

Importance of the forest state in estimating biomass losses from tropical forests: combining dynamic forest models and remote sensing

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Abstract. Disturbances, such as extreme weather events, fires, floods, and biotic agents, can have strong impacts on the dynamics and structures of tropical forests. In the future, the intensity of disturbances will likely further increase, which may have more serious consequences for tropical forests than those we have already observed. Thus, quantifying aboveground biomass loss of forest stands due to ~~tree mortality~~ system mortality (hereafter biomass loss rate) is important for the estimation of the role of tropical forests in the global carbon cycle. So far, the long-term impacts of altered ~~tree mortality~~ system mortality on rates of biomass loss have been described little.

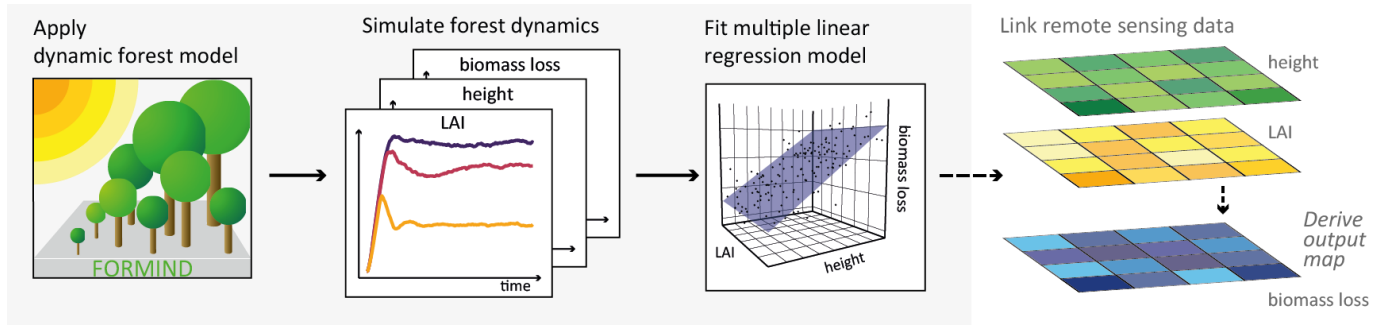
This study aims to analyse the consequences of long-term elevated ~~tree mortality~~ system mortality rates on forest dynamics and ~~biomass loss~~ biomass loss rate. We applied an individual-based forest model and investigated the impacts of permanently increased ~~tree mortality~~ system mortality rates on the growth dynamics of humid, terra firme forests in French Guiana. Here, we focused on biomass, leaf area index (LAI), forest height, productivity, forest age, quadratic mean stem diameter, and ~~biomass loss~~ biomass loss rate. Based on the simulation ~~s data~~, we developed a multiple linear regression model to estimate biomass loss ~~rates~~ es of forests in different successional states from the various forest attributes.

The findings of our simulation study indicated that increased ~~tree mortality~~ system mortality altered the succession patterns of forests in favour of fast-growing species, which ~~changed~~ increased the ~~old-growth~~ forests' gross primary production, though net primary production remained stable. ~~Tree mortality~~ The stem mortality intensity rate had a strong influence on the functional species composition and tree size distribution, which led to lower values in LAI, biomass, and forest height at the ecosystem level. We observed a strong influence of a change in ~~tree mortality~~ system mortality on ~~biomass loss~~ biomass loss rate. Assuming a doubling of ~~tree mortality~~ system mortality rate, ~~biomass loss~~ biomass loss rate increased (from 3.2% y⁻¹ to 4.5% y⁻¹) at equilibrium; avg. over simulated time 0—300. We also obtained a multidimensional relationship that allowed for the estimation of ~~biomass loss~~ biomass loss rates from forest height and LAI. Via an example, we applied this relationship to

remote sensing data of LAI and forest height and mapped ~~biomass loss~~ biomass loss rates for French Guiana. We estimated a mean biomass loss rate of 3.2% per year.

35 The approach described here provides a novel methodology for quantifying ~~biomass loss~~ biomass loss rates, taking the successional state of tropical forests into account. Quantifying biomass loss rates may help to reduce uncertainties in the analysis of the global carbon cycle.

Keywords. FORMIND forest model, MODIS MCD15A2H Version 6 LAI, Lidar, forest height map, French Guiana, biomass loss map, biomass turnover time



40 1. Introduction

Tropical forests represent an important pool in the global carbon cycle, as they store approximately 55% of the amount of global forest carbon (471 ± 93 PgC) in their living biomass (Pan et al., 2011). Intact tropical forests assimilate an average of 0.96 ± 0.46 PgC of carbon per year (Hubau et al., 2020). This carbon sink behaviour of tropical forests has considerably reduced the growth rate of atmospheric carbon dioxide (Friedlingstein et al., 2019; Le Quéré et al., 2016). However, the carbon
45 assimilation capacity of forests is affected by ~~tree mortality~~stem mortality due to disturbances, which can cause rapid, extensive carbon loss (Chambers et al., 2013; Fisher et al., 2008; Körner, 2003; Pugh et al., 2019; Seidl et al., 2014). Increased ~~tree mortality~~stem mortality due to disturbances has been related to a reduction in the carbon sink of tropical forests (Brienen et al., 2015; Hubau et al., 2020). A number of studies have discussed different climate-controlled mortality drivers, such as temperature (Clark et al., 2010), vapour pressure deficit (Trenberth et al., 2014), drought (Fauset et al., 2019; Phillips et al.,
50 2010), and wind-throw (Chambers et al., 2009; Magnabosco Marra et al., 2016; Marra et al., 2014; Negrón-Juárez et al., 2010, 2018; Rifai et al., 2016; Silvério et al., 2019). In addition, mechanical disturbances, such as insect calamities (Coley and Kursar, 2014), fires (Barlow et al., 2003; Brando et al., 2014; Slik et al., 2010), and lianas (Ingwell et al., 2010; Nepstad et al., 2007a; Wright et al., 2015), may also lead to increased ~~tree mortality~~stem mortality. The expected increase in the frequency and intensity of those disturbances may result in an overall increase in ~~tree mortality~~stem mortality and its associated
55 physiological mechanisms (McDowell et al., 2018). Higher levels of ~~tree mortality~~stem mortality thus present a major risk to climate mitigation efforts (e.g., REDD+: Reducing Emissions from Deforestation and Forest Degradation), as reductions in carbon assimilation rates and a decrease in the carbon stocks of tropical forests could counteract attempts to compensate for climate change (Gumpenberger et al., 2010; Körner, 2017; Le Page et al., 2013).

Mortality is a complex process because the causes leading to tree death can be diverse. Trees can die naturally from senescence
60 or from forest disturbances which may be abrupt or continuous and may have abiotic or biotic, allogenic or autogenic, as well as extrinsic or intrinsic causes (Franklin et al., 1987; McDowell et al., 2018). Furthermore, drivers of ~~tree mortality~~stem mortality often occur in combination, so the primary factors of death are not obvious (Franklin et al., 1987; McDowell et al., 2018). ~~Tree mortality~~Stem mortality leads to temporal changes in stand structure, tree species composition, and releases of resources, particularly biomass (Franklin et al., 1987; Hülsmann et al., 2018). Consequently, tree death affects important forest
65 growth processes, including tree growth and establishment, which are influenced by species-specific competition strategies (Snell et al., 2014) as well as by environmental and competitive factors such as light availability (Kuptz et al., 2010; Poorter, 1999; Uriarte et al., 2004). The influence of ~~tree mortality~~stem mortality on forest growth dynamics is determined by the disturbance intensity, which can range from the temporary loss of vitality to the mortality (Kindig and Stoddart, 2003) of individual trees, forest stands, and entire landscapes. Finally, ~~tree mortality~~stem mortality events are heterogeneously
70 distributed such that spatial patterns can be scattered or clustered (Franklin et al., 1987). Empirical studies have already

analysed the effects of short-term disturbances (i.e., intra-annual or over a few years) on increases in tropical ~~tree mortality~~[system mortality](#) (e.g., Barlow et al., 2003; Brando et al., 2014; Chambers et al., 2009, 2013; Doughty et al., 2015a; Holzwarth et al., 2013; Magnabosco Marra et al., 2016; Marra et al., 2014; McDowell et al., 2018; Negrón-Juárez et al., 2010, 2017; Nepstad et al., 2007b; Phillips and Brienen, 2017; Slik et al., 2010; Stovall et al., 2019; Wright et al., 2015). Nevertheless, using
75 empirical studies that are limited in space and time, it is difficult to quantify the long-term effects of permanently increased ~~tree mortality~~[system mortality](#) levels and to assess the consequences of such alterations on the dynamics, the structures, and the successional states of forests. Also, new remote sensing technologies offer enhanced potential for measuring the vertical and horizontal structures of forests at country to global scales (e.g., Bi et al., 2015; Hall et al., 2011; Lefsky et al., 2002, 2005; Myneni et al., 2015; Simard et al., 2011; Le Toan et al., 2011). Remote sensing products have previously been used for large-
80 scale identification of ~~tree mortality~~[system mortality](#) following disturbances (e.g., Pugh et al., 2019; Senf and Seidl, 2020); however, the estimation of ~~biomass loss~~[biomass loss rates](#) due to ~~tree mortality~~[system mortality](#) for forests at different states remains still uncertain.

