## Answers to the comments from Referee #1 1 2 We would like to thank you for the time you spent on our manuscript and for you constructive 3 comments. 4 5 6 This manuscript discriminates between the drivers of woody growth at close to 1,000 site.years. Because some possible drivers – GPP, water stress, ... -7 are not directly measured in the experimental plots, a process-based model is used to 8 estimate these drivers instead. Key conclusions are a hierarchy of all possible drivers 9 according to their ability to explain either inter-annual woody growth variability or spatial 10 woody growth variability. This hierarchy is – in my view – the main novelty brought by 11 this manuscript, and a welcomed one. The method underlying this classification is 12 13 partly questionable: the authors claim that the statistical models they use are robust to collinearity in explanatory variables but this claim is not sufficiently justified (see 14 15 detailed comments). However, even if collinearity ends up being a problem, this would only add uncertainty to the results and would not greatly reduce the interest of the 16 17 manuscript. As a result, and although some details could be improved, I think the manuscript is readable and worth being published. 18 We addressed your concern regarding collinearity implications in your dedicated specific 19 20 comment. 21 In the following, we address each of your specific comments. 22 23

based simulations (e.g. p2214 l. 9, p2218 l3) disturbed me for a long time as I sought

The claim that field measurements are combined with process-

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26 to distinguish between what results came from the model structure and what really 27 was observed on-site. Then I remembered that simulations are only used when mea-28 *surements are not available, and that the dependent variable – woody growth – is* measured (the confusion coming from the fact that the model also simulates woody 29 growth). Unless I missed something, the study is mainly empirical, and only supple-30 mented by process-based modelling to estimate some possible drivers of woody growth 31 32 when more direct methods are not applicable. Therefore, although I do not dispute the "combination", I would recommend clarifying this by removing the term "combination" 33 throughout the text, and highlighting that modelling is only here to estimate explanatory 34 variables (or drivers) when these can't be measured. 35 36 Our study is indeed mainly empirical. We only used process-based modelling to estimate 37 potential drivers that could not be directly measured in the field, as stated in your comment. We 38 clarified this point in the abstract and in the manuscript by removing the word "combination" in the description of our approach. Rather, we briefly explained why process-based modelling was 39 40 used to complement field measurements (e.g., 126, 1133, 1420). 41 42 P2215 l4 This 60% figure does not directly come from Pan et al (2011). Their estimate 43 of 2.6 PgC yr-1 divided by AR5's estimate of 8.3 PgC yr-1 for fossil fuel emissions would land around 30%. How do you get this 60%? 44 We evocated the gross carbon uptake by the established forests and tropical regrowth forests, 45 which amount to 73 PgC over the period 1990 to 2007, "equivalent to 60% of cumulative fossil 46 emissions in the period" (Pan et al., 2011). We acknowledge that this statement was confusing as 47 48 1) we did not explicitly mentioned in the manuscript that we addressed gross carbon uptake 2) Literature usually reports the figures that you mentioned i.e. the proportion of anthropogenic C 49 50 emission captured in established forests. We therefore modified our sentence based on Pan et al.

persistent carbon sink for decades, sequestering almost 30% of the world's total anthropogenic C 52 emissions between 1990 and 2007 (Pan et al., 2011)." (151-53) 53 54 55 P2215 l26 An overview of the mechanisms potentially underlying cambial activity other than C availability would be useful here (they are provided later in the discussion). 56 We added a short description of the cell processes into play. "Source control of wood growth is a 57 58 mechanism that has been questioned by several authors, who argue that cambial activity is more 59 sensitive than C assimilation to several environmental stressors (Fatichi et al., 2014). In particular, the decrease in cell turgor that occurs because of water stress strongly affects cell 60 61 division and expansion (Woodruff and Meinzer, 2011) before there is any strong reduction in the gas exchange (Muller et al., 2011; Tardieu et al., 2011). Similarly, cell division is affected by low 62 temperatures before it is affected by photosynthesis (Körner, 2008). The onset of cambial activity 63 64 is also known to be highly responsive to temperature (Delpierre et al., 2015; Kudo et al., 2014; 65 Lempereur et al., 2015; Rossi et al., 2011) and, in turn, may partly determine annual cell 66 production and wood growth (Lupi et al., 2010; Rossi et al., 2013)." (173-82) 67 68 P2219 l16-22 The key features (e.g. list of explanatory variables) of the empirical model and the allometric function should be provided in the manuscript itself (not in the SM). 69 70 We added a description of the empirical model, with a focus on its parametrization. "The entire 71 stand tree CBH distribution was reconstructed from the CBHs of the sampled trees using an empirical tree competition model (Deleuze et al., 2004). This model stipulates that only trees with 72 73 a CBH above a given threshold ( $\sigma$ , the minimum circumference needed to gain direct access to

(Table 3) as follows "Inventory-based estimates indicate that established forests have been a

sunlight), have a significant growth. Overstory trees then have an annual basal area growth rate that is proportional to their size, according to a slope coefficient,  $\gamma$ . Following the work of Guillemot *et al.* (2014), the model was calibrated annually, beginning at year (n) of the core sampling and used iteratively to reconstruct the past stand CBH growth. The  $\sigma$  parameter was first defined using an empirical relationship with the maximum CBH of the stand tree distribution from year (n). The  $\gamma$  parameter was then adjusted using the tree rings measured on the sampled trees in year (n-1). The parameterized model was finally used to predict the basal area increments of all the trees in the distribution, and consequently the tree CBH distribution in the year (n-1). A detailed description of the iterative process can be found in Supplement S2 and in Guillemot *et al.* (2014)." (1173-185)

Because of the number of relations involved, it was difficult to provide all the species-specific allometric equations in the manuscript. Rather, we reported in the manuscript all the references that we used (1190-192). The allometric functions used in our work and their parametrizations are extensively described in appendix S3 and S4.

- 89 P2219 l22 What does "historical basal area" mean? Basal area measured before the
- 90 measurement years? Since when?
- 91 The "historical stand basal area" is the past stand basal area, that we calculate using the past
- 92 trajectory of the stand CBH distribution obtained thanks to the empirical model described above.
- 93 We clarified this point in the manuscript (1186-188).

- 95 P2220 l6-12 Although SWC and SNA can unquestionably be put in the "plot fertility"
- ocategory, putting LAI and LNC in the same category is debatable as these variables

97 depend among others on management and stand age as well as fertility.

98 We acknowledge that the relation between LNC and plot fertility is debatable. We chose to

include this variable because some studies report a significant link between LNC and soil type

(see for example Le Maire et al, 2005, Tree Physiology 25, 859-872, Figure 4). However we did

not found a significant link between C allocation to growth and any of the variable included in

the "fertility" category, possibly because they are poor proxy of the actual plot fertility. LAI was

not used to characterize plot fertility in the statistical analyses (1205-206).

P2222 11-4 How are these water stress indices defined? As water stress ends up

being high in the hierarchy of drivers of woody growth, it is important to explain it is

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108 The CASTANEA model simulated the daily soil water balance, based on a bucket soil sub-model

with 2 layers (a top soil layer and a total soil layer that includes the top soil layer, Dufrêne et al.,

110 (2005)). WS\_intgp was then used to quantify the intensity of water stress by summing the reduc

index on a daily basis (Granier et al., 1999).

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$$reduc_{t} = \max \left(0, \min \left(1, \frac{SWC_{t} - SWC_{wilt}}{0.4 \times \left(SWC_{fc} - SWC_{wilt}\right)}\right)\right)$$

where SWC<sub>t</sub> is the soil water content on day t (mm), SWC<sub>wilt</sub> is the soil water content at the

wilting point (mm) and SWC<sub>fc</sub> is the soil water content at field capacity (mm).

 $WS\_per_{gp}$  is the number of days of the current growth period during which the soil water content

was less than 60% of the soil water holding capacity (Table 2, modified from Mund et al.,

119 (2010)). Water stress indices were also calculated for the entire preceding year (lagged effect of 120 water stress,  $WS_int_{v-1}$  and  $WS_per_{v-1}$ ). 121 We added this description to the manuscript (1249-261). 122 123 P2222 l24-26 See general comment. The statistical models used are apparently: Pear-124 son's correlations, multiple linear regression, and random forest. The robustness of the 125 first two to collinearity in explanatory variables is not justified (and indeed, it would be 126 difficult to justify it) and the justification provided for random forest (p 2224 l23-25) is 127 that the collinearity problem is diminished by the random selection of variables participating in the classification. I'm clearly not an expert of this algorithm, but I fail to 128 see how random selection reduces collinearity: if two variables (eg. GPP and water 129 stress) are strongly correlated in the sample, selecting sometimes one and sometimes 130 131 the other will not diminish the fact that they both tend to provide the same clusters and 132 that therefore it's difficult to say whether it's GPP or water stress which explains why the clusters are "different". I acknowledge that I may be missing some background here, 133 but so will many readers so your claim should either be better justified or dropped. If 134 dropped, then the results should be discussed with the collinearity problem in mind. 135 We acknowledge that the presentation of the empirical models regarding their ability to deal with 136 correlation among predictors was misleading. Pearson's correlation and multiple linear 137 138 regressions are indeed not able to handle highly correlated variables. Consequently, we changed 139 the manuscript and we highlighted that the correlations among the explanatory variables that 140 were used in this study generally were lesser than 0.7, which has been reported as the level above which collinearity begins to severely affect model performance (Dormann et al., 2013). One 141 142 exception was for the important correlation among components of the tree C balance (because

NPP = GPP - Ra). As a consequence the tree C balance components were introduced one at a

time in the models, as previously specified in section 2.4.2. The effectiveness of Random forest methods in identifying "true" predictors among a large number of correlated candidate predictors have been empirically (i.e. comparing RF to traditional models) highlighted in many studies in the last decade (e.g. Archer and Kimes, 2008; Cutler et al., 2007; Genuer et al., 2010). The reason often given is that the iterative sample of predictors allows for an independent evaluation of each variable importance. As a consequence, the ranking of variable importance obtain in RF appears to be a more robust procedure than traditional model selection in the face of collinearity, and RF "does not suffer some of the shortcomings of traditional variable selection methods, such as selecting only one or two variables among a group of equally good but highly correlated predictors" (Cutler et al., 2007). However to our knowledge, the theoretical causes of the effectiveness of RF, along with the limit of its application, remain debated (see Gregorutti et al., 2015 arXiv:1310-5726). These controversies are however beyond the scope of this paper and our field of science. We clarified the section 2.4.1. with regards to the ability of models to deal with collinearity and we added the mainlines of the above discussion (1280-295). P2225 l10-15 How variable is Ra? If GPP and NPP are both correlated with AWBI and Ra is not, an obvious possibility is that Ra is broadly constant in CASTANEA. Can you rule this out? The coefficient of variation of the simulated annual NPP, GPP and Ra is  $10.8\% \pm 3$ ,  $7.4\% \pm 2$ ,  $6.8\% \pm 3$ , respectively (the values are mean CV  $\pm$  standard deviation among sites). NPP, GPP and Ra appear to have comparable annual variability, which lead us to think that it is unlikely that the low Ra-AWBI correlation would be a consequence of the low variability of Ra. Previous

