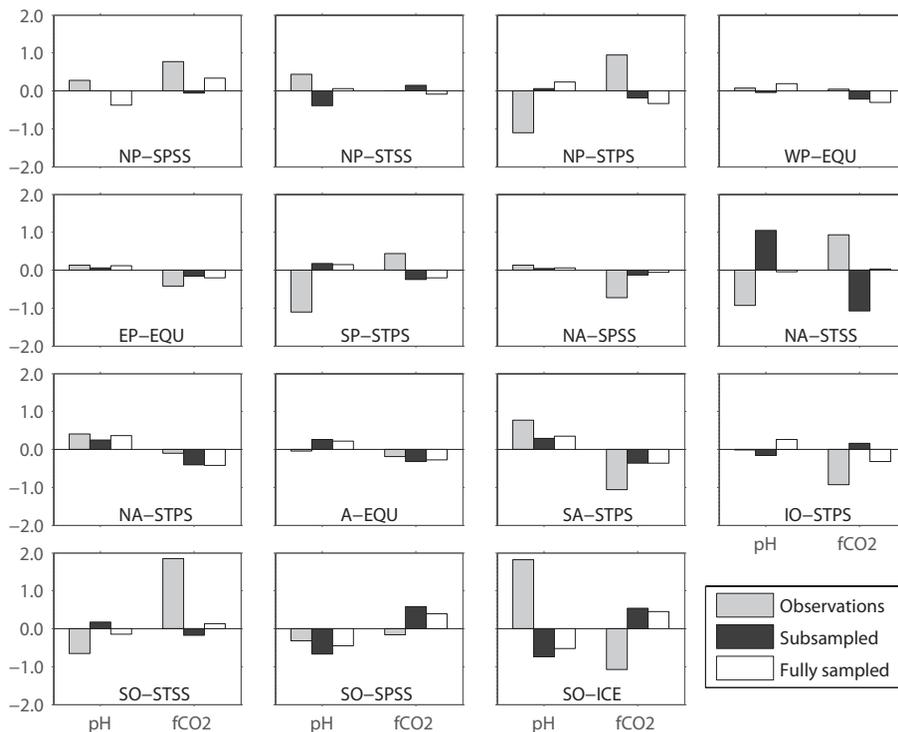
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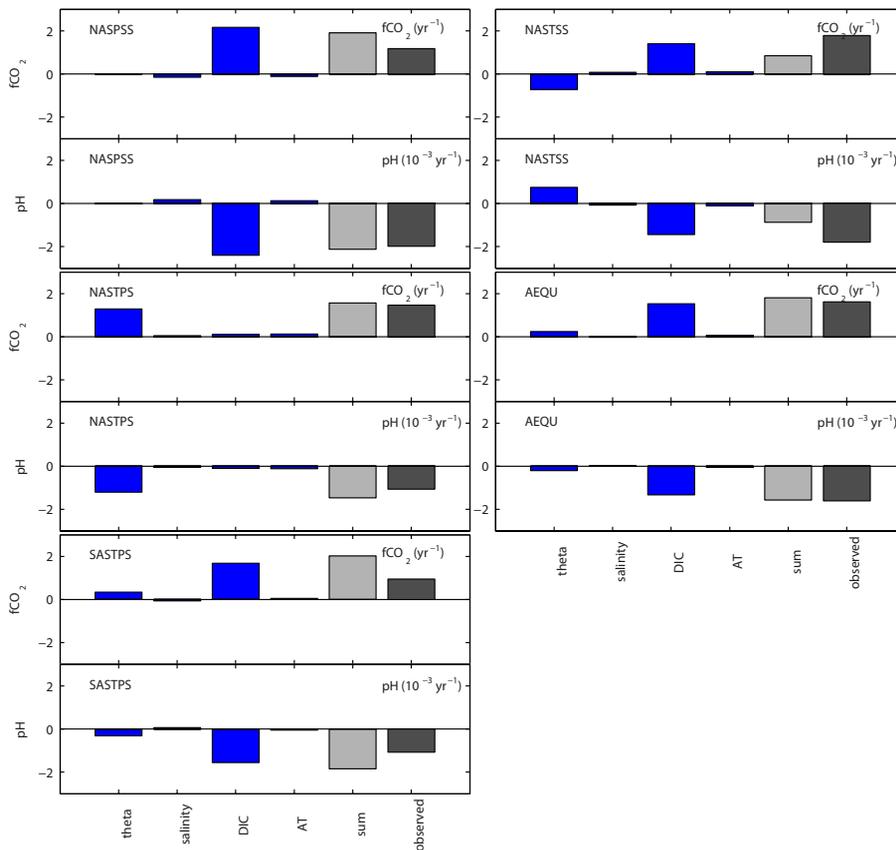


**Figure 7.** The residual between the actual trends in pH and  $f\text{CO}_2$  and the sum of the four decomposition parts (SSS, SST, DIC, ALK). In gray is the residual for the observations, in black the residual for the sub-sampled model output, and in white the residual for the fully-sampled model output.