In this context, individual-based forest gap models offer an approach by which to analyse forest dynamics (Botkin et al., 1972; Bugmann, 2001; Bugmann et al., 2019; Fischer et al., 2016; Shugart, 2002). Individual-based forest models are parameterised
85 with forest inventory data to allow for the investigation of forest growth dynamics over longer periods. By simulating the growth, establishment, mortality, and competition among trees within a forest, these models can contribute in estimating the biomass gain and loss of tropical forests (e.g., Hiltner et al., 2018, 2021; Maréchaux and Chave, 2017). As a result of gap formation after tree falling (Fischer et al., 2016; Huth et al., 1998), simulation areas consist of a mosaic of forest stands on which the vertical and horizontal structures and dynamics of forests in different successional states are modelled (Botkin, 1993;
90 Botkin et al., 1972; Bugmann, 2001; Fischer et al., 2016; Shugart, 1984). Structural state variables describing successional states of forests, such as tree size distributions and functional tree species compositions, play a major role in the estimation of the carbon budgets of forest stands and entire landscapes (Bohn and Huth, 2017; Fischer et al., 2018, 2019; Rödig et al., 2017, 2018, 2019; Rüger et al., 2020). Successional state variables of forests can be derived on large spatial scales (e.g., country to global levels) through a combination of individual-based forest gap modelling and remote sensing (Rödig et al., 2017, 2019;
95 Shugart et al., 2015, 2018), as this allows for a quantification of the spatial variation in forest structure due to ~~tree mortality~~[system mortality](#) (Rödig et al., 2017). The combination of individual-based forest gap models and remote sensing methods may also provide information on the spatial distribution of the annual rates of aboveground biomass loss due to ~~tree mortality~~[system mortality](#) (hereafter ~~biomass loss~~[biomass loss rate](#)).

The aims of this study are to investigate the impacts of permanently increased ~~tree mortality~~[system mortality](#) rates on forest
100 dynamics, to provide a framework for estimating ~~biomass loss~~[biomass loss rates](#) in terra firme forests at different successional states, and to derive a sample map of biomass loss for an entire country (i.e., French Guiana). This biomass loss map represents an application example to demonstrate the synergetic benefits of linking an artificial dataset derived from an individual-based forest model with remote sensing data. Here, we address the following research questions in detail:

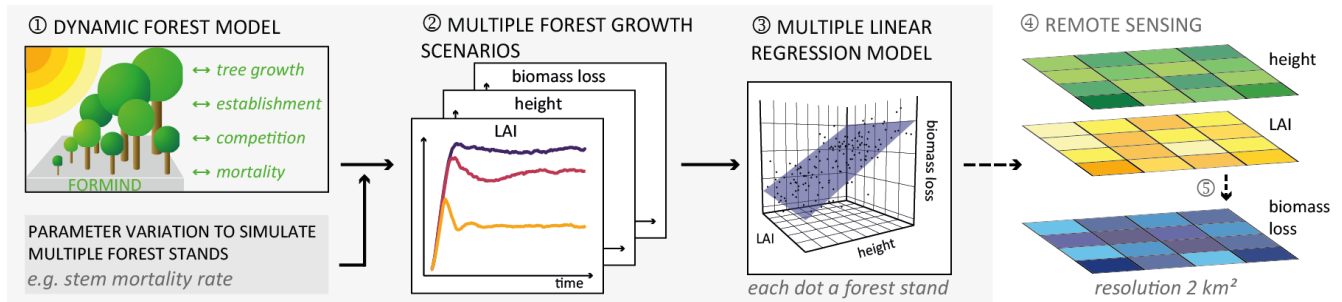
1. What are the consequences of permanently increased ~~tree mortality~~system mortality levels-rates on the dynamics of forest attributes (e.g., aboveground biomass, forest height, gross primary production, net primary production, leaf area index, quadratic mean stem diameter, mean forest age, and ~~biomass loss~~biomass loss rates) in tropical forests?
2. Can the ~~biomass loss~~biomass loss rates of tropical forests be estimated using various forest attributes that can also be derived from remote sensing?

2. Materials and Methods

110 We applied the ‘terra firme’ version of the dynamic individual-based forest model FORMIND (Fischer et al., 2016; Hiltner et al., 2018; Köhler and Huth, 2004) and simulated the effects of long-term increased ~~tree mortality~~system mortality levels on the dynamics of multiple forest attributes (Figure 1). This artificial dataset of forest dynamics covers a wide range of possible forest states such as the variability in tree species composition, successional state, and tree size distribution. We assume that we can use it to partially cover almost every state of forest stands in French Guiana (so-called forest factory approach, see

115 Bohn et al., 2017). We included aboveground biomass (hereafter biomass), mean forest height, gross primary production (GPP), net primary production (NPP), leaf area index (LAI), biomass turnover time (τ_B), quadratic mean stem diameter (QMD), mean forest age, and rate of biomass loss (m_{AGB}) in our assessment. We analysed all of these forest ~~stand~~-attributes in relation to the intensity of increased ~~tree mortality~~system mortality. Each simulated forest ~~stand~~-used in the analysis has the area of one hectare, with the forest states of each hectare differing from each other in each simulated time step and scenario. Then, we

120 developed a multiple linear regression model by testing different forest state attributes, such as LAI and forest height, as proxy variables. In addition, we derived a sample map for the biomass loss rate and biomass residence time of an entire region by using values for forest height and LAI obtained from satellite products. Simulated terra firme forests of French Guiana served as a case study.



125 **Figure 1:** Framework developed for estimating ~~biomass loss~~biomass loss rates by linking a dynamic forest model and remote sensing. 1. A forest model was applied to (2.) simulate the succession dynamics of forest stands of various forest attributes, such as LAI, forest height, and ~~biomass loss~~biomass loss rate, in a set of different ~~tree mortality~~system mortality scenarios (results used to answer research question 1). A simulated forest stand has the area of one hectare, with the forest states of each simulated hectare differing at each simulation time step and scenario. 3. Then, we developed a multiple linear regression model to the simulated forest states with LAI and forest height as proxy variables and ~~biomass loss~~biomass loss rate as the response (results to answer research

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question 2). 4. In addition, we applied the multiple linear regression model to remote sensing maps containing the values of the investigated forest attributes (LAI and forest height) to (5.) derive a sample map of biomass loss.

2.1 Study region

135 The study region is French Guiana, 95% of which is covered by humid, lowland terra firme forests (Hammond, 2005; Stach et al., 2009). These forests are characteristic for the Guiana Shield (Grau et al., 2017). The forests are generally species-rich, with an average of 150 to 200 tree species per hectare (Gourlet-Fleury et al., 2004), and are dense in biomass stock (Johnson et al., 2016; Rödiger et al., 2017; Saatchi et al., 2011).

2.2 Forest model FORMIND

2.2.1 Model description

140 To analyse the forest dynamics under the impacts of different levels of disturbance, we applied the ‘terra firme parameterisation’ of the forest model FORMIND v3.2 (Fischer et al., 2016) and took relevant parameter values from Hiltner et al. (2018), including tree growth, mortality, and establishment (see Supplements Tab. S1 and Tab. S2). FORMIND is an individual-based forest gap model that describes forest dynamics, tree growth, and changes in forest structures on a simulation area (1 hectare to multiple square kilometres) consisting of 20 m by 20 m patches that interact with each other (see Fig. S1),
145 where trees are not positioned explicitly. Our assumption in the ‘terra firme parameterisation’ of FORMIND is that forest stands also have no explicit position.

Every tree with a stem diameter at breast height (DBH) ≥ 0.1 m was simulated considering the following main processes at annual time steps: tree growth, establishment, mortality, and competition for light and space. The biomass gain of a tree results from the difference between photosynthetic production and respiratory losses (Fischer et al., 2016; Hiltner et al., 2018, 2021).

150 In the model, ~~tree mortality~~stem mortality is a key driver of forest dynamics. ~~Free mortality~~Stem mortality increases if the space for canopy expansion is limited, which depends on a tree’s position within the forest stand (self-thinning by crowding), whether tree growth is reduced (growth-dependent), and whether surrounding trees die after large trees fall (gap formation). Finally, each tree is subject to a ~~tree mortality~~stem mortality rate, which is stochastic. Here, we modified the ~~tree mortality~~stem mortality rate (eq. 1) to induce heterogeneity in the horizontal and vertical forest structures (i.e., tree size distribution and
155 functional species composition) of terra firme forests in French Guiana. Our study does not focus on short-term disturbances, but on the effects of long-term changes (> 100 years) in the intensity of ~~tree mortality~~stem mortality. Possible factors altering ~~tree mortality~~stem mortality rates in the forest model include environmental drivers, such as ~~extreme climate events, forest fires, wind throw, and diseases~~sustained elevated temperatures, -and reduced- altered precipitation shifts regimes, and reduced soil water availability. The ~~tree mortality~~stem mortality rate refers to individual trees at the stand level and is not stand-
160 replacing.

For the generic forest model parameterization of French Guiana's terra firme forests, tree species were classified into eight plant functional types (PFTs) according to species-specific traits, i.e., the maximum incremental rates of DBH and maximum tree height. We assume here that major parts of the terra firme forests can be characterised on the basis of three functional species groups: light-requiring species, species with intermediate light requirements, and shade-tolerant species. This functional species diversity is considered to be sufficient to capture forest succession dynamics in tropical forests (Fischer et al., 2018; Rödig et al., 2017; Rüger et al., 2020). Detailed model descriptions can be found in Fischer et al. (2016), in Hiltner et al. (2018), and online at www.formind.org.

