

# Methane efflux from an American bison herd

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**Abstract.** American bison (*Bison bison* L.) have recovered from the brink of extinction over the past century. Bison reintroduction creates multiple environmental benefits, but impacts on greenhouse gas emissions are poorly understood. Bison are thought to have produced some 2 Tg year<sup>-1</sup> of the estimated 9-15 Tg year<sup>-1</sup> of pre-industrial enteric methane emissions, but few contemporary measurements have been made due to their mobile grazing habits and safety issues associated with direct measurements. Here, we measure methane and carbon dioxide fluxes from a bison herd on an enclosed pasture during daytime periods in winter using eddy covariance. Methane emissions from the study area were negligible in the absence of bison (mean  $\pm$  standard deviation =  $-0.0009 \pm 0.008 \mu\text{mol m}^{-2} \text{s}^{-1}$ ) and were significantly greater than zero,  $0.048 \pm 0.082 \mu\text{mol m}^{-2} \text{s}^{-1}$  with a positively skewed distribution, when bison were present. We coupled bison location estimates from automated camera images with two independent flux footprint models to calculate methane efflux on a per-animal basis, which varied from  $55 \mu\text{mol s}^{-1} \text{bison}^{-1}$  to  $62 \mu\text{mol s}^{-1} \text{bison}^{-1}$ . Per-animal methane efflux estimates were also sensitive to uncertainties in bison location, and sensitivity analyses suggest that a conservative uncertainty estimate is on the order of 22%. Combined with conservative uncertainty estimates of the eddy covariance measurements themselves, we arrive at methane flux estimates of  $76 \pm 21 \text{ g CH}_4 \text{ bison}^{-1} \text{ day}^{-1}$  when using the Hsieh et al. (2000) model and  $86 \pm 24 \text{ g CH}_4 \text{ bison}^{-1} \text{ day}^{-1}$  when using the Kljun et al. (2015) model, similar to eddy covariance measurements of methane efflux from a cattle feedlot during winter. Annual measurements are ultimately necessary to determine the full greenhouse gas burden of bison grazing systems. Our observations point to the need for direct comparisons of methane emissions from conventional and alternate grazing systems using eddy covariance and demonstrate the potential for using eddy covariance to measure methane efflux from non-domesticated animals.

## 30 1 Introduction

The American bison (*Bison bison* L.) was hunted to near extinction during European expansion across North America (Flores 1991, Isenberg 2000, Smits 1995). Fewer than 100 reproductive individuals existed on private ranches in the United States during the late 19<sup>th</sup> Century from an original population of 30 – 60 million (Hedrick, 2009). The current bison population of about 500,000 is due to the collective efforts of sovereign Indian tribes, government agencies, and private  
35 landowners (Gates et al., 2010; Sanderson et al., 2008; Zontek, 2007), all of whom have spurred a growing interest in bison reintroduction. The bison population is likely to further increase, increasing the incentive for researchers and land managers to understand the environmental impacts of their expansion.

The ecological role of bison has become better understood as populations have recovered (Allred et al., 2001; Hanson 1994; Knapp et al., 1999). Bison feed preferentially on grasses (Plumb and Dodd, 1993; Steuter and Hiding, 1999) and enhance  
40 forb diversity as a result (Collins, 1998; Hartnett et al., 1996, Towne et al., 2005). They tend to graze in preferred meadows during winter and search broadly for the most energy-dense forages in summer (Fortin et al., 2003), often in areas which have recently burned (Allred et al., 1991; Coppedge and Shaw, 1998; Vinton et al., 1993). Bison also need not migrate to follow the ‘green wave’ of fresh vegetation during spring green-up like other ungulates; rather, their vigorous grazing tends to stimulate plant growth and create fresh, nutrient-rich foliage (Geremia et al., 2019). Combined, these observations suggest  
45 that bison select for forage quality rather than quantity which likely impacts their efflux of methane – which all ruminants emit – because ruminant methane emission is related to the cellulose and hemicellulose intake of their diet (Moe and Tyrrell, 1979) and feed quality (Hammond et al., 2016). It remains unclear how much methane results from the cellulose-rich grass-dominated diet of bison given their preference for fresh foliage and if management for bison may increase or diminish the greenhouse gas burden of ruminant-based agriculture.

50 Atmospheric methane concentrations have been rising at an accelerated rate since 2016 for reasons that remain unclear (Nisbet et al., 2019) and there is an urgent need to improve our understanding of its surface-atmosphere flux. Between 30 and 40 percent of current anthropogenic methane emissions are due to enteric fermentation in livestock (Kirschke et al., 2013) and the greenhouse gas burden of cattle alone is some 5 Pg of carbon dioxide equivalent per year (Gerber et al., 2013; FAO, 2017). Recent studies have revised methane emission estimates from livestock upward by over 10% (Beauchemin et al., 2008; Thornton and Herrero, 2010; Wolf et al., 2017), further emphasizing their critical role in global greenhouse gas  
55 budgets (Reisinger and Clark, 2017). Reducing unnecessary greenhouse gas emissions is a global imperative for Earth system management and reducing enteric methane sources is seen as a promising approach to do so (Boadi et al., 2002; DeRamus et al., 2003; Herrero, et al., 2016; Hristov et al., 2013; Johnson and Johnson, 1995; Moss et al., 2000).

Bison in North America are thought to have been responsible for some 2.2 Tg year<sup>-1</sup> (Kelliher and Clark, 2010; Smith et al.,  
60 2016) of the 9-15 Tg year<sup>-1</sup> of pre-industrial enteric methane emissions (Thompson et al., 1993; Chappellaz et al., 1993; Subak, 1994). Enteric CH<sub>4</sub> emissions from wild ruminants in the United States in the pre-settlement period comprised nearly 90% of current CH<sub>4</sub> emissions from domesticated ruminants assuming an historic bison population size of 50 million

(Hristov, 2012), further demonstrating the importance of bison to methane fluxes in the past. The current and future contribution of non-domesticated ungulates to methane fluxes are uncertain (Crutzen et al., 1985). Previous approaches used inventory approaches or scaling equations that were not derived using methane efflux measurements from bison; the only direct bison methane flux observations that we are aware of measured 30 L per kg dry food intake (17 g methane per kg dry food intake) in one-year-old penned female bison fed alfalfa pellets (Galbraith et al., 1998), more than elk (*Cervus elaphus*) and white-tailed deer (*Odocoileus virginianus*) on a dry matter intake basis and similar to dairy cattle fed high maize silage (Hammond et al., 2016). Cattle methane emissions tend to be greater when fed alfalfa than grass (Chaves et al., 2006) such that existing published values may not represent an accurate estimate of the methane efflux from bison in a natural field setting, which has not been measured to date.

Here, we measure methane flux from a bison herd on winter pasture using the eddy covariance technique (Dengel et al., 2011; Felber et al., 2015; Prajapati and Santos, 2018; Sun et al., 2015). We use flux footprint analyses combined with bison locations determined using automated cameras to estimate methane flux on a per-animal basis and discuss observations in the context of eddy covariance methane flux measurements from other ruminants.

## 2 Methods

### 2.1 Study site

The study site is a 5.5-hectare pasture on the Flying D Ranch near Gallatin Gateway, Montana, USA (45.557, -111.229) on a floodplain immediately west of the Gallatin River (Figure 1). Daily high temperatures average 1.6 °C and daily low temperatures average -11.5 °C at Bozeman Yellowstone International Airport (BZN), located 24 km north-northeast of the site, during the November – February measurement period. BZN records an average of 18.2 mm of precipitation per month during November – February, almost entirely as snowfall. A herd of 39 bison entered the pasture on November 17, 2017 and left on February 3, 2018. The mean (standard error) bison weight measured by the landowners on November 16, 2017 before bison entered the pasture was  $329 \pm 28$  kg and the bison varied in age from 0.5 to 7.5 years old (Table S1). Bison consumed a mixture of perennial grasses grown *in situ* that was supplemented by perennial grass hay grown in nearby fields (Table S2) delivered every three days on average (Table S3).

### 2.2 Instrumentation

A 3 m tower was installed near the center of the study pasture during November 2017 (Figure 1) and surrounded by electric fencing to avoid bison damage. Four game cameras (TimelapseCam, Wingscapes, EBSCO Industries, Inc., Birmingham, AL, USA) were mounted to the tower and pointed in cardinal directions. Two additional game cameras were mounted near the pasture edge facing the tower. Cameras captured images every five minutes and an example of an individual image from the south-facing camera located on the northern edge of the study pasture is shown in Figure 2. Bison locations at the half-

hourly time interval of the eddy covariance measurements were estimated by manually attributing bison locations to squares in a 20 m grid overlaid on the pasture area (Figure 1). The 20 m grid size represents the grid that we felt that we were able to attribute bison locations given features of the field that could be identified by camera. The bison location approach introduces uncertainty, and we test the sensitivity of per-animal methane efflux estimates to bison location estimates as described in the *Spatial Uncertainty* section below.

Incident and outgoing shortwave and longwave radiation and thereby the net radiation were measured using a NR01 net radiometer (Hukseflux, Delft, The Netherlands) mounted 1.5 meters above ground level. A SR50 sonic distance sensor (Campbell Scientific Inc., Logan, UT, USA) was installed at 1.3 m to gauge snow depth, and air temperature and relative humidity were measured at 2.25 meters using a HMP45C probe (Vaisala, Vantaa, Finland). Average 0–30 cm soil moisture and temperature were collected using CS650 probes (Campbell Scientific). Meteorological variables were measured once per minute, and half-hourly averages were stored using a CR3000 datalogger (Campbell Scientific).

Three-dimensional wind velocity was measured using a CSAT-3 sonic anemometer (Campbell Scientific) at 2.0 m above the ground surface. Carbon dioxide mixing ratios were measured at 10 Hz using a LI-7200 closed-path infrared gas analyzer (LI-COR Biosciences, Inc.) with inlet placed at the same height as the center of the sonic anemometer. Methane mixing ratios were measured at 10 Hz using a LI-7700 open-path infrared gas analyzer (LI-COR Biosciences, Inc., Lincoln, NE, USA) with the center of the instrument likewise located at 2.0 m and a 22 cm horizontal offset from the sonic anemometer; open- and closed-path infrared gas analyzers for eddy covariance have similar performance in field settings (Detto et al., 2011; Deventer et al., 2019). We use the atmospheric convention in which flux from biosphere to atmosphere is positive. Measurements were made during winter daytime hours from 0700 to 1700 local time to avoid depleting the battery bank and to ensure sufficient light to estimate bison location using game cameras. Flux measurements began on November 14, 2017 and ended on February 14, 2018.