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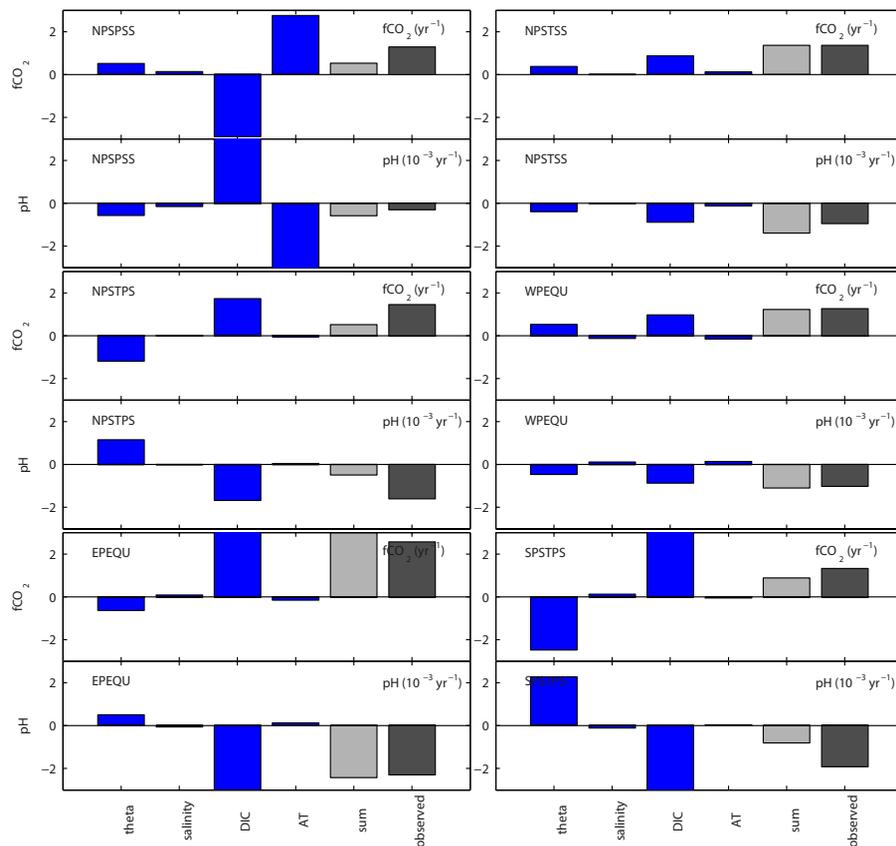
**Figure 8.** The long term trends in  $f\text{CO}_2$  and pH from Fig. 2 decomposed into the contributions from SST, SSS, alkalinity, and DIC. Also shown is the sum of the four contributions and the actual observed trend. Note that for pH the trend has been multiplied by 1000 for easier visualization.

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## Trends and drivers in global surface ocean pH over the past three decades

S. K. Lauvset et al.



**Figure 9.** The long term trends in  $f\text{CO}_2$  and pH from Fig. 3 decomposed into the contributions from SST, SSS, alkalinity, and DIC. Also shown is the sum of the four contributions and the actual observed trend. Note that for pH the trend has been multiplied by 1000 for easier visualization.

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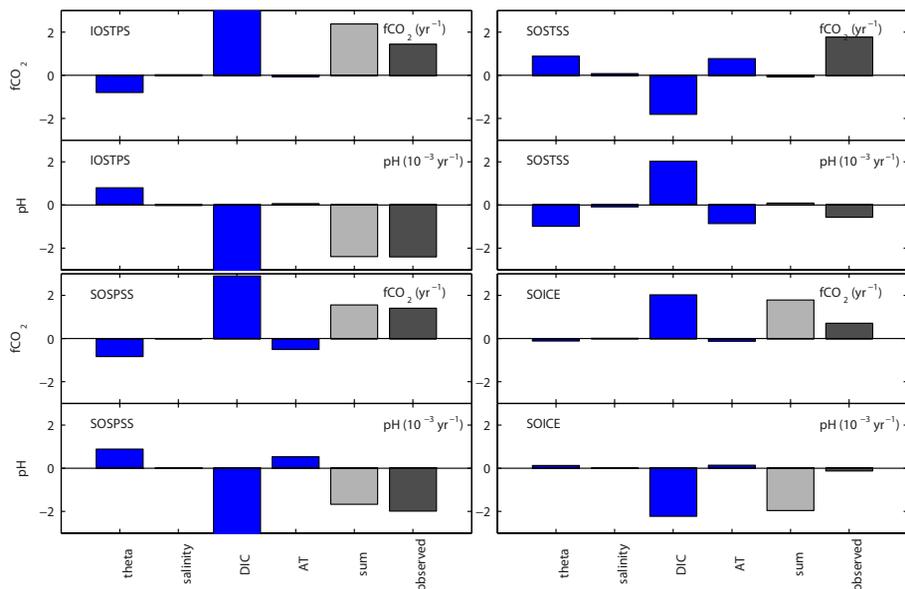
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## Trends and drivers in global surface ocean pH over the past three decades

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**Figure 10.** The long term trends in  $f\text{CO}_2$  and pH from Fig. 4 decomposed into the contributions from SST, SSS, alkalinity, and DIC. Also shown is the sum of the four contributions and the actual observed trend. Note that for pH the trend has been multiplied by 1000 for easier visualization.

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