2.2.2 Simulation settings

To investigate the effects of different ~~tree mortality~~ stem mortality intensities on the dynamics and the structure of terra firme forests, we developed seven simulation scenarios: a baseline scenario and six scenarios with permanently altered ~~tree mortality~~ stem mortality rates (Tab. 1). The baseline scenario based on observed ~~tree mortality~~ stem mortality rates (m_{bl}), ~~which-~~ ~~This~~ was computed by averaging the ~~species-~~ specific ~~background~~ ~~tree mortality~~ stem mortality rates of all PFTs (Tab. S2). To obtain ~~tree mortality~~ stem ~~background~~ mortality rates for ~~all-the~~ scenarios, the baseline's ~~tree mortality~~ stem ~~background~~ mortality rate (~~$m_{p,bl}$~~) was multiplied by a factor (f) for each scenario (sc), resulting in the following equation:

$$m_{p,sc} = f \cdot m_{p,bl}, \quad \text{with } f \in \left\{ \frac{1}{4}, \frac{1}{3}, \frac{1}{2}, 2, 3, 4 \right\}. \quad (1)$$

This resulted, ~~in combination with the other modes of modelled mortality (see Chap. 2.2.1),~~ in different ~~forest disturbance~~ stem mortality intensities ~~rates~~ at the stand level per simulated scenario (Tab. 1). The scenario with $f = 1$ represented the baseline scenario. In this study, we simulated forest stands with an area of one hectare and consisting of interacting patches on which the forests grow (patch size 20 m · 20 m; [see Fig. S1](#)). The one-hectare stands extended over a total simulation area of 16 hectare per scenario. Forest dynamics were computed in an annual time step that started in year 0 on bare ground and ended after 300 years. In the baseline scenario, a forest stand reached equilibrium after 210 years.

Table 1: Average ~~tree mortality~~ stem mortality ~~intensities~~ rate per simulation scenario ~~plus-and~~ specification (see eq. 1). ~~Background mortality rate $m_{p,sc}$ is the unweighted average over the mortality model-parameters of eight PFTs in the forest model. The resulting average stem mortality rate is is-the ---ratio of dead trees to the total number of trees in a simulated forest stand (averaged over the entire simulation time).~~

Factor (f)	Background mortality rate ($m_{p,sc}$) (y^{-1})	Resulting A <u>average forest</u> tree mortality <u>stem mortality</u> rate (m_{st}) (y^{-1})	Specification
1/4	<u>0.0032</u>	<u>0.0390</u> 0.003225	Low impact
1/3	<u>0.0043</u>	<u>0.0413</u> 0.0042957	
1/2	<u>0.0065</u>	<u>0.0432</u> 0.00645	
1	<u>0.0129</u>	<u>0.0523</u> 0.0129	Baseline
2	<u>0.0258</u>	<u>0.0723</u> 0.0258	
3	<u>0.0387</u>	<u>0.0845</u> 0.0387	

From the model outputs of all scenarios, we analysed the average development of multiple forest attributes (averaged over 16 hectares), such as aboveground biomass (AGB), LAI, forest height (mean height of the tallest three trees per 40 m · 40 m; Rödig et al., 2017; Simard et al., 2011), gross primary production, net primary production, quadratic mean stem diameter (square root of the sum of squared stem diameters per tree divided by the number of trees in a stand), mean forest age (arithmetic mean age of the 25 oldest trees per simulated 1 hectare area, selecting the oldest tree per patch), and [biomass loss](#) [biomass loss rate](#) (m_{AGB}), which we defined as annual proportion of dead biomass (AGB_{dead}) to total stand biomass (AGB_{total}):

$$m_{AGB} = AGB_{dead} \cdot AGB_{total}^{-1}. \quad (2)$$

In our forest model, we use a non-linear relationship between the AGB_t of tree (t) and its stem diameter (D_t):

$$AGB_t = \frac{\pi}{4} \cdot D_t^2 \cdot H_t \cdot F_t \cdot \frac{\rho_t}{\sigma_t}, \quad (3)$$

where H_t is the tree height, F_t is a form factor, ρ_t is the wood density, and σ_t is the fraction of aboveground biomass attributed to the stem (Fischer et al., 2016). Then, the tree biomasses are summed up to yield the total biomass, AGB_{total} . [We also calculated the stem mortality rate \(\$m_{SN}\$ \) as the ratio of the number of dead trees to the number of total trees in a forest stand at each simulation time step.](#)

In addition, we computed the [time period](#) over which each forest attribute reached the stable state (hereafter: equilibrium time) as well as the mean biomass turnover times (τ_B) with τ_B averaged over all successional states (simulated years 0 – 300). According to Carvalhais et al. (2014), turnover times can be defined as the ratio of the biomass stock to the flux (i.e., influx or outflux) of biomass. However, biomass outflux is not yet observable over large spatial scales (Thurner et al., 2016). Therefore, it was defined that biomass outflux equals biomass influx for forests in equilibrium (Carvalhais et al., 2014). Transferred to our study, the stock corresponds to the total biomass, influx to NPP, and outflux to dead biomass. Therefore, the following holds true for forests in equilibrium:

$$\tau_B = stock \cdot flux^{-1} = AGB_{total} \cdot NPP^{-1} = AGB_{total} \cdot AGB_{dead}^{-1}. \quad (4)$$

[In our model, NPP of a stand is the sum of NPP values of all trees within this stand. of a tree NPP is calculated as is the difference between GPP and autotrophic respiration \(it can be summed up from tree to stand level\).](#) The τ_B value can also be calculated from eqs. 2 and 4 as the reciprocal of the [biomass loss](#) [biomass loss rate](#):

$$\tau_B = 1 \cdot m_{AGB}^{-1} \quad \text{for } m_{AGB} > 0. \quad (5)$$

Using eq. 4, we calculated τ_B by taking forest succession into account.

2.3 Derivation of a multiple linear regression model to estimate ~~biomass loss~~biomass loss rates

215 To estimate the ~~biomass loss~~biomass loss rates, we analysed a number of forest simulations which produced a large number of different forest stands ~~each of (each~~-1 ha (~~averaged over~~ 16 ha simulation areas) in different successional states (per simulated year) with unique functional species compositions and tree size distributions. Thus, we generated a total of 33,600 terra firme forest stands. We assumed that the rate of biomass loss can be related to other forest attributes (e.g., biomass, LAI, forest productivity, and forest height). For a multiple linear regression model, the temporal and spatial components are not
220 important since forest states are considered independently of either. ~~We acknowledged that when fitting linear regression models, it is important that the proxy variables do not perfectly~~strongly correlate. We tested this by using a covariance matrix with the Pearson's correlation coefficient of the proxy variables (Tab. S4). Then, ~~W~~we tested different statistical models using different combinations of the proxy variables. The best estimate was made by a multiple linear regression model which describes variations in m_{AGB} as a function of two proxy variables, ~~LAI~~LAI, and forest height (Tab. S3). We estimated m_{AGB} (in
225 units of y^{-1}) as follows:

$$m_{AGB} = \beta_H \cdot H + \beta_L \cdot L + \varepsilon, \quad (6)$$

where H is the forest height, L is the LAI (-), ε is the error term, and β_i are the regression coefficients of the i^{th} forest attributes. The intercept was set to zero, as the ~~biomass loss~~biomass loss rate is expected to be zero when both LAI and forest height equal zero. Moreover,

230 To test whether the linear regression model is robust, we simulated additional scenarios with altered productivity rates. Based on this new data together with the previous data, we fitted an alternative multiple linear regression model. Similarity between both linear multiple regression models implied high robustness of the original model. For further information, see the supplemental material (Fig. ~~S8-S10~~ and Fig. ~~S9-S11~~).

2.4 Estimation of the country-wide ~~biomass loss~~biomass loss rates

235 2.4.1 Input maps

To estimate forest height, we used a global map in the WGS-84 geographical projection with a pixel size of approximately one kilometre (Simard et al., 2011; Fig. ~~S45~~.a). To create an LAI map, we used 139 data layers from the MCD15A2H Version 6 Moderate Resolution Imaging Spectroradiometer (MODIS) Level 4 with a pixel size of 500 m and averaged the LAI values between 2004-01-31 and 2006-12-31 to reduce the overall LAI variance (Myneni et al., 2015). We harmonised and stacked
240 the two input maps by first projecting the LAI map onto the coordinate reference system of the forest height map using the Geospatial Data Abstraction Library for French Guiana (www.gdal.org). Resampling was conducted with the bilinear method. The spatial aggregation of the LAI map (Fig. ~~S54~~.b) was performed by calculating the mean values of pixels whose centres were within 1-kilometre cells of the forest height map.

2.4.2 Output maps

245 The biomass loss ~~ratio~~rate was estimated for each pixel by applying the multiple linear regression model (eq. 6) to the two input maps (see Fig. S54). ~~The values of the two satellite-derived input maps were harmonized to meet the assumptions of the simulation~~We compared the density distributions of both input data sets with the ranges of FORMIND's simulation results (LAI and forest height (H)). Then, we corrected the input maps against the simulated data. First, we reduced the remotely sensed LAI values by 0.4 following the rationale that no understory trees (DBH < 10 cm) are considered in the forest model. 250 Second, we subtracted 3.5 m from the remotely sensed height values because maximum tree height is constrained in FORMIND by height allometric functions (see parameter h_{max} in Tab. S2 and). ~~No correction factors were required for the extrapolations~~ (see Fig. S7). We chose correction values in a way that the most frequent simulated forest states were congruent with the most frequent remote sensing observations and used the corrected input maps for our workflow. The biomass loss ~~values~~rates were then averaged over a pixel size of 2 km². We simulated forest stands of one hectare with the forest model. 255 This fine resolution allowed us to scale up to the accuracy of remote sensing products. We used a resolution of 2 km for the final biomass loss map, although the input data are available in 0.5 km (LAI, Myneni et al., 2015) and 1 km (forest height, Simard et al., 2011). Our regression model estimated negative biomass loss ~~ratios~~rates for a small portion of pixels, which were excluded from the biomass loss map. This was mainly the case for pixels without forest cover (cf. Fig. S5), according to a land use map (see Fig. S6) published by Stach et al. (2009) ~~(see Fig. S6)~~. To create the biomass turnover time map, we 260 computed the reciprocal of each pixel for the biomass loss map (see eq. 4).