Bison are dangerous and will charge humans. Their presence complicated data retrieval and game camera upkeep; some high-frequency flux measurements were overwritten and cameras shut down during exceptionally cold periods, resulting in missing measurements. Simultaneous flux and photographic data were obtained for the January 7, 2018 to February 13, 2018 period excluding January 10, 2018 when instruments were obstructed by snowfall. Flux data without accompanying game camera footage were obtained for the periods from November 14 through 29, 2017 and December 31, 2017, through January 6, 2018.

## 2.3 Flux calculations

Methane and carbon dioxide fluxes were calculated using EddyPro (LI-COR Biosciences, Lincoln, NE, USA). Standard double rotation, block averaging, and covariance maximization with default processing options were applied. Spike removal was performed as described by Vickers and Mahrt (1997) and spikes were defined as more than 3.5 standard deviations from the mean mixing ratio for carbon dioxide and more than 8 standard deviations from the mean mixing ratio for methane given

125 the expectation of intermittent methane spikes from the bison herd. The default drop-out, absolute limit, and discontinuity  
tests were applied using the default settings following recommendations by Dumortier et al. (2019), and the Moncrieff et al.  
(1997) and Moncrieff et al. (2004) low and high-pass filters were applied. The Webb-Pearman-Leuning correction (Webb et  
al., 1980) was applied to calculate methane efflux using the open-path LI-7700 sensor. Estimates of storage flux in the 2 m  
130 airspace below the infrared gas analyzers were assumed to be minor and excluded from the flux calculation. Flux  
measurements for which the quality control flag was greater than 1 following Mauder and Foken (2004) (see also Foken et  
al., 2004) were discarded and the net effect of all corrections when bison were present was a methane flux reduction of 14%.  
Measurements that exceeded an absolute value of  $1 \mu\text{mol m}^{-2} \text{s}^{-1}$  for the case of methane flux and  $20 \mu\text{mol m}^{-2} \text{s}^{-1}$  for the case  
of carbon dioxide flux were discarded following an analysis of the probability distribution of observations. We tested the  
sensitivity of flux measurements to the friction velocity ( $u^*$ ) to see if measurements made under conditions of insufficient  
135 turbulence should be excluded from the analysis despite the daytime-only flux measurement approach.

## 2.4 Flux footprint modelling

The eddy covariance flux footprint was calculated using the approach of Hsieh et al. (2000) extended to two dimensions  
following Detto and Katul (2006). Such analytical footprint models have been found to give minimally biased estimates of  
point-source fluxes in field settings (Dumortier et al., 2019). We performed the footprint analysis on a grid of 1 m pixels and  
140 aggregated values to the  $20 \times 20$  m grid to which the bison locations were estimated (Figure 1); Figure 3 demonstrates an  
example of a flux footprint for a single half-hourly period.

To further characterize the uncertainty in our per-animal methane flux estimates, described next, we also applied the flux  
footprint parameterization method of Kljun et al. (2015) aggregated to the same  $20 \times 20$  m grid. The Kljun et al. (2015)  
model performs well in point-source experiments (Heidbach et al., 2017) and is widely used by the flux community.

145 The momentum roughness height ( $z_{0m}$ ) is required by both footprint models. Instead of assuming a constant  $z_{0m}$  over snow of  
0.001 m (Andreas et al., 2004), we followed the approach of Baum et al. (2008) who calculated a unique  $z_{0m}$  for each half-  
hour eddy covariance measurement for a cattle feedlot system by rearranging the wind profile equation:

$$z_{0m} = \frac{z - d}{\exp(ku/u_* + \psi_m)} \quad (1)$$

Where  $z$  is measurement height,  $u$  is wind speed,  $k$  is the von Karman constant, and  $\psi_m$  is the correction factor for  
atmospheric stability, here following Brutsaert (1982). The zero-plane displacement ( $d$ ) for a field with obstacles is  
150 calculated following Verhoef et al. (1997):

$$d = z - \frac{z(1 - \exp(-\sqrt{42a}))}{\sqrt{42a}}. \quad (2)$$

where  $a$  is the frontal area index of the obstacles (Raupach, 1994), here bison:

$$a = \frac{nbh}{s}. \quad (3)$$

The calculation of  $a$  uses the number of animals ( $n = 39$ ), the size of the pasture ( $S$ , m<sup>2</sup>), and the average breadth ( $b$ , m) and height ( $h$ , m) of the animals. We used established relationships for beef cattle as a function of weight (ASABE, 2006) given the lack of similar equations for bison.  $h$  was adjusted upward by 50% such that the height of adult males better-matched average values of fully-grown bison on the order of 1.8 m. The methane source location was assumed to be near the ground per the typical posture of bison assuming that most methane efflux in ruminants is from erucation.

## 2.5 Per-bison methane flux estimation

Given that mean methane emissions were not significantly different from zero in the absence of bison – as detailed in *Results* – we assume that observed methane emissions are due to bison in the flux footprint. The relative contribution of bison to each half-hourly eddy covariance measurement was calculated by expanding the approach of Dumortier et al. (2019) (see also Prajapati and Santos (2019)) for multiple point sources. From the definition of the footprint function, the measured density of a scalar  $X$ ,  $F_X$ , for our study area of  $8 \times 12$  grid cells (Figure 1) is:

$$F_X = \sum_{i=1}^8 \sum_{j=1}^{12} F_{ij} \phi_{ij} \Delta x_{ij} \Delta y_{ij} \quad (4)$$

where  $\phi_{ij}$  is the value of the footprint function in grid cell  $ij$  (here per 400 m<sup>2</sup>) and  $x$  and  $y$  are the dimensions of the 20 m grid cells. Dumortier et al. (2019) considered a known point source from a single cell,  $f_x$ , such that:

$$f_x = \frac{F_X}{\phi_{ij,source}} \quad (5)$$

where  $\phi_{ij,source}$  is the value of the footprint function at the source location. We have  $n = 39$  sources (i.e. bison) that are free to wander to any grid cell  $ij$ . We also have no basis for identifying individual bison given the resolution of the cameras, noting that this is possible using higher-resolution cameras (Merkle and Fortin, 2013) or GPS instruments. We also have no basis for determining if the methane sources of individual bison are different using our approach, so we must assume that methane efflux from each bison is equal. Under these assumptions we can write:

$$\langle f_x \rangle = \frac{F_X}{\sum_{i=1}^8 \sum_{j=1}^{12} n_{ij} \phi_{ij} \Delta x_{ij} \Delta y_{ij}} \quad (6)$$

Where  $n_{ij}$  is the number of bison in grid cell  $ij$  (i.e. per 400 m<sup>2</sup>) and  $\langle f_x \rangle$  is the average flux per bison. We only adopt this approach for calculating average methane efflux per bison as measured carbon dioxide fluxes in the absence of bison were greater than zero. Methane efflux values less than  $-200 \mu\text{mol bison}^{-1} \text{s}^{-1}$  and greater than  $300 \mu\text{mol bison}^{-1} \text{s}^{-1}$  were treated as outliers and excluded based on an analysis of the probability distribution of observations.

## 2.6 Spatial uncertainty

The location of bison in the pasture was approximated visually by identifying the position of bison in relation to static cues in the study area using five-minute photographs. Observations were then aggregated to half-hourly flux measurement

180 periods. This approach results in spatial uncertainty in bison location, especially due to movements within half-hourly periods and due to the size of the animals themselves with reference to the grid (Figure 1). Both will result in a greater spatial distribution of bison locations than represented in the maps of bison distributions that were created. To arrive at a conservative estimate of uncertainty in per-bison methane flux estimates, we explored the sensitivity of flux estimates to maps of bison location that were more distributed in space. To do so we used two-dimensional Tikhonov Regularization (Tikhonov and Arsenin, 1977), a classic mathematical technique to solve ill-posed problems, here the challenge of estimating the best spatial distribution of bison with intermittent observations.

185 To briefly describe the motivation for using Tikhonov Regularization for our case, consider the extreme distributions of potential bison locations: all are located in a single grid cell (a Dirac delta function) or all are perfectly aggregated across the field (an uninformed prior). The true distribution likely exists between these cases, especially given intermittent bison movements. The estimates of their location using cameras provides an initial guess of the true location. We assume that the true number of each bison in each pixel is likely to be similar to those measured in adjacent pixels because the bison movements were usually minor and because of the uncertainty that exists when attempting to associate bison to a particular pixel. Two-dimensional Tikhonov Regularization can provide an estimate of the true distribution of elements (here bison) given the constraints that the distribution is bounded and that adjacent pixels are likely similar to a given pixel (Stoy and Quaife, 2015).

We use a form of Tikhonov Regularization to create spatial disaggregation of each bison distribution map ( $\alpha'$ ) following Stoy and Quaife (2015):

$$\alpha' = \alpha(I + \gamma^2 B^T B)^{-1} \frac{\sigma^2}{\psi(\gamma^2)} - \mu_\alpha + \mu_{\alpha'} \quad (7)$$

195 Here,  $\alpha$  is the measured distribution map with mean  $\mu_\alpha$  (the number of bison per pixel) and variance  $\sigma$ ,  $I$  is the identity matrix,  $B$  represents the constraint that neighboring elements should be similar by requiring a first difference of zero in the cardinal directions of the map,  $\gamma$  is the Lagrange Multiplier, and  $\psi(\gamma^2)$  is a normalization term equal to the variance of  $\alpha(I + \gamma^2 B^T B)^{-1}$ . Large values  $\gamma$  constrain each pixel to be near the overall mean such that the bison distribution map is smoother across space.

200 We applied the Tikhonov Regularization approach to each bison distribution map using Lagrange multipliers that ranged from 0.1 to 4 as demonstrated for a single half-hour period in Figure 4; note that the simulation with a Lagrange multiplier of 4 results in a simulation where bison are widely distributed across segments of the field and amounts to a highly conservative estimate of their location. Our results are subject to simultaneous uncertainties in footprint and bison location in addition to the eddy covariance methane flux measurements themselves, which range from 6 – 41% for half-hourly fluxes and 7 – 17% for long-term sums (Deventer et al., 2019). We use 17% as a representative uncertainty of eddy covariance sums as we are primarily concerned with providing a conservative assessment of uncertainty using our approach. We suggest strategies for reducing uncertainty in the *Discussion* section.