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studies report that GPP has a preponderant role in the annual variability of net ecosystem

167 productivity of European forests (e.g. Delpierre et al., 2012), which could explain its strong link with annual growth. We reported the CV values in the manuscript (1361). 168 169 170 P2226 "between" sites is more appropriate than "among" I think. There is a typo in line 171 172 5: "did not revealed". We rephrased this sentence (1375). 173 174 P2228 15-10 One of your key explanations is species difference. Would species differ-175 ence reconcile these cited works? More generally, how do your results help in solving 176 the apparent paradox between these cited works? 177 Our study was based on 4 species with contrasted autecology, in order to be representative of the 178 main European biomes. Consequently these species face different environmental conditions, 179 180 described in Table 1 and in the Supplement. A limit of this approach is that it is not possible to 181 evaluate whether the differences highlighted among species (especially the ranking in variable 182 importance, Fig. 4) result from the environment of from species-specific (genetic) growth 183 response. Common-garden experiment could be very relevant in this respect, but will require 184 studying seedlings and not mature trees. More generally, the paradox that we report in C flux – 185 growth agreements is not solved in this study: our aim was to highlight C allocation drivers across 49 sites, but the relative impacts of these drivers has not been evaluated at the site level. 186 187 We only argue that in the modelling framework that we present in section 4.4., the simulated growth is subject to the combined controls of C supply and changes in C allocation due to 188 endogenous adjustments and/or modulations of sink activity. These controls result from distinct 189

processes, which are independently represented in the modelling framework. The relative

influences of the various processes, i.e., the simulated growth causalities, are thus likely to vary both spatially and temporally, depending on the environmental conditions faced by trees. Our approach has therefore the potential to shed light on the contrasted results reported by correlative studies (1569-575). The processes reported in our work should now be implemented in an integrated allocation scheme as part of a process-based model. Using such a model at contrasted sites to quantify the relative importance of each process in the resulted simulated annual growth should allow bringing more light on this paradox, as already initiated by recent studies (http://www.biogeosciences-discuss.net/12/2745/2015/bgd-12-2745-2015.pdf and Schiestl-Aalto, Kulmala, Mäkinen, Nikinmaa, & Mäkelä, 2015, New Phytologist). Table 4 Which data has been centered and scaled? Why? The variables introduced into the linear models were centred and scaled such that their normalized coefficient estimates indicated the relative influence of the predictors on the AWBI (1311).

## **Answers to comments from Referee #2**

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We would like to sincerely thank you for the time you spent on our manuscript and for your insightful comments and suggestions. Because you pointed some misspelling in your comments, the entire manuscript was edited for correct English language and style. Very well written and presented study, though the interpretation of the results at times could be viewed as overreaching. The main issue is that the C source is modeled, and therefore the relationship (or lack thereof) between C source and biomass increment are highly dependent on the accuracy of the modeled C source. Unfortunately there is no data available to estimate the accuracy of the modeled C source at the studied sites (with the exception of Puechabon, though no validation is presented here). The authors claim that the model has been widely validated at European sites. Of course practically all models are, but the quality of their extrapolation to other sites remains subject to the idiosyncrasies of those sites. The issue is that authors repeatedly highlight the lack of relationship between source dynamics and biomass increments as evidence for a sink limitation, but do not acknowledge that there could be errors in the modeled C source that are responsible for the lack of relationship. Indeed the modeled C source is regularly presented in a way that could lead the reader to believe the authors are presenting observed C source. The C source should be referred to as the modeled C source at all times, and the impact of

the potential disjoint between modeled and actual C source should be discussed in an open and non-defensive way. Regardless of the above, the findings and approach used in this study are a novel and valuable contribution to a growing area of interest. I have no doubt it will be of interest to the readers of Biogeosciences. We acknowledge that process-based modelling is a source of uncertainty that was not properly discussed in the first version of our paper. As a consequence, we clarified throughout the text whether the considered explanatory variable was simulated or measured (e.g., 1354, 1360, 1366, 1392). We additionally discuss in section 4.4. the potential impact of model uncertainty on our results and the fact that the quality of our simulations remain subject to the idiosyncrasies of the sites used in this work. Among other considerations, we acknowledge that "A third factor that hampered the ability of our empirical models to explain the annual growth variability is the potential disagreement between the CASTANEA outputs that were used as explanatory variables and the corresponding actual drivers. [...] In particular, a number of past disturbances such as insect outbreaks, windthrow or unreported commercial thinning could have temporarily induced large discrepancies between the actual and simulated C fluxes (Grote et al., 2011; Hicke et al., 2012). The error that is attributable to model performance unfortunately remains unknown because of the absence of EC measurements at our study sites (except for the Puéchabon site, see Delpierre et al., 2012)." (1583-594). However, as stated in the paper, the uncertainty of the simulated C fluxes was considered in both spatial and temporal analyses using a bootstrap procedure (Chernick, 2011): all linear models were fitted 1000 times, randomly sampling at each iteration the C fluxes values within the root

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mean square error of the CASTANEA simulations (Appendix S9) to obtain for each variable a

parameter estimate distribution. We finally retained explanatory variables with parameter estimate distributions excluding zero value in a bilateral 5% probability level. Consequently, the results discussed in the paper proved to be significant even when considering the reported uncertainty of CASTANEA flux simulations (1321-326, 1.349-350, 1409). In the following, we address each of your specific comments. Detailed comments: Page 2215 Line 4: 60% of the fossil fuel emissions. Please check this. The airborne fraction ranges between 30-40%, meaning oceans and terrestrial ecosystems together take up about 60%. Estimates suggest the ocean accounts for 30%, which would leave 30% for global ecosystems (which includes all ecosystem types, not just forests). We evocated the gross carbon uptake by the established forests and tropical regrowth forests, which amount to 73 PgC over the period 1990 to 2007, "equivalent to 60% of cumulative fossil emissions in the period" (Pan et al., 2011). We acknowledge that this statement was confusing as 1) we did not explicitly mentioned in the manuscript that we addressed gross carbon uptake 2) Literature usually reports the figures that you mentioned i.e. the proportion of anthropogenic C emission captured in established forests. We therefore modified our sentence based on Pan et al. (Table 3) as follows "Inventory-based estimates indicate that established forests have been a persistent carbon sink for decades, sequestering almost 30% of the world's total anthropogenic C emissions between 1990 and 2007 (Pan et al., 2011)." (151-53) We chose the term "established forests" rather than "global ecosystems" because it makes more sense in the context of our work,

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278 as Pan et al. reported that "within the limits of reported uncertainty, the entire terrestrial C sink is 279 accounted for by C uptake of global established forests". 280 Page 2216: 281 282 Line 7: Constant – constrained. 283 We followed your suggestion and corrected the manuscript (184). 284 Line 7: This sentence structure is confusing. Perhaps rephrase as: "The above exper-285 imental evidence suggests that growth is mostly limited by the direct effects..." 286 287 We followed your suggestion and corrected the manuscript (185-86). 288 Line 15: The statement that rising CO2 could increase the terrestrial sink only if growth 289 290 is a source limited process is misleading. There are multiple pathways through which 291 CO2 can increase the sink. For example, increased root exudates under elevated CO2, and increased allocation to roots in general, is now widely accepted. This is thought to 292 alleviate nutrient limitation. Water use efficiency also increases under elevated CO2. 293 294 This can alleviate water limitation, thus enhancing sink strength. 295 We acknowledge that our statement was naive and we accordingly corrected the sentence: 296 "The extent to which wood growth is under source or sink control is of paramount importance for predicting how trees will respond to global changes and specifically how increasing atmospheric 297 298 CO<sub>2</sub> will affect forest productivity and the future terrestrial C sink." (189-91)

299 Line 19: the future of forest? Perhaps be more specific. 300 We further explained our statement in a new sentence:" The implementation of the respective 301 302 roles of source and sink controls on growth in TBMs is therefore a substantial challenge for modellers, because it may determine our ability to project future forest C sink, diebacks and 303 distributions" (193-95). 304 305 Page 2217 306 Line 13: the intensity effects of which 307 308 We corrected the manuscript and clarified this sentence. "The woody biomass increment 309 therefore appears to be under the control of multiple factors. The effects of these drivers are expected to strongly vary in space and time." (1114-116). 310 311 Line 26: in forests 312 We followed your suggestion and corrected the manuscript (1127). 313 314 Page 2218 315 Line 5: I would suggest increasing the font on this schematic. It will be quite small in 316 the final print version. 317 318 We increased the font of the schematic for a better readability.

320 Line 25: 'allowed to extensively assessing'. Please revise for proper language use. 321 We corrected the sentence as follows: "This hybrid approach allowed us to assess and disentangle 322 the effects of previously reported environmental and endogenous drivers of C allocation to wood growth (Fig. 1)." (1154-155). 323 324 Page 2220 Line 10: used to rank 325 326 We rephrased this sentence (1204). 327 328 Page 2221 329 Line 6-7: The level of CASTANEA agreement with observed interannual variability in 330 the Delpierre et al., 2012 study is indeed impressive. Are we to believe, however, that 331 the model does similarly well for site for which there is no calibration data such as the 332 ones included in this study? It is quite likely that when it was first applied to the sites in the Delpierre et al. study it did not do well at all, until some site specific characteristics 333 334 were accounted for by adjusting parameters. This is the weak point of this study – we have no way of knowing if the model does a good job of reproducing variability in NPP 335 at the studied sites. If it does not do a good job, then it is no surprise that modeled C 336 337 source diagnostic variables were not found to be related to actual carbon allocation. 338 We acknowledge that the use of process-based modelling is a source of uncertainty that was not 339 properly discussed in the first version of our work and that remain unknown in the absence of EC 340 measurements at our study sites. Dedicated discussion of this point can be found in our above

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general answer.

343	Line 23: 2 day resolution
344	We followed your suggestion and corrected the manuscript (1241).
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346	Page 2222
347	Line 12: State which carbon fluxes.
348	The considered C fluxes are GPP, NPP and Ra. We clarified this point in the sentence (1270).
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350	Line 15: The age related trend.
351	We rephrased this sentence (1274).
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353	Line 22: I'm not sure collinearity is the right word here, as the relationships are not
354	necessarily linear. Perhaps covariance?
355	We agree that covariance is better here, as a more general statement (1281). Thank you for this
356	suggestion.
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358	Line 23: could hamper
359	We followed your suggestion and corrected the manuscript (1282).
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361	Page 2223
362	Line 5: gathered – grouped.
363	We followed your suggestion and corrected the manuscript (1296).