We tested the reliability of the mapped ~~biomass loss~~biomass loss rate in the underlying input maps for the LAI and forest height via a sensitivity analysis regarding variations of $\pm 30\%$ as compared to the original input maps. We examined two methods to quantify the sensitivity of our study results concerning the input data. In a first experiment we ~~changed~~modified the values of the two input ~~data~~maps by exactly +30% and -30% each (LAI and forest height). I, in a second experiment we 265 ~~changed~~sampled the input data for the LAI and forest height by randomly ~~values~~ using a uniform distribution. These values ranged between +30% and -30% of the original input values. This resulted in ~~four~~six new input maps (LAI_{+30%}, H_{+30%}, LAI_{-30%}, H_{-30%}, H_{random}, LAI_{random}). We then applied the multiple linear regression model (see eq. 6, ~~Tab. 2~~) to ~~all~~some possible pairwise combinations of these ~~four~~six new input maps (i.e., LAI_{+30%}-H_{+30%}, LAI_{-30%}-H_{-30%}, LAI_{+30%}-H_{-30%}, LAI_{-30%}-H_{+30%}, H_{random} - LAI_{random}) to obtain ~~four~~five uncertainty maps. We calculated Δm_{AGB} per pixel, which is defined as the difference between 270 each of these ~~four~~five new uncertainty maps and the biomass loss map obtained from the original input maps. Thus, Δm_{AGB} represents the variation in the rate of biomass loss given 30% variation in the input variables. Furthermore, we compared our biomass loss map with forest plot data, provided by Brienens et al. (2015), and with map data of Johnson et al. (2016). Please refer to the supplementary material for details on the computer software used in this study.

3 Results

275 3.1 Influence of increased ~~tree mortality~~ system mortality on forest succession dynamics

To analyse the influences of varying ~~tree mortality~~ system mortality intensities, we simulated succession dynamics, which were affected by competition among individual tree species belonging to species groups. Here, we show that successional stages can be differentiated based on the development of the total stand biomass (Fig. 2). After 40 years of forest succession, the simulated stand biomass peaked at $500 \text{ t}_{\text{ODM}} \text{ ha}^{-1}$ (ODM: organic dry matter). This peak in stand biomass was caused by a high GPP of the pioneer species ($\text{GPP}_{\text{pioneer}} = 83 \text{ t}_{\text{ODM}} \text{ ha}^{-1}$; Fig. S24.a). After the early successional stage (years 0 – 40), the stand biomass fell slightly until year 100 due to the rapidly declining pioneer biomass, while the biomasses of other species increased (mid successional stage; Fig. 2). After 100 years, the stand biomass stabilised at approximately $420 \text{ t}_{\text{ODM}} \text{ ha}^{-1}$ (average over years 100–300), while the functional species composition reached a steady state after only 210 years (Fig. 2). In the late successional stage (gap dynamics), climax species and species with intermediate light requirements fixed five times more carbon in biomass than pioneer species (NPP of baseline scenario; Fig. S2.b).

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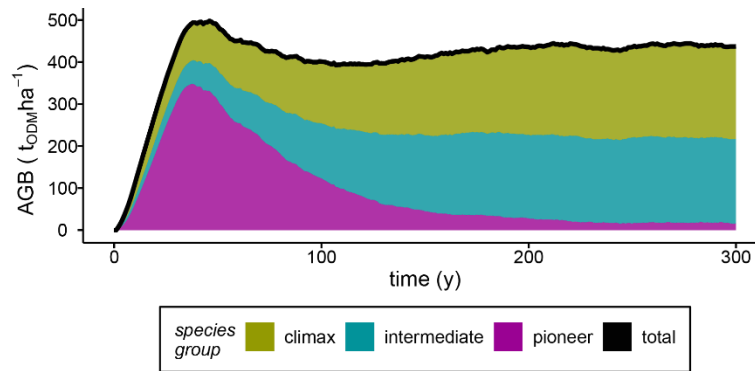
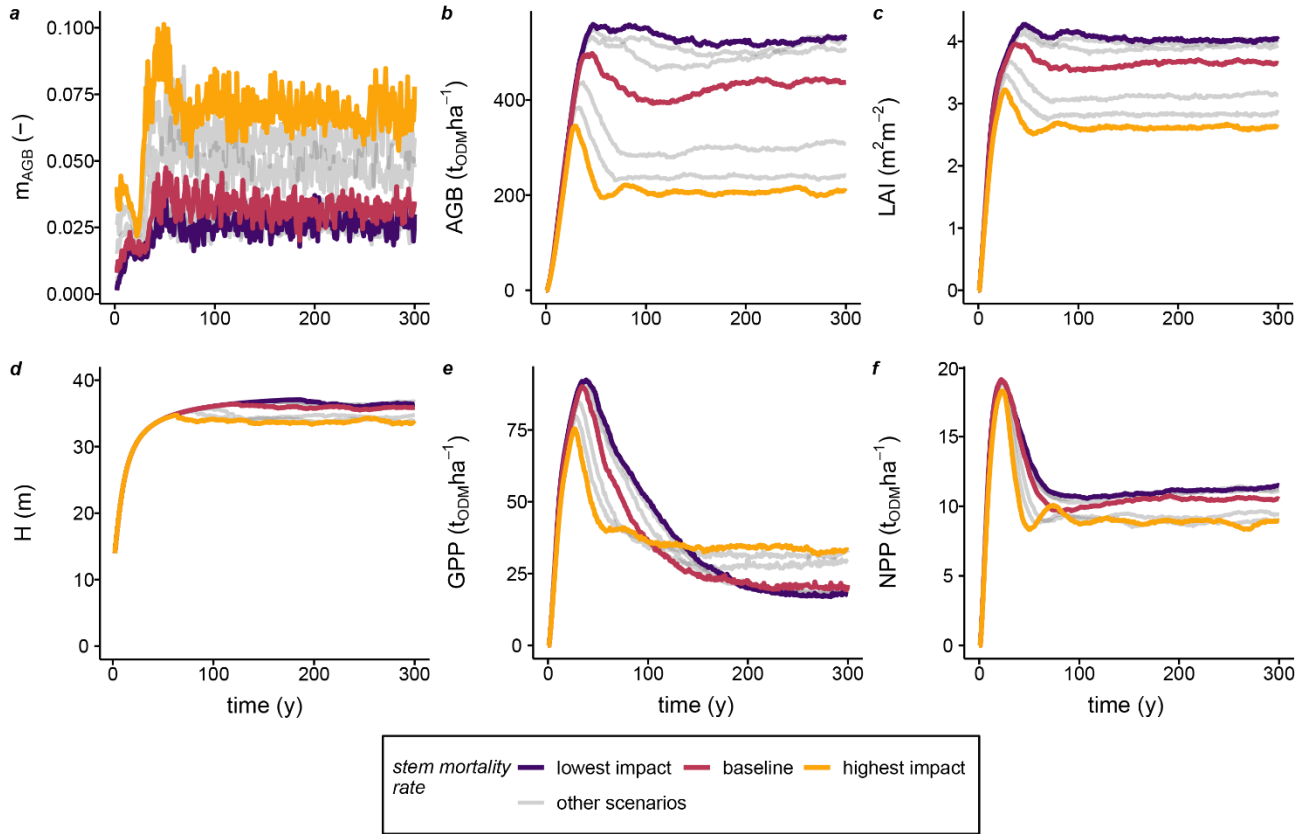


Figure 22: The baseline scenario's aboveground biomass (AGB) per species group and the total biomass for simulations of terra firme forests. (ODM: organic dry matter).

Our simulation results reveal a sensitive response of the biomass loss rate to increased ~~tree mortality~~ system mortality intensities (Fig. 3.a). At higher ~~tree mortality~~ system mortality levels, higher biomass loss with greater variance emerged and the biomass loss rate was on average greater than the stem mortality rate (see Fig. S4). At the level with the highest ~~tree mortality~~ system mortality, a peak in the biomass loss rates occurred at approximately 0.12 y^{-1} during the early phase of forest succession before levelling off at a value of 0.08 y^{-1} in the steady state (Fig. 3.a). Due to the higher biomass loss rates, the light climate in the forest stand changed (Fig. 3.c). The pioneer species were able to establish quickly in forest gaps. Hence, the GPP of the pioneer species was highest among all species groups (Fig. S2.a), which also affected the productivity of the total stand (Fig. 3.e – 3f, Fig. 4.a). Despite the distinctly higher GPP values obtained in the case of higher biomass loss rates, NPP did not change distinctly among the different scenarios (Fig. 3.f). Thus, the ~~tree mortality~~ system mortality intensity-rate had a strong influence on the species composition (e.g., higher pioneer GPPs; Fig. S2.a) and forest structure (e.g., QMD and mean stand age; Fig.