### 3. Results

#### 210 3.1 Meteorology

Air temperature averaged  $-2.8$  °C and soil temperature averaged  $-0.3$  °C during the measurement period (Figure 5A). Incident shortwave radiation ranged between 100 and 400  $\text{W m}^{-2}$  during peak daylight hours (1000-1400 hours local time) across the study period, and clear conditions were common except for four weeks beginning in mid-December (Figure 5B). Snow depth within the tower enclosure increased from 0.15 m to nearly 0.4 m in late 2017 and decreased to 0.1 m beginning  
215 in late January 2018 (Figure 5C) noting that snow outside of the electrified tower enclosure was often trampled (see Figure 2). The mean (median) wind direction was  $221^\circ$  ( $208^\circ$ ) during periods when visible imagery of bison locations was available and eddy covariance measurements passed quality control checks (Figure 6).

#### 3.2 Gas flux

Half-hourly methane fluxes averaged  $0.048 \pm 0.081$   $\mu\text{mol m}^{-2} \text{s}^{-1}$  (mean  $\pm$  standard deviation) and carbon dioxide fluxes  
220 averaged  $1.6 \pm 1.4$   $\mu\text{mol m}^{-2} \text{s}^{-1}$  when bison were present (Figure 7), noting again that measurements were made only during daytime periods. Methane flux in the absence of bison averaged  $-0.0009 \pm 0.008$   $\mu\text{mol m}^{-2} \text{s}^{-1}$  and carbon dioxide flux averaged  $0.64 \pm 1.0$   $\mu\text{mol m}^{-2} \text{s}^{-1}$ , significantly lower than when bison were present ( $P < 0.001$  for both  $\text{CH}_4$  and  $\text{CO}_2$ ).  $\text{CO}_2$  flux was significantly related to methane flux and explained 52% of its variance when bison were present but only 7% when they were absent (Figure 8).  $\text{CO}_2$  flux was significantly and positively related to air and soil temperature across the entire  
225 measurement record ( $P < 0.001$  in both cases), but methane flux was not. There were no significant temporal patterns of methane flux during the daytime periods investigated here, and neither incident nor net radiation were related to methane fluxes. When bison were present, methane flux was not significantly different at the  $P < 0.05$  significance level during days when feed was delivered ( $0.051 \pm 0.083$   $\mu\text{mol m}^{-2} \text{s}^{-1}$ ) and days when it was not ( $0.035 \pm 0.10$   $\mu\text{mol m}^{-2} \text{s}^{-1}$ ) ( $P = 0.075$ ). Methane flux was significantly and positively related to friction velocity in the absence of bison at  $u^*$  values greater than 0.2  
230  $\text{m s}^{-1}$  ( $P = 0.003$ ) but not positively related to  $u^*$  values less than 0.2  $\text{m s}^{-1}$ , indicating that flux measurements were unrelated to friction velocity values commonly associated with insufficient turbulence (Figure 9A). Carbon dioxide flux was not related to  $u^*$  in the absence of bison (Figure 9B) but negative values were observed at  $u^*$  values greater than 0.45  $\text{m s}^{-1}$ . Given these observations, we did not apply a  $u^*$  filter to our eddy covariance measurements, which were made only during daytime periods. We discuss potential reasons for the observed increase in methane flux and negative  $\text{CO}_2$  flux with high  
235 values of  $u^*$  in the *Discussion* section.

#### 3.3 Bison location and methane efflux

Timelapse camera footage yielded usable imagery for 444 half-hourly periods of which 245 half-hourly periods had available eddy covariance observations and of which 177 had eddy covariance measurements that passed quality control criteria. Bison tended to aggregate in an area on the west side of the pasture near the location where supplemental hay was



240 often provided (Figure 10A). They intermittently visited the area north of the tower in mornings and afternoons and  
intermittently made sporadic mass movements to the southernmost edge of the field near its gate during midday periods  
(Figure 10B-D).

Bison were located within the 90% flux footprint 40% of the time (Figure 11). There was an average of eight (seven bison  
within the 90% flux footprint of the Hsieh et al. (2000) (Kljun et al. (2015)) models which increased to both footprint  
245 models which increased to 21 (20), respectively, when excluding observations with no bison (Figure 11). Per-bison methane  
emission estimates when using the Hsieh et al. (2000) footprint model had a mean ( $\pm$  standard error) of  $55 \pm 0.96$   
 $\mu\text{mol bison}^{-1} \text{ s}^{-1}$  and a median of  $29 \mu\text{mol bison}^{-1} \text{ s}^{-1}$  as a result of the positively skewed measurement distribution (Figure  
12A). These estimates are 11% lower than per-bison methane emission estimates from the Kljun et al. (2015) footprint  
250 estimates are sensitive to flux footprint methodology. Per-animal flux estimates are also sensitive to the estimates of their  
location within the field; mean methane flux estimates ranged from  $43 - 58 \mu\text{mol bison}^{-1} \text{ s}^{-1}$  when applying the Hsieh et al.  
(2000) model and  $50 - 75 \mu\text{mol bison}^{-1} \text{ s}^{-1}$  when applying the Kljun et al. (2015) model after spatial smoothing using  
Tikhonov Regularization. These estimates are up to 22% different than the mean methane flux estimates that were generated  
by taking the maps at face value. If we adopt 22% as a conservative uncertainty estimate due to spatial uncertainty and 17%  
255 as a conservative uncertainty estimate of long-term methane flux sums (Deventer et al., 2019) for a combined uncertainty of  
28%, we arrive at a daily per-bison methane flux estimate of  $76 \pm 21 \text{ g CH}_4 \text{ bison}^{-1} \text{ day}^{-1}$  when using the Hsieh et al. (2000)  
footprint model and  $86 \pm 24 \text{ g CH}_4 \text{ bison}^{-1} \text{ day}^{-1}$  when using the Kljun et al. (2015) footprint model.

#### 4 Discussion

260 The eddy covariance flux footprint analysis coupled to bison location estimates from automated camera images resulted in a  
mean (median) methane flux of  $55$  ( $29$ )  $\mu\text{mol bison}^{-1} \text{ s}^{-1}$  when applying the Hsieh et al. (2000) footprint model and  $62$  ( $43$ )  
 $\mu\text{mol bison}^{-1} \text{ s}^{-1}$  when applying the Kljun et al. (2015) footprint model. Measurements were made during daytime periods in  
winter and are sensitive to estimates of bison location (Figure 12). If we naively assume that methane flux from bison varies  
negligibly across the full diurnal and seasonal range, a notion that needs to be substantiated, our measurements with  
265 conservative uncertainty estimates correspond to  $26 \pm 8$  ( $31 \pm 9$ ) kilograms of methane per bison per year when applying the  
Hsieh et al. (2000) (Kljun et al. (2015)) model, noting that methane emissions from cattle have been observed to be on the  
order of 10-17% higher in summer than winter (Todd et al., 2014; Prajapati and Santos, 2018; Prajapati and Santos, 2019)  
but lower in evenings if animals eat less during these times (Gao et al., 2011). The mean weight of the study bison herd was  
329 kg, similar to the 300 kg buffalo that is assumed to emit  $55 \text{ kg year}^{-1}$  in the 2006 IPCC report (IPCC 2006) noting that  
270 dairying buffalo cows are estimated to have higher methane emissions than other buffalo (C6ndor et al. 2008). The study  
herd here comprised numerous pregnant females (Table S1) that have higher metabolic requirements. Previous estimates of  
methane emissions from range cattle are on the order of 60 kg per year per animal (Hogan, 1993), about twice as large as the  
mean per-bison methane flux calculated here. Values were instead similar to the lower range measurements from young

275 heifers feeding on ryegrass of  $88 \text{ g CH}_4 \text{ animal}^{-1} \text{ day}^{-1}$  (Hammond et al., 2016) and wintertime measurements of beef cattle  
in a feedlot on the order of  $75 \text{ g CH}_4 \text{ animal}^{-1} \text{ day}^{-1}$  (Prajapati and Santos, 2019). In other words, while there is no evidence  
from our measurements that bison have more or less methane efflux than typical values reported for cattle, it is critical to  
make full year-round methane flux measurements with uncertainty to understand the seasonal course of bison methane efflux  
to establish defensible annual sums.

280 Below, we discuss potential reasons for the relatively low bison methane emissions observed here as well as a strategy for  
reducing uncertainty in eddy covariance measurements of methane efflux from grazing non-domesticated ruminants.

#### 4.1 Bison methane and carbon dioxide efflux in response to environmental variables

Methane flux was not related to air or soil temperature but was related to  $u^*$  – especially at relatively high values of  $u^*$  – in  
the absence of bison (Figure 9). These observations are consistent with a potential pressure pumping mechanism for trace  
gases through snow at higher wind speeds (Bowling and Massman, 2011) although it is unclear why this relationship exists  
285 for methane flux and not carbon dioxide flux as is frequently found in snow-covered conditions (Rains et al., 2016). Carbon  
dioxide flux at high values of  $u^*$  was negative indicating net  $\text{CO}_2$  uptake by the biosphere, which is unlikely in our study site  
during winter, suggesting that values with excessively high  $u^*$  may need to be filtered, but with only five observations of  
 $\text{CO}_2$  flux less than zero it is unclear how to apply such a filter in our case.

Insufficient evidence exists in our data record to attribute observed methane efflux to the onset of freezing conditions in soil  
290 (Mastepanov et al., 2008). We note that extensive snow trampling (e.g., Figure 2) likely resulted in a situation where snow  
depth (Figure 5C) and its insulating effect on soil temperature (Figure 5A) varied across the field and therefore differed from  
snow and soil measurements taken within the instrumentation enclosure. Regardless, mean methane flux when bison were  
absent,  $-0.0009 \mu\text{mol m}^{-2} \text{ s}^{-1}$ , was nearly two orders of magnitude less than the mean methane flux when bison were present,  
 $0.041 \mu\text{mol m}^{-2} \text{ s}^{-1}$ . Whereas we cannot exclude – and in fact expect – positive and negative background methane fluxes from  
295 non-bison sources in a grassland in winter in the vicinity of a riparian area (Figure 1, Merbold et al., 2013; McLain and  
Martens, 2006; Mosier et al., 1991), these appear to be minor compared to the  $\text{CH}_4$  flux attributable to bison (Figures 7 and  
8). Bison are associated with a distinct methane flux signature as shown by the immediate decline of methane fluxes  
following their removal from the study pasture (Figure 7) and strong relationship with carbon dioxide flux (Figure 8) given  
the common source of respiration and most enteric methane losses from the mouths of ungulates. Methane flux was related  
300 to carbon dioxide flux when bison were present or absent (Figure 8), suggesting both soil and ruminant sources (and in the  
case of methane sinks) of both gases (Baldocchi et al., 2012; Motte et al., 2019).