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366	Page 2224
367	Line 14: important – large.
368	We followed your suggestion and corrected the manuscript (1333).
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370	Page 2225
371	Line 7: The elementary components of the simulated seasonal forest C balance. It
372	should be made clear in all instances when referred to Castanea output that this is
373	indeed model output.
374 375	We clarified in all instances whether the considered variables is simulated or not (e.g., 1354, 1360 1366, 1392).
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377	You are not simulating C balance, as the forest C balance also
378	includes heterotrophic respiration. Please revise throughout the manuscript.
379	We acknowledge that this statement was confusing because we only simulated tree-atmosphere C
380	fluxes. We therefore changed all the occurrences of "forest C balance" referring to our study to
381	"tree C balance" (e.g., 1237, 1281, 1286, 1534).
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383	Page 2226 Line 7: "no retained models included" Please state whether this is for all
384	species or just the temperate oaks.
385	This statement is for all species. We clarified this point in the manuscript (1377).

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       Page 227
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       Fig 5: These partial dependencies are very interesting. It would help the reader if the
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       footer contained information on how they were derived.
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       We added to the Figure caption a short technical description of partial plot in random forest: "
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       The marginal effect of a given variable X was obtained by fixing the value of X and averaging the
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       RF predictions over all the combinations of observed values for the other predictors in the dataset
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       (Cutler et al., 2007). The marginal predictions were collected over the entire range of X in the
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       training data using a regular grid."
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       Page 2227
       Line 25: "Our results have far reaching...." This is unnecessary.
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       We removed this sentence (1423).
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       Page 2228
       Line 23: fluxes – flux
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       We followed your suggestion and corrected the manuscript (1445).
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       Page 2232
       Line 21: It is disingenuous to cite a paper over a decade old in support of the claim
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       that current models do not simulate the IAV of growth well. There are multiple studies
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407 that show quite accurate simulation of tree ring IAV using GPP driven modeling (e.g., http://www.biogeosciences.net/11/6711/2014/bg-11-6711-2014.html). 408 409 We acknowledge that recent studies report very satisfactory simulations of annual growth based on C source modelling. We rather would like to emphasize that there is a risk of "getting the right 410 answers for the wrong reason" (Fatichi et al. 2014) because of the high correlation usually found 411 between GPP and cambial activity. Because the simulated fundamental processes are different, 412 even if they obtain similar performances against observations, this could be of great importance 413 414 for productivity projections under climate change. We therefore modified our sentence: "This C-415 centric perspective overlooks the possibility of sink control of growth and thus ignores results such as those presented in this study and those of earlier local studies (reviewed by Fatichi et al. 416 417 2014). Consequently, this perspective possibly hampers the ability of TBMs to project future forest productivity (Fatichi et al. 2014)." (1547-550). 418 419 Line 20-25. You need to add another explanation here – the possibility that modeled C 420 421 source is not accurate enough. Even if your model is globally applicable with absolute 422 confidence as it is presented here, it will still be unable to account for forest distur-423 bances such as insect outbreaks, and various extreme events. This is a very important 424 issue and must be discussed. 425 We added a further discussion of the possible implications of process-based modelling on our 426 results, based on your suggestions. A dedicated discussion of this point can be found in our above

general answer.

428	The dynamic of the annual carbon allocation to wood in
429	European forests is consistent with a combined source-sink
430	limitation of growth: implications for modelling
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451 452 Correspondence to: J. Guillemot (joannes.guillemot@gmail.com) 453 454 **Abstract** 455 The extent to which forest-wood growth is limited by carbon (C) supply (i.e., source control) or 456 457 by cambial activity (i.e., sink control) will strongly determines the responses of trees to global changes. HoweverNevertheless, the physiological processes that are responsible for the 458 459 limitingation of forest growth are still under debated. The aim of this study is was to evaluate the key drivers determinants of the annual carbon C allocation to wood along large soil and climate 460 461 regional gradients in for five tree species representative of the main European forest biomes 462 (Fagus sylvatica, Quercus petraea, Quercus ilex, Quercus robur and Picea abies). 463 Combining field measurements and process based simulations at 49 sites (931 site years), we assessed the stand biomass growth dependences at both inter site and inter annual scales. The 464 drivers of stand biomass growth were assessed on both inter-site and inter-annual scales. Our 465 dataset included field measurements performed at 49 sites (931 site-years) and it was 466 complemented with process-based simulations when possible explanatory variables could not be 467 directly measured. Specifically, the relative influences of forest-tree C balance (source control), 468 469 direct environmental control (water and temperature controls of sink activity) and allocation adjustments related to age, past climate conditions, competition intensity and soil nutrient 470 471 availability on growth were quantified. 472 The inter-site variability in the stand C allocation to wood was predominantly driven by an-agerelated decline. The direct control effects of temperature or and water stress on sink activity (i.e., 473 effects independentindependently from their effects on the C supply) exerted a strong influence 474 475 on the annual stand woody growth in all of the species considered, including deciduous temperate species. The lagged effect of the past environmental conditions (e.g., the previous year's water 476 stress and low C uptake) was a significantly affected driver of the annual C allocation to wood. 477

The CCarbon supply appeared to strongly limit growth only in deciduous temperate deciduous

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species.

We provide an evaluation of the spatio-temporal dynamics of the annual earbon-C allocation to wood in European forests. Our study supports the premise that European forest growth is under subject too complex control\_processes that includeing both source and sink limitations. The relative influences of the different growth drivers strongly vary with time and across\_years and spatial ecological gradients. We suggest a straightforward modelling framework with which to implement these combined forest growth limitations into terrestrial biosphere models.

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## 1 Introduction

Forests play a critical role in the global carbon (C) cycle. Inventory-based estimates show indicate that established forests have been a persistent carbon sink forin recent decades, sequesteringuptaking almost 360% of the eumulative world's total anthropogenic C\_fossil emissions between 1990 and 2007 (Pan et al., 2011). The fate of the entering sequestered C strongly relies is highly dependent on the C dynamics in trees, which determines its-the residence time of C in the forest ecosystems. Despite its importance for the future of the terrestrial C sink ecosystem respiration remains poorly understood (Brüggemann et al., 2011). In particular, there has been considerable amount of debates have arisen from regarding the physiological mechanisms that driveing the increment of the forest woody biomass (Palacio et al., 2014; Wiley and Helliker, 2012). The fraction of the assimilated C stored in woody biomass can be inferred by combining biometric measurements to-with estimates of the C exchanges between the ecosystem and the atmosphere, based on the Eddyeddy Covariance (EC) technique (Babst et al., 2014; Litton et al., 2007; Wolf et al., 2011). Global meta-analyses of such dataset (thatgathering included data from various different biomes and different species) have revealed a strong correlation between the observed gross primary production (GPP) and the woody biomass increment (Litton et al., 2007; Zha et al., 2013). Accordingly, growth has for long been thought as to be C limited, because of the through an hypothesized causal link between C supply and growth (i.e., source control, Sala et al. 2012). The environmental factors that have been reported to affect growth (soil water content, temperature, nutrient content, light and CO<sub>2</sub>) were therefore supposed to operate through their effects on photosynthesis and respiration fluxes. This C-centric paradigm underlies most of the C allocation rules formalized in the terrestrial biosphere models (TBMs) that are currently used to evaluate the effectsoutcome of global changes on forests (Clark et al., 2011; Dufrêne et al., 2005; De Kauwe et al., 2014; Krinner et al., 2005; Sitch et al., 2003).

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SThe source control of tree-wood growth is a mechanism that has been questioned by several authors, who argueing that cambial activity is more sensitive than C assimilation to a panel of several environmental stressorses (Fatichi et al., 2014). In particular, the decrease in cell turgor that occurs because of water stress strongly affects cell division and expansion (Woodruff and Meinzer, 2011) before there is any strong reduction in the gas exchange (Muller et al., 2011; Tardieu et al., 2011). Similarly, cell division is affected by low temperatures before it is affected by photosynthesis (Körner, 2008). ; including water deficit (Muller et al., 2011; Tardieu et al., 2011) and low temperature (Körner, 2008). The onset of cambial activity is also known to be highly responsive to temperature (Delpierre et al., 2015; Kudo et al., 2014; Lempereur et al., 2015; Rossi et al., 2011) and, which in turn, may partly determine annual cell production and forest wood growth (Lupi et al., 2010; Rossi et al., 2013). Finally, the quality and quantity of available soil nutrients, particularly nitrogen (N), could affect growth independently of their impacts on C assimilation, because of due to the relatively constrained stoichiometry of the tree biomass (Leuzinger and Hättenschwiler, 2013). These experimental evidences These studies gave ways to an alternative understanding of tree C dynamic where suggest that growth is mostly—limited by the direct effects of environmental factors (i.e., sink control). However, numerous key environmental factors (e.g., nutrients, temperature and water) affect both sink and source activities, and it is thus difficult to unseramble determine whether wood growth is more related to carbon C supply or to the intrinsic environmental sensitivity of cambium functioning to the environment (Fatichi et al., 2014). The extent to which forest wood growth is under\_a-sourceor-a-sink- control is of paramount importance to-for predicting how trees will respond to global changes and s Specifically, risinghow increasing atmospheric CO<sub>2</sub> will affecteould possibly increase forest productivity, and hence the and the future terrestrial C sink, only if growth is a source limited process. The implementation of the respective roles of source- and sink- controls on growth in TBMs is therefore a huge-substantial challenge for modellers, because it will-may determine our ability to project the future of forest C sink, diebacks and distributions (Cheaib et al., 2012; Fatichi et al., 2014; Leuzinger et al., 2013).

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The allocation of assimilated C within forest ecosystems is a complex, integrative process that can be described on the basis of by several non-exclusive principles (Franklin et al., 2012), among which are including i) allometric scaling, ii) functional balance and iii) evolution-based optimal responses. i) Allometric The allometric scaling principle is based on the assumption that biophysical laws determine C partitioning among within the different forest compartments. Current knowledge Important allocation constraints include include vascular network optimality (Enquist, 2002) and functional homeostasis in water transport (Magnani et al., 2000) as main allocation constraints. ii) The fFunctional balance theory principle suggests a preferential allocation to that the organ responsible for acquiring sition of the most limiting resource is preferentially allocated C. In lineConsistent with this principle, a higher C allocation to fine roots at the expense of <u>C allocation to</u> woody growth has been reported <u>onfor</u> poor or dry soils (Chen et al., 2013; Keyes and Grier, 1981)., In addition, along with a possibly lausible higher greater allocation to root symbionts and exudates at the expense of biomass production has also been reported (Vicca et al., 2012). iii) Finally, the optimal response principle postulates that allocation maximizes a-fitness proxy in a fixed environment. This hypothesis of fitness maximization is in lineagrees with the previously assumed idea that a dynamic reserve pool acting as temporary storage, possibly at the expense of growth-expense, to allow-promote long-term tree survival-of trees (Chapin et al., 1990; Sala et al., 2012). Indeed, time lags between C uptake and growth have been reported (Gough et al., 2009; Richardson et al., 2013). The optimal response theory principle is consistent with severalalso explains welllong known life history traits, such as the preferential allocation to reproduction in ageing plants, that which could lead to the observed agerelated declines in woody biomass allocation (Genet et al., 2010; Thomas, 2011). Woody The woody biomass increment therefore appears to be under the control of a panel of driversmultiple factors. T-he which intensity effects of these drivers are expected to strongly vary in space and time. As a consequenceConsequently, local studies have reported contrasted agreementsconflicting relationships between the C supply and woody growth (Gielen et al., 2013; Richardson et al., 2013), ranging from no significant relationships (Mund et al., 2010; Rocha et al., 2006) to close relationshipstight links on seasonal (Babst et al., 2014; Granier et al., 2008; Zweifel et al., 2010) or annual (Ohtsuka et al., 2009; Peichl et al., 2010; Zweifel et al., 2010) time scales. Determining the key processes that affecting woody growth onat the different spatio-temporal scales is a necessary step to explain unify these apparently contradictory results inusing a common framework. Moreover, investigations should be conducted at the species level, as because phylogeny may strongly constrain forest functioning (Carnicer et al., 2013; Drobyshev et al., 2013) and induce different contrasted growth determinisms determinants between among taxa (Genet et al., 2010).

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There is a gap between the knowledge obtained from lessons drawn from global studies ofregarding universal C allocation rules in forests and our understanding of the cell processes underpinning that underlie cambial activity; currently, this gap appears to be the eurrent mainprimary obstacle toward a betterto a more complete understanding of forest-wood growth drivers. In this regard, species-specific studies that evaluateing the dynamic of C partitioning to annual woody growth along soil and climate regional gradients constitute a key missing linkwould be highly useful but are lacking. Unfortunately, advances are so far precluded by thethere is a scarcity of datasets that combineing EC and growth measurements from the same sites (Luyssaert et al., 2007). Here, we released this constraint by combining complementing stand and soil measurements at from a permanent plot network of 49 forest sites and with process-based simulations of forest annual and seasonal tree C balance (Fig. 1) over 49 forest sites. Simulations were performed using a process-based model (CASTANEA, Dufrêne et al. 2005) that was thoroughly validated using against EC data from throughout over—Europe ((Dufrêne et al., 2005)(Davi et al., 2005; Delpierre et al., 2009, 2012) and was applied here with using site-specific parameterizations parameters. By rRelating biometric measurements to different-variables linked tothat explain the C source and sink activity, we evaluated the key drivers of the annual C allocation to stand woody growth in five species that are representative of the main European forest biomes: Fagus sylvatica, Quercus petraea and Quercus robur for temperate deciduous broadleaf forests<sub>is</sub>, Picea abies, found infor high-latitude and high-altitude evergreen needleleaf forests; and Quercus ilex, an Mediterranean evergreen broadleaf species from Mediterranean forests. Specifically, the relative influence of forest-annual and seasonal (from one month to the year) tree C balance (source control), direct environmental control (water and temperature controls effects onf sink activity) and allocation adjustments related to age, past climate conditions, competition intensity and soil nutrient availability on tree growth were considered (Fig. 1). We aimed to (1) quantify the relative contributions of source- and sink-

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controls <u>in\_to</u> the spatio-temporal dynamic of forest woody growth across a wide range of environmental contexts and (2) provide <u>transferable</u>-information <u>to helpthat can be used to refine the representation of forest growth causalities in TBMs.</u>

#### 2 Materials and methods

We based our analyses on three complementary data sources: field measurements, climatic variables from atmospheric reanalysis (Vidal et al., 2010) and process-based simulated simulation data. This hybrid approach allowed to us to extensively assessing and disentangling disentangle the effects of previously reported environmental and endogenous drivers of C allocation to woody growth (Fig. 1).

#### 2.1 Study sites and field data

Our investigation We gathered field measurements from 48 plots from the French Ppermanent Pplot Nnetwork for the Mmonitoring of Fforest Eecosystems (RENECOFOR, Ulrich, 1997) and from the heavily instrumented Puéchabon tower flux site (Martin-StPaul et al. 2013). The Location and general climatic features of these plots are showngiven in Fig. 2 and Table 1., a eComplete site description of the sites is available in AppendixSupplement -S1.

## 2.1.1 Growth measurements and historical stand growth reconstruction

Growth measurements eonsisted of were obtained by two methods: *i)* Delendrochronological sampling, in which: 12 to 30 overstorey trees per plots were cored to the pith at breast height with an incremental borer. Cores were collected—(in 1994 in-at the RENECOFOR sites and in 2008 at the Puéchabon site; (Lebourgeois 1997; J.M. Ourcival, *unpublished data*). The—tree circumferences at breast height (CBHs) and total heights were also measured. The Aaverage stand age was inferred from the tree ring series. *ii)* Forest inventories, in which: extensive CBH

surveys were conducted on every plot on in a 0.5 ha area of every plot (Cluzeau *et al.* 1998; Gaucherel, Guiot & Misson 2008; J.M. Ourcival, unpublished data).

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Tree ring series were combined to—with the CBH surveys to reconstruct the historical CBHs of every tree on the plots (over 8 to 43 years, AppendixSupplement\_S1). The entire stand tree CBH distribution was reconstructed from the The scaling from the sampled trees to the entire stand tree CBHs of the sampled trees using—distribution involved—an empirical tree competition model (Deleuze et al., 2004). This model stipulates that only trees with a CBH above a given threshold ( $\sigma$ , the minimum circumference needed to gain direct access to sunlight), have a significant growth. Overstory trees then have an annual basal area growth rate that is proportional to their size, according to a slope coefficient,  $\gamma$  Following the work of Guillemot *et al.* (2014), the model was calibrated annually, beginning at year (n) of the core sampling and used iteratively to reconstruct the past stand CBH growth. The  $\sigma$  parameter was first defined using an empirical relationship with the maximum CBH of the stand tree distribution from year (n). The  $\gamma$  parameter was then adjusted using the tree rings measured on the sampled trees in year (n-1). The parameterized model was finally used to predict the basal area increments of all the trees in the distribution, and consequently the tree CBH distribution in the year (n-1). A detailed description of the iterative process can be found in Supplement S2 and in Guillemot *et al.* (2014).

The inferred past trajectory of the stand CBH distribution was used to calculate the historical number of stems (*numstem*, Table 2) and stand basal area, which we considered to be a proxy for within-stand competition intensity (*SBA*, Table 2, Kunstler *et al.* 2011). ealibrated annually (Guillemot *et al.* 2014, Appendix S2). The historical total woody stand biomass was then also calculated (Supplement S3) using species-specific tree level allometric functions (Bontemps et al., 2009, 2012; Dhôte and Hercé, 1994; Seynave et al., 2005; Vallet et al., 2006) (Appendix S3) and wood density models (Bouriaud et al., 2004; Wilhelmsson et al., 2002; Zhang et al., 1993). For *Q. ilex*, we used the appropriate function from (Rambal et al., 2004) to calculate the stand woody biomass from CBHs. Ppast annual woody biomass increments (AWBIs) were then inferred (AppendixSupplement -S4).

Historical stand basal area was additionally considered as a proxy for within stand competition intensity (SBA, Table 2, Kunstler et al. 2011), along with the number of stems (numstem, Table 2).

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## 2.1.2 Measurements of stand characteristics

<u>The sS</u>tand measurements included <u>the soil</u> water holding capacity (SW<u>H</u>C), leaf area index (LAI), leaf <u>Nnitrogen</u> content (LNC) and soil nutrient availability (SNA). <u>The SWH</u>C was estimated <u>via the from</u> soil depth and texture measured <u>aton</u> two soil pits per plots (Brêthes and Ulrich, 1997). <u>The LAI</u> was estimated from litter collection (Pasquet, 2002), and <u>the sunlit LNC</u> was determined annually <u>foron</u> 8 trees <u>betweenfrom</u> 1993 <u>andto</u> 1997 (Croisé et al., 1999).

SNA was assessed <u>as through</u> the <u>soil's C:Nearbon (C):nitrogen (N) soil</u> biomass ratio, the absolute value of the <u>soil</u>-cation-exchange capacity and <u>the-its</u> per\_cent base saturation (Ponette, 1997). These soil indices were measured at 3 depths (0 to -10, 10 to -20, 20 to -40 cm) and <u>were</u> used to <u>categorize theranked</u> soil plots into three nutrient classes, from low to high nutrient availability (<u>AppendixSupplement</u> -S5). <u>The SNA, SWHC</u> and LNC were used to characterize plot fertility in the statistical analyses (Table 2).

670 2.2 Climate data

The following Mmeteorological variables at the hourly temporal scale (with 8-km spatial resolution) were obtained from the SAFRAN atmospheric reanalysis (Vidal et al., 2010): global radiation, rainfall, wind speed, air humidity and air temperature. Temperature, which was associated related to the averaged altitudes of the SAFRAN cells, was corrected using plot-specific elevation measurements (considering assuming a lapse rate of 0.6\_°K per 100\_m, AppendixSupplement -S1). These variables were used for the climate forcing of in the CASTANEA model (Dufrêne et al. 2005, see the followingnext section). In addition, two annual temperature indices were derived used as proxies offer winter frost damage and low temperature stress during the growingth period (frost and templing, respectively, Table 2).