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S3), which led to lower LAI, biomass, and mean forest height values at the ecosystem level than those of the reference (Fig. 300 3.b – 3.d). In addition, structural changes gave rise to modified forest stand dynamics, with unique succession patterns depending on the intensity of the disturbance due to [tree mortality](#) [system mortality](#).

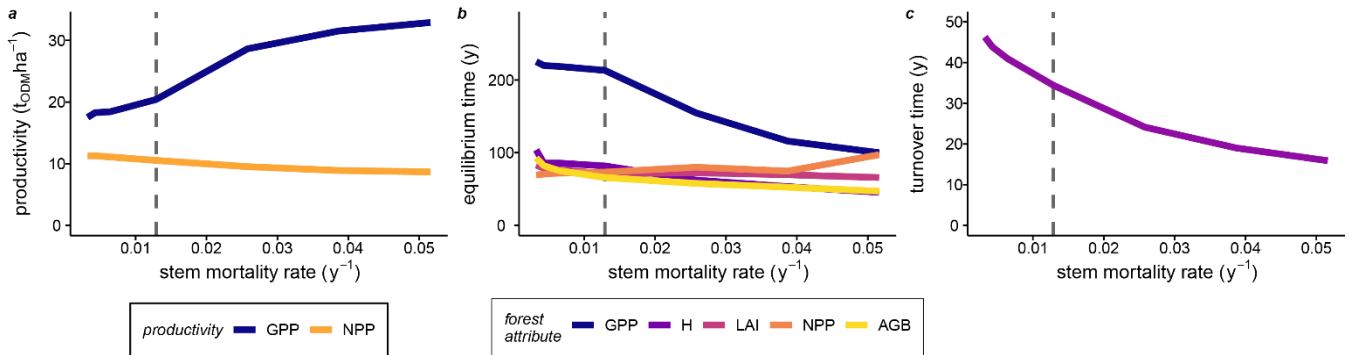


305 **Figure 33:** Simulation results of the (a) [biomass loss](#) [biomass loss rate](#) (m_{AGB}), (b) aboveground biomass (AGB), (c) leaf area index (LAI), (d) mean forest height (H), (e) gross primary production (GPP), and (f) net primary production (NPP) of terra firme forest stands under different [tree mortality](#) [system mortality](#) intensities. Grey lines indicate the entire set of scenarios under varying [tree mortality](#) [system mortality](#) rates (eq. 1) (ODM: organic dry matter; for further details see [Tab. 1](#), Fig. S2 and Fig. S3).

Furthermore, we analysed how the [tree mortality](#) [system mortality](#) level affected the time needed to reach equilibrium (Fig. 4.b). GPP responded particularly sensitively and inversely proportionally to the [tree mortality](#) [system mortality](#) [intensity](#) [rate](#), showing 310 a strong decrease with rising [tree mortality](#) [system mortality](#) levels. In contrast, other forest attributes, such as biomass and NPP, had altogether shorter equilibrium times than that of GPP, responding inversely proportionally to the [tree mortality](#) [system mortality](#) [mortality](#) level.

Finally, we evaluated the effect of increasing [tree mortality](#) [system mortality](#) rates on the turnover time of biomass (eq. 5) in the forest stands while taking forest succession into account (Fig. 4.c). The biomass turnover time τ_B was more than halved at a

315 four-fold higher ~~tree mortality~~~~system mortality~~ ~~intensity~~~~rate~~ as compared to the baseline ($\tau_{B,(f=1)} = 34$ y, $sd_{B,(f=1)} = 12$ y; $\tau_{B,(f=4)} = 46$ y, $sd_{B,(f=4)} = 43$ y; $\tau_{B,(f=1/4)} = 15$ y, $sd_{B,(f=1/4)} = 5$ y). Important forest properties are profoundly affected if the functional species composition, tree size distribution, or forest dynamics are changed.



320 **Figure 44: Influence of ~~tree mortality~~~~system mortality~~ levels in simulated terra firme forests on (a) mature forests' mean GPP and mean NPP (averages over simulation years 250 – 300), (b) the time until the forest attributes reached equilibrium states, and (c) the mean biomass turnover times, represented as the reciprocal values of ~~the biomass loss~~ ~~rate~~ (cf. eq. 5; averages over years 0–300). Dashed lines indicate the baseline scenario (GPP: gross primary production, NPP: net primary production, H: forest height, LAI: leaf area index, AGB: aboveground biomass, ODM: organic dry matter).**

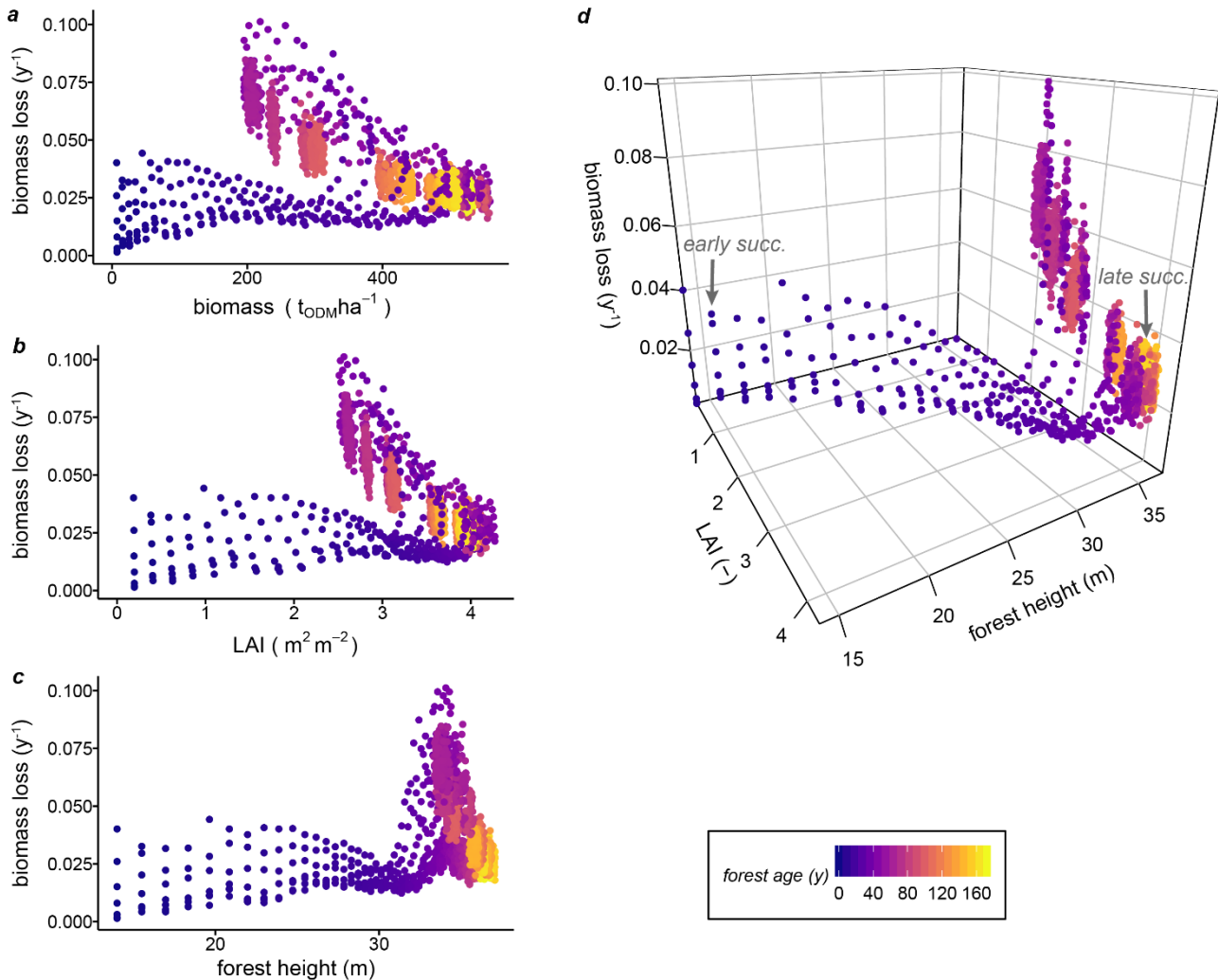
325 3.2 Estimation of ~~biomass loss~~~~biomass loss rates~~ using single and multiple forest attributes

In a further analysis, we assessed how ~~biomass loss~~~~biomass loss rates~~ can be derived from different proxy variables, such as the mean forest height and LAI. Including forests at different successional states, we tested the relationships between several single forest attributes and ~~the rate of~~ biomass loss but did not find distinct relationships (Fig. 5.a – 5.c; Tab. S3: regression model types 2 – 6). The ~~biomass loss~~~~biomass loss rates~~ showed a widely scattered range of values and thus unclear relationships to all single forest attributes during the early successional stage (forest age < 20 years; Fig. 5.a – 5.c). For instance, the LAI values of less-disturbed, old-growth forests (i.e., LAI = 4 during the gap dynamics stage ~~of mature forests~~) indicated similar biomass losses to forests in the early stages of succession. Relationships between single forest attributes and ~~biomass loss~~~~biomass loss rate~~ ~~can be imagined~~ seemed to be statistically significant (Tab. S3: regression model types 2 – 6), however, this was not useful because the relationships are clearly not linear (Fig. 5.a – 5.c). ~~for forests older than 20 years, but the regression statistics were not convincing (see Tab. S3: regression model types 2 – 6).~~ The relationships are strongly influenced by forest age and ~~tree mortality~~~~intensity~~~~system mortality rate~~. The linear regression models using only one proxy variable already show high significance (Tab. S3: Model types 2 – 6: p value < 0.01). For instance, the linear model of forest height correlates positively with biomass loss (Tab. S3: Model type 2), which is plausible since forest height is an indicator of the successional stage (see Fig. 3.d). As forest height increased, more tall trees of climax species died, resulting in proportionally more deadwood biomass compared to forests in early successional stages, in which many pioneer species with lower biomass were

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present and the forest height was lower (see Fig. 3.d; Fig. S2). Also, the linear model with LAI as proxy variable showed high significance (Table S3: model type 3), yet LAI was positively correlated with biomass loss, which seems implausible. The LAI is an indicator of disturbance intensity due to tree mortality. As tree mortality increased, the LAI and mean forest age decreased (Fig. 3.c; Fig. S3.b). This suggests that model type 3 still includes effects that are not explained by LAI alone, making the expected negative correlation of LAI with biomass loss a false positive.



