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Assessing seasonality of boreal coniferous forest CO₂ exchange by estimating biochemical model parameters from micrometeorological flux observations

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Abstract

The biochemical seasonality of the northern boreal coniferous forests was investigated by means of inversion modelling using eddy covariance data. Eddy covariance data was used to optimize the biochemical model parameters. Our study sites consisted of three Scots pine (I. Pinus sylvestris) forests and one Norway spruce (I. Picea abies) 5 forest that were located in Finland and Sweden. We obtained temperature and seasonal dependence for the biochemical model parameters: the maximum rate of carboxylation $(V_{c(max)})$ and the maximum rate of electron transport (J_{max}) . Both of the parameters were optimized without assumptions about their mutual magnitude. The values obtained for the biochemical model parameters were similar at all the sites dur-10 ing summer time. To describe seasonality, different temperature fits were made for the spring, summer and autumn periods. During summer, average J_{max} across the sites was 54.0 μ mol m⁻² s⁻¹ (variance 31.2 μ mol m⁻² s⁻¹) and $V_{c(max)}$ was 12.0 μ mol m⁻² s⁻¹ (variance 6.6 μ mol m⁻² s⁻¹) at 17°C. The sensitivity of the model to LAI was also studied. Simulation runs were done to study the effect of the seasonality implemented in 15

the model using different temperature fits. The impact of seasonality implemented in was 15%, which corresponded to an increase of 2°C in air temperature.

1 Introduction

According to scenarios of future climate, the boreal forest zone is expected to experience larger increase in temperature than other regions (Trenberth et al., 2007). Comparison across the latitudinal spread of boreal forest is therefore important, so that it is possible to predict how the northern forests will behave in the future, and what effects the future climate might have on their carbon balance. In our study we have assessed this by investigating boreal sites at different latitudes and comparing them.

²⁵ In order to obtain estimates for large-scale carbon sinks, it is important that the global and regional models are parameterized using a sufficiently good method. Large-





scale models often use photosynthesis parameters that have been estimated at the leaf level and then scaled to the canopy level (Sellers et al., 1996). Wang et al. (2006) pointed out that this is not the most reliable way, and therefore it is essential to also do parameterization of photosynthesis models also on the larger scale, taking advantage of the widespread eddy covariance flux tower network.

A biochemical model based on a mechanical description of photosynthesis was developed by Farquhar et al. (1980) and will henceforth be referred to as the biochemical model. The biochemical model is widely used in modelling on various scales (Juurola et al., 2005; Knorr and Kattge, 2005). It has three important parameters: the maximum carboxylation rate ($V_{c(max)}$), the maximum potential electron transport rate (J_{max}) and the efficiency factor for the use of light (q). Global models do not usually take into account the seasonality of the temperature-dependent photosynthesis parameters. These have been shown to be affected by seasonality (Dang et al., 1998; Xu and Baldocchi, 2003; Han et al., 2006).

¹⁵ boreal forests (Wang et al., 2006).

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In cold climate regions, the photosynthesis in forests does not immediately reach its full capacity at the beginning of the active season. It may take several weeks before the damage caused by low winter temperatures are fully repaired (e.g., Pelkonen and Hari, 1980). Thus, the transition period from winter dormancy to full photosynthetic capacity plays a significant role in altering the carbon balance of northern boreal coniferous

- 20 plays a significant role in altering the carbon balance of northern boreal coniferous stands (Bergh et al., 1998). A field study by Bergh and Linder (1999) of Norway spruce concluded that the spring recovery was mainly controlled by mean air temperature and severe night frosts. The importance of mean air temperature for spring recovery was also shown by Tanja et al. (2003). In their study they used eddy covariance data from
- ²⁵ boreal sites; their objective was to find the average air temperature that raises the photosynthesis to a level of 20% of the maximum summertime fluxes.

Our aim was to study the seasonality and transition periods of northern forests by means of biochemical model parameters estimated from micrometeorological observations. Model parameters were deduced from CO_2 flux observations by inverting a



canopy photosynthesis model. Our motivation was to study whether the phenomenology of larger-scale models can be improved, and also to assess how a warmer climate will affect the carbon balances of northern boreal forests.

- We parameterized a canopy-scale model that was upscaled from the leaf level in order to obtain the parameters $V_{c(max)}$ and J_{max} for four different coniferous forest sites, all located in the boreal zone. The parameterization results for the different sites were intercompared. The relations between the three biochemical model parameters were assessed, as well as the relations between air temperature and the biochemical model parameters, over the whole growing season. Also, it was investigated whether temper-
- ature indices could be used in improving the seasonality of the biochemical parameters in the modelling. In addition, the effect of night frosts was studied at different sites together with how its effect on the delay in the spring recovery was different for different values of the temperature sum. Finally, some simulation runs were conducted in order to study how large an impact the modelled seasonality and future climate conditions will have on GPP.

2 Materials and methods

2.1 Measurement sites and measurements

We studied four conifer forests that were located in the boreal zone. Two sites, Kenttärova (67°59'N, 24°15'E) and Sodankylä (67°21'N, 26°38'E), are situated in northern Finland and in the northern boreal zone (Solantie, 2005). The Sodankylä site is a Scots pine (*Pinus Sylvestris*) forest (total Leaf Area Index, LAI 3.6). Kenttärova is a homogenous Norway spruce (*Picea abies*) forest (LAI 6.6). The Scots pine-dominated forest at Hyytiälä (61°51'N, 24°17'E, LAI 8.0) is in the southern boreal zone in Finland. The southernmost site of Norunda (60°50'N, 17°28'E) is in the hemi-boreal zone in forest (LAI 13.5). Annual mean temperatures range from –1.7°C in the northern boreal



zone to 5.5° C in the hemi-boreal zone. More detailed descriptions of the sites are to be found in Lindroth et al. (2008).

We used at least two years of data from each of these sites: Hyytiälä 2000–2001, Sodankylä 2001–2004 and Norunda 1999, 2001 and 2002. For these three sites the year 2001 was used for the model parameterization. To study the spruce forest of Kenttärova, we used data for the years 2003–2006, the latter year 2006 being used for parameterization.

Net fluxes of momentum, sensible and latent heat and carbon dioxide fluxes were measured by the eddy covariance method using fast-response sonic anemometers and closed-path IRGAs. Measurement heights were at least three metres above the highest trees. Other meteorological variables measured included Photosynthetic Photon Flux densities (PPFD), air temperature, relative humidity, air pressure and precipitation. The global radiation and reflected radiation were used to calculate albedo.

2.2 Modelling and data analysis

¹⁵ A canopy-level gas exchange model was parameterized by inversion from the canopy CO_2 flux data. The leaf level CO_2 gas exchange model was based on a formulation first introduced by Farquhar et al. (1980) and Farquhar and von Caemmerer (1982) with later modifications (De Pury and Farquhar, 1997). The biochemical model has a RuBP regeneration-limited rate (A_j) and a Rubisco activity-limited rate (A_c) , of which the minimum of the two governs the photosynthesis.

 A_i -assimilation is

$$A_j = J \frac{c_j - \Gamma_*}{4(c_j + 2\Gamma_*)} - R_d$$

and A_c –assimilation is described as

$$A_{c} = V_{c(\max)} \frac{c_{i} - \Gamma_{*}}{k_{c} (1 + o/k_{o}) + c_{i}} - R_{d}$$

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(1)

(2)

In these two equations $V_{c(\max)}$ is the maximum rate of carboxylation, R_d is the rate of non-photorespiratory respiration, k_c and k_o are the Michaelis-Menten constants for CO_2 and O_2 , Γ^* is the CO_2 compensation point in the absence of non-photorespiratory respiration, o is the oxygen concentration in chloroplasts (assumed constant) and c_i is the carbon dioxide concentration inside chloroplasts. J is the potential electron transport rate that is a function of incident irradiance, the light use efficiency factor (q), the

convexity of the light response curve and J_{max} , the maximum rate of electron transport. The temperature dependence of Γ_* was taken from Brooks and Farquhar (1985) and the temperature dependencies of the Michaelis-Menten constants were adopted from

¹⁰ Farquhar et al. (1980) and Harley and Baldocchi (1995). $V_{c(max)}$ and J_{max} have an Arrhenius-type temperature dependency (Harley and Baldocchi, 1995):

$$f_T = f_0 \exp\left[\frac{E_f(T - 298.15)}{298.15RT}\right]$$

where f_0 denotes the parameter ($V_{c(\max),o}$ or $J_{\max,o}$), E_f is the activation energy, R is the gas constant and T is temperature in Kelvin.

¹⁵ The Ball-Berry conductance (Ball et al., 1987) was used in conjunction with the biochemical model for describing the stomatal conductance.

The model and its up-scaling are described in detail in Thum et al. (2007). In order to up-scale the leaf-level model to the canopy scale, a vertical profile of the leaf area distribution was constructed individually for each site with the help of the beta distribution

(Wu et al., 2003). The forest canopy was divided into four layers, with equal biomass in each layer. The radiative transfer by Sellers (1985) was used for radiation calculations. The biochemical model parameters were assumed to decrease with height above the ground proportionally to the percentual PPFD, similarly to the nitrogen content (Sellers et al., 1992; Kull and Jarvis, 1995). The leaf layers were also separated into sunlit and shaded parts, according to Thornley (2002).

Before we can use NEE data from eddy covariance measurements for calculation of photosynthesis parameters, we have to subtract the respiration components.

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(3)





Needle respiration values for the two sites located south of the Arctic Circle were estimated from the Hyytiälä shoot chamber measurements (Kolari at al., 2007). For Sodankylä, too, the needle respiration was estimated from shoot chamber measurements (Thum et al., 2007). For Kenttärova, the needle respiration was taken from the literature (Stockfors and Linder, 1998), using needle dry weight data measured at the site (Steinbrecher et al., 1999).

First, the needle respiration was subtracted from the night-time flux measurement, after which the soil respiration was fitted to the night-time flux measurement data using the temperature response presented by Lloyd and Taylor (1994). Air temperature was used, since it had a more continuous time series than soil temperature. The fittings were made to biweekly data sets, and both of the two parameters (respiration at 10°C and the activation energy) were fitted. At Hyytiälä a response function based on both air and soil temperature introduced by Markkanen et al. (2001) was used, since a continuous time series in soil temperature was available at Hyytiälä and the fit yielded

15 good results.

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The soil and needle respiration dependencies on temperature and the vertical biomass distribution were different from site to site, but other model parameters, e.g., radiation parameters and light use efficiency q, were kept constant in this study (see Thum et al., 2007).

²⁰ A parameterization year was chosen, and temperature responses for the biochemical model parameters J_{max} and $V_{c(max)}$ were obtained from the inversed CO₂ flux data. For estimating the biochemical parameters, a procedure introduced by Lloyd et al. (1995) was used: The measured CO₂ flux points from late morning were used in the inversion. Measurements with light levels between 600 μ mol m⁻² s⁻¹ and 800 μ mol m⁻² s⁻¹ were used to obtain the temperature response for the parameter J_{max} . To estimate the pa-

rameter $V_{c(max)}$ measurements at higher light levels and the fit found for the parameter J_{max} were used. Inversed parameter values were plotted as a function of temperature, and Arrhenius-type fittings (Eq. 3) were performed. Further details of the optimization process are presented in Thum et al. (2007).

The biochemical model parameters were solved by inversion for all the years studied, and daily averaged values of them were used in conjunction with the values of the temperature-related indices to study the seasonality of the different forests. To study whether parameterizations are also applicable during other years, model runs for dif-

⁵ ferent years were done. In these runs, various different temperature responses were temporally kept the same as they were in the parameterization year. The applicability of the summertime parameterization at Hyytiälä to other sites was investigated.

We also improved the modelling of seasonality by using temperature indices. The so-called changeover dates are the days when the temperature fit for one period is

- switched to that for the next, e.g., from spring to the summer period. The spring recovery of forests is more dependent on temperature than on the calendar date. Moving these changeover dates according to temperature might improve modelling results during other years compared to keeping the dates the same as those in the parameterization year. To test this, we used the temperature sum, which is the sum of positive daily average temperatures (Solantie, 2004) and the five-day floating average temperature
 - (Tanja et al., 2003).

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Simulation runs were conducted for one year at Sodankylä. We compared the control run GPP having realistic Sodankylä data with model runs separately as follows: without seasonality, with a doubled CO_2 concentration, with a two-degrees warmer temperature and with both of the last two effects combined.

We studied the sensitivity of the inversed model parameters to the LAI, since LAI is very important factor in photosynthesis (Lindroth et al., 2008). The effect of LAI on the modelling results was investigated by changing its value in 2001 at Hyytiälä by 20%. We also studied the functioning of the model with other values of the biochemical parameters. In fact, there are three important parameters involved: as well as the above-mentioned J_{max} and $V_{c(\text{max})}$, the internal light use efficiency q also plays an important part in the model. In many models it is $V_{c(\text{max})}$ and q that are estimated, while J_{max} is just assumed to have a certain relationship to $V_{c(\text{max})}$ (Wang et al., 2003; Verbeeck et al., 2006).

In the Eq. (1) q is involved in the RuBP regeneration-limited rate, and controls the potential electron transport rate J. At low light levels, it is the parameter q that determines the RuBP-regeneration-limited rate, and therefore the lower light limit for estimating J_{max} was set at 600 μ mol m⁻² s⁻¹. Since q and J_{max} appear in the same equation, their estimation might influence one another. In order to study this possibility, we simultaneously minimized all three biochemical parameters for Hyytiälä, and looked at what effect this had on their values.

We also assumed a constant relationship between $J_{\rm max}$ and $V_{c({\rm max})},$ using the equation

¹⁰
$$J_{\text{max}} = 29.1 \mu \text{molm}^{-2} s^{-1} + 1.64 V_{c(\text{max})}$$

from Wullschleger (1993) and estimated J_{max} with this equation from the inversed $V_{c(max)}$ values. In this examination q was kept constant. We looked at the ratio of the two parameters with this kind of estimation. In addition to these values, we studied the performance of the model with J_{max} and $V_{c(max)}$ taken from shoot chamber measurements made at Värriö, which is located in Finnish Lapland (Aalto et al., 2002).

3 Results

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3.1 The parameterization at different sites

We obtained plots for temperature dependencies from model inversion. After analyzing these results and the functioning of the model, using multiple temperature dependen-

²⁰ cies for the parameters appeared as a feasible step forward. These different responses changed according to the season, and different time periods were used for each of the sites. At all sites, the inversed spring values for parameter $V_{c(max)}$ were at a lower level than the summertime values, and therefore another fit was made for the spring time. For three sites, different fittings for J_{max} were also performed separating the spring and ²⁵ summer periods.

(4)



The fitting parameters are shown in Table 1 together with the values of the biochemical parameters at 17°C. In the summertime fits, the values for J_{max} vary between 48.8 and 61.7 μ mol m⁻² s⁻¹, which is not a large variation, considering the latitudinal spread between the different sites. The value of the parameter $V_{c(max)}$ at 17°C had a relatively

larger fluctuation in summertime: between 10.4 and 15.8 μ mol m⁻² s⁻¹. 5

In Fig. 1a and b the temperature responses of J_{max} and $V_{c(max)}$ and their fittings are displayed for the northernmost site, Kenttärova, in the year 2006. In all the figures the temperature responses were extrapolated for the temperature range shown in the figure, even though the temperature for the springtime dependencies did not cover this whole range. For the Kenttärova site, it was feasible to produce two temperature

responses for both of the biochemical parameters. The springtime fitting period was May, when photosynthesis started. The estimates of J_{max} and $V_{c(max)}$ found in May were lower than those during summer, and they still had a temperature dependency. Some J_{max} late summer points were also low, but better modelling results were not obtained by using the spring fit during that time. 15

The temperature responses of the parameters and their fittings for Sodankylä are shown in Fig. 2a and b. For this site, the year 2001 was used for the seasonal temperature-dependence parameterization. Two fits were made for the parameter J_{max} . The changeover date was determined by the change in magnitude of the daily values and the simulation results of the model. The change between the spring and 20 summer regimes was estimated to occur on 4 June. In the parameterization of $V_{c(max)}$, a similar spring fit was made, as well as an additional fit for the transition period from 4 to 24 June.

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Using the chronologically-extensive Hyytiälä database, it was feasible to divide the growing season of 2001 into three periods for J_{max} and four for $V_{c(max)}$ according to the season. These are shown in Fig. 3. For the parameter J_{max} , a separate fit was made for springtime until the end of April, and again for the autumn beginning on 1 September. During the spring, the values were hardly temperature dependent at all, whereas the autumn values slowly rose with temperature. The $V_{c(max)}$ results were not temperature

dependent either in April or May, but in the latter month the values were larger. The autumn values increased with temperature, but slower than the summer values. With these fits the model worked best when compared to the measurements.

- For Norunda the inversed points were quite scattered (Fig. 4). There were no early spring values for the parameter J_{max} and therefore it was not possible to perform a fit for this time period separately. A single fit for the whole growing season was therefore used for J_{max} . During March and the first half of April the observations of $V_{c(max)}$ were at a lower level, and a separate fit for this time period was done. For the rest of the growing season from 20 April onwards only one fit for $V_{c(max)}$ was made.
- ¹⁰ Model runs for different years were conducted. In Table 2, r^2 values are shown for the model performance during the whole growing season in the various years studied. The model also succeeded similarly in those years that were not used for the parameterization. Since similar summertime values for the parameters were obtained (Table 1), their applicability across the four sites was examined. Parameter values from ¹⁵ Hyytiälä in summertime were applied to the other sites during summer. The r^2 values
- from these simulations are also shown in Table 2. There is not much variation in the results for the summertime runs between the sites' own parameterization and Hyytiälä's values. Only at Sodankylä are the r^2 values then slightly smaller.
 - 3.2 Linking seasonality to temperature indices and modelling
- ²⁰ In Fig. 5a the daily values of the inverted biochemical parameters J_{max} and $V_{c(max)}$ for Kenttärova in the year 2001 are shown. The CO₂ flux measured by eddy covariance method is shown in Fig. 5b. The five-day floating average temperature (5Dave) the minimum temperature are also shown (Fig. 5c). The same parameters, data and fiveday average temperature for Sodankylä, Hyytiälä and Norunda, respectively in Figs. 6–
- 8 are represented. For the biochemical parameters similar values were obtained at all the sites. At Norunda and Hyytiälä. the first night frost of the autumn caused a large drop in the biochemical parameters (Figs. 7a and 8a).





We investigated the relation between the temperature sum and the inversed biochemical model parameters. The temperature sum is defined as the sum of positive daily average temperatures. At Hyytiälä it was observed that, when the temperature sum was below 270°C d, night frosts resulted in a lowering of $V_{c(max)}$ to early springtime values. Later-occurring night frosts did not have such large effect, only causing some reduction in the rate of increase of $V_{c(max)}$. At Norunda similar behaviour was found. The temperature sum rose above 270°C d during mid-May at Hyytiälä and during late April–early May at Norunda.

At the Sodankylä and Kenttärova sites, severe night frosts occurred until the end of May. At such northern sites some night frosts may even take place during August, as happened at Sodankylä during 2001 and 2004. When the temperature sum was below 200°C d, the night frosts seemed to have some effect on the biochemical parameters at these two sites. Crossing this threshold occurred during late May–early June in both places.

As a result of the optimizations the sites obtained several temperature responses for the biochemical parameters, as was seen in Figs. 1–4. The time periods for these temperature responses were based on the inversion results during the parameterization year. Since different years are not similar e.g. in respect to the spring recovery, this kind of parameterization caused, for example, an underestimation of the fluxes in Sodepleding in the environment of 0000 which was warmen then the provise environ.

²⁰ dankylä in the springtime of 2002, which was warmer than the previous spring. In order to investigate whether temperature indices can be useful here, we linked them with the changeover dates of temperature responses and investigated the resulting simulated CO₂ fluxes.

We used 2002 at Sodankylä and 2000 at Hyytiälä as our test years. The five-day ²⁵ running temperature and the temperature sum were used to as proxies to describe the photosynthetic state of the vegetation (Tanja et al., 2003).

The changeover dates of the temperature responses of the parameterization year 2001 in Sodankylä were 4 June and 25 June. Taking the temperature sums on these two days and locating them in the year 2002 resulted in the dates 27 May and 12 June.

When the model was run with these new changeover dates, the modelling results were improved.

Using the five-day running temperature it was feasible to locate the first changeover date, which was 29 May, but it was not possible to define the other date. The temperature limit was crossed many times. This occurred for the first time on 30 May; trying this as the second changeover date caused overestimation of the fluxes. Using May 29 as the first changeover date did improve modelling results, however.

At Hyytiälä there were three changeover dates, two of them taking place during spring time. The first changeover date for temperature dependencies in the parameterization year 2001 was May 1 and the second 1 June. With the temperature sum, the corresponding dates in 2000 were 27 April and 25 May. Applying these dates for changing the temperature responses yielded better modelling results. At Hyytiälä there was also a separate fitting period for the autumn time starting, in 2001, on 1 September. The temperature sum approximated the corresponding change to be on 6 September in 2000. It was not possible to see if using this new date improved modelling results

¹⁵ In 2000. It was not possible to see if using this new date improved modelling residuces because there was a data gap in the flux measurements at this time.

The five-day running average temperature estimated the first changeover date in spring 2000 to be 21 April. This approximation gave even better results than the earlier estimate by the temperature sum. It was not possible to obtain the second changeover

- date because that temperature limit was crossed many times. During autumn the estimation was taken from the last crossing of the temperature limit (12.3°C), on 3 September. Again, it was not possible to compare the modelling results due to the data gap, but it was interesting that the two methods gave estimates for the changeover date in autumn that were very close to each other.
- 25 3.3 Effect of climate warming and seasonality in the simulation experiments

A control run was performed using realistic data for one year at Sodankylä. The GPP for the control run was 190.1 g C m⁻². In the first experiment the CO₂ concentration was doubled. This caused an increase of 20% in GPP. In the second experiment the



temperature was raised by two degrees. The dates for the switching of the temperature fits was made according to the method described in the previous section, using the temperature sum. The GPP was calculated with temperature fits for the biochemical model parameters levelling off at 20°C and also with temperature fits continuing to

- ⁵ increase after this point. The first approach gave a 13% addition to GPP and the latter a 16% increment. The last experiment included increases in both temperature and CO_2 concentration. Without the upper threshold for temperature fits for the model parameters there was a 41% addition to GPP, while with the limiting conditions the increment was 37%.
- ¹⁰ The effect of the seasonality included in the model via different temperature responses in spring and summer was also investigated by simulation. When only the summer time temperature fits for the model parameters were employed throughout the year, the GPP increased by 15%.
 - 3.4 Sensitivity analysis
- 15 3.4.1 Effect of LAI

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In the sensitivity analysis, the effect that LAI had on the modelled CO₂ fluxes was studied. The sensitivity analysis was performed using data from Hyytiälä in 2001. At Hyytiälä the LAI was $8 \text{ m}^2/\text{m}^2$. The LAI was increased and decreased by 20%. First was examined how the changes impacted the fluxes without re-parameterization of J_{max} and $V_{c(\text{max})}$.

As LAI was decreased by 20%, the vegetation in lower layers obtained more light, yet the amount of assimilating biomass was decreased. This resulted in a lowering of the modelled CO_2 flux. On a bright summer day around noon the modelled CO_2 flux maximum was 23% lower compared to the model result with the measured LAI.

Next the re-parameterization was conducted by inversion; this caused an increase in both of the parameters. There was no other systematic change in the modelling results compared to the original model. However, more scattering was introduced, and on





bright summer days the CO_2 flux was estimated to be up to 27% more than the model result with the original LAI.

When LAI was increased by 20%, the CO_2 flux was also systematically underestimated without re-parameterization, by approximately 9%. Even though there was more biomass, more attenuation of the incoming radiation occurred. A quarter of the biomass was located in the lowest layer, which was not receiving much light. After reparameterization of J_{max} , the model results were, on the average, close to the original ones. The highest difference between the two was approximately 14%.

3.4.2 Different ways to estimate J_{max}

- ¹⁰ According to the literature, the biochemical model parameters are considered to have a relatively constant ratio to each other, the relation $J_{max}/V_{c(max)}$ usually being around 1.68 to 2.0 (Medlyn et al., 2002b; Leuning 2002) at 25°C. Since the temperature dependencies of J_{max} and $V_{c(max)}$ are dissimilar this ratio is dependent on temperature: at 20°C their ratio has been estimated to be 2.68 (Leuning, 1997).
- ¹⁵ Since our estimations for the biochemical parameters, and especially for their relationship, differed from the values stated in the literature, they were examined more closely. Since their mutual relationship is considered to be relatively constant, even though the magnitudes of the parameters may vary considerably, this was considered important. When three parameters $(J_{max}, V_{c(max)}, q)$ were optimized simultaneously, the light use efficiency q remained at a quite constant level of 0.14. This was the same value as that estimated earlier based on the literature values (Aalto, 1998). The new J_{max} and $V_{c(max)}$ values were close to earlier estimates. This result gave us confidence in our previous estimates by inversion.

Closer inspection of the inversed J_{max} values led to the conclusion that the A_j and A_c – limited rates were alternating as was expected. The A_c – limited rate was dominant in very good light conditions, while photosynthesis was limited by A_j at lower light levels. We used Eq. (4) (Wullschleger, 1993) to estimate J_{max} at Hyytiälä in 2001. The new J_{max} values are shown in Fig. 9a. Here we have the values for J_{max} that were obtained,



when $V_{c(\text{max})}$ was optimized and J_{max} was then estimated by Eq. (4). When compared to the inversed results displayed in Fig. 3a, it was noticed that, at temperatures above 15°C, these two different approaches gave values of the same magnitude. At lower temperatures the difference was more pronounced. At around 10°C, Eq. (4) gave values in the range 35 to 40 μ mol m⁻² s⁻¹, whereas our inversion results were around 18 to 24 μ mol m⁻² s⁻¹. Equation (4) is based on results at temperatures at which the A/C_i -curves used in the parameterization were measured; these results do not include any measurements at temperatures below 20°C (Wullschleger, 1993).

Equation (4) is based on the relationship between J_{max} and $V_{c(max)}$ of 109 plant species. From our inversion we obtained the value of $10.4 \,\mu$ mol m⁻² s⁻¹ for $V_{c(max)}$ at 17° C at Hyytiälä. By inversion we got $48.8 \,\mu$ mol m⁻² s⁻¹ for J_{max} . Placing the Hyytiälä value for $V_{c(max)}$ in Eq. (4) produced $46.2 \,\mu$ mol m⁻² s⁻¹, which is close to our inversed estimation. This led to a ratio 4.4 between the parameters, however, which is different from that found in the literature.

¹⁵ We studied three different methods to estimate the parameters. In the first approach, the inversed $V_{c(max)}$ values were used, and J_{max} was estimated using Eq. (4). In the second approach, we used the parameterization made in Värriö for Scots pine using shoot chamber measurements (Aalto et al., 2002). The third approach was the inversion modelling of the eddy covariance data for both of the parameters. Three dif-²⁰ ferent temperature relationships for J_{max} are displayed in Fig. 9a), i.e., our summertime estimate from the inversion, Eq. (4) results and the fit obtained from chamber measurements. In Fig. 9b the fits from inversion and chamber measurements for $V_{c(max)}$ are shown.

When we used parameter values from the chamber measurements, the modelled fluxes were greatly overestimated during April and May. During the summer months the simulated fluxes underestimated the measurements. In September and October the CO₂ fluxes were slightly exaggerated by the model. This behaviour can be seen from the temperature fits in Fig. 9. At low temperatures the fit from the chamber is higher than the fits from the EC measurements, especially during the spring months,

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whose fits are shown in Fig. 3. At higher temperatures the fit is lower than the inverted fit, and as a result the warm days are underestimated.

Using the estimation for J_{max} from Eq. (4) resulted in similar results to the inverted J_{max} , except that during April the simulated flux is overestimated. This was caused by the difference in the estimations of J_{max} at low temperatures.

4 Discussion

The values of the biochemical parameters (J_{max} and $V_{c(max)}$) were comparable at all sites, the parameterizations performed for the different sites yielded surprisingly similar results (Table 1). Originally, more variation in the results with latitude and different species was expected. However, according to Bergeron et al. (2007), there was not much variability in three boreal black spruce forests located in different climatic zones regarding the temperature responses of gross ecosystem production and respiration on a monthly timescale. Medlyn et al. (2005a) studied three different coniferous sites and found that the largest difference in net ecosystem productivity was caused by soil respiration, with needle respiration also playing a role.

When comparing our values to the literature, we need to multiply our parameters by π , in order to convert from total leaf area to the projected leaf area (Stenberg et al., 1995).

The results for Scots pine forest in Sodankylä were compared to literature in Thum et al. (2007). The J_{max} values were in quite good accordance in low temperatures with the literature but highly exaggerated in temperatures above 15°C. The $V_{c(max)}$ values were close to the literature values. The same applies to the two more southern Scots pine sites. For Sodankylä J_{max} in 17°C was 193.8 μ mol m⁻² s⁻¹, Hyytiälä J_{max} was 153.3 μ mol m⁻² s⁻¹ and in Norunda 170.1 μ mol m⁻² s⁻¹, as shown in Table 1 and multiplied by π . For $V_{c(max)}$ we had 49.6 μ mol m⁻² s⁻¹ in Sodankylä, 32.7 μ mol m⁻² s⁻¹ in Hyytiälä and 35.2 μ mol m⁻² s⁻¹ in Norunda. The literature values for the biochemical parameters are shown in Table 3.

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The values in Table 1 are shown for 17°C, since this is a common summertime temperature in northern Finland. Usually the values of biochemical model parameters are given for 25°C, but since this is an uncommon temperature at many of the sites, it was not possible to make a reliable fitting in that temperature range. From leaf chamber measurements in northern Finland at Värriö it was only possible to estimate the parameter J_{max} below 20°C (Aalto, 1998). In estimations of J_{max} in Finland, the fitted temperature response has been found to have an optimum: in Aalto (1998) the optimum J_{max} value was below 20°C and in Wang et al. (1996) the optimum was below 25°C.

- In the literature (Wullschleger, 1993), there is a difference between the parameter values for Norway spruce and Scots pine, the latter giving considerably higher estimates. For Norway spruce, Wullschleger (1993) has given two estimates from two different measurements, shown in Table 3, measured at higher temperatures than ours. In this inversion for Kenttärova we estimated J_{max} to be 161.2 μ mol m⁻² s⁻¹ during summer time at 17°C. For 1/2
- ¹⁵ summer time at 17°C. For $V_{c(max)}$ we obtained 33.0 μ mol m⁻² s⁻¹ during summer. The literature values are shown in Table 3. Compared to the biochemical parameters measured by Grassi et al. (2001) our J_{max} value was large and our $V_{c(max)}$ estimation was on the same scale. Estimates for Flakaliden in Sweden are of the same magnitude (Roberntz and Stockfors, 1998).
- ²⁰ The springtime temperature dependencies were at a lower level at all sites. At Hyytiälä it was found possible to obtain more temperature responses for the parameters than at the other sites. This might be due to the more extensive time series or to some influence by the deciduous trees on the site. At Hyytiälä the model performed better with separate fits e.g. for $V_{c(max)}$ in April and May.
- The seasonality of the parameters has been observed in many deciduous species (Wilson et al., 2001; Xu and Baldocchi 2003; Kosugi et al., 2003) and also in evergreen broadleaf species (Kosugi and Matsuo, 2006). In a similar eddy flux data inversion as in this study done by Wang et al. (2006), no seasonality in the biochemical parameters of the conifer forests was observed, even though Hyytiälä was one of the sites studied.

The seasonality of the biochemical model parameters in conifer forests has been observed in some measurements (Rayment et al., 2002; Medlyn et al., 2002a; Han et al., 2004).

The relationship between the two biochemical parameters was not as represented in the literature: in our measurements J_{max} had larger values at higher temperatures. This caused the model to function poorly when using literature values for the parameters, but with our own estimates it worked as expected. However, in the literature it has also been suggested that the ratio might be influenced by growth temperatures (Kattge and Knorr, 2007) and that the ratio is affected by seasonality (Xu and Baldocchi, 2003). Often in literature the ratio is kept constant and only $V_{c(max)}$ is optimised (Wang et al., 2006; Verbeeck et al., 2006). Seasonal variation in the ratio might thus be neglected

(Verbeeck et al., 2008). Modelling the seasonality is still a challenge for vegetation models today (Sasai et al., 2007; Ricciuto et al., 2008; Harrison et al., 2008).

The parameterizations obtained were also applicable for other years at each site. ¹⁵ More importantly they were suitable for all four sites during the summertime. Inversion modelling was used in the parameterizations. Many different factors might thus influence these results, e.g. the descriptions of respiration, radiative transfer and the conductance module. These might provide reasons why our ratio between J_{max} and $V_{c(max)}$ as well as certain other of our results differed when compared to the literature. ²⁰ However, since the same method was used at all sites, the similar summertime results indicate suitability of the same biochemical parameters in these forests.

In order to get an insight into the carbon balances of forests in the future, knowledge of the respiration fluxes and their seasonal behaviour is essential (Law et al., 2002; Falge et al., 2002; Medlyn et al., 2005a). Since our modelling contained only a rough estimate of the respiration, we can only discuss the photosynthesis side in more detail.

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We divided the biochemical model with simple upscaling into small parts in order to obtain values for the model parameters. The more complicated models do not necessarily outstand the simpler ones (Lawrie and Hearne, 2007). There are e.g. models based on light use efficiency that yield very good estimations for GPP (Mäkelä et al.,

2008) and they often provide better results than process-based models (Moffat et al., 2007). These models are not however useful for scenario runs since increase of CO_2 is not included in them (Verbeeck et al., 2008). Therefore simple models using the biochemical approach are useful.

- According to the simulation runs, the effect of seasonality via the different temperature responses of the model parameters has a quite remarkable effect on GPP. Incrementing both temperature and CO₂ concentration enhances GPP. It is not a straightforward matter, however, that the model assumptions will hold in the actual conditions of the scenarios. An increase of temperature might lead to some acclimation of the biochemical model parameters (Kattge and Knorr, 2007; Way and Sage, 2008), so it
- is not obvious that we can use our parameterization as it is. An enhancement of CO_2 concentration has been reported to cause changes in the anatomy and photosynthetic capacity of the Scots pine needles (Luomala et al., 2003; Luomala et al., 2005), as well as changes in stomatal conductance in plants (Ainsworth and Rogers, 2007). However, Barnaschi et al. (2006) and Leokev et al. (2006) suggest that the effects of CO
- ¹⁵ ever, Bernacchi et al. (2006) and Leakey et al. (2006) suggest that the effects of CO₂ concentration increase have been overestimated.

5 Conclusions

We studied four different forest stands in the boreal zone and obtained estimates of the seasonality in the biochemical parameters via inverse modelling using eddy covariance data. For the biochemical model parameter J_{max} we were able to acquire spring and summertime temperature responses separately for all the sites except for Norunda, for which we got only a summertime temperature dependency. The springtime temperature responses were at a lower level and did not increase as strongly with temperature as the summertime temperature dependencies. For the biochemical model parameter $V_{c(max)}$, we were able to obtain both springtime and summertime tem-

perature responses for all the sites. At Hyytiälä and Sodankylä it was even possible to describe the spring recovery in two phases. Hyytiälä also had a separate temperature



dependency for $V_{c(max)}$ during the autumn-time. The summertime parameterizations for Hyytiälä were applicable at all sites. Our estimation method by using eddy covariance data enables using data from cold spring period, which is often left unmeasured with leaf chambers.

⁵ With the help of temperature indices, better modelling results can be attained by binding the changeover dates of the parameters' temperature fits to temperature. The temperature sum is useful in this context, as is also the five-day average temperature, on condition that it is uniquely defined.

In the future, the warming of the climate will be more pronounced in higher latitudes (Trenberth et al., 2007) and will thus affect the boreal forests, that play an essential role in the global carbon balance (Schulze, 2006). Studying the behaviour of the boreal forests facing this warming is thus of the utmost importance. Larger-scale parameterizations are needed for the models, and phenomenology is important for assessing the carbon balances of the northern areas. In our simulation runs we noticed that the relative change in GPP due to seasonality was as large as the impact of a 2°C increase in

tive change in GPP due to seasonality was as large as the impact of a 2°C increase in air temperature. The increment of GPP was also larger with increases in both temperature and CO₂ concentration together than as the sum of their influences calculated separately.

In the future climate the role of respiration will also be important, but this was not assessed in this study. An increase in the ambient CO₂ concentration might also change feedbacks from the vegetation (Janssens et al., 2005) and e.g. frost hardiness (Repo et al., 1996).

Modelling measured eddy covariance fluxes contains many sources of uncertainties (Medlyn et al., 2005b; Rannik et al., 2006) but inverse modelling of these fluxes can ²⁵ be used to obtain important results (Reichstein et al., 2003) and eddy covariance measurements can be used in model parameterization (Hollinger et al., 2004; Verbeeck et al., 2008). It was interesting to notice that in this study we can obtain seasonallyresolved temperature responses for biochemical parameters if we have an extensive enough dataset, as we had at Hyytiälä. We were also able to capture the decrease



in the photosynthetic capacity during autumn, as has been observed in e.g. Repo et al. (2006).

Our model was sensitive to changes in the leaf area index. As the LAI was changed, the radiative transfer and description of the forest stand played a significant role. It was therefore not straightforward to separate the various effects. It would be good to implement various different radiative transfer models throughout the whole model and investigate their impact on the modelling results.

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Table 1. The fitting parameters f and E for J_{max} and $V_{c(max)}$ from Eq. (3) for all the time periods in the parameterization year. Also the value of the biochemical model parameter at 17 °C shown for each fit.

Site	f	E	value at 17 °C
Parameter $J_{max}(\mu mol m^{-2} s^{-1})$			
Sodankylä			
Spring (1 May–3 June)	101.5	82 928	40.3
Summer (4 June-30 September)	142.1	75 031	61.7
Kenttärova			
Spring (May)	71.0	72 676	31.6
Summer (June–August)	101.4	61 309	51.3
Norunda			
Spring and summer	144.0	88 057	54.1
(April–October)			
Hyytiälä			
Spring (April)	6.2	9 852	5.6
Summer (May–August)	100.5	64 890	48.8
Autumn (September)	43.1	34 256	29.5
Parameter $V_{c(max)}(\mu mol m^{-2} s^{-1})$			
Sodankylä			
Spring (1 May-3 June)	13.9	75 477	6.0
June (4 June–24 June)	23.6	86 446	9.9
Summer (25 June–August)	40.3	84 152	15.8
Kenttärova			
Spring (May)	13.0	68 024	6.1
Summer (June–August)	21.5	64 911	10.5
Norunda			
Spring (1 March–19 April)	2.6	31 626	1.9
Summer (20 April–30 September)	25.4	73 591	11.2
Hyytiala			
Spring (April)	1.5	310.0	1.6
Spring II (May)	4.6	14 381	3.9
Summer II (June–August)	25.0	/ 8 648	10.4
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Table 2. The r^2 values (model vs. measurement) for the whole growing season (r_{GS}^2) and summertime with each sites' own parameterization (r_{ST}^2) and r^2 for each site using the summertime parameterization of Hyytiälä ($r_{ST,HYY}^2$). P in parenthesis after the year denotes the year used for parameterization.

Site	Year	r_{GS}^2	$r_{\rm ST}^2$	r ² _{ST,HYY}
Norunda	1999	0.63	0.62	0.62
	2001 (P)	0.59	0.58	0.57
	2002	0.58	0.57	0.55
Hyytiälä	2000	0.75	0.76	0.76
	2001 (P)	0.81	0.78	0.78
Kenttärova	2003	0.62	0.58	0.57
	2004	0.63	0.61	0.60
	2005	0.61	0.59	0.56
	2006 (P)	0.62	0.61	0.60
Sodankylä	2001 (P)	0.70	0.75	0.69
	2002	0.61	0.65	0.60
	2003	0.60	0.64	0.53
	2004	0.67	0.71	0.65
	2005	0.59	0.70	0.63

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Species	$J_{\rm max} (\mu { m mol} { m m}^{-2} { m s}^{-1})$	$V_{c(\max)} \ (\mu \text{mol m}^{-2} \text{ s}^{-1})$	Reference	Title I	Page
Scots Pine	121	46	Wullschleger (1993)	Abstract	Introduction
	95.7	44.8	Aalto et al. (2002)		
	118	60.3	Wang et al. (1996)	Conclusions	References
	314.9	121.9	Medlyn et al. (1999)		
Norway Spruce	32	12	Wullschleger (1993)	Tables	Figures
	17	6	Wullschleger (1993)		
	57.8	23.4	Grassi et al. (2001)	14	►I
	63	37	Roberntz and Stockfors (1998)		
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Table 3. Literature values for the biochemical model parameters.

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Fig. 3. The temperature responses of the biochemical model parameters J_{max} (a) and $V_{c(\text{max})}$ (b) for Scots pine dominated forest at Hyytiälä in 2001.





















Fig. 6. Various parameters for Sodankylä in 2001 as a function of day-of-the-year. (a) Inverted daily values of the biochemical parameters J_{max} (red symbols) and $V_{c(max)}$ (magenta symbols); (b) eddy covariance flux measurements; (c) five-day average air temperature (magenta line) and minimum air temperature (blue line).





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Fig. 7. Various parameters for Hyytiälä in 2001 as a function of day-of-the-year.(a) Inverted daily values of the biochemical parameters J_{max} (red symbols) and $V_{c(max)}$ (magenta symbols); (b) eddy covariance flux measurements; (c) five-day average air temperature (magenta line) and minimum air temperature (blue line).





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Fig. 9. The symbols in **(a)** denote the J_{max} values obtained by inversion using the relationship represented in Eq. (4) between J_{max} and $V_{c(\text{max})}$ at Hyytiälä. In **(b)** the symbols denote the $V_{c(\text{max})}$ values obtained by inversion at Hyytiälä. The solid line shows the fit from the previous inversion during the summer period (same as symbols in (b)), while the dashed lines show the temperature dependencies obtained from chamber measurements (Aalto et al., 2002). The fits from chamber measurements are shown over a larger temperature range that to which they were fitted.