## 2.3 Process-based simulated simulation data

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We used the CASTANEA model for to simulateing an ensemble of diagnostic variables that are related linked to the C source and the sink activity of the forest stands. The eco-physiological process-based <u>CASTANEA</u> model <u>CASTANEA</u> aims at to simulateing earbon <u>C</u> and water fluxes and stocks of an even-aged monospecific, same-aged forest stand on at the a rotation time scale. The hourly stand-atmosphere C fluxes simulated predicted by the CASTANEA model on an hourly basis have been thoroughly validated using against EC data from throughout over Europe (Davi et al., 2005; Delpierre et al., 2009, 2012). Importantly, the biophysical hypotheses that were formalized in thise model are able to reproduce the interplay of the complex mechanisms interplay leading that lead to the inter-annual variability in the stand C balance variability (Delpierre et al., 2012); modelling this interplay, which has been recognized as a ssubstantial trong challenge for TBMs (Keenan et al., 2012). A complete description of CASTANEA is given provided in Dufrêne et al. (2005), with and subsequent modifications from are described in Davi et al. (2009) and Delpierre et al. (2012). For the purpose of the present study, CASTANEA was parameterized with site-specific SWHC and LNC values. Measured The measured LAI and total woody biomass were used to initialize the model simulations. The model's ability to reproduce the annual variability in LAI and the forest growth at regional scale has been recently validated (Guillemot et al., 2014). Nevertheless, the annual sstanding woody biomass was nonetheless forced on each yearto conform to the observed values, because as the model was here used for diagnostic purposes in this study.

An ensemble Several groups of variables was were simulated and aggregated on an annual basis (Table 2):

1. The elementary components of the forest-tree C balance. These components included the gross primary productivity (GGPP), and autotrophic respiration (Ra), along with and the net balance (i.e., net primary productivity, NPP = GPP - Ra). For a given year y, we aggregated the hourly simulated C fluxes over different seasonal time periods, with starting days that ranged from 30 to 190 and ending days that ranged from 190 to 350, at a 2-day resolution. For a given year y, we aggregated the hourly simulated C fluxes over different seasonal time periods including starts ranging from day 30 to 190, and ends ranging from day 190 to 350, with a 2-days resolution. Carbon The C fluxes were also

summed i) over for the species-specific biomass growth periods reported in the literature  $(GPP_{gp}, Ra_{gp} \text{ and } NPP_{gp}, AppendixSupplement S10S6})$  and ii) forover the whole entire preceding year (y-1) as a proxy of fine forest C status induced by past climate conditions (lagged effect,  $GPP_{y-1}$ ,  $Ra_{y-1}$  and  $NPP_{y-1}$ ).

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2. Bioclimatic water stress indices. These indices included: the intensity and the duration of the water stress (WS int<sub>gp</sub> and WS per<sub>gp</sub>, respectively, Supplement S7) over the during species-specific growingth periods that have been reported in the literature (Supplement S6), were evaluated (WS\_int<sub>gp</sub> and WS\_per<sub>gp</sub>, respectively, Appendix S7). The CASTANEA model simulated the daily soil water balance, based on a bucket soil submodel with 2 layers (a top soil layer and a total soil layer that includes the top soil layer, Dufrêne et al., (2005)). WS\_int<sub>gp</sub> was then used to quantify the intensity of water stress by summing the reduc index on a daily basis (Granier et al., 1999).

$$reduc_{t} = \max \left(0, \min \left(1, \frac{SWC_{t} - SWC_{wilt}}{0.4 \times \left(SWC_{fc} - SWC_{wilt}\right)}\right)\right)$$

where  $SWC_{\underline{i}}$  is the soil water content on day t (mm),  $SWC_{\underline{wilt}}$  is the soil water content at the wilting point (mm) and  $SWC_{\underline{fc}}$  is the soil water content at field capacity (mm).

WS  $per_{gp}$  is the number of days of the current growth period during which the soil water content was less than 60% of the soil water holding capacity (Table 2, modified from (Mund et al., (2010)). Water stress indices were also calculated over-for the whole entire preceding year-(y-1) (lagged effect of water stress,  $WS_int_{y-1}$  and,  $WS_per_{y-1}$ ).

The onset of the biomass growth (camb\_onset). We used in this work-a new growth\_-onset module (N. Delpierre, N. K. Martin StPaul, A. David, unpublished results(David, (2011):
 N. Delpierre and N. K. Martin-StPaul, unpublished results) based on a temperature sum trigger (AppendixSupplement -S8).

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## 2.4 Statistical analyses

#### 2.4.1 General overview

The sStatistical analyses were conducted in three complementary steps in for each studied species. (1) Wwe calculated the correlation between of the annual woody biomass increments (AWBIs) and the C fluxes (GPP, NPP and Ra) aggregated seasonally (from 1 month to the one year) to evaluate the direct agreement relationship between between the C supply and annual biomass growth changes. (2) The dependences of the AWBIs dependences toon the C source and the sink activity were evaluated on an inter-site spatial (inter site) scale in order to highlight determine the influence of the site characteristics on biomass growth. The relationship between the age and C Age related trend in C allocation to woody biomass was also evaluated in this step. By using the age differences among sites, we covered a large our chronosequence included a large range of ages (including stands that ranged in age from approx-imately 30 to 150 years-old, Table S1). (3) Finally, the drivers of AWBI dependences were assessed finally assessed at temporal scale temporally to highlight determine the factors that were responsible for variability in the inter-annual biomass growth.

#### drivers responsible for the inter annual growth biomass variability.

Because numerous keymany environmental factors affect both forest sink and source activities, there may be a strong collinearity covariance between among the forest tree C balance and proxies of environmental stress proxies (Fatichi et al., 2014) that could hampered the inferential interpretation power of classical statistical tests (Graham, 2003). However, the explanatory variables used in this study generally had correlation coefficients of less than 0.7, the level above which collinearity begins to severely affect model performance (Dormann et al., 2013). One exception was the correlation of components of the tree C balance (because NPP = GPP – Ra). Consequently, the tree C balance components were introduced one at a time into the models. In addition, temporal growth dependencies were evaluated using the random forest (RF) learning method (Breiman, 2001). A number of studies have empirically demonstrated the effectiveness of RF at identifying the "true" predictors among a large number of correlated candidate predictors (e.g., Archer and Kimes, 2008; Cutler et al., 2007; Genuer et al., 2010). The statistical models used in the study were consequently selected for their reliability in the presence of confounded

Mis en forme: Espacement automatique entre les caractères asiatiques et latins, Espacement automatique entre les caractères asiatiques et les chiffres variables (see below). EThe explanatory variables considered in our spatial and temporal analyses are presented in Table 2 and Fig. 1. Analyses were conducted with the R software (R Development Core Team 2013), using the packages lme4\_(Bates et al., 2007), randomForest (Liaw and Wiener, 2002) and MuMIn\_(Barton and Barton, 2014)). Because *Quercus petraea* and *Quercus robur* are difficult to distinguish in the field and have a high hybridization rate (Abadie et al., 2012), these two species have been were gathered grouped in the analyses and are hereafter collectively referred to as "temperate oaks" in the following.

## 2.4.2 Correlations between growth and C fluxes

Pearson correlations between the AWBIs and simulated C fluxes aggregated in over different seasonal time periods were assessed calculated separately for each site. The highest median correlation value for eachs per species wasere retained and tested against zero using Wilcoxon signed rank tests. Critical correlations (i.e., the threshold values for a significant difference with the retained maximum correlation) were determined in order to evaluate the sensitivity of the correlation values to changes in the C flux aggregation aggregation periods.

#### 2.4.3 Spatial Drivers of spatial variations in biomass growth dependences

The <u>drivers of spatial variations in biomass</u> growth <u>dependences</u>-were evaluated using <u>a selection</u> of multiple regression <u>models using ans based on the</u> information-theoretic approach (Burnham and Anderson, 2002). <u>The AWBIs</u> and the considered explanatory variables were averaged <u>perfor each plots</u>. The variables <u>entering introduced into</u> the linear models were centered and scaled <u>sucho</u> that the <u>values of their</u> normalized coefficient estimates indicated the relative influence of the predictors on <u>the AWBI</u>. <u>The eE</u>lementary components of <u>the forest-tree</u> C balance (NPP, GPP and Ra) were introduced one at a time into the models. For each species, multiple regression <u>modelss that</u> containing <u>ed</u> all possible combinations of the explanatory variables were fitted. <u>The mM</u>odels were compared using the second\_-order Akaike <u>information</u> criterion (AICe), and all models with an Akaike weight of at least 1% of the best approximating (lowester AICe) model were considered to be plausible (Burnham and Anderson, 2002). <u>Ultimately</u>, <u>We-we</u> retained

finally the variables that appeared in at least 95% of the selected models. Models fitted on using *P. abies* data were restricted to a maximum of 3 explanatory variables because of due to the low amount of datasmall sample size (n=6, Table 1). *Q. ilex* (n=1) was not considered in the spatial analyses. The uncertainty of the simulated C fluxes was considered assessed in the analyses using a bootstrap procedure (Chernick, 2011): all linear models were fitted 1000 times, and randomly sampling at each iteration, the C fluxes values were randomly sampled within the root mean square error of the CASTANEA simulations (Appendix Supplement –S9) to obtain for each variable a parameter estimate distribution for each variable. We finally retained the explanatory variables with parameter estimate distributions that excludeding the zero value at a bilateral two-tailed 5% probability level of 5%.

## 2.4.4 Drivers of tTemporal variations in biomass growth dependences

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The A temporal analysis was conducted on the standardized AWBI series: a double-detrending process was applied for to each series; based on an initial linear regression model, followed by fitting of a cubic smoothing spline with a 50% frequency response cut-off (Mérian et al., 2011). For analysing the temporal dependences variations in of biomass growth we used an Random Forest (RF) learning method (Breiman, 2001), which was made possible because of the important large amount of datasample size (n = 931 site-years). The RF learning method is a nonparametric method that is used to rank the contribution of different explanatory variables and evaluate their marginal effects on a variable of interest without assuming an a priori dependence forms. The principle of RF method is to combined 500 binary decision trees that were built using bootstrap samples from the initial dataset. Decisions The decisions trees aimed to reduce the heterogeneity of the explained variable in the resulting branches. For each of the 500 built-trees, the data that were not involved in the tree construction wereare used for validation. The tree predictions and errors are-were then averaged to provide the final RF results. Consequently, The RF method does not overfit and does not need or require cross-validation (Cutler et al., 2007). A subset of explanatory variables is was randomly chosen at each node, thus reducing the effect of collinear variables on the output. The RF method was used to select important variables that explained related to the growth temporal variability in biomass growth (Genuer et al., 2010). Variable selection relied on permutation importance, i.e., the existence of an increaseing inof the global mean square error when a given variable <u>wasis</u> randomized in the validation subsamples. The forms of the dependences <u>was were illustrated by discussed thanks to partial</u> dependence plots (graphical depiction of the marginal effect of a given variable, Cutler *et al.* 2007). We used this information (variable selection and dependence forms) to test for the significance of the temporal AWBI dependences within the linear model. <u>The uUncertaintyies on in</u> the simulated C fluxes <u>were was considered</u> in the linear model<u>ss</u>, <u>following the procedure described in as for the spatial growth dependence</u> analysis <u>section</u>.