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Figure 5: Dependence of biomass loss rates from single attributes (a) biomass, (b) LAI, (c) forest height versus (d) the multiple forest attributes, LAI and forest height, of simulated forest states, including all tree mortality system mortality scenarios. Each dot represents a terra firme forest stand with a unique forest structure (i.e., tree size distribution and functional species composition). All simulated years are included in the analyses. The colours of the dots show the mean forest ages characterising the successional states (ODM: organic dry matter).

For forests older than 20 years, Fig. 5.d illustrates a three-dimensional relationship between LAI, forest height, and the rate of biomass loss. Only when combined in a multiple linear regression model did the LAI (L) and forest height (H) explain the biomass loss rates of forests at different successional states well ($R^2 = 0.94672$, $RMSE = 0.00999$, $p\text{-value} \leq 0.01$; Fig. 6; Tab. S3):

$$m_{AGB} = 0.004166 \cdot H - 0.030614 \cdot L + \varepsilon. \quad (7)$$

The LAI negatively influenced biomass loss rates, whereas forest height is positively correlated with biomass loss. The obtained residuals were normally distributed around the expected value ($E(m_{AGB}) = 0.0$; Fig. S56.b), and depending on the estimated biomass loss rates, the residuals were homoscedastic with almost no trend (Fig. S56.c). Specifically, the residuals attributed to the LAI contained no remaining trend for forest stands older than 40 years and the ones attributed to forest height contained only a little remaining trend, which is age-dependant (cf. Fig. S74). So we can prove that forest height and LAI can be used as proxy variables to estimate biomass losses.

The one-to-one comparison of biomass loss rates for the simulated forest stands estimated by the multiple linear regression model (cf. eq. 7) versus those simulated within the dynamic forest model fit well. However, the fitted multiple linear regression model cannot describe the biomass loss rates of all forest stands. There is a small number of stands with simulated low but predicted high biomass loss rates. These are forests of simulated age between 0 – 20 years which are also accompanied by relatively low values of LAI and forest height (cf. Fig. 5.d and Fig. S86). Other forest attributes, such as GPP and NPP, were not included in the multiple linear regression model because they did not improve the estimation of the biomass loss rates substantially (Tab. S3).

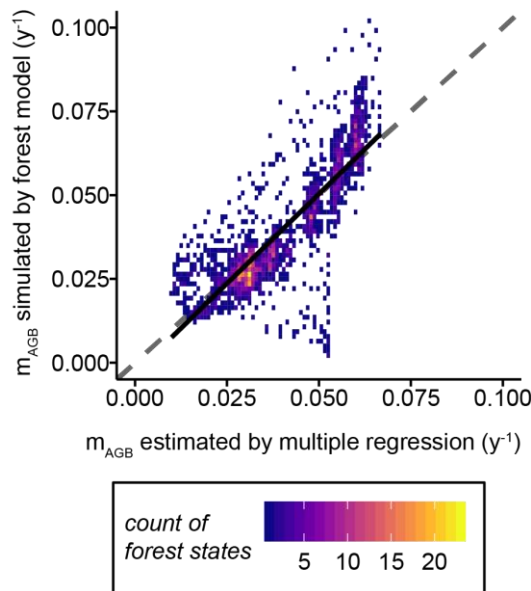


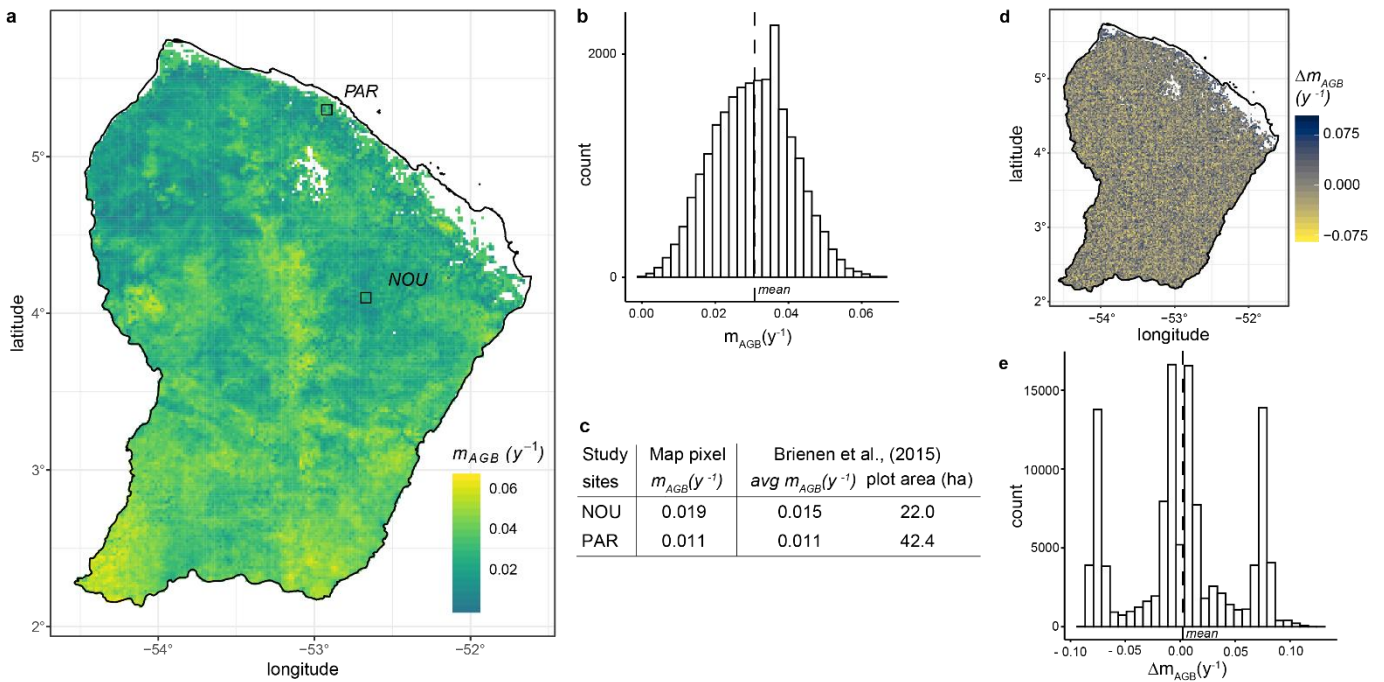
Figure 6: One-to-one density plot of biomass loss rates simulated by the dynamic forest model versus biomass loss rates estimated using a multiple linear regression model, with the forest height and leaf area index as proxy variables (eq. 7, Tab. S3). The

375 dashed line shows the line of perfect fit. Each dot represents a forest stand with a unique forest structure (i.e., tree size distribution and functional species composition) while the colours represent the density distributions of the combinations. The black solid line indicates the mean deviations of the ~~biomass loss~~biomass loss rate simulated with the forest model from the estimated ones ($m_{AGB,DFM} = 1.0688 \cdot m_{AGB,LM} - 0.0031 + \epsilon$, $R^2 = 0.6905$, $RMSE = 0.0099$, $p\text{-value} < 0.01$). (m_{AGB} : biomass loss ~~rate~~; ~~DFM~~: dynamic forest model; ~~LM~~: multiple linear regression model).

3.3 Estimation of biomass loss ~~rates~~ from remote sensing by deriving a sample map

380 By combining simulated forest states with the maps of LAI and forest height obtained via remote sensing (Myneni et al., 2015; Simard et al., 2011), we derived a biomass loss map for terra firme forests of French Guiana (Fig. 7.a). Based on this sample map, we obtained a mean biomass loss rate of 0.0312 y^{-1} (standard deviation of 0.01 y^{-1} ; Fig. 7.b). The values of ~~biomass loss~~biomass loss rate varied among regions, with higher values in the southern part of the country and lower rates in the northern part of the country. The highest biomass loss rates can be observed in the centre and at the southwestern country borders ($m_{AGB} > 0.06$). Such high values resulted from a combination of tall forest height together with low LAI values (Fig. S86). In the region surrounding the Paracou and Nourage sites, the biomass loss rates had values of 0.0124 y^{-1} and 0.0189 y^{-1} , respectively, which agree well with the mean biomass loss rates we derived from the empirical data of Brienen et al.'s (2015) study ($m_{AGB,Par} = 0.011 \text{ y}^{-1}$, $sd_{Par} = 0.127 \text{ y}^{-1}$; $m_{AGB,Nou} = 0.015 \text{ y}^{-1}$, $sd_{Nou} = 0.027 \text{ y}^{-1}$; Fig. 7.c). The sensitivity analysis revealed the dependence of the mapped biomass loss on the quality of the input data (Fig. 87.d, Fig. 7.e). The sensitivity is moderate (i.e., Δm_{AGB} is small) if the LAI and forest height change uniformly by a certain amount (Fig. S14X). If the changes in the LAI and forest height are contrary, the sensitivity of the mapped biomass loss is high (i.e., Δm_{AGB} is large), though we assume that a contrarian change in the input data rarely occurs (Fig. S14X). Considering forests at all successional states, we derived another sample map for biomass residence time (Fig. S12) calculated from the reciprocal of ~~the~~ biomass loss ~~rate~~ (eq. 5).