It is important to note that potential methane fluxes from bison manure may have been dampened by freezing conditions but  
are an important methane source during warmer conditions in ruminant grazing systems (Dengel et al., 2011). Manure is  
thought to contribute a nontrivial portion ( $10\text{-}14 \text{ Tg CH}_4 \text{ yr}^{-1}$ ) of total global ruminant methane efflux ( $77 \text{ Tg CH}_4 \text{ yr}^{-1}$ ,  
305 Johnson and Ward 1996; Moss et al., 2000) noting that some farm-scales studies arrive at lower percentages (Taylor et al.,

2017). Though we did not observe higher methane efflux early in the study period when soil temperature was above freezing nor temperature sensitivity of methane efflux in the presence or absence of bison, it is important to note that field-scale methane efflux may be diminished by the thermal environment of manure in our measurements such that methane efflux is lower than would be expected during warmer seasons.

#### 310 **4.2 Bison spatial distribution and movement**

Ruminant behavior is critical to track to estimate field-scale efflux (de la Motte et al., 2019). The spatial distribution of bison in the study pasture varied from morning to midday and afternoon (Figure 10). It is difficult to infer from the available data whether the study bison are more active during morning and evening hours in the pasture environment like cattle (Gregorini 2012). Supplemental hay was made available to the bison approximately 50 meters west of the tower and increases in the  
315 frequency of bison appearance there are likely associated with the animals' preferred feeding times after dawn and before dusk, but observed methane flux did not vary as a function of time of day (e.g. Dengel et al. 2011) and was not significantly different during days when hay was provided and when it was not, noting that the animals were free to also graze on vegetation within the pasture. Regardless, ruminant methane flux measurements are simpler to make when animals congregate (Coates et al., 2017; Tallec et al., 2012) as was often observed here (e.g. Figures 2, 10 & 11). Aggregation  
320 behavior in our study bison herd was often upwind of the eddy covariance tower (Figures 6 & 10) and resulted in more overlap between flux footprint and bison location than would have occurred if bison locations were randomly distributed throughout the study area, emphasizing the importance of tower placement in eddy covariance studies of grazing systems. Spatial uncertainties in bison location interact with uncertainties in flux footprint modelling for methane source attribution (Figure 12). Footprint models of the type used here have been found to accurately estimate point sources of trace gas flux  
325 (Heidbach et al., 2017; Dumortier et al., 2019), but it is important to note that footprint modelling techniques play a large role in the spatial attribution of observed fluxes of ruminant trace gas flux (Felber et al., 2015). Prajapati and Santos (2018), for instance, found that an analytical model (Kormann and Meixner 2001) predicted flux footprint areas five to six times larger than did an approximation of a Lagrangian dispersion model (Kljun et al., 2002), such that footprint model uncertainty is a major source of uncertainty for measuring methane flux from multiple point sources as we also found here.

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#### **4.3 Future directions for greenhouse gas accounting in ruminant grazing systems**

Methane efflux cannot be completely removed from ruminant grazing systems; some 4.6 – 6.2% of gross energy intake is lost as methane in cattle, sheep and goats worldwide (Johnson and Ward 1996) with cattle often falling on the higher end of the observed range (Lassey et al., 1997). But there are other aspects of bison ecology that merit consideration when  
335 designing greenhouse gas-cognizant grazing systems. For example, cattle tend to graze close to water more frequently than bison do (Allred et al., 2011) with unclear consequences for riparian vegetation, water quality, and potential methane efflux

from cattle wallows. Cattle also tend to graze for longer periods than bison (Plumb and Dodd, 1993) and it is unclear if there is an associated consequence for methane efflux. Future work should consider the large inter-animal variability in methane efflux (Lassey et al., 1997), possibly using advanced techniques for identifying individual animals through photographs (Merkle and Fortin, 2013) or tracking devices (Felber et al., 2015). Animal age and size are also important factors in ruminant methane efflux (Jiao et al., 2014) and individual tracking may improve our estimates of this variability in a field setting.

Adding seasonal foraging behavior, estimating emissions from individual animals, and addressing seasonal and inter-annual variability and trends in forage nutrition are likely to further improve prediction of methane emissions from grazing systems (Moraes et al., 2013). Advanced eddy covariance algorithms for are also likely to improve flux estimates on short time scales noting that non-stationary bursts have not been found to create systematic bias in methane budgets measured over longer time periods using eddy covariance (Göckede et al., 2019). Of these, advanced footprint attribution techniques like Environmental Response Functions designed to create improved maps of surface-atmosphere fluxes (Metzger et al., 2013; Xu et al., 2017) may be uniquely applicable to the challenging case presented by grazing systems with mobile point sources and intermittent biogeochemical hotspots created by animal waste. Going forward, increases in atmospheric carbon dioxide concentrations are likely to decrease forage quality (Jégo et al., 2013), resulting in higher leaf carbon to nitrogen ratios and which is expected to increase ruminant methane emissions (Lee et al., 2017), all else being equal. Understanding greenhouse gas fluxes from ruminants is therefore likely to be even more important in the future. An ongoing interest in bison reintroduction and ungulate ecology coupled with established micrometeorological measurement techniques will help us understand the present and future role that bison and alternative grazing systems play in the Earth system.

### **Supplemental Information**

The land managers provided information that describes bison age, sex, weight, and pregnancy status (Table S1) and the composition (Table S2) and the delivery schedule (Table S3) of hay.

360

### **Acknowledgements**

PCS acknowledges support from the U.S. National Science Foundation awards DEB-1552976 and OIA-1632810, the USDA National Institute of Food and Agriculture Hatch project 228396, the Multi-State project W3188, the Graduate School at Montana State University, and the University of Wisconsin – Madison. Funding for the LI-7700 methane analyzer used in this work was provided to JED by NSF-EPSCoR award EPS-1101342 and Montana State University. Daniel Salinas, Gabriel Bromley, Zheng Fu, and James Irvine provided technical assistance. This work could not have been completed without permission of Turner Enterprises, Inc. and the assistance of Carter Kruse and Danny Johnson.

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### **Code/Data availability**

370 Eddy covariance and micrometeorological data have been submitted to Ameriflux for publication at  
https://ameriflux.lbl.gov/sites/siteinfo/US-Tur.

### Author contributions

PCS designed the study with AC, JD, and WK and wrote the manuscript with all coauthors. AC collected data and analyzed  
375 it with PCS and TG. NK assisted with the footprint analysis.

### Competing interests

The authors declare no competing interests.

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Tables

590 **Table S1: The sex, age, and pregnancy status of the study bison with weight measured on November 16, 2017 shortly before they entered the pasture on November 17, 2017. Bison were assumed to be born on June 1 of the birth year by the landowners such that animals born in 2017 were nearly 6 months old when measurements began.**

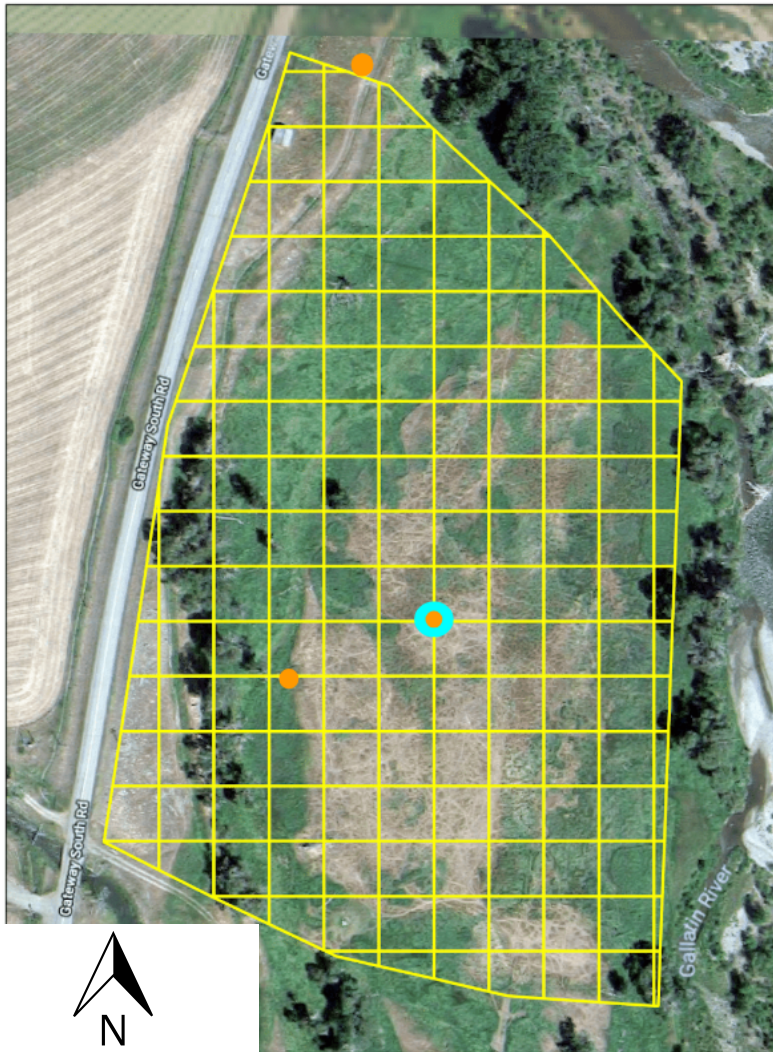
Sex	Age	Weight (kg)	Pregnant
F	7.5	467	Y
F	7.5	419	Y
F	7.5	428	Y
F	7.5	479	Y
F	7.5	510	Y
F	7.5	476	Y
F	7.5	492	Y
F	7.5	454	Y
F	7.5	567	Y
F	7.5	476	Y
F	7.5	497	Y
F	7.5	460	Y
F	7.5	443	Y
F	7.5	435	Y
F	7.5	426	Y
F	7.5	476	Y
F	7.5	411	Y
M	5.5	646	
M	5.5	701	
F	3.5	381	Y
F	3.5	410	Y
F	1.5	334	
F	0.5	110	
F	0.5	144	
M	0.5	160	
F	0.5	166	
M	0.5	138	
M	0.5	152	
M	0.5	147	
M	0.5	183	
F	0.5	96	
M	0.5	208	
M	0.5	104	
M	0.5	163	
F	0.5	127	
M	0.5	136	
M	0.5	165	
M	0.5	126	
F	0.5	127	

**Table S2: Composition of the first cut and second cut hay provided as a supplement to the study bison herd.**

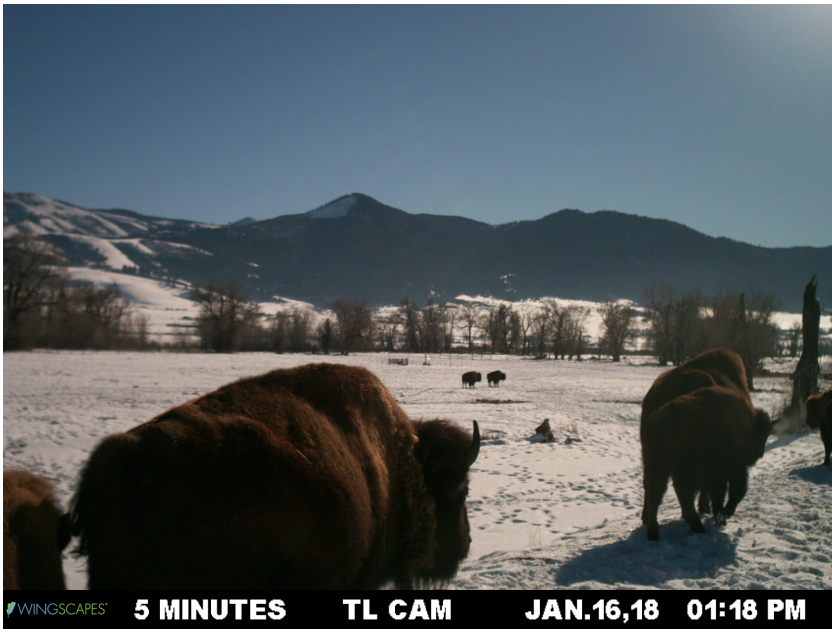
Variable (% unless otherwise noted)	First cut	Second cut
Crude Protein	9.7	17.2
Acid detergent fiber	47.9	38.3
Total digestible nutrients	48.9	59.7
Calcium	0.8	1.51
Phosphorus	0.2	0.21
Magnesium	0.21	0.32
Potassium	1.92	2.06
Sulfur	0.15	0.32
Sodium	<0.011	0.028
Zinc (mg/kg)	14	15
Iron (mg/kg)	66	61
Manganese (mg/kg)	60	56
Copper (mg/kg)	7	9

**Table S3: The number of bails of first cut and second cut hay (Table S2) delivered to the bison pasture.**

Date	First cut	Second cut
Nov. 17, 2017	2	2
Nov. 20, 2017		2
Nov. 22, 2017	1	2
Nov. 25, 2017		2
Nov. 27, 2017	2	2
Nov. 29, 2017		2
Dec. 1, 2017		2
Dec. 3, 2017		
Dec. 5, 2017	2	2
Dec. 8, 2017	2	2
Dec. 12, 2017		2
Dec. 15, 2017	2	2
Dec. 19, 2017	2	2
Dec. 21, 2017	2	2
Dec. 26, 2017	2	2
Dec. 28, 2017	2	2
Dec. 31, 2017	2	2
Jan. 2, 2018	2	2
Jan. 5, 2018	2	2
Jan. 8, 2018	2	2
Jan. 11, 2018	2	2
Jan. 15, 2018	2	2
Jan. 18, 2018	2	2
Jan. 22, 2018	2	2
Jan. 26, 2018		2
Jan. 27, 2018	2	
Jan. 29, 2018	1	1
Jan. 31, 2018	1	1
Feb. 3, 2018	2	2



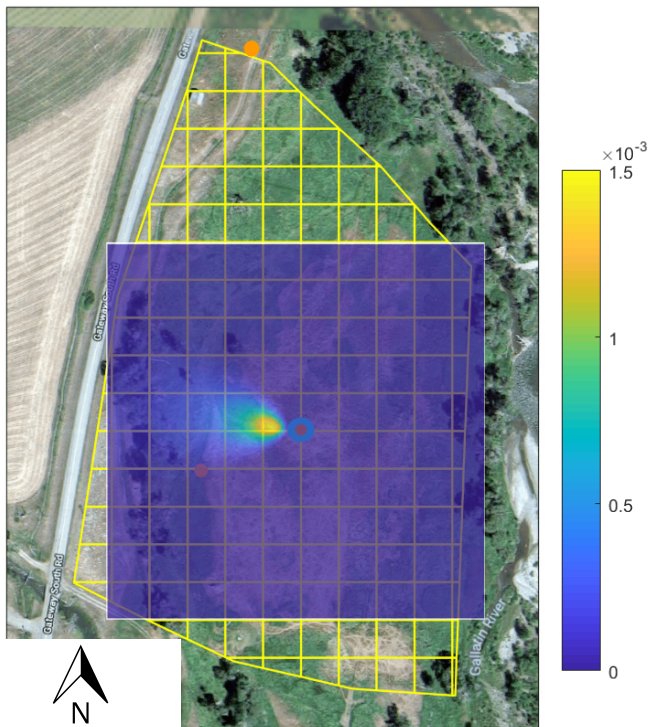
**Figure 1:** The study site near Gallatin Gateway, MT (45.557, -111.229). Bison locations are mapped within the 20-meter grid here superimposed in yellow. The tower location is in cyan and game camera locations are indicated in orange. Background image: Google, Maxar Technologies and the USDA Farm Service Agency ©2018.



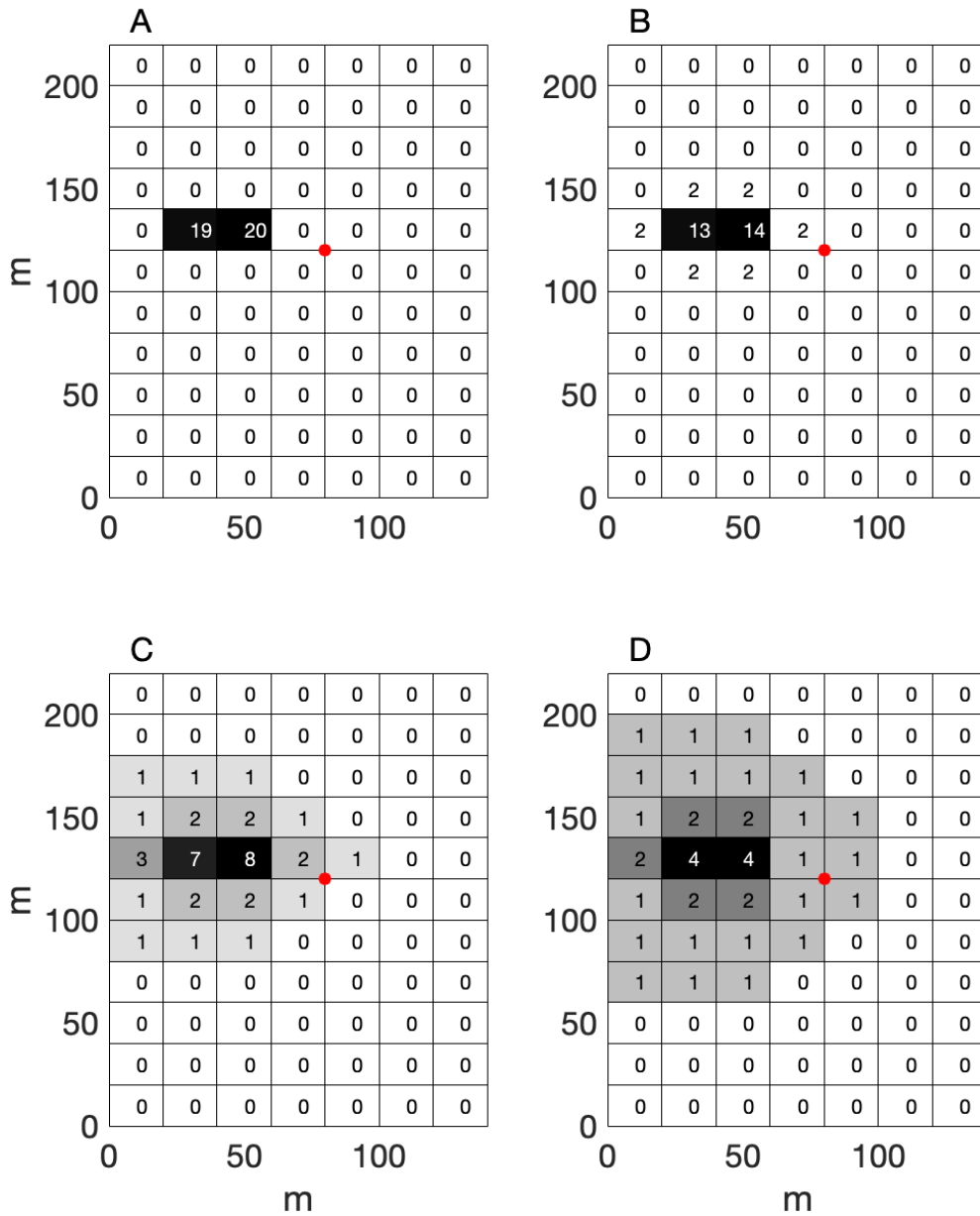
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Figure 2: A sample image of bison as viewed from the south-facing time-lapse camera located to the north of the study area (Figure 1). The eddy covariance installation is visible toward the center of the study site.

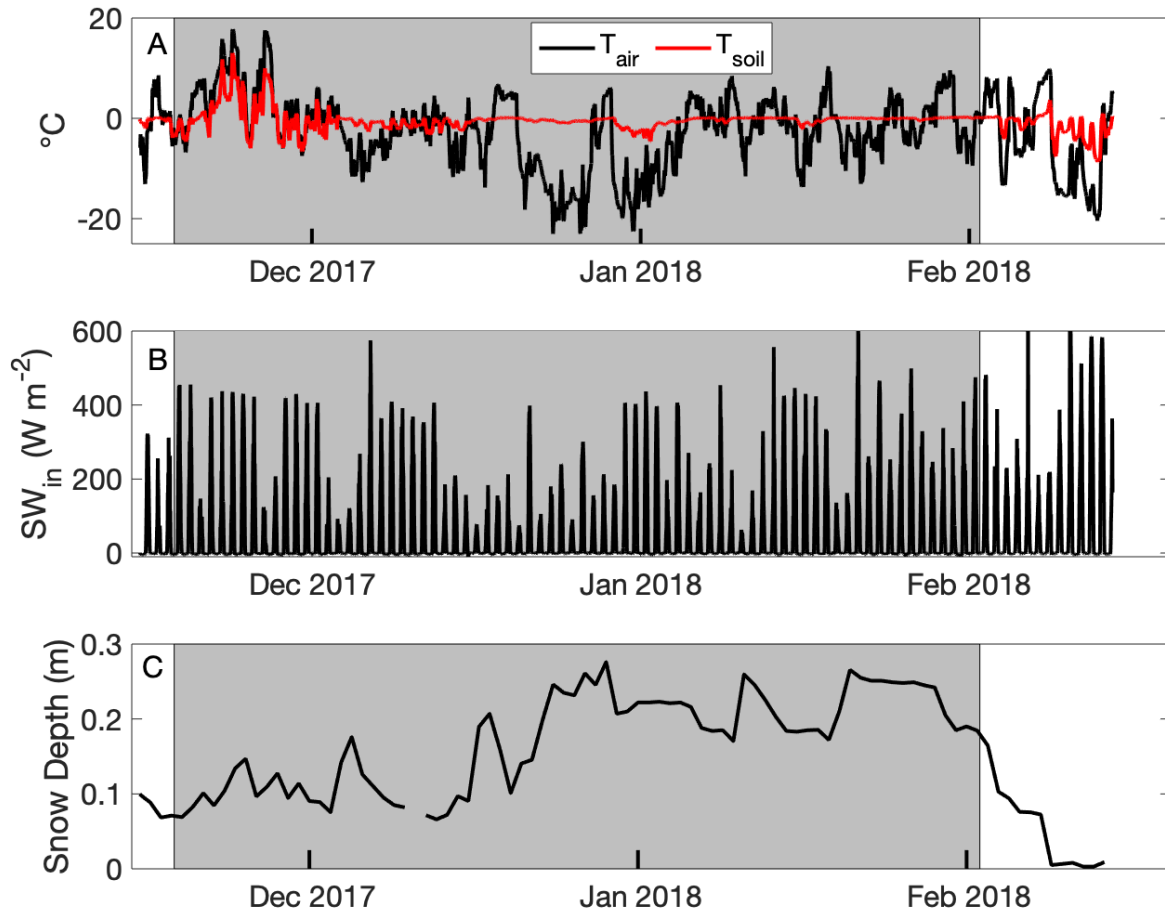




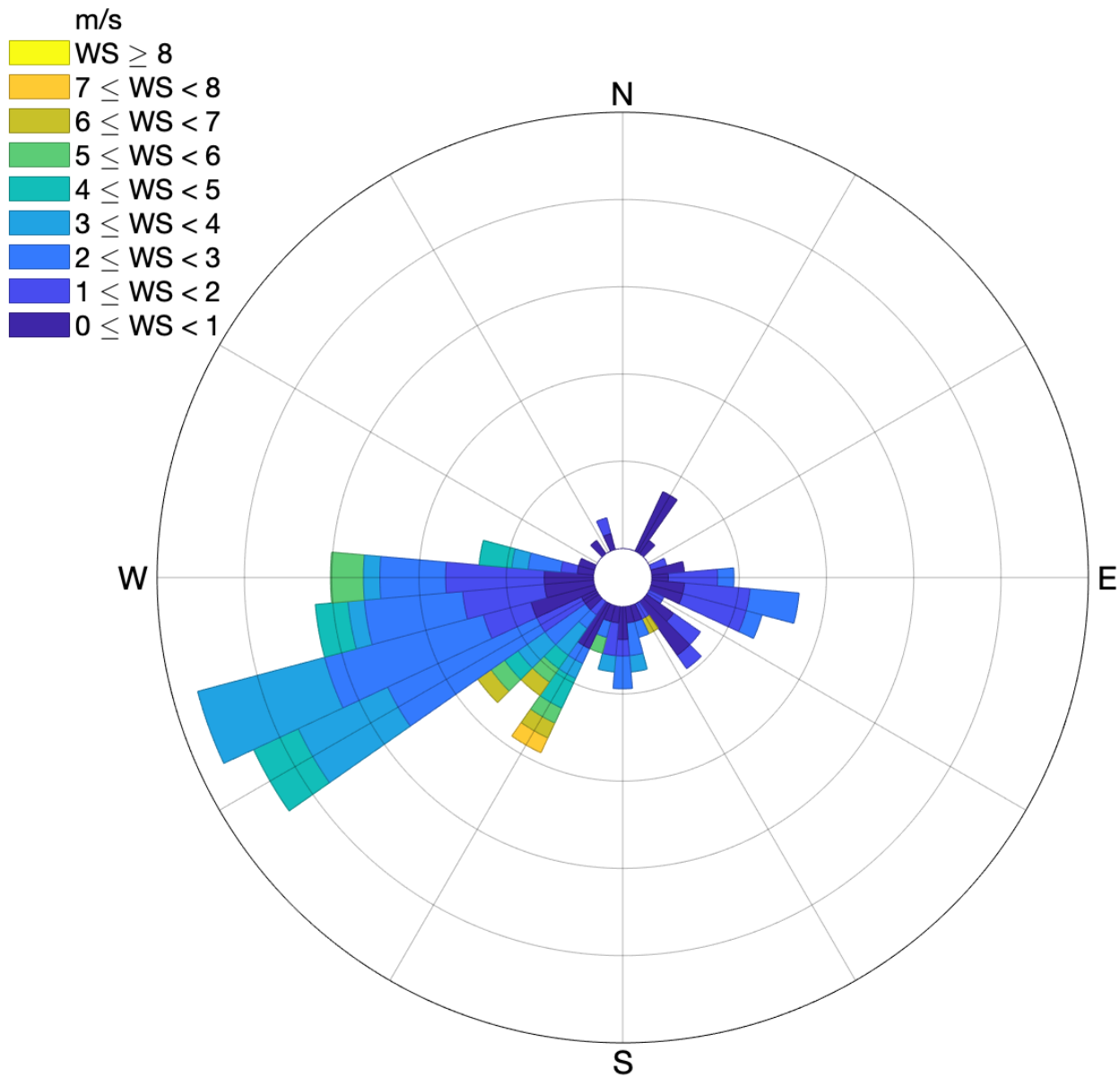
610 **Figure 3: An eddy covariance flux footprint calculated following Hsieh et al. (2000) and Detto and Katul (2006) at 1 m resolution for a single 30-minute interval superimposed on the study field (Figure 1). The fraction of the footprint in each grid box is shown in the legend is summed for each 20 m pixel to calculate the contribution of each pixel to the total flux. Background image: Google, Maxar Technologies and the USDA Farm Service Agency ©2018.**



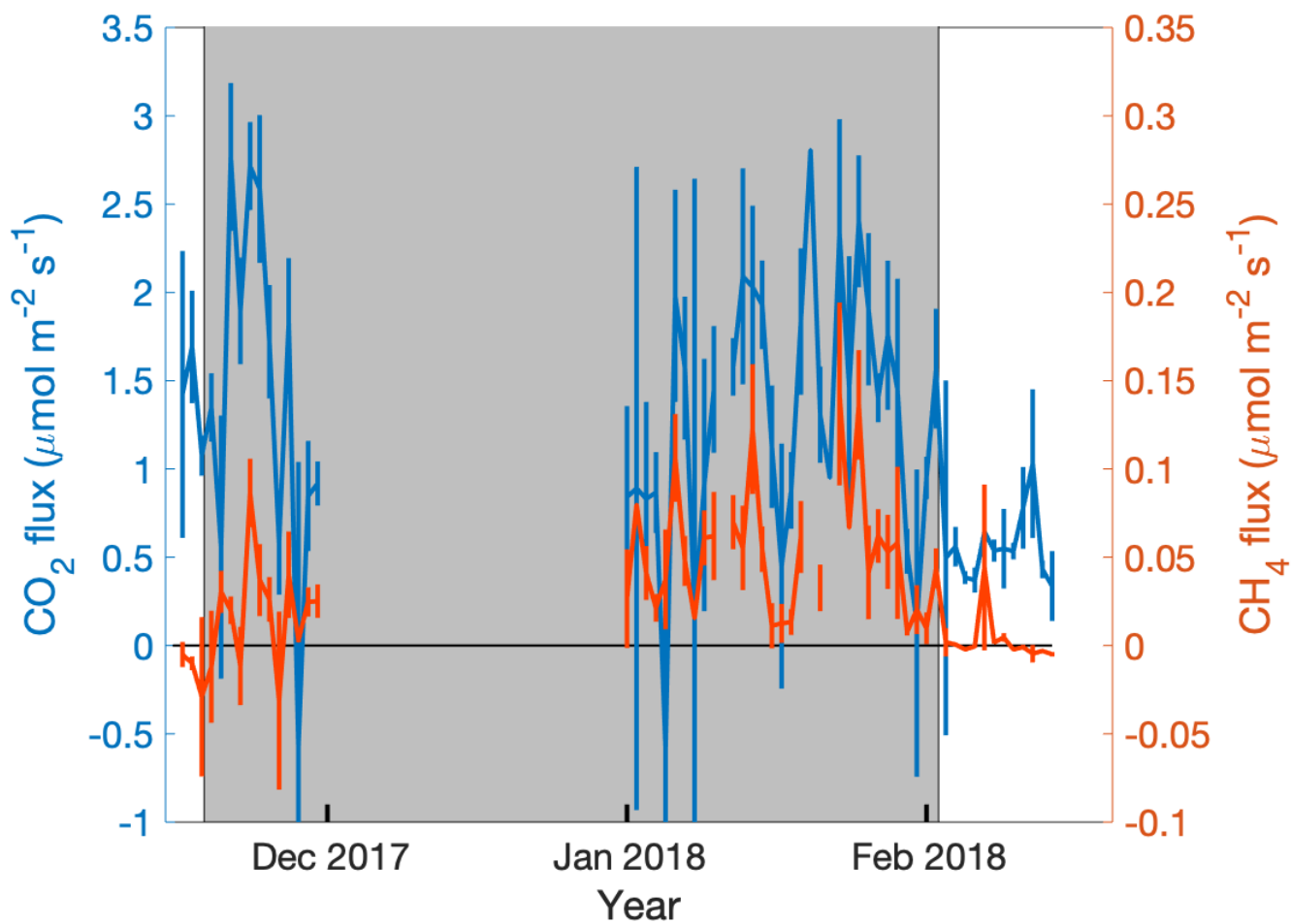
615 **Figure 4: The number of bison observed per grid cell at 16:30 on January 11, 2017 (A) and a distributed bison location map generated using two-dimensional Tikhonov Regularization with a Lagrange multiplier of 0.1 (B), 1 (C), and 4 (D).**



620 **Figure 5: (A) Air ( $T_{\text{air}}$ ) and soil temperature ( $T_{\text{soil}}$ ), (B) incident shortwave radiation ( $\text{SW}_{\text{in}}$ ), and (C) snow depth from a micrometeorological tower enclosed within an electric fence on a bison pasture near Gallatin Gateway, Montana, USA. Bison were present in the pasture during the interval bounded by the grey background.**

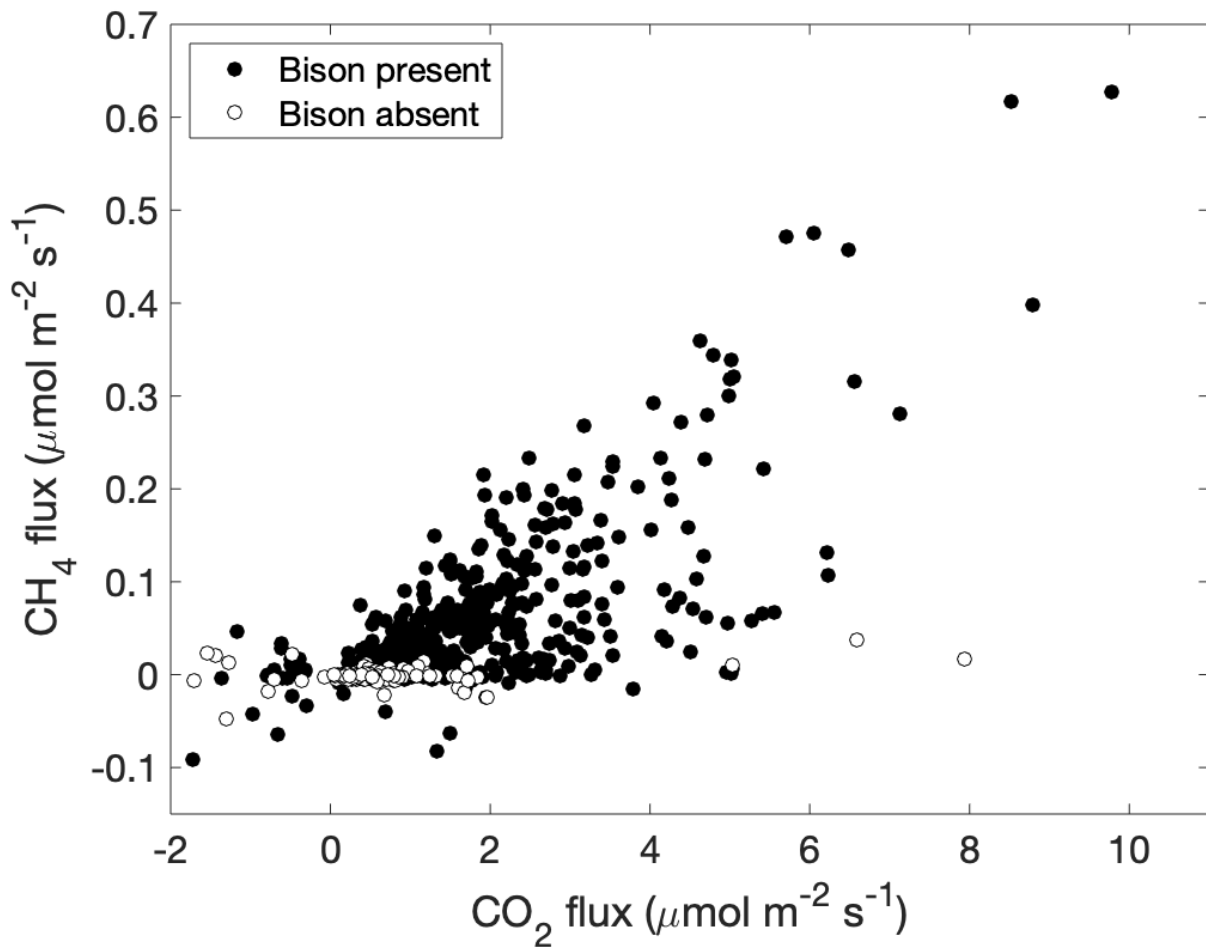


**Figure 6: A wind rose following Pereira (2020) for periods when eddy covariance measurements and bison location measurements were available. WS: wind speed.**



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Figure 7: The daily mean and standard error carbon dioxide and methane fluxes with standard error during daytime hours (0700-1700) from a pasture near Gallatin Gateway, MT, USA. The gray background denotes the interval during which bison were present on the study site.



630 Figure 8: The relationship between carbon dioxide and methane fluxes from the study pasture is shown for periods when bison were present (filled circles) and when bison were absent (open circles).

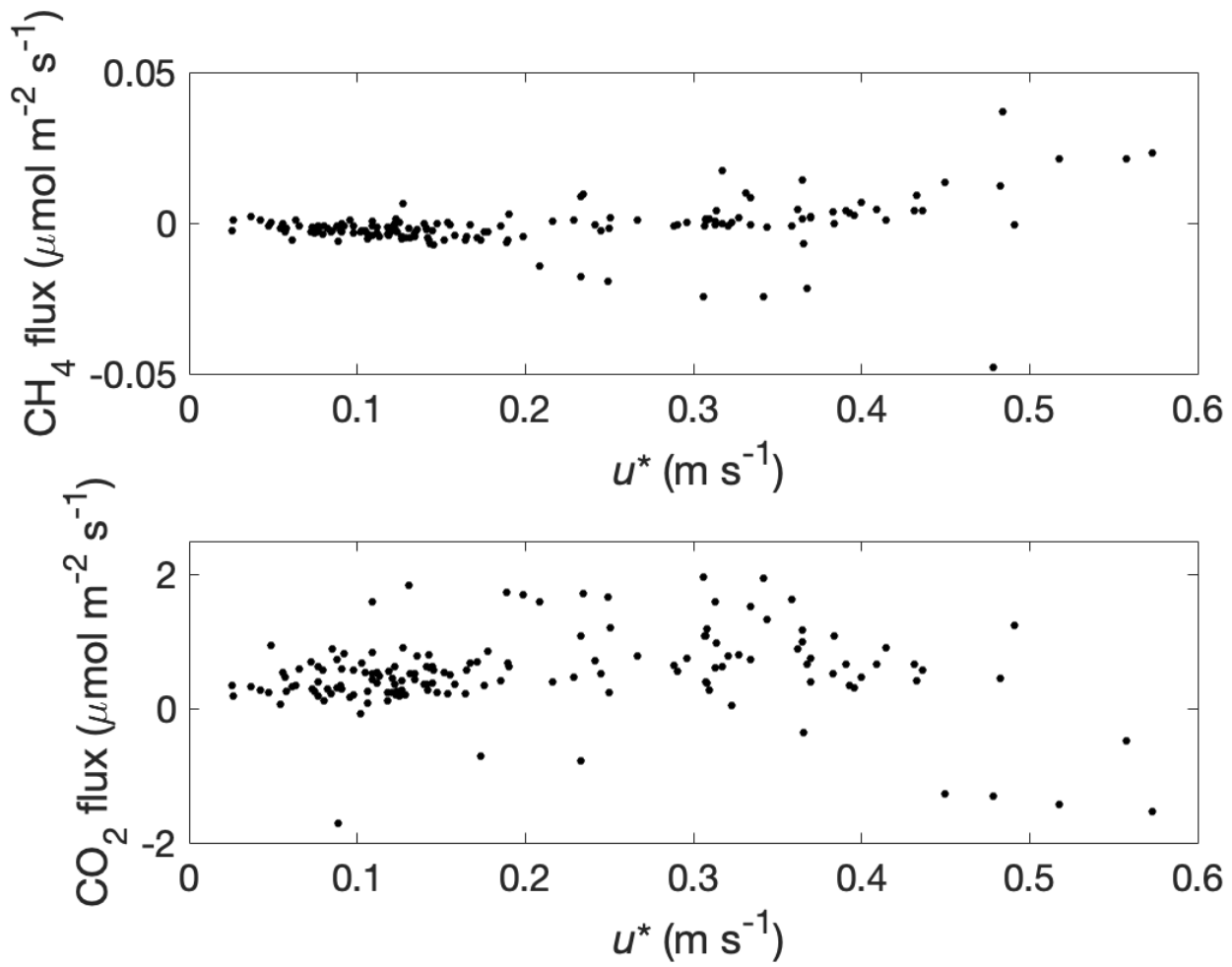
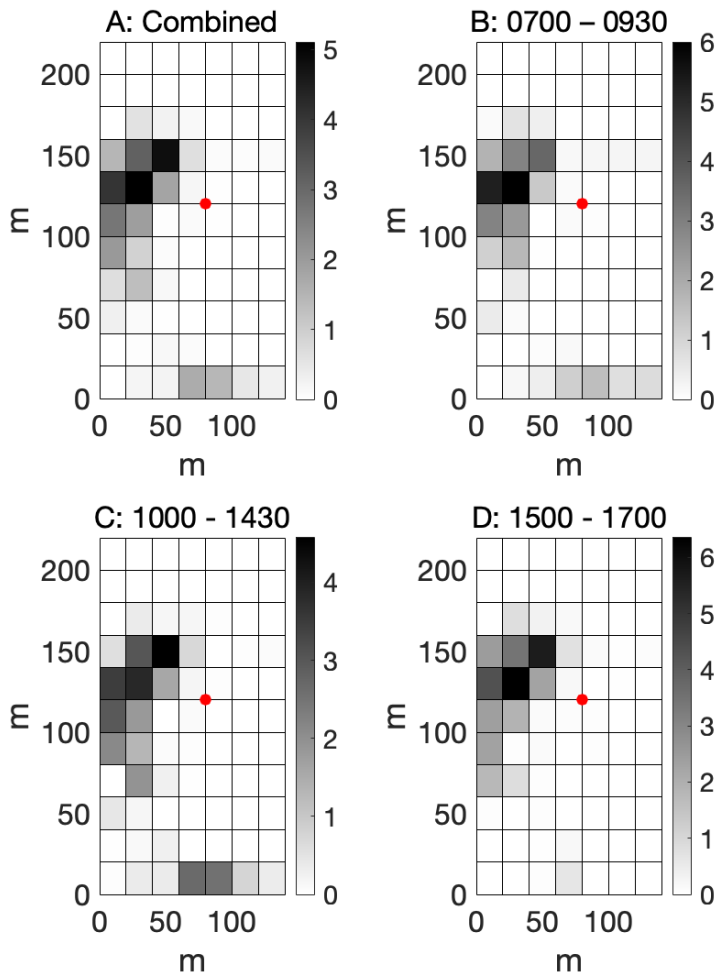


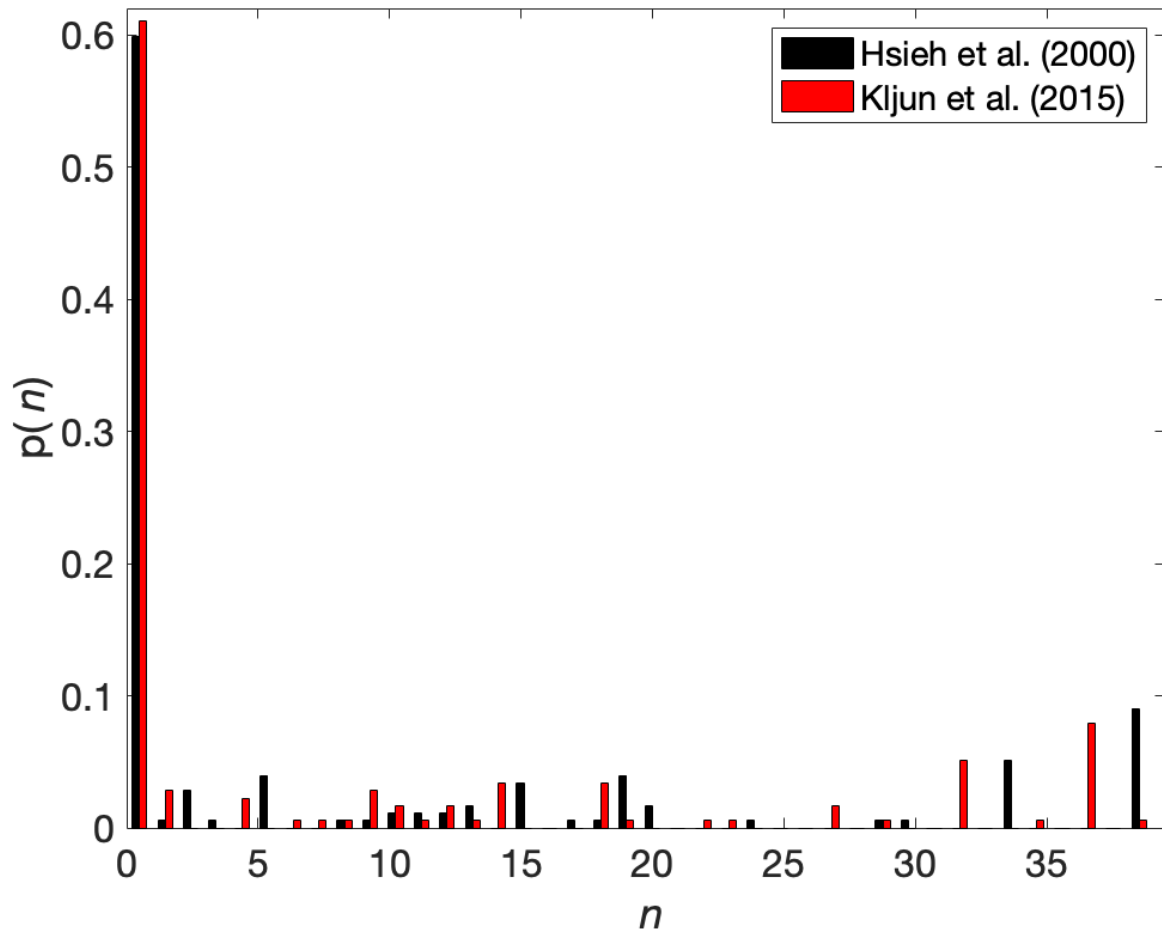
Figure 9: Methane (A) and carbon dioxide (B) fluxes as a function of friction velocity ( $u^*$ ) when bison were absent from the study pasture.



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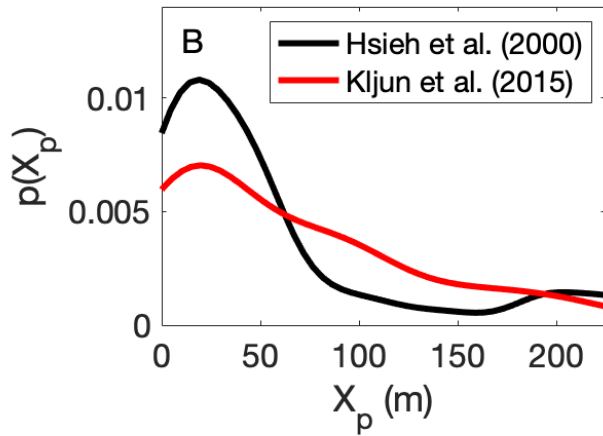
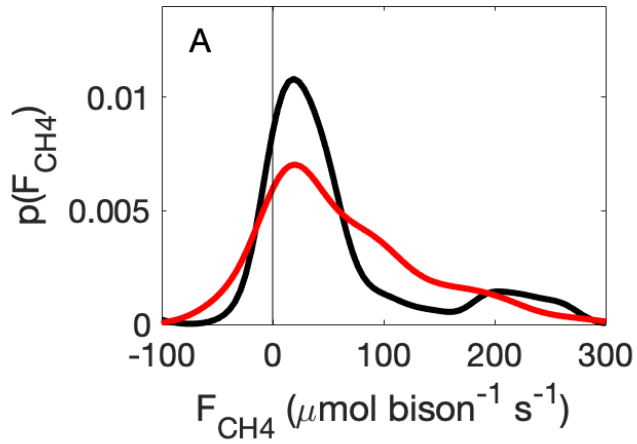
Figure 10: Average proportional bison density for three periods of the day. Each colored pixel represents a 20-meter grid square, red dots denote the location of the eddy covariance tower, and subplot titles refer to local time. Color denotes average proportion of bison present in each grid cell for the 39-animal herd.



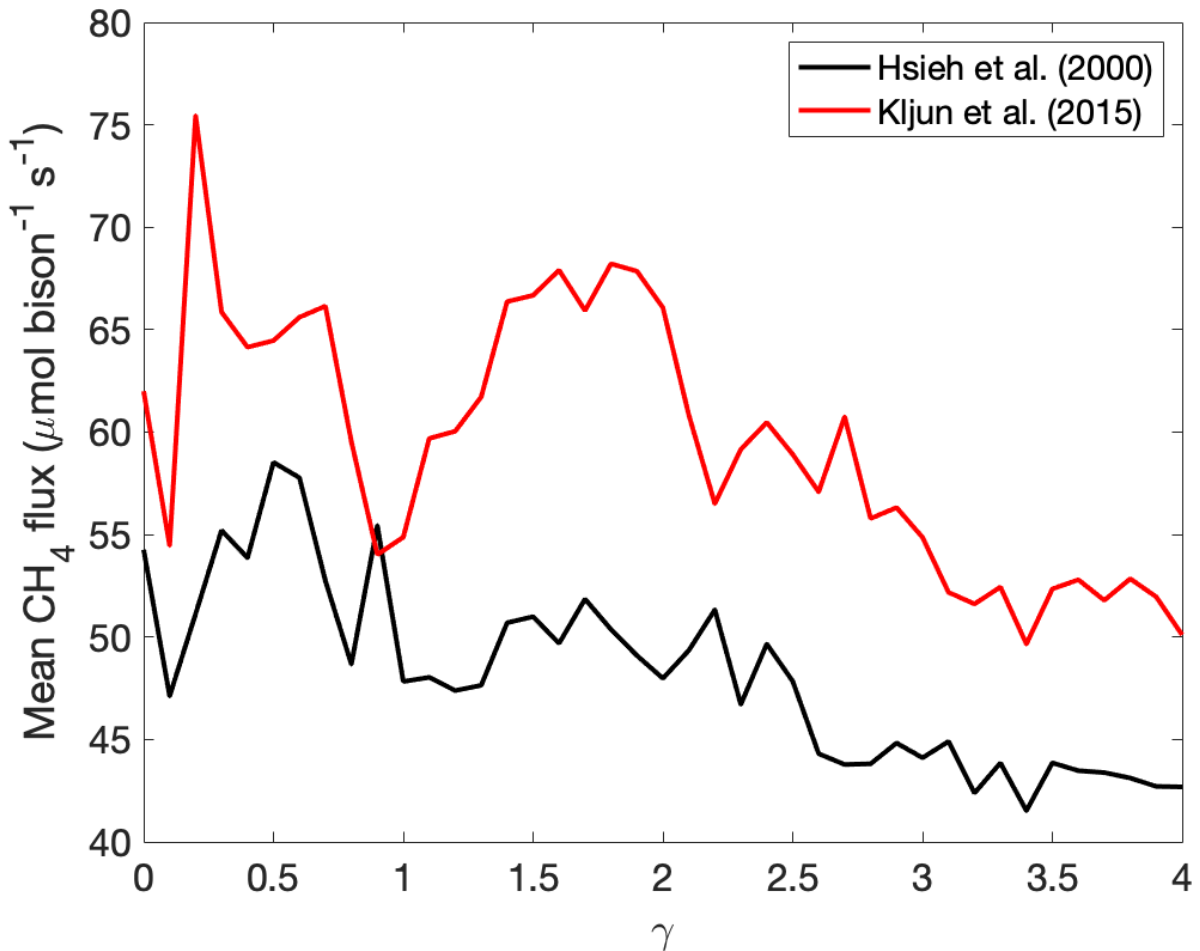


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Figure 11: The probability ( $p(n)$ ) of the number of bison ( $n$ ) in the 90% flux footprint for the Hsieh et al. (2000) and Kljun et al. (2015) footprint models for periods when flux measurements were made and camera imagery was available.



645 **Figure 11: Kernel density estimates of the distribution ( $p$ ) of (A) methane efflux ( $F_{\text{CH}_4}$ ) on a per-bison basis and (B) the peak ( $X_p$ ) of the source-weight function for half-hourly flux footprints derived by the Hsieh et al. (2000) and Kljun et al. (2015) flux footprint models.**



650 Figure 12: Mean methane efflux on a per-bison basis as a function of spatial smoothing of bison location estimates using the two-dimensional Tikhonov Regularization approach described in Stoy and Quaife (2015) for different values of the Lagrange multiplier  $\gamma$  and the footprint models of Hsieh et al. (2000) and Kljun et al. (2015).