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#### 3 Results

# 3.1 Agreement Relationship between woody biomass growth and carbon C fluxes

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The elementary components of the simulated seasonal forest-tree C balance differed in terms of their relationships<del>revealed contrasted agreement</del> with the inter-annual variability in of the AWBI (Table 3). The simulated sSeasonal gross primary productivity (GPP) and net primary productivity (NPP) were strongly linked to AWBIs with a comparable agreement between species. However, the simulated autotrophic respiration (Ra) revealed had weak and often nonsignificant correlations relationships with the AWBIs across the 49 studied plots. The highest strongest correlations were obtained for flux aggregation periods that i) were generally consistent within a species for GPP and NPP but different for Ra and ii) strongly differed among species (Table 3). The coefficients of variation of the simulated annual NPP, GPP and Ra across the 49 studied sites were  $10.8\% \pm 3$ ,  $7.4\% \pm 2$ , and  $6.8\% \pm 3$ , respectively. Notably, the GPP and NPP in temperate deciduous species were summed from the beginning of May to the beginning of August or and September, in temperate oaks and F. sylvatica, respectively. The longest GPP and NPP aggregation periods were obtained in-for P. abies (from the beginning of February to mid-September), and the shortest period were found in for Q. ilex (from the beginning of July to mid-August). Minor (less than 20 days) changes in the flux aggregation period associated to-with the maximum simulated flux-AWBI correlation usually marginally affected marginally the Mis en forme : Anglais

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correlation values (<u>AppendixSupplement\_S6S10</u>). As a <u>consequenceConsequently</u>, aggregation periods <u>differing that were</u> less than 13 days <u>different (start or endeither in terms of their starting or ending dates</u>) from the values reported in Table 3 were generally not significantly lower than the maximum values (see the critical values presented in <u>AppendixSupplement\_S6S10</u>).

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#### 3.2 Spatial dynamics of carbon C allocation to woody biomass growth

The amonginter-sites variability in biomass growth variability was well explained by the selected multiple regression models (R<sup>2</sup>≥0.6). We highlighting highlighted that species varied in terms of their inter-site dependences contrasted growth dependences among species (Table 4). The simulated CCarbon supply during the growth period (GPP<sub>gp</sub>, Table 2) was positively related correlated with biomass growth in F. sylvatica and P. abies, whereas there was no significant relationshiptemperate oaks did not revealed a significant link between the average AWBI and photosynthesis among sites for temperate oaks (Fig. 3A). Notably, no-the final retained models did not includeed NPPgp or Ragp for any species. Stand The stand age appeared aswas an important driver of growth biomass growth driver in temperate oaks and F. sylvatica. Accordingly The, stand age explained a substantial part-portion of the AWBI: to-C supply ratio in all species, although the relationship was not significant in for P. abies (Fig. 3B). The fraction of C sequestered in woody biomass decreased with stand age (Table 4, Fig. 3B) and was halved reduced by half in temperate oaks and F. sylvatica for stands that were from between 50 to and 150 years of age (from 0.3 to 0.13 and from 0.25 to 0.1, respectively). Additionally, we reported identified a significant and positive effect of stand basal area on both AWBI (Table 4) and the AWBI: to- $GPP_{gp}$  ratio (*data not shown*) in temperate oaks.

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## 3.3 Temporal dynamics of carbon allocation to woody biomass growth

The ranking of the <u>drivers of</u> biomass growth <u>drivers using the RF based variable</u> <u>importance obtained using the RF algorithm</u> <u>highlighted strongly indicated that the contrasted</u> temporal AWBI dependences <u>varied</u> among species (Fig. 4). The growth of temperate deciduous

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species appeared to bewas under a more complex environmental control than the growth of P. abies and Q. ilex, with several variables explaining a substantial part portion of the annual <u>variability in AWBI</u> annual variability (Fig. 4A,\_B). Simulated C supply  $(GPP_{gp})$  was strongly related to the AWBI in of temperate oaks and F. sylvatica, and, to a lesser extent, in P. abies (Fig. 4A,B,C), with positive marginal effects (Fig. 5\_a,e,h). The duration of wWater stress duration overduring the study growth period (WS\_perep) was the predominant driver of the AWBI variability in of Q. ilex, and was also strongly related to the growth of in temperate deciduous species. Low temperatures over during the growing growth period (templim<sub>en</sub>) was most substantially affected the most important dependence in P. abies and also explained a part-portion of the variability in AWBI variability inof temperate oaks. The simulated wwater and temperature stress indices had negative and quasi-linear marginal effects on the AWBI (Fig 5). Finally, environmental lagged effects contributed substantially to explain the AWBI variability in all species: the water stress intensity of the previous year (WS\_int<sub>y-1</sub>) was reported to affected the growth in of F. sylvatica and Q. ilex, whereas the simulated C supply of the previous year (GPP<sub>y</sub>. 1) affected temperate oaks and P. abies stands. Lagged effects generally revealed threshold in marginal dependences, with a significant negative effect on AWBI only for under high water stress or low C supply (Fig. 5). The effects of the retained variables (Fig. 4) was were evaluated in-via multiple regression models thats useding dummy variables to test for the significance of slope changes when thresholds appeareds on partial plots (Fig. 5). The medodels explained approx-imately 20% of the variability in the AWBI variability forof temperate oaks and P. abies, and approximately, 40% of the variability for in F. sylvatica and Q. ilex (Table 5). All of the explanatory variables had significant effects, but templim was not retained in the models for temperate oaks after the bootstrap procedure that accounteding for the uncertainty of the C flux simulations. We reported observed significant slope changes in the slopes of the effect of  $GPP_{\nu,I}$ <u>in-on</u> temperate oaks and <u>in-the</u> effect of  $GPP_{gp}$  <u>in-on</u> P. abies (Table 5). The models <u>including</u> with NPP<sub>gp</sub> and NPP<sub>y-I</sub> variables revealed the same AWBI dependences that as the models described above, but with a reduced explanatory power, whereas The models including with Rapp and  $Ra_{y-1}$  <u>variables</u> were not significant (*data not shown*).

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### 4 Discussion

This study quantified the C that is annually allocated annually to the woody biomass increment in for five species that are representative of the main European forest biomes. Combining By complementing field measurements from a permanent plot network with relimited database and process-based modellingsimulations, our approach released circumvented the constraint limitation imposed byof EC data scarcity to and characterized the annual partitioning of C into woody biomass at 49 sites (931 site-years). We were thus able to highlight identify the species-specific drivers of the spatio-temporal dynamics of the C-allocation of C to woody growth along large ecological gradients. Our results have far reaching implications for the representation of biomass growth causalities in TBMs.

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# 4.1 On correlatingThe correlation between the forest tree Cearbon balance to and woody biomass growth

Relating EC-based estimation estimates of forest C balance and biometric measurements of woody biomass growth has been the concern-focus of an increasing number of local studies. These studies results are expected to can enhance our understanding of the ecosystem C dynamics but have so far provided conflicting conclusions so far. Indeed, tThe reported correlation between woody biomass growth and forest C gain indeed rangeshas been reported as both-from non-significant correlations (Mund et al., 2010; Richardson et al., 2013; Rocha et al., 2006) to-and highly significant tight links (Babst et al., 2014; Ohtsuka et al., 2009; Peichl et al., 2010; Zweifel et al., 2010). Accordingly, the agreements relationships between AWBI and C fluxes reported in this study strongly varied among sites in for each of the species studied species (Table 3). Nevertheless, the aAnnual woody biomass increment was nonetheless consistently related to  $GPP_{gp}$  and  $NPP_{gp}$ , and only marginally to  $Ra_{gp}$  in-for thea majority of sites (Table 3). (Babst et al., (2014) Babst et al. (2013) reported a similar dependence of biomass growth to onthe C fluxes at 5 sites that spanneding a wide range of latitude range acrossin Europe. The authors attributed this result to a common sensitivity of C assimilation and biomass growth to the water balance. Our results also support the view that biomass growth and forest tree C balance components are under the control of distinct but partially correlated processes (Beer et al., 2007;

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Fatichi et al., 2014); these processes-that may or may not induce consistent annual changes, depending on the environmental conditions faced by trees. For F. sylvatica and temperate oaks, Mmaximum correlation values were observed in F. sylvatica and temperate oaks for corresponded to flux aggregation periods that were consistent with the previously reported phenology of the woody biomass increment (Table 3, Michelot et al. 2012, Appendix Supplement -S10). (Babst et al., (2014) and; Granier et al., (2008) accordingly similarly reported high agreementsclose relationships between the AWBI and forest C fluxes that were summed until the time of cessation of growth—cessation (August/September). The Ffluxes aggregation periods were, however, strongly incoherent with not related to the timing of woody growth in Q. ilex and or P. abies (Cuny et al., 2012; Lempereur et al., 2015), which indicates that the inter-annual variation in the AWBI is not always solely (or even primarily, e.g., and in some cases (here in Q. ilex and P. abies) not primarily, dependenting on the C fuelled derived from photosynthesis. Specifically, the agreement between the observed annual growth and a short period of C flux aggregation occurring in early summer that we was reported in for Q. ilex is in line with corresponds to the major effect of growth cessation on the annual biomass increment, which that hasve been attributed to a drought-induced limitation of cambial activity at the Puéchabon site (Lempereur et al., 2015)M. Lempereur, N. K. Martin-StPaul, J.M. Ourcival et al., unpublished data). The processes that underlieving the relationship of the long flux aggregation period related to and the annual biomass increment in of P. abies possibly may involves include the effect of late winter temperature on cambium phenology (Rossi et al., 2011). Overall, our results suggest that using growth-flux correlation coefficients when investigating either the source -limitation of growth or the seasonality of C allocation to woody tissue-biomass can lead to misleading conclusions.

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# 4.2 <u>BThe between-site</u> variability in <u>the carbon C</u> allocation to woody biomass growth is related to ontogeny and competition intensity

We highlighted an age-related decline of in the C partitioning to woody biomass at spatial scale in all three species (Fig. 3B). This result had previously been reported observed in F. sylvatica stands using measurements of the main C compartments along a chronosequence (Genet et al., 2010). Different Several non-exclusive processes can explain this age-related trend. The lincreases of in tree height is are associated with an increases in the hydraulic resistance of

xylem, which hydraulic resistance that may lead to a declines in the turgor of living cells with and result in potentially negative consequences on cambial activity (Woodruff et al., 2004). This constraint may result in a height-related sink-limitation of growth (Woodruff and Meinzer, 2011), which is consistent with in line with our reportresults. Concurrently Additionally, life-history traits, such as a predominant greater emphasis oneffort toward reproduction in aged older stands, could also be involved... although However, the interactions between of growth and reproductiveen mechanisms are still under debate (Hoch et al., 2013; Thomas, 2011) and remain have yet to be properly represented in TBMs. Only the GPP component of the simulated forest tree C balance was retained in the final models (Table 4), thereby indicating that the an increase of in maintenance respiration with greater standing biomass was probably most likely did not involved contribute toin the age-related decline in biomass growth (Drake et al., 2011; Tang et al., 2014). While the Although height-related hydraulic constraints on C assimilation associated with tree height haves been thought suggested as to be an important driver (Ryan et al., 2006; Tang et al., 2014), recent local studies have evidenced suggested that changes in demography and stand structure -rather than decrease in tree functioning-may primarily explain the age-related decline observed in stand woody growth (Binkley et al., 2002; Xu et al., 2012). Our results suggests that changes in the C allocation should also be considered, as because no mortality occurred in the ourstudied plots over during the measurement period (data not shown). We additionally evidenced identified a significantly higher C partitioning to woody biomass in temperate oak stands with high-greater competition intensity (i.e., high stand basal area, Table 3). To date, rReports regarding the effect of competition on C allocation dynamics are so far conflicting (Litton et al., 2007) showing no large and consistent effect and suggest no significant or consistent effect. Besides Moreover, we found no significant effect of soil nutrient availability on the C allocation dynamics was evidenced along the studied ecological gradient whereas a recent meta-analysis reported that this factor affects positively affects C partitioning to forest biomass at on the global scale (Vicca et al., 2012). The RENECOFOR network only includes relatively fertile sites (AppendixSupplement –S5), which could putatively explain this the apparent tension between our results but remains putative at this stage and the conclusions of the meta-analysis. Therefore, mMore studies are thus required to further specifyelucidate the contributions of the different various drivers to the variation in the C partitioning to woody biomass along local to global environmental gradients.on scales that range from local to global.

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## 4.3 <a href="IThe-inter-annual variability">IThe-inter-annual variability in woody biomass growth is consistent with combined source-sink limitations</a>

Water or and temperature stresses exerted a significant direct control on the inter-annual variation of woody biomass growth (i.e., independently from their effects on C assimilation) in-for\_every species and biomes (Table 5 and Fig. 4 and 5). Cambial growth has been reported to be inhibited at lower water stress levels than photosynthesis (Muller et al., 2011; Tardieu et al., 2011). Indeed, the drought-induced decrease in cell turgor strongly affects the cell divisions (Woodruff and Meinzer, 2011) and cell wall expansion (Cosgrove, 2005; Lockhart, 1965) before gas exchange modulation comes into play. Similarly, there is evidence that cell growth processes, such as cell division, are affected earlier by low temperaturemore sensitive than photosynthesis to low temperatures (Körner, 2008). While Although these evidences findings documented the plausible mechanisms of the sink control of biomass growth at the cellular scale, there is still considerable debate as toregarding whether the sink or the C source actually limit the growth of the world's forests (Palacio et al., 2014; Wiley and Helliker, 2012). The typically observed large C reserve pools (Hoch et al., 2003; Würth et al., 2005) have been interpreted as a consequence of an overabundant C supply and thus as an evidence of a sink control of tree growth (Körner, 2003). However, recent works <u>have</u> suggested that a source limitation of growth may be compatible with large C reserve pools if part of this mobile C is sequestered rather than stored (Millard and Grelet, 2010) or if C storage is an active tree response to environmental stress (Dietze et al., 2014; Wiley and Helliker, 2012). Using an alternative methodology (i.e. a methodology that is not based on C storage measurement) our results suggest that sink limitation has a significant effect on the annual woody biomass growth in of five species that are representative of contrasted different European biomes, including including deciduous temperate forests. As-Because sink limitation implies that there are periods with significant C supply but no growth, this our results also corroborates recent empirical studies reporting that reported a significant role of growth duration in the annual variability of tree radial increment (Brzostek et al., 2014; Cuny et al., 2012; Lempereur et al., 2015). AWe additionally, we reported observed that past environmental constraints significantly affected the C partitioning to woody growth in everyfor each species and biomes (Table 5 and Fig. 4 and 5). The lagged effect of the previous year's low C supply (GPP<sub>y</sub>. 1) possibly indicates a preferential C allocation to storage at the expense of growth in trees facing that face C reserve pool depletion (Bansal and Germino, 2008; Wiley et al., 2013). In support of this finding, Richardson et al. (2012) accordingly reported a strong agreement relationship between the AWBI and the EC-based estimation estimate of the previous year's earbon-C supply in a mature maple stand. The important detrimental effect of low-a previous year's low C supply on the temperate oak woody growth (Fig. 4) may be related to its growth phenology, because as this species relieds on C reserves to achieve a large part of the its annual biomass growth before prior to leaf expansion in the spring (Barbaroux et al., 2003). The lagged effect of high water stress intensity in on F. sylvatica and Q. ilex (Fig. 4) may be linked to the pastprevious droughtinduced mortalities of buds or fine roots (Leuschner et al., 2001; López et al., 2003). Indeed, prebuilt buds are thought to strongly regulate the following year's cambial activity of the next year (Delpierre et al., 2015; Palacio et al., 2012; Zweifel et al., 2006) and a recent meta-analysis highlighted concluded that C is preferentially allocatedion to fine roots at the expense of woody growth in stands facing that face constraining environments (Chen et al., 2013). Finally, our results suggest that C supply (GPP<sub>gp</sub>) is an important driver of the annual woody biomass woody growth in temperate deciduous forests (Daudet et al., 2005). GPP was the component of the simulated forest-tree C balance that was most closely related to the annual variability of in growth; this result indicates GPP's important role in explaining the annual variability in which is in line with its preponderant role in the annual variability of the net ecosystem productivity of European forests- (Delpierre et al., 2012). Overall, our findings support the premise that forest woody biomass growth is under asubject to complex control processes that includeing both source and sink limitations, following the principle of Liebig's law: while although numerous processes potentially influence woody growth, the stand growth at a given site and a given time <u>year</u> is <u>predominantly</u> limited <del>only</del> by the most constraining factor. The C (source) limitation of growth can thus only occur when other factors are non-limiting (Fatichi et al., 2014), which a situation that is expected to be rare in strongly constrained environment such as Mediterranean or mountainous areas (Fig. 4).

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### 4.4 Toward an integrated modelling framework

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Most of the models that are currently used to project the outcome of global changes on forests represent woody growth as a fraction coefficient of the total C uptake (i.e., source control of growth, De Kauwe et al. 2014). This carbonC-centric perspective overlooks the possibility of-a sink control of growth, contradicting and thus ignores results such as those presented in this study and those of earlier local studies evidences found in this and earlier local studies (reviewed by Fatichi et al. 2014). Consequently, this perspective possibly hampers the ability of TBMs to project future forest productivityand possibly (Fatichi et al. 2014), explaining the low ability of TBMs to simulate the inter-annual variability of woody growth (Le Roux et al., 2001). On the basis of our analysis of the spatio-temporal dynamics of C allocation to woody growth at on a regional scale, we here suggest a straightforward, way toward a combined source and -sinkdriven forest growth modelling framework (Fig. 6). In this viewframework, a potential sitespecific allocation coefficient is first defined to represent the effect of soil fertility on the C allocation to wood (Vicca et al., 2012). In a second step, this coefficient is adjusted to the physiological state of the stand using by accounting for the dependences of the C allocation to-on ontogeny, competition intensity and lagged environmental stressesstressors. Lagged environmental stressors are represented by a negative effect on the previous year's water stress index and low C uptake on the allocation coefficient. Inclusion of tThe age-related decline in forest growth has been highlighted as an important modelling goal (Magnani et al., 2000; Zaehle et al., 2006), but However, this effect has thus far has so far been incorporated in models by including implemented through a negative effect of the increasing stand respiration on the forest tree C balance, which contradictsing current knowledge (Ryan et al., 2006). The resulting annual allocation coefficient is finally modulated by the direct water and temperature constraints on growth. In a similar fashion, Leuzinger et al. (2013) accordingly improved the LPJ Lund-Potsdam-Jena (LPJ) terrestrial biosphere model predictions at-for the cold tree-line using by accounting for representation of the direct temperature limitation of growth. Our approach can be seen as an intermediate step toward a more mechanistic representation of C allocation to woody biomass (Hölttä et al., 2010; Schiestl-Aalto et al., 2015). It further synthesizes our the current knowledge of regarding forest growth dependences and has the potential to unify in the same modelling framework seemingly contradictory observations within a single modelling framework. The simulated growth is indeed under subject to the combined controls of C supply

and changes in C allocation due to endogenous adjustments and/or modulations of the sink activity (Fig. 6). These controls result from distinct processes, which are independently represented in the modelling framework. The relative influences of the different various processes, i.e., the simulated growth causalities, are thus likely to vary both spatially and temporally, depending on the environmental conditions faced by trees. Our approach has therefore the potential to possibly shedding light on the contrasted results reported by correlative studies. While Although the value is comparable to those of previous reports studies (Lebourgeois et al., 2005; Mérian et al., 2011), the proportion of the annual growth variability that was explained by our approach remained was moderate (Table 5). Plausible explanations of this result include: i) unreported management interventions that could may have skewed the historical stand growth reconstruction and ii) potentially important growth drivers that were not considered here, such as changes in C partitioning due to mast seeding (Mund et al., 2010), genetic differentiation among tree populations (Vitasse et al., 2014) or allometry-mediated tree acclimation to drought (Martin-StPaul et al., 2013). A third factor that hampered the ability of our empirical models to explain the annual growth variability is the potential disagreement between the CASTANEA outputs that were used as explanatory variables and the corresponding actual drivers. Although we argued that i) the CASTANEA model has been thoroughly validated at many EC sites from throughout Europe and ii) the presented growth dependences demonstrated their robustness against the reported uncertainties of the CASTANEA simulations, the quality of the simulations was limited by the idiosyncrasies of the sites we examined in this study. In particular, a number of past disturbances such as insect outbreaks, windthrow or unreported commercial thinning could have temporarily induced large discrepancies between the actual and simulated C fluxes (Grote et al., 2011; Hicke et al., 2012). The error that is attributable to model performance unfortunately remains unknown because of the absence of EC measurements at our study sites (except for the Puéchabon site, see Delpierre et al., 2012). Despite this additional uncertainty, the combined use of field measurements and process-based modelling allowed us to present the first species-specific evaluation of annual C allocation to growth along regional environmental gradients. ONonetheless, our results suggest that implementing the presented C allocation dependences in TBMs will allow-refineing the projections of the outcome of global changes on forest growth, with and have implications on for the predicted evolution of the forest C sink, forest diebacks and tree species distributions (Cheaib et al., 2012).

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**Table 1.** Climate of the study sites. ETP: annual Penman - Monteith potential evapotranspiration; Precip.: annual precipitation; Temp.: annual temperature. Values are site averages ± standard deviation among sites.

Climatic features of the studied sites. ETP is annual Penman — Monteith potential evapotranspiration, Precip. is annual precipitation, Temp. is annual temperature. Values are site averages ± standard deviation among sites.

Species	number of plots	number of site- years	elevation (m)	ETP (mm)	Precip. (mm)	Temp. (°C)	Source
F. sylvatica	16	313	565 ± 326	1010 ± 121	1001 ± 133	$10.1 \pm 0.98$	RENECOFOR
Q. petraea / Q. robur	26	484	193 ± 76	999 ± 71	821 ± 96	$10.7 \pm 0.63$	RENECOFOR
P. abies	6	101	1056 ± 313	933 ± 44	1559 ± 340	$7.1 \pm 1.4$	RENECOFOR
Q. ilex	1	43	270	1417	907	13.4	Puéchabon site

<u>Table 2. Description of explanatory variables.</u> The "Type" category indicates the source of the data: field measurement (M), SAFRAN climate database (C) or CASTANEA simulation (S). The "Scale" categories indicate whether the variable was considered in the spatial (S) and temporal (T) analyses.

Table 2. Description of the explanatory variables considered in the analyses. The type category indicates the source of the data: measurement (M), SAFRAN climate database (C) or CASTANEA simulations (S). Scale categories indicate the variables considered in spatial (S) and/or temporal (T)

ľ	IDs	Description	Unit	Type	Scale
ı	age	Stand age	years	M	S
	AWBI	Annual woody biomass increment	gC.m <sup>-2</sup>	M	ST
	SBA	Stand basal area	m²	M	S
	camb_onset	Onset of the cambial activity	day of the year	S	T
	$GPP_{gp}$	Gross primary production of the current (y) growth period	gC.m <sup>-2</sup>	S	ST
	$GPP_{y-1}$	Gross primary production of the previous (y-1) year	gC.m <sup>-2</sup>	S	T
	frost	Sum of the average daily temperatures below -2°C during the last winter (year y-1 and y)	°C	C	ST
	LNC	Leaf nitrogen content	$gN.gDM^{\text{-}1}$	M	S
	$NPP_{gp}$	Net primary production of the current (y) growth period	gC.m <sup>-2</sup>	S	ST
	$NPP_{y-I}$	Net primary production of the previous (y-1) year	gC.m <sup>-2</sup>	S	T
	numstem	Stem density	number. ha <sup>-1</sup>	M	S
	SNA	Class of soil nutrient availability (1: low, 2: medium, 3: high)	unitless	M	S
Ì	SW <u>H</u> C	Soil water holding capacity	mm	M	S
	templim <sub>ep</sub>	Number of days of the current (y) growth period with an average temperature below 6°C	number of days	С	ST
	$Ra_{gp}$	Autotrophic respiration of the current (y) growth period	gC.m <sup>-2</sup>	S	ST
	$Ra_{y-1}$	Autotrophic respiration of the previous (y-1) year	gC.m <sup>-2</sup>	S	T
	$WS\_per_{gp}$	Number of days of the current (y) growth period with a soil water content below 60% of the soil water holding capacity	number of days	S	ST
	WS_per <sub>y-1</sub>	Number of days of the previous (y-1) year with a soil water content below 60% of the soil water holding capacity	number of days	S	Т
	$WS\_int_{gp}$	Water stress intensity index over the current (y) growth period	unitless	S	ST
	$WS\_int_{y-1}$	Water stress intensity index of the previous (y-1) year	unitless	S	T

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Table 3. Relationships of annual wood growth and the components of the seasonal forest carbon balance: NPP, GPP and Ra. The *start* and *end* terms (day of the year) indicate the carbon flux period that yielded the maximum value for the median of the growth-flux correlations among sites. The r term is the maximum obtained for the median of the site-specific Pearson correlation coefficients; values that are significantly different from 0 are indicated (\* indicates P < 0.05 and \*\* indicates P < 0.001). The  $\sigma$  term is the standard deviation of the Pearson correlation values among sites.

Species	GPP			Ra			NPP					
	start	end	r	$\sigma$	start	end	r	$\sigma$	start	end	r	$\sigma$
F. sylvatica	124	258	$0.62^{**}$	0.18	96	200	-0.29*	0.33	126	262	$0.58^{**}$	0.24
Q. petraea / Q. robur	136	214	$0.59^{**}$	0.25	98	192	0.31*	0.24	130	214	$0.50^{**}$	0.28
P. abies			$0.52^{**}$		78	348	0.11	0.52	32	200	$0.49^{**}$	0.29
Q. ilex	186	226	0.60		36	256	-0.26		186	226	0.58	

Table 4. Spatial dependences of annual wood growth: multiple regression estimates. The data have been centred and scaled.  $GPP_{gp}$  is the GPP of the growth period, age is the average age of the stand, and SBA is the stand basal area (Table 2). Values: estimates [F values]. All estimated values differed significantly from 0 (P < 0.001). All variables were retained in the bootstrap procedure (see main text).

Species	Estimates				
	$GPP_{gp}$	age	SBA	P	adj. R <sup>2</sup>
Q. petraea / Q. robur		-8.88×10 <sup>-1</sup> [39.5]	4.27×10 <sup>-1</sup> [19.5]	<10 <sup>-4</sup>	0.69
F. sylvatica	5.07×10 <sup>-1</sup> [59.4]	-6.96×10 <sup>-1</sup> [61.6]		<10 <sup>-4</sup>	0.88
P. abies	8.25×10 <sup>-1</sup> [8.6]			0.04	0.60

Table 5. Temporal dependences of annual wood growth: multiple regression estimates. The data have been centred and scaled.  $GPP_{gp}$  is the GPP of the growth period, WS  $per_{gp}$  is the water stress index of the growth period, WS  $int_{y-1}$  is the water stress index of the previous year, and  $templim_{gp}$  is the low temperature index of the growth period (see Table 2). D1 and D2 are dummy variables (D1 = 0 if  $GPP_{gp} < 1400$  gC m<sup>-2</sup>; otherwise, D1 = 1. D2 = 0 if  $GPP_{y-1} < 1550$  gC m<sup>-2</sup>; otherwise, D2 = 1; see Fig. 5). The ρ term is the parameter of the first-order autoregressive process that was used to model the temporal autocorrelation of the within-stand errors. Values: estimates [F values]. Estimated values that are significantly different from 0 are indicated (\* indicates P < 0.05, \*\* indicates P < 0.01, and \*\*\* indicates P < 0.001). A Δ index indicates that the variable was not retained in the bootstrap procedure (see main text).

Estimates	Species			
	Q. petraea /	F. sylvatica	P. abies	Q. ilex
	Q. robur			
$GPP_{gp}$	3.26×10 <sup>-1</sup> ***	4.87×10 <sup>-1</sup> ***	2.4×10 <sup>-1</sup> *[3.5]	
$WS\_per_{gp}$	-1.09×10 <sup>-1</sup> **	-2.04×10 <sup>-1</sup> ***		-5.8×10 <sup>-1</sup> ***
$WS\_int_{y-1}$		-2.37×10 <sup>-1</sup> ***		-2.2×10 <sup>-1*</sup> [6.3]
$GPP_{y-1}$	$3.82\times10^{-1}$ * [3.3]		-4×10 <sup>-1</sup> ** [3.2]	
templim <sub>gp</sub>	-9.60×10 <sup>-2</sup> **Δ		-1.26 *** [3.5]	
D1			-2.4×10 <sup>-1</sup> ***	
D2	-3.9×10 <sup>-1</sup> ** [0.8]			
$D1 * GPP_{gp}$			1.33 ** [8.2]	
$D2 * GPP_{y-1}$	-4×10 <sup>-1</sup> ** [6.4]			
ρ	0.61	0.68	0.52	0.44
P	<10 <sup>-4</sup>	<10 <sup>-4</sup>	7.7.10 <sup>-3</sup>	<10 <sup>-4</sup>
adj. R²	0.21	0.42	0.20	0.43

## Figure captions

Figure 1. The conceptual framework and the three sources of data (field measurements, climate reanalysis, and process-based simulations) used in the analyses.

Figure 2. Locations of the study sites.

Figure 3. **Spatial dependences of annual wood growth.** A: Relationship of the AWBI and the GPP of the growth period ( $GPP_{gp}$ ) averaged over sites. B: Age-related decline of the C partitioning to AWBI (AWBI /  $GPP_{gp}$ ).

Figure 4. Temporal dependences of annual wood growth: the roles of explanatory variables from RF classification. Variable importance is expressed as the percentage of the importance of the top-ranked explanatory variable. The variable identifiers (IDs) are listed in Table 2. The coloured variables were retained in subsequent analyses.

Figure 5. Temporal dependences of annual wood growth: marginal effects of each explanatory variable on the annual wood growth. The lines represent smoothing splines with 50% frequency response cut-offs. The coloured areas indicate the 95% confidence intervals. The 5% and 95% data quantiles (grey areas) were not considered in the discussion. The marginal effect of a given variable *X* was obtained by fixing the value of *X* and averaging the RF predictions over all the combinations of observed values for the other predictors in the dataset (Cutler et al., 2007). The marginal predictions were collected over the entire range of *X* in the training data using a regular grid.

Figure 6. Modelling framework for a combined source- and sink-driven representation of C allocation to wood growth.

Figure 1. The conceptual framework and the three sources of data (field measurements, climate reanalysis, process-based simulations) used for the analyses.

### Figure 2. Location of the studied sites.

Figure 3. Spatial dependences of the annual woody growth. A: relationship between the annual woody biomass increment (AWBI) and the gross primary productivity of the growth period (GPP<sub>sp</sub>) averaged over sites. B: Age-related decline of the earbon partitioning to AWBI (AWBI / GPP<sub>sp</sub>).

Figure 4. Temporal dependences of the annual woody growth: importance of explanatory variables from random forest classification. Importance is expressed as % of the importance of the explanatory variable ranked first. Variable IDs are given in Table 2. Colored variables were retained in the following analyses.

Figure 5. Temporal dependences of the annual woody growth: partial dependence describing the marginal effect of each explanatory variable on annual woody growth. Lines are smoothing splines with 50% frequency response cut off. Colored areas are 95% confidence intervals. The part of the dependences corresponding to the 5% and 95% data quantiles (grey areas) were not considered in the discussion.

Figure 6. Modelling framework for a combined source-sink-driven representation of the carbon allocation to woody growth.