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395 **Figure 7: (a) Map of biomass loss for terra firme forests in French Guiana (~ 2 km resolution) and (b) its histogram. The dashed line in b indicates the estimated country-wide mean (3.21% with a standard deviation of 1.0%). The black squares in the map show the**
 400 **locations of forest plots at Paracou (PAR) and Nourage (NOU), from which census-data was used to compare estimated and field-based biomass loss values. (c) The census data originate from Brien et al., (2015). (d) Sensitivity analysis for the mapped biomass loss rates of terra firme forests in French Guiana (cf. Fig. 7.a) and (e) its histogram. The values for the input maps of leaf area index (LAI) and forest height were randomly sampled using a uniform distribution ranging between $\pm 30\%$ from their original values (cf. Fig. S4). Δm_{AGB} represents the variation in biomass loss rates given this variation in the input variables. For further results about the uncertainty analysis performed, see Fig. S7, Fig. S8, Fig S9, Fig. S10, Fig. S11, and Fig. S14.**

4 Discussion

4.1 Mechanism of tropical forests in dealing with the increasing intensity of ~~tree mortality~~ system mortality

In this study, we analysed dynamics in tropical forests in relation to ~~tree mortality~~ system mortality. We demonstrated that most of the analysed forest stand attributes (biomass, forest height, LAI, GPP, biomass loss rate, QMD) had specific responses during succession. Moreover, we showed that the rate of biomass loss is strongly affected by succession dynamics as well as by the ~~tree mortality~~ system mortality intensity rate. The period until the stand's equilibrium was reached differed in duration among each simulation scenario. Additionally, the mean turnover time of biomass, i.e., the reciprocal value of biomass loss rate (eq. 5), varied considerably.

There were multiple reasons for the unique succession patterns of each forest attribute. Succession dynamics are influenced by assimilation rates (e.g., photosynthesis rates, light requirements) and physiognomic characteristics (e.g., maximum stem diameter increment rates, maximum heights, and wood densities), both of which are specific to each species group (Hiltner et al., 2018). Functional traits are crucial in simulations of the succession dynamics in forests because they determine the competitiveness of species groups (Fischer et al., 2018; Rüger et al., 2020).

The relationship between successional stages and ~~tree mortality~~ stem mortality rate has been investigated in empirical studies to estimate mortality in tropical forests (Aubry-Kientz et al., 2013a; Chambers et al., 2013; Doughty et al., 2015b; Holzwarth et al., 2013). Aubry-Kientz et al. (2013) introduced a method that estimated the ~~tree mortality~~ stem mortality probability of terra firme forests at Paracou. Similar to our results, they found that the ~~tree mortality~~ stem mortality probability depends on the successional stages of the forests as well as on the functional traits of species, such as the specific leaf area, wood density, stem diameter increment, and potential height.

Interestingly, we observed similar NPP values at different ~~tree mortality~~ stem mortality levels for forests in equilibrium. Erb et al. (2016) argued that the NPP of vegetation is effectively independent of the ~~tree mortality~~ stem mortality intensity rate, which is ~~confirmed~~ supported by our results. The observed stability of NPP under different disturbance regimes can be explained by shifts within the functional species compositions and tree size distributions. Pioneer species, which typically have lower wood densities (Chave et al., 2009; Zanne et al., 2009) and lower potential heights than those of slow-growing climax and intermediate species (Hiltner et al., 2018), store less carbon in their living biomasses. Since pioneer species grow faster, they can bind as much carbon per time as slow-growing climax species. Therefore, at the forest stand level, higher ~~tree~~ stem mortality ~~levels~~ rates result in similar NPP values as those observed with lower ~~tree~~ stem mortality ~~levels~~ rates, although the individual trees show different growth behaviours. Our simulation results show how the carbon storage of forests in equilibrium changes across different levels of ~~tree mortality~~ stem mortality intensity rates, despite constant levels of NPP. Instead, our findings indicate that carbon storage depends on the functional species composition. At high ~~tree mortality~~ stem mortality rates (e.g., a high-impact scenario), more pioneer trees of a younger age were present in the forest stands. Thus, to achieve a high forest carbon storage capacity, there is a trade-off between large, old, and less productive trees (e.g., climax species) and smaller, younger, and more productive trees (e.g., pioneer species).

4.2 Performance of the regression model for estimating ~~biomass loss~~ biomass loss rates

One of the main findings of this study is that the simulated biomass loss rates of terra firme forests can be estimated using multiple linear relationships among several forest attributes. The premise was that all forest attributes used could be provided by remote sensing and could give information about the forest structure and productivity. We recognized the relationship between ~~biomass loss~~ biomass loss rates to LAI and forest height when fitting many different statistical models with different simulated forest attributes (Tab. S3). If ~~tree mortality~~ stem mortality rates increased, this led to a higher ~~biomass loss~~ biomass loss rates. However, it is impossible to directly infer ~~tree mortality~~ stem mortality rates from biomass loss rates because forest

structural state variables differed for each simulated forest stand, depending on its successional stage (Bohn and Huth, 2017; Rödig et al., 2017). For example, the stem number distribution of dying trees is not evenly distributed across tree size classes (Aubry-Kientz et al., 2013b; Holzwarth et al., 2013; Muller-Landau et al., 2006; Rowland et al., 2015).

In our approach for identifying appropriate forest attributes to infer ~~biomass loss~~biomass loss rates, we considered results from empirical studies that have investigated ~~tree mortality~~stem mortality in tropical forests (Aubry-Kientz et al., 2013b; Esquivel-Muelbert et al., 2019; Stovall et al., 2019). Esquivel-Muelbert et al. (2020) investigated the ~~tree mortality~~stem mortality of the Amazon by using empirical data to show that stem diameter growth rate and tree size are strong predictors. Fast-growing species with low wood densities are at a higher risk of mortality, whereas the effect of tree size varies. Aubry-Kientz et al. (2013) used functional traits, such as potential tree height and specific leaf area, to estimate the probability of ~~tree mortality~~stem mortality. Based on large-scale remote sensing observations, tree height was identified as an important predictor of ~~tree mortality~~stem mortality during drought, with large trees having twice the mortality rate of small trees, while environmental drivers (i.e., temperature, soil water, and competition) controlled the intensity of the height-mortality relationship (Stovall et al., 2019). The results of these studies underline the importance of productivity (e.g., increment rates and tree size), biomass, and functional characteristics (e.g., wood densities, potential stem diameter increment rates, leaf areas, and potential tree heights) of trees or tree species in the context of ~~tree mortality~~stem mortality. In our forest model, such characteristics are included in the derivation of ~~tree mortality~~stem mortality rates of specific PFTs (cf. Tab. S1, S2). In forest gap models, forest structural state variables, such as the stem number distribution, tree size distribution, and functional species composition, of the dying trees emerge, rather than get specified as input parameters (Botkin et al., 1972; Bugmann, 2001; Shugart, 2002). This fact is a useful model behaviour for estimating the biomass loss rate of simulated forest stands and, moreover, holds true for the derivation of other forest attributes, which we considered when fitting our regression model. Besides the LAI and forest height, we tested GPP, NPP, and biomass as proxy variables for the rate of biomass loss. On the forest stand level, however, these variables did not improve the performance of the multiple linear regression model substantially. These results suggest using forest height and LAI as proxy variables to estimate the ~~biomass loss~~biomass loss rates of forest stands. Despite the simplicity of the multiple linear regression model, meaning that we included only two proxy variables, its statistical performance proved to be robust (cf. eq. 6; Tab. S3, Fig. S65, ~~Fig S7~~Fig. S8, Fig. S9). Thus, it was possible to derive ~~biomass loss~~biomass loss rates from LAI and forest height of simulated data for forests in different successional states. It was important that the signs of the regression coefficients of our linear model plausibly reflected the relationships that were observed in the field. In the regression model, forest height was directly proportional, and LAI was indirectly proportional to the ~~biomass loss~~biomass loss rates of the forest stands. For example, tall forests with low LAI values resulted in high biomass loss rates (cf. Fig. S86). We would like to note that to further improve the linear regression model (e.g., further minimise the residual's remaining trend), additional proxy variables would need to should be included and effects of environmental factors on simulated forest states should be investigated in more detail. In this study, we tested various forest attributes available as

remote sensing products (cf. Tab. S3, Tab. S4). We decided to use the simplest possible linear regression model (in terms of the number of included proxies) that estimated biomass loss rates best.

Using a forest model to derive the relationships among different forest attributes has several advantages. First, the simulated LAI and forest height data were generated mechanistically, integrating a broad spectrum of information about forest dynamics and successional states emerging from different physiological processes. This can lead to a lower level of noise in the simulation data compared to that in the observed field data. Nevertheless, forest models also include stochastic processes, including ~~tree mortality~~system mortality rates and establishment (Bugmann, 2001; Fischer et al., 2016; Shugart, 2002). By using plant functional types to simulate forest dynamics, we reduced the possible uncertainties in species traits. Simplifications allow for a transferability of the regression analysis to forests with similar characteristics and succession states. These simplifications also enabled the estimation of the biomass loss rates of terra firme forests across the entirety of French Guiana. With the approach pursued here, it is possible to derive regression models for estimating ~~biomass loss~~biomass loss rates in other locations worldwide. Forest model simulation results contain structural information about the conditions of forests in different successional states, allowing the data to be used as training data for the development of statistical regression models. Whether LAI and forest height are also suitable as proxy variables of the biomass loss rates of other forest types remains to be investigated.

4.3 Mapping of ~~biomass loss~~biomass loss rates of terra firme forests in French Guiana

We combined remote sensing maps of forest height (Simard et al., 2011) and LAI (Myneni et al., 2015) with forest modelling to derive a sample map of biomass loss rates in French Guiana. In doing so, we presented an innovative approach for estimating ~~biomass loss~~biomass loss rates in tropical forests. A comparison of estimated ~~biomass loss~~biomass loss rates with census-based values for two sites showed reasonable similarity, although in perspective, it would be important to further validate such maps using more field data (not available to us at present). In another comparison of biomass loss ~~values~~rates obtained for French Guiana with census-based values for the entire Guiana Shield (i.e., French Guiana, Suriname, Guyana, northern Brazil, eastern Venezuela; Johnson et al., 2016), our estimate is about 50% higher, though it is noteworthy that Johnson et al. (2016) estimated the rate of biomass loss for the entire Guiana Shield, with higher values on average in French Guiana. Capabilities for improved projections of biomass loss rates are indispensable in the context of improved estimates of the role of tropical forests in the global carbon cycle (Anderegg et al., 2020; Friedlingstein et al., 2019; Friend et al., 2007; IPCC, 2014). Remote sensing by airborne and satellite-based instruments provides large-scale data on forests, such as the forest height (Simard et al., 2011) and LAI (Myneni et al., 2015). However, remote sensors can record measurements only at certain time points; hence, the successional stages of forest variables are uncertain in remotely sensed data. Such forest dynamics can be simulated by individual-based, dynamic forest models. A combination of remote sensing data and forest models therefore has the potential to improve predictions of large-scale ecosystem dynamics (Plummer, 2000; Shugart et al., 2015).

Forests can be in different successional stages due to disturbances that influence forest height and LAI (Dubayah et al., 2010; Kim et al., 2017). In the forest height and LAI maps, disturbed regions can be detected visually (cf. Fig. S45); these regions have been identified as disturbed areas in other studies (Asner and Alencar, 2010; Piponiot et al., 2016a; Stach et al., 2009).
510 Such areas include disturbed areas in the flood plains of lakes and rivers, along the coast, near roads and settlements, and in the secondary forests of French Guiana, where the forest height and the crown coverage is, on average, lower than that in primary forests (Piponiot et al., 2016a; Stach et al., 2009; forest height map from Simard et al. (2011) in Fig. S54).

Remotely sensed products often include uncertainties. In this study, we demonstrated the sensitivity of the sample biomass loss map to variations in the LAI and forest height maps (Fig. 7.d, Fig. S14). Accuracy of the input remote sensing data is
515 beneficial. Balanced deviations of LAI and forest height (e.g., LAI +30% and H + 30%) result in smaller deviations in ~~biomass~~ biomass loss rates than opposite deviations (e.g., LAI +30% and H -30%). However, small-scale fluctuations in the LAI (e.g., on the individual tree level) were not captured due to the coarse resolution of the MODIS data (500 m). However, by using an individual-based dynamic forest model, the small-scale processes that manifest as variations in the LAI and forest height are accounted for in the simulations. Because simulated forest structures were the basis upon which the regression model
520 was derived, the multiple linear regression model (eq. 7) accounts for successional states and small-scale dynamics of forests. It is plausible that the LAI alone is not representative of the forest successional state. Therefore, we analysed forest succession (resulting from model-inherent processes), taking species diversity and interactions between trees into account, and we linked this information to the remote sensing products. Such approaches have already been successfully carried out in several studies (Rödig et al., 2017, 2018, 2019; Shugart et al., 2015). We obtained a good proxy for the successional state and, thus, for
525 ~~biomass loss~~ biomass loss rates, only when forest height was also included in the analysis. It may also be the case that the forest height estimated by FORMIND and the forest height mapped by Simard et al., (2011) at the 1-km scale (Fig. S5.a) are subject to a certain bias. Simard et al. (2011) used the mean height of the three tallest trees to validate the GLAS data at the basal footprint level, though not to validate the gridded 1-km data product, which tends to be shorter and less variable. The gridded product is based on a biome-level random forest model using other auxiliary data (e.g., percentage of tree cover, precipitation, altitude, temperature, conservation status), so that the variation in a 1-km resolution of forest height does not necessarily reflect the simulated variation in forest structure, rather it reflects the predicted variation based on biome-level correlations with other factors.

4.54.4 Introduction of an alternative method for estimating biomass turnover time

Information on the carbon balance of forests is important for quantifying the biomass accumulation rates of trees. Various
535 studies have estimated the turnover time of biomass, which we defined here as the reciprocal value of the rate of biomass loss, in forests worldwide (Carvalhais et al., 2014; Erb et al., 2016; Pugh et al., 2019). Carvalhais et al. (2014) were the first to estimate biomass turnover times for forests in equilibrium from biomass and GPP (cf. eq. 4: $\tau = AGB_{total} \cdot NPP^{-1}$). For the French Guiana region, the authors estimated biomass turnover times of approximately 20 to 40 years and discussed that

disturbances can shorten the biomass turnover time by increasing biomass loss rates. Our study quantifies the extent to which ~~tree mortality~~ stem mortality leads to high ~~biomass loss~~ biomass loss rates and thus to short biomass turnover times.

Erb et al. (2016) observed decreases caused by land use in the biomass turnover time. They found turnover times of 20 to 30 years for the French Guiana region, which are similar to our results (Fig. 4.c). Pugh et al. (2019) showed that stand-replacing disturbances also shortened the biomass turnover times. We found that the biomass turnover time is strongly affected by succession dynamics and ~~tree mortality~~ stem mortality intensities ~~rates~~. For our full simulation data set, we found a mean biomass turnover time of ~~32-37~~ years (Fig. S13). We derived an alternative framework to estimate the turnover time from ~~biomass loss~~ biomass loss rates. This framework allows both turnover time and rate of biomass loss to be modelled in a simple way, considering succession dynamics and disturbances due to ~~tree mortality~~ stem mortality. ~~This method can be applied to forests in equilibrium and to forests in which the early stages of succession are emerging due to disturbances and logging.~~

4.6.4.5 Limitations and Outlook

Our simulation results revealed complex relationships between ~~tree stem~~ mortality rate and biomass loss rate. The growth stage of a tree evidently has an effect on ~~tree mortality~~ stem mortality, which often results in a U-shaped relationship of ~~tree mortality~~ stem mortality as a function of the tree size distribution in a forest (Aubry-Kientz et al., 2013b; Holzwarth et al., 2013; Muller-Landau et al., 2006). With regard to tree age, it is more likely that the youngest and oldest trees will die (Aubry-Kientz et al., 2013b; R uger et al., 2011) due to intense competition for light and space between the juvenile trees in the understorey and the senescence of the old trees in the canopy layer. Such mortality processes are often represented in forest models (Bugmann et al., 2019). Although empirical mortality algorithms which mechanistically describe the main causes of ~~tree mortality~~ stem mortality and their effects on entire ecosystems (e.g., self-thinning, ~~dying-death~~ of trees by crushing, and growth-dependent mortality) have already been developed, other causes of ~~tree mortality~~ stem mortality with unclear signals are often summarised as stochastic processes (Bugmann et al., 2019; H ulsmann et al., 2017, 2018). In our study, biomass loss rates at the stand level arose from different mortality processes that occurred at the tree level (competition due to crowding, ~~dying-death~~ of other trees by crushing, growth dependency, gap formation, and stochastic ~~tree mortality~~ stem mortality). Compared to the U-shaped ~~tree mortality~~ stem mortality distribution across stem diameter classes, the biomass loss rates of a forest stand depended in more complex ways on the functional species composition and the levels of carbon fluxes (GPP and NPP). ~~We analysed the relationships between different levels of stem mortality rate with biomass loss rate, GPP, NPP, and biomass stock. It would be interesting to explore simulation results for different modes of stem mortality in future studies.~~

In our study, the effects of disturbances were represented in a simplified manner by modifying the mortality rates of trees. We analysed the effects of permanently increasing the ~~tree mortality~~ stem mortality rates in the studied forests. However, it is also necessary to consider the effects of discrete or continuously changing disturbance patterns (e.g., Barlow et al., 2003; Brando et al., 2014; Chambers et al., 2009, 2013; Doughty et al., 2015a; Holzwarth et al., 2013; Magnabosco Marra et al., 2016; Marra et al., 2014; McDowell et al., 2018; Negr on-Ju arez et al., 2010, 2017; Nepstad et al., 2007b; Phillips and Brienen, 2017; Slik

et al., 2010; Stovall et al., 2019; Wright et al., 2015). The impacts of single, discrete disturbