



1 **When trees don't act their age: size-deterministic tree-ring**  
2 **standardization for long-term trend estimation in shade-tolerant**  
3 **trees**

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15 **Abstract**

16 With increasing awareness of the consequences of climate change for global ecosystems, the  
17 focus and application of tree-ring research has shifted to reconstruction of long-term climate-  
18 related trends in tree growth. Contemporary methods for removing the biological growth-trend  
19 from tree-ring series (standardization) are ill-adapted to shade-tolerant species, leading to biases  
20 in the resultant chronology. Further, many methods, including regional curve standardization  
21 (RCS), encounter significant limitations for species in which accurate age estimation is difficult.  
22 In this study we present and test two tree-ring standardization models that integrate tree size in  
23 the year of ring formation into the estimation of the biological growth-trend. The first method,  
24 dubbed size deterministic standardization (SDS), uses tree diameter as the sole predictor of the  
25 growth-trend. The second method includes the combined (COMB) effects of age and diameter.  
26 We show that both the SDS and COMB methods reproduce long-term trends in simulated tree-  
27 ring data better than conventional methods – this result is consistent across multiple species.  
28 Further, when applied to real tree-ring data, the COMB method is more parsimonious than its  
29 than RCS. We recommend the inclusion of tree size in the year of ring formation in future tree-  
30 ring standardization models, particularly when dealing with shade-tolerant species, as it does not  
31 compromise model parsimony and allows for the inclusion of unaged trees.

32 **1 Introduction**

33 Tree-rings have long-served as a record of environmental change in forest ecosystems. Early  
34 dendrochronological studies used tree-ring chronologies from climate sensitive species to  
35 elucidate the dynamics of growth-climate relationships and reconstruct climate anomalies from  
36 periods before the existence of instrumental records. However, with increasing awareness of the  
37 consequences of climate change for global ecosystems, the focus and application of tree-ring



38 research has shifted to reconstruction of low-frequency climate related trends in tree growth  
39 (Gedalof and Berg 2010, Boisvenue and Running 2006, Jacoby and D'Arrigo 1997). As it stands,  
40 previous optimism regarding the benefits of carbon fertilization for forest growth (Battipaglia et  
41 al. 2012, Norby et al. 2005) has been quelled by a lack of consistent evidence in real forests.  
42 While many studies have noted increases in long-term growth rates over time in temperate  
43 forests (Gedalof and Berg 2010, Huang et al. 2007, Martinelli 2004) others suggest no change  
44 (Giguère-Croteau et al. 2019, Camarero et al. 2015, Granda et al. 2014, Silva et al. 2010,  
45 Peñuelas et al. 2011). Further, in boreal and drought prone species, growth decline (Chen et al.  
46 2017, Dietrich et al. 2016, Girardin et al. 2012, Silva and Anand 2013) and increased mortality  
47 (Herguido et al. 2016, Liang et al. 2016) in response to climate stress have been prevalent.  
48 Central to all these studies is the assumption that long-term growth-trends can be accurately and  
49 unbiasedly estimated from tree-ring data.

50 As it stands, accurate estimation of long-term growth-trends in forests may be limited by poorly  
51 adapted tree-ring standardization (age-trend removal) methods (Briffa et al. 1996) and  
52 inappropriate sampling methods (Nehrbass-Ahles et al. 2014, Brien et al. 2012). Early  
53 standardization methods (i.e. conservative detrending) were designed to maintain high-frequency  
54 variation in tree-ring series and discard long-term, low-frequency variation. It is accepted that  
55 these methods are inappropriate for estimating long-term climate related growth-trends (Briffa  
56 1992); however, they are still used in situations where contemporary standardization methods are  
57 not applicable due to restrictive data requirements (e.g. Villalba et al. 2012, Gedalof and Berg  
58 2010, Geoff Wang et al. 2006). More recently, the use of regional curve standardization (RCS),  
59 and its many variants, as well as the conversion of tree-ring widths to basal area increments  
60 (BAI) have become commonplace (Peters et al. 2015). But, due to the difficulties in separating

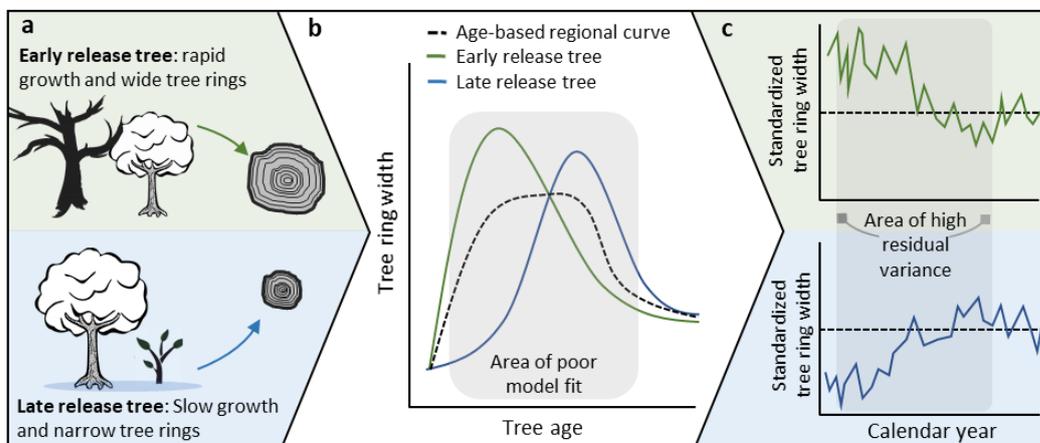


61 climate related trends that vary on long time scales from those related to biological tree growth  
62 and/or succession-related environmental change, neither of these methods are likely to produce  
63 accurate estimates of external forcing when trees from only a single age/size class are sampled  
64 (Brienen et al. 2012, Briffa and Melvin 2011). While increased awareness of sample biases has  
65 led to better prescriptions for study design (see Nehrbass-Ahles et al. 2014, Brienen et al. 2012),  
66 systematic tests of the ability of these models to accurately reproduce long-term trends are  
67 limited (e.g. Sullivan et al. 2016, Peters et al. 2015, Esper 2010).

68 RCS remains the standard method for estimating long-term growth-trends in tree-ring data  
69 (Helama et al. 2017). However, the standard RCS approach encounters large limitations for  
70 many species in which accurate age estimation is difficult. Additionally, we suggest the inherent  
71 assumption of RCS that biological growth-trends are sufficiently determined by tree age may not  
72 be appropriate in all species. More specifically, this assumption is problematic for shade-tolerant  
73 trees. Shade-tolerant species exhibit relatively low low-light mortality and thus can persist in  
74 forest understories for variable amounts of time before release from overstory light suppression.  
75 In these cases, traditional age-deterministic models exhibit high variance, and thus low precision,  
76 in the period following tree establishment and leading up to the age when most trees have been  
77 released from suppression (Fig. 1). This period of ill-fit means that trees which are released  
78 relatively early (or late) from light suppression will exhibit inflated (or deflated) growth relative  
79 to the chronology. As a result, the final chronology will show less agreement than would be  
80 expected in a shade-intolerant species. Even more problematic, if trees are sampled according to  
81 minimum size thresholds, the youngest trees in the chronology are likely to be early-release trees  
82 leading to an artificial inflation of modern growth rates in the final chronology. While  
83 modifications to traditional RCS that address variance in contemporaneous growth rates and



84 regional environmental conditions have been prevalent in shade-intolerant species (see Helama  
85 et al. 2017) there has been little to no focus on the improvement of standardization techniques  
86 specific to shade-tolerant tree species.



87 **Figure 1: (a) In shade-tolerant species young trees are stochastically released from low-light suppression in**  
88 **the understory. (b) Since release from suppression is not strictly related to tree age, widely used communal**  
89 **age-trend models (RCS) poorly model tree growth in the period following establishment and leading up to the**  
90 **age when most trees have been released from suppression. (c) Poor model-fit in this period implies that the**  
91 **biological growth-trend is not entirely removed from individual series and leads to high residual variance**  
92 **when standardized tree-ring series are aligned according to calendar year.**

93 Alternatively, in the field of forest growth and yield modelling, size, rather than age,  
94 deterministic predictive growth models are ubiquitous. It is well understood that tree size  
95 regulates the capacity for resource acquisition, namely, light (Canham et al. 2004), water and  
96 nutrients (Homann et al. 2000), resource allocation (Lehnebach et al. 2018) and metabolic costs  
97 (West et al. 2001). As such, the notion of radial growth being deterministic according to size  
98 rather than age is logical from both a physiological and ecological perspective. We propose that a  
99 size-deterministic model for tree-ring standardization may be more appropriate than traditional  
100 RCS for shade-tolerant tree species. The application of size-deterministic models has been  
101 limited, with few examples of tree size in a given year being incorporated into BAI models (e.g.  
102 Marqués et al. 2016, Camarero et al. 2015, Nock et al. 2011, Martínez-Vilalta et al. 2008) and  
103 even fewer of uniquely size-based tree-ring models (e.g. Bontemps and Esper 2011). Further,



104 there have been no systematic evaluations of the ability of size-based models to accurately  
105 estimate long-term trends in tree-ring series.

106 We present two tree-ring standardization models that integrate tree size in the year of ring  
107 formation into estimation of the biological growth-trend. The first model uses tree diameter as  
108 the sole predictor of the communal growth-trend while the second includes the combined effects  
109 of both age and diameter. It follows that the objective of this study is to determine the efficacy of  
110 both models in estimating long-term growth-trends in their resultant tree-ring chronologies. First,  
111 we use modelled tree-ring data from shade-tolerant and intolerant species to make explicit the  
112 inappropriateness of age-based models for shade-tolerant trees. Further, we investigate the  
113 performance of size-based models relative to contemporary standardization methods in the  
114 presence of size thresholds in tree sampling. Last, we apply the developed models to tree-ring  
115 data from shade-tolerant temperate species to evaluate model performance relative to  
116 contemporary methods on the basis of model-fit and chronology quality statistics.

## 117 **2 Methods**

### 118 **2.1 Model formulation**

119 Traditional RCS makes two assumptions about tree growth. First that trees of the same species in  
120 a given region exhibit a common growth-trend as they age, and second, that growth of an  
121 individual tree in a given year is thus a product of its age and common climatic or environmental  
122 forcing in that year (Esper et al. 2003, Briffa et al. 1992). We present a variant of the RCS  
123 method that uses tree size, measured by diameter at breast height (DBH), in the year of ring  
124 formation as the primary determinant of the common biological growth-trend. As with RCS we  
125 assume that the relationship between expected growth and tree size is non-linear and can be  
126 approximated for a region from a sufficiently large sample of trees from the species in question.



127 Further, we assume that using a sample of trees from a range of size/age classes ensures  
128 estimation of the common trend is not confounded by underlying low-frequency climate or  
129 environmental forcing in the chronology (Brienen et al. 2012). The size-based regional curve  
130 model, hereafter referred to as the **size deterministic standardization (SDS)** model, takes the  
131 following form:

$$132 \quad (1) \ E(RW_{y,i}) = B_o + f_l(DBH_{y,i}) + e_{yi}$$

133 Where  $E(RW_{y,i})$  represents the expected ring width of a given tree ( $i$ ) in year ( $y$ ), and  $f_l$   
134 represents a non-linear function relating DBH of a given tree ( $i$ ) in year ( $y$ ) to  $E(RW_{y,i})$ . The non-  
135 linear relationship is estimated using penalized spline fitting techniques in a generalized additive  
136 model (GAM). Under this paradigm the model residuals ( $e_{yi}$ ) represent individual standardized  
137 ring width indices and, by extension, individual tree response to climatic or environmental  
138 forcing. Annual model residuals subject to a robust mean, thus, represent the final standardized  
139 chronology. This approach differs slightly from traditional RCS, whereby occasionally  
140 standardized ring width indices are produced by division of raw data by the expected value.  
141 Calculation of standardized ring width indices by subtraction from the expected value, as in the  
142 case of residuals, is now commonly used as it tends to reduce bias in the resultant chronology  
143 (Helama et al. 2004) and eases in the formulation of more complex tree-ring standardization  
144 models. However, unlike division methods, the subtraction method does not provide any  
145 stabilization of variance in the resulting residuals; as such, it may be necessary to use a  
146 stabilization procedure (i.e. log transformation, power transformation) on raw ring width data  
147 beforehand.

148 Tree size in a given year can be estimated by outside-in or inside-out techniques. If the pith of a  
149 tree is present in the core (or reasonably close to)  $DBH_y$  is a simple summation of all previous



150 ring widths since the year of origin, multiplied by two. Alternatively, if the pith is missed, DBH<sub>y</sub>  
151 can be calculated via subtraction of more modern ring widths (multiplied by two) from the  
152 inside-bark diameter. In this case inside-bark diameter is calculated as the measured DBH minus  
153 bark thickness (multiplied by two), where bark thickness can be directly measured or estimated  
154 using species-specific allometric equations (e.g. Stayton and Hoffman 1970).

155 Similar to the model formulation for SDS, RCS models were estimated with GAMs of the  
156 following form:

$$157 \quad (2) \ E(RW_{yi}) = B_o + f_1(\text{Age}_{yi}) + e_{yi}$$

158 Where  $\text{Age}_{yi}$  is the age of an individual tree in a given year and the resultant standardized tree-  
159 ring indices are derived from model residuals ( $e_{yi}$ ).

160 In addition, a more complex model that integrated independent size and age effects was also  
161 evaluated for comparison. This model, hereafter referred to as the **combined model (COMB)**,  
162 took the following form:

$$163 \quad (3) \ E(RW_{yi}) = B_o + f_1(\text{Age}_{yia}) + f_2(\text{DBH}_{yi}) + e_{iy}$$

164 In a large variety of long-lived tree species, accurate age estimation (pith sampling) is difficult or  
165 impossible; rendering traditional RCS or combined models inappropriate for all trees sampled.

166 To address this issue, the above model incorporates unaged trees. Here  $f_1$  represents the non-  
167 linear function relating age to expected ring width for the subset of all trees that are aged (ia). In  
168 this model, ring widths from unaged trees are assigned arbitrary ages which do not contribute to  
169 the linear approximation of the smooth term for Age (i.e.  $f_1(\text{Age}_{yia})$ ) but these trees still contribute  
170 to the smooth term for size  $f_2(\text{DBH}_{yi})$ . Syntax for missing data in GAMs followed the protocol



171 provided in mgcv (Wood 2011). In this study all GAMs were fit using the mgcv package (Wood  
172 2011) in the R statistical program (v.3.5.0).

173 In addition to the models presented above we investigated two more contemporary  
174 standardization methods; conservative detrending (CD) and BAI. Conservative detrending  
175 describes functions (i.e. negative exponentials, straight lines) or flexible splines fit to individual  
176 tree-ring series. In this study we use spline fitting techniques rather than modified negative  
177 exponentials as they are more appropriate for shade-tolerant tree species. Alternatively, BAI  
178 attempts to remove biological growth-trends by converting ring widths from individual trees to  
179 estimates of annual basal area growth. For simplicity, untransformed BAI was used to compile  
180 chronologies for this study. Both CD and BAI methods were applied using the dplR package  
181 (Bunn 2008) in R.

## 182 **2.2 Simulated tree-ring data**

183 To evaluate the efficacy of each standardization method in detecting long-term trends, we  
184 simulated tree-ring data using a well-established gap-phase model. The SORTIE-ND model was  
185 chosen over other similar gap-phase models as it better emulates understory light conditions and  
186 low-light mortality both of which are central to the notion of age being an inappropriate  
187 determinant of growth in shade-tolerant species. For simplicity, a 100% sugar maple (*Acer*  
188 *saccharum*) dominated stand was simulated as sugar maple is a model shade-tolerant species that  
189 grows in self-replacing stands. All living trees (>5 cm dbh), (n=3657) in the final year of the  
190 model run were used for further analysis. Additionally, to elucidate our claim that age-  
191 deterministic growth estimation is more problematic in shade-tolerant species, we completed a  
192 similar SORTIE simulation for the shade-intolerant species white pine (*Pinus strobus*). Again,  
193 the stand was 100% white pine, standard model parameters were used, and the simulation was



194 run for 1000 years. All living trees (>5 cm dbh), (n=7362) in the final year of the model run were  
195 used for further analysis. Additional details regarding model parameters for the SORTIE  
196 simulations are provided in the supplementary materials (Suppl. S1).

197 To simulate a low-frequency climate related growth-trend, a logistic trend was added to raw tree-  
198 ring width of individual trees produced by both SORTIE simulations. The logistic trend  
199 simulated an initial rapid increase in growth and subsequent levelling off that aimed to represent  
200 a period of carbon fertilization and eventual acclimation. The logistic model was applied to the  
201 last 100 years of growth and took the following form, where  $RW_t$  represents ring widths with the  
202 simulated long-term trend and  $RW_r$  are raw ring widths:

$$203 \quad (4) \quad RW_{t_{yi}} = RW_{r_{yi}} + \frac{0.6 * \overline{RW}_i}{1 + ae^{-r * y}}$$

204 The logistic trend parameters ( $r$ ,  $a$ ) were chosen such that increases in growth did not exceed 5%  
205 of individual average tree growth per decade.

206 Sixty trees were randomly selected, without replacement, from the simulated tree populations  
207 and subject to each of the five standardization methods (SDS, RCS, COMB, CD, BAI). Model  
208 residuals (in the case of RCS, SDS and COMB), and standardized (CD) or transformed (BAI)  
209 tree-ring widths were compiled into an annual mean chronology using Tukey's biweight robust  
210 mean. The resultant chronologies were then tested for significant correlation with the logistic  
211 growth-trend using Spearman's rank correlation coefficient. This process was bootstrap  
212 resampled 100 times to produce confidence intervals for correlation coefficients.

213 To examine the effect of minimum size sampling thresholds on the accuracy of long-term trend  
214 reconstruction by each of the standardization methods, we completed the same analysis on trees  
215 from the simulated populations that exceeded certain size thresholds. The thresholds employed



216 were 10 cm DBH, which represented a practical minimum size threshold for sampling, and 30  
217 and 50 cm DBH which represented thresholds for mature and dominant trees, respectively. The  
218 mean Spearman's rho for all detrending methods and sampling thresholds were compared using  
219 two-way ANOVA and post-hoc tests. Further, two-way ANOVA compared the effect of model  
220 choice on Spearman's rho between species (sugar maple and white pine).

### 221 **2.3 Real tree-ring data**

222 We evaluated the appropriateness of the SDS, COMB and RCS models for use in real tree-ring  
223 data from shade-tolerant species. We collected tree-ring data from seven mature sugar maple  
224 dominated stands in Ontario, Canada (Table 1). Further, tree-ring data sets from the shade-  
225 tolerant species red spruce (*Picea rubens*) were obtained from the DendroEcological Network  
226 database (<https://www.uvm.edu/femc/dendro>), (Table 1). Red spruce was chosen as it had  
227 sufficient replication across studies in the database. Descriptions of the sampling strategies and  
228 data processing methods for all sites considered are provided in either the supplementary  
229 materials (Suppl. S2) or in their respective references (i.e. Kosiba 2013, Kosiba 2017). Data was  
230 considered suitable for this study if age and DBH estimates were provided and if a minimum 20  
231 trees per site and species were sampled. All cores in which pith offset was estimated to be greater  
232 than 10 years were considered unaged. The SDS, RCS, and COMB models were fit to tree-ring  
233 data from all site-species combinations and the resultant chronologies were compiled with a  
234 robust mean. In all cases models were fit to log-transformed ring widths, as it increased residual



235 homoskedasticity. For simplicity and ease of model comparison we did not fit CD or BAI models  
 236 to the real tree-ring data set.

237 Model fits from the SDS, RCS and COMB methods were compared according to Akaike  
 238 information criterion (AIC) and percent variance explained ( $R^2$ ). Since model comparison via  
 239 AIC requires equal sample sizes, reduced data SDS ( $SDS_{red}$ ) and COMB ( $COMB_{red}$ ) models,  
 240 which only included aged trees, were also fit. These reduced data models have no practical  
 241 application but allow for direct AIC comparison between the RCS,  $COMB_{red}$  and  $SDS_{red}$  models.

**Table 1:**

Location, sample size, chronology length and source of tree ring data sets used in this study.

Species	Site (code)	Longitude (°)	Latitude (°)	N. trees total	N. trees aged	Length of chronology	Source
Sugar maple ( <i>A. saccharum</i> )	Toobee Lake (TB)	46.7459	-82.8668	79	67	1750-2015	This study
	Wolf Mtn. (WM)	46.7390	-82.8467	22	18	1827-2015	...
	Roosevelt Road (RS)	47.2852	-79.7063	20	11	1792-2016	...
	Raven Lake (RL)	45.3309	-78.6339	31	19	1864-2015	...
	Freezy Lake (FR)	45.2998	-78.4329	20	11	1887-2015	...
	Kakakise Lake (KK)	46.0554	-81.3317	22	7	1773-2016	...
	Mt. Zion Road (MT)	46.4000	-83.7004	29	15	1777-2015	...
Red spruce ( <i>P. rubens</i> )	Mt. Mansfield (MTM)	44.3750	-73.8750	111	109	1769-2011	Kosiba et al. (2016)
	Burnt Mtn. (BNT)	44.2068	-72.3515	40	40	1891-2010	Kosiba et al. (2013)
	Mt. Carmel (CAR)	43.7709	-72.9205	41	41	1795-2010	...
	Mt. Ellen (ELL)	44.1656	-72.9221	42	42	1824-2010	...
	Mt. Equinox (EQU)	43.1487	-73.1273	89	89	1857-2010	...
	Mt. Greylock (GRY)	42.6738	-73.1575	44	44	1911-2010	...
	Hubbard Brook (HUB)	43.9577	-71.7350	89	89	1885-2010	...
	Killington Mtn. (KIL)	43.6146	-72.8088	104	103	1742-2010	...
	Mt. Mansfield (MAN)	44.5106	-72.8297	57	57	1767-2010	...
	Mt. Moosilauke (MOO)	44.0056	-71.8215	54	54	1760-2010	...
Mad River Glen (MRG)	44.1932	-72.9232	36	36	1927-2010	...	

242 Further, we calculated chronology quality statistics including: mean interseries-correlation,  
 243 expressed population signal (EPS) and signal-to-noise ratio (SNR), for all chronologies.



244 Differences between model fit statistics and quality indices among models were tested using a  
245 linear mixed-effect modelling (LME) approach whereby, model error was specified according to  
246 site. This approach is analogous to traditional repeated-measures ANOVA but allows for contrast  
247 analysis between models.

### 248 **3 Results**

#### 249 **3.1 Comparisons of methods in simulated data**

250 Bootstrapped correlations for chronologies produced by each standardization method are  
251 displayed in Figure 2a for the sugar maple and Figure 2b for white pine.

252 The sugar maple two-way ANOVA suggested a significant effect of both standardization model  
253 ( $p < 0.001$ ) and minimum size sampling threshold ( $p < 0.001$ ) on average correlation. Across all  
254 sampling thresholds the SDS model produced chronologies with the highest mean correlation  
255 ( $\bar{r}_s = 0.972 \pm 0.024$ ). Tukey HSD contrasts suggested mean correlation for the SDS model was  
256 significantly higher than all other models ( $p < 0.001$ , all cases) except for the COMB model  
257 ( $\bar{r}_s = 0.969 \pm 0.023$ ), ( $p = 0.993$ ) which produced the second highest mean correlation. BAI produced  
258 the third highest mean correlation ( $\bar{r}_s = 0.954 \pm 0.103$ ) which was not significantly different from  
259 the combined model ( $p = 0.61$ ) but was significantly higher than correlations from the two  
260 remaining models (RCS and CD), ( $p < 0.001$ ). The CD ( $\bar{r}_s = 0.720 \pm 0.153$ ) and RCS  
261 ( $\bar{r}_s = 0.925 \pm 0.054$ ) models produced the lowest and second lowest correlations respectively; both  
262 were significantly different from each other and all other models ( $p < 0.01$ , all cases).

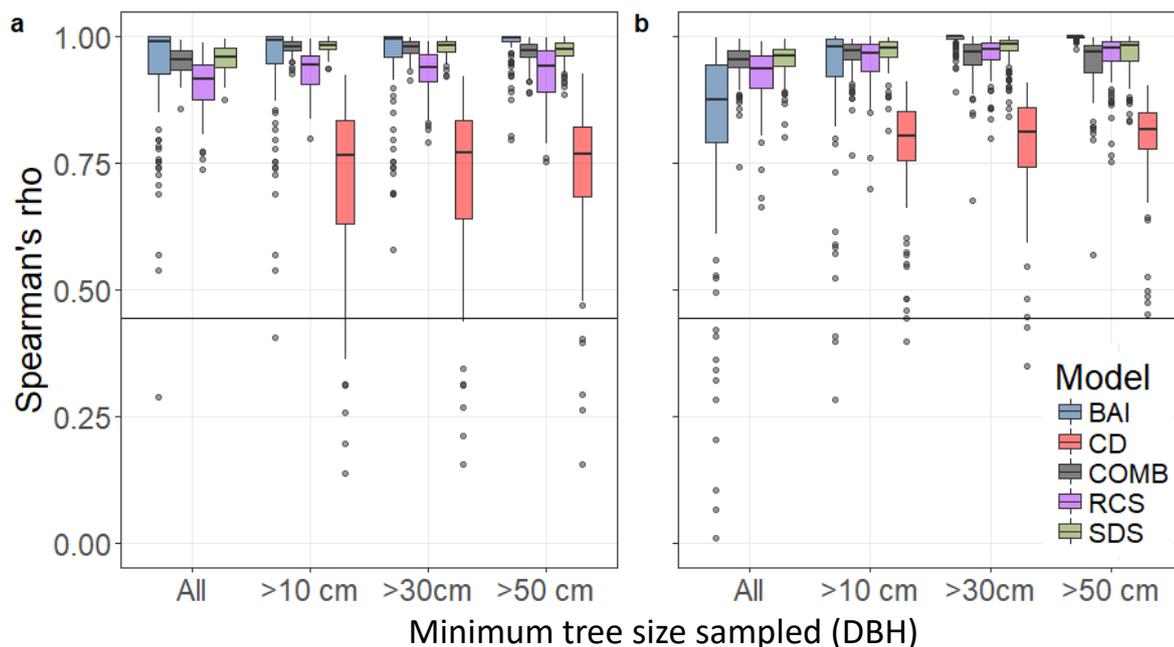
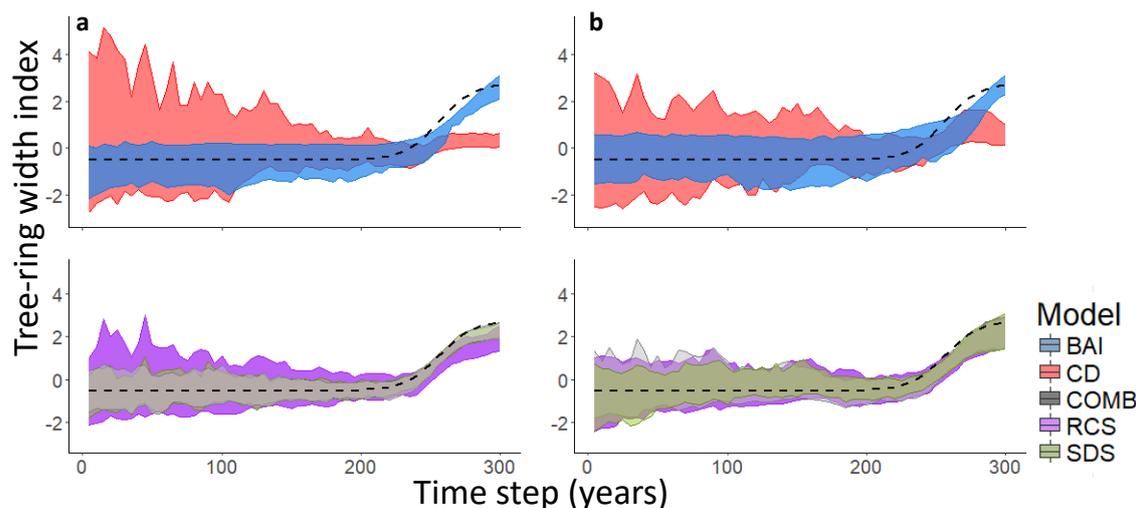


Figure 2: Spearman's rank correlation between chronologies produced by each of the five standardization methods and the imposed logistical trend in simulated (a) sugar maple and (b) white pine tree-ring data. Horizontal axis denotes minimum tree size (DBH) thresholds for sampling from the population.

263 Contrasts suggested that the significant effect of minimum size threshold was driven by  
264 significant differences in correlations between BAI produced chronologies among samples. As  
265 evident in Figure 2a, BAI chronologies performed best when size thresholds exceeded 50 cm  
266 DBH. At this threshold BAI chronologies produced significantly higher correlations than when  
267 all trees were sampled ( $p=0.002$ ) and marginally significantly higher correlations than when trees  
268  $>10$  cm DBH were sampled ( $p=0.054$ ). Contrastingly, the SDS, RCS, COMB and CD  
269 chronologies produced similar correlations across all minimum size thresholds. Bootstrapped  
270 sugar maple chronologies produced by each of the standardization methods are provided in  
271 Figure 3a.



**Figure 3: 95% confidence intervals for chronologies produced by each standardization method in SORTIE simulated (a) sugar maple and (b) white pine tree-ring data (Top: BAI, CD; Bottom: SDS, RCS, COMB). Confidence intervals were obtained via bootstrap resampling (rep=100) of 60 trees (>10 cm DBH) from the SORTIE simulated populations. Dotted lines indicate the logistic trend that was added to the raw tree-ring data. For ease of comparison all chronologies and the simulated trend were centred and scaled before plotting.**

272 The white pine two-way ANOVA also suggested a significant effect of both standardization  
273 model ( $p < 0.001$ ) and minimum size sampling threshold ( $p < 0.001$ ) on average correlation. Across  
274 all size thresholds the SDS ( $\bar{r}_S = 0.965 \pm 0.035$ ), COMB ( $\bar{r}_S = 0.954 \pm 0.037$ ), and RCS  
275 ( $\bar{r}_S = 0.949 \pm 0.049$ ) models produced the highest correlations respectively. Among these, none  
276 were significantly different from each other. Again, chronologies produced by CD produced the  
277 lowest correlations of all models ( $\bar{r}_S = 0.785 \pm 0.074$ ), ( $p < 0.001$ , all cases). BAI produced  
278 chronologies ( $\bar{r}_S = 0.922 \pm 0.177$ ) performed significantly worse than SDS, COMB and RCS  
279 chronologies and significantly better than CD ( $p < 0.001$ , all cases). Tukey HSD contrasts  
280 suggested that the significant effect of minimum size threshold was again driven by significant  
281 differences in correlations between BAI produced chronologies among samples. As evident in  
282 Figure 2b, BAI chronologies performed significantly better when sampling size thresholds

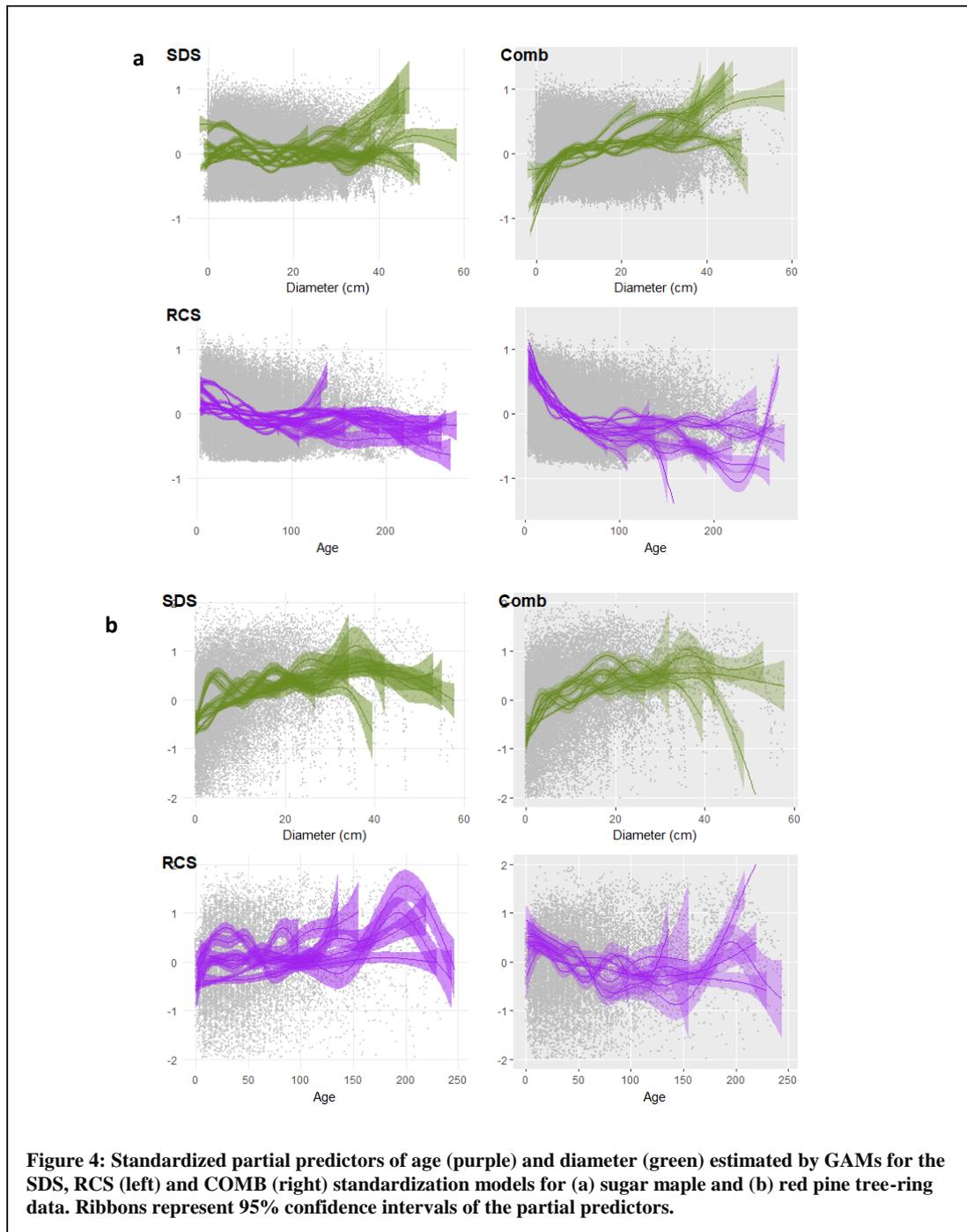


283 exceed 30 cm DBH ( $p < 0.001$ , all cases). However, for BAI chronologies increased size  
284 thresholds of 50 cm DBH did not produce significantly higher correlations relative to the 30 cm  
285 DBH threshold ( $p = 1.000$ ). Bootstrapped white pine chronologies produced by each of the  
286 standardization methods are provided in Figure 3b.

287 When comparing mean correlations produced by each method between sugar maple and white  
288 pine, two-way ANOVA suggested a significant interaction between species and method  
289 ( $p < 0.001$ ). Tukey HSD contrasts suggested this effect is driven by significant differences in  
290 correlations produced by RCS, BAI and CD between species. More specifically, across all size  
291 thresholds RCS and CD produced significantly higher correlations in white pine relative to sugar  
292 maple ( $p = 0.011$ ,  $p < 0.001$ , respectively). Whereas, BAI produced significantly higher  
293 correlations in sugar maple ( $p < 0.001$ ).

### 294 **3.2 Comparisons of methods in real tree-ring data**

295 The standardized partial predictors estimated for each of the SDS, RCS and COMB models fit to  
296 the sugar maple and red spruce data are provided in Figure 4. In general, SDS and RCS models  
297 suggested rather flat relationships between tree age and size with in average growth (log-  
298 transformed) in both species. However, when the effects of age and tree size were considered  
299 together in COMB models both showed more distinct relationships with average growth. In the  
300 COMB models average tree growth appeared to increase rapidly with tree size and eventually  
301 decline as trees exceed 50 cm DBH in both species. Average growth appeared to decline with  
302 increasing age before leveling off around 100 years. In both cases the COMB models showed  
303 better agreement among sites in small/young trees than the simpler models.



304

305 In sugar maple, Tukey HSD contrasts suggested no significant differences in the meaningful AIC  
306 comparisons after controlling for site differences (Fig. 5a). However,  $R^2$  produced by the COMB

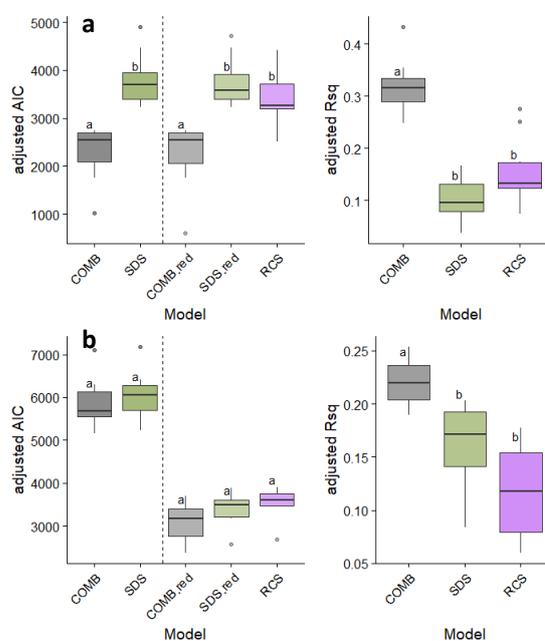


307 models were significantly higher than those produced by both the SDS ( $p=0.039$ ) and RCS  
308 ( $p<0.001$ ) models, but SDS and COMB were not significantly different from each other  
309 ( $p=0.177$ ), (Fig. 5a). LMEs did not suggest a significant effect of model choice on any of the  
310 chronology quality statistics (mean interseries-correlation, SNR, EPS).

311 In red spruce, the COMB model fits exhibited  
312 significantly lower AIC values than the SDS model  
313 ( $p<0.001$ ) after controlling for site differences (Fig.  
314 5b). Similarly, the COMB<sub>red</sub> model fits exhibited  
315 significantly lower AIC values than the SDS<sub>red</sub>  
316 ( $p<0.001$ ) and RCS ( $p<0.001$ ) models, while the  
317 SDS<sub>red</sub> and RCS were not significantly different  
318 from each other ( $p=0.706$ ), (Fig. 5a). Similarly,  $R^2$   
319 produced by the COMB models were significantly  
320 higher than those produced by both SDS ( $p<0.001$ )  
321 and RCS ( $p<0.001$ ), which were not different from  
322 each other ( $p=0.114$ ), (Fig. 5a). As with sugar  
323 maple LME did not suggest a significant effect of  
324 model choice on any of the chronology quality  
325 statistics.

## 326 4 Discussion

### 327 4.1 Size-vs age-deterministic models for long-term trend reconstruction



**Figure 5: Adjusted AIC and  $R^2$  values for fits of standardization models applied to (a) 7 sugar maple and (b) 11 red spruce tree-ring chronologies. Site-specific error was removed via a mixed modelling approach. Significant differences ( $\alpha=0.05$ ) between models are indicated with small letters (a, b). Dashed vertical lines signify meaningful pairwise AIC comparisons (i.e. COMB vs. SDS).**



328 Using simulated tree-ring data, from the shade-tolerant species sugar maple, we have shown that  
329 standardization models which include tree size in the year of ring formation (SDS, COMB)  
330 produced chronologies that retain long-term/low-frequency variation better than those produced  
331 by models that only include age as a predictor (RCS). Alternatively, in the shade-intolerant  
332 species white pine, chronologies produced by the SDS, RCS and COMB models showed no  
333 significant difference in their estimation of long-term trends. Further, our analysis suggests that  
334 the traditional RCS method performed significantly worse in the shade-tolerant species, sugar  
335 maple, than in shade-intolerant white pine.

336 The finding that size-based standardization models performed well in simulated tree-ring data is  
337 not surprising given that the SORTIE model calculates annual tree growth as function of tree  
338 size. Thus, the underlying growth-trend would be well-approximated by a flexible curve  
339 estimated on the basis of tree size. As such, we use these results solely to elucidate the  
340 problematic nature of age-based standardization methods for shade-tolerant species. SORTIE's  
341 use of diameter, rather than age, as a determinant of tree growth is not arbitrary; it is well  
342 established that tree metabolic processes are directly related to size (West et al. 2001).  
343 Additionally, there is little evidence for a unique effect of age on tree growth that is independent  
344 of size (Munné-Bosch 2007 (and within)). With the exception of dendrochronological models,  
345 the vast majority of individual tree growth and process models are indeed size-based. It follows  
346 that the ubiquitous use of age or calendar year in tree-ring standardization methods (RCS, signal-  
347 free standardization, C-method, CD, Hugesshoff curves) is a practice born out of convenience  
348 rather than physiological consideration. As such, we agree with previous accounts that this  
349 assumption may be especially problematic in shade-tolerant trees where age and size may not be  
350 perfectly correlated (Peters et al. 2015, Bontemps and Esper 2011).



351 Unfortunately, all systematic comparisons of tree-ring standardization methods in real tree-ring  
352 data (e.g. Sullivan et al. 2016) are limited by their inability to validate long-term trends estimated  
353 by chronologies. Instead, we evaluate standardization models on the basis of model parsimony.  
354 We have shown that in the shade-tolerant species red spruce, COMB models are significantly  
355 more parsimonious (estimated by AIC) than simpler models (RCS, SDS). Further, the COMB  
356 models explain more variance (estimated by  $R^2$ ) in tree-ring data regardless of differences in  
357 underlying sample sizes. Overall, our results are conservative relative to similar comparisons  
358 performed by Nock et al. (2011) in tropical tree species of varying shade-tolerance. Nock et al.  
359 (2011) note that LMEs of BAI that included tree diameter had more support than those that  
360 included age. In line with discussion above, Nock et al. (2011) attribute this finding to size being  
361 a more important determinant of light capture as it relates to tree height and crown size (King et  
362 al. 2005). Further, in both red spruce and sugar maple we have shown that tree size and age  
363 exhibit stronger relationships with average growth when their unique effects are estimated  
364 simultaneously in COMB models rather than alone in SDS and RCS models, respectively. This  
365 result is interesting given the high correlation expected between these variables and it may  
366 explain why COMB models explained significantly more variance than each of the simpler  
367 models. Given the relatively weak trends shown in predictors from both the SDS and RCS,  
368 models we suggest that low-frequency variance related to the underlying biological growth-trend  
369 may be retained in these chronologies.

370 Regardless of differences in model fits, the implications for the resultant chronologies remain  
371 conservative (Fig. S.2). Similarly, in comparison of RCS and SDS chronologies in common  
372 beech (*Fagus sylvatica* L.) Bontemps and Esper (2011) note both chronologies exhibit similar  
373 annual variations. The resultant chronology is more likely to be influenced by sample size of the



374 underlying tree population than by choice of standardization model. Tree age can be difficult or  
375 impossible to accurately estimate for some trees. In contrast, annual tree size can be reliability  
376 estimated from DBH and tree-ring measurements more ubiquitously. We note that in this study  
377 only 66% of sugar maple trees could be accurately aged. Since unaged trees are likely to be the  
378 oldest trees in the chronology, it follows that RCS chronologies may exhibit poor sample  
379 replication (especially in early years) and may be significantly shorter than those produced by  
380 SDS or COMB models. This has obvious implications for data quality and suitability.  
381 Considerably problematic is the “segment length curse” whereby, almost all standardization  
382 methods are ill-equipped to estimate long-term trends on time scales greater than or equal to the  
383 length of the chronology itself (Cook et al. 2005). Excessively short RCS chronologies are  
384 therefore limited in their application. A large advantage of SDS and COMB models is that they  
385 can incorporate otherwise inadmissible tree-ring data.

#### 386 **4.2 BAI and CD methods for long-term trend reconstruction**

387 The finding that CD did not produce accurate long-term trends in simulated tree-ring data is  
388 consistent with our expectations (Peters et al. 2015, Briffa et al. 1992). We maintain CD should  
389 be avoided if the goal is long-term reconstruction from tree-ring data.

390 BAI chronologies accurately reproduced long-term trends in simulated tree-ring data. However,  
391 our analysis suggests BAI is less reliable when small/young trees are sampled. This was in line  
392 with Peters et al. (2015) who note high accuracy and sensitivity of BAI chronologies to imposed  
393 long-term trends, but that BAI is likely to produce erroneous trends when the underlying trend is  
394 of low signal, as would be the case for young/small trees that have low BAI rates and low  
395 climate sensitivity. As presented here, the BAI method imparts a strict relationship between tree  
396 size and growth. It has been suggested that this relationship may not account for the entire



397 biological growth-trend (Peters et al. 2015). Accordingly, we caution future studies in their  
398 interpretation of BAI trends in low-signal tree-ring series. Alternatively, other studies have  
399 explicitly modelled size and/or age effects on BAI using a mixed-effect modelling approach (e.g.  
400 Marqués et al. 2016, Camarero et al. 2015, Nock et al. 2011, Martínez-Vilalta et al. 2008). We  
401 suggest this approach may better account for species- and site-specific factors that influence  
402 expected growth rates, leading to more accurate estimates of long-term trends in the resultant  
403 chronology. While our findings regarding the importance of inclusion of size in tree-ring  
404 standardization models are presented in the context of raw tree-ring width models, they are also  
405 directly relevant to explicit models of BAI.

#### 406 **4.3 Other considerations and future research**

407 It is important to note that the goal of this study was not to explicitly test the effect of sample  
408 biases (i.e. modern sample bias, selection bias, etc.) on trend reconstruction, but instead to assess  
409 reliability across different underlying sampling distributions. There is now considerable evidence  
410 to suggest that the long-standing practice of sampling only dominant trees or trees exceeding a  
411 minimum size threshold within a stand leads to considerable bias in the resultant chronology  
412 (Nehrbass-Ahles et al. 2014, Brienen et al. 2012, Briffa and Melvin 2011). This bias is consistent  
413 across standardization methods (Nehrbass-Ahles et al. 2014). We maintain that in cases of long-  
414 term trend reconstruction, stands should be sampled according to the underlying stand age/size  
415 distribution, either through use of fixed-plots or random tree selection, regardless of the  
416 standardization procedure used.

417 Our study has suggested that the choice of standardization model (SDS, RCS, COMB) has no  
418 discernable effect on indices of chronology quality (EPS, SNR, interseries-correlation). We  
419 suggest this finding is a result of the chosen species exhibiting low climate sensitivity (Phipps



420 1982) and thus, low common signal in the chronology. As such we do not regard this finding as  
421 failure of any of the standardization models. We suspect more conclusive results would be found  
422 in climate sensitive species. Given the underlying physiological justification of the models  
423 presented here, we have no reason to suggest they are not broadly applicable to species of all  
424 shade-tolerance levels. We recommend future studies investigate the applicability of SDS and  
425 COMB models to both raw tree-ring width and BAI data in wider range of species. That said,  
426 shade-tolerant and broadleaf species, and their applicable standardization procedures, are  
427 underrepresented in dendrochronological studies (Zhao et al. 2019). We advocate for continued  
428 refinement of tree-ring standardization procedures that are relevant to the ecological questions  
429 they aim to address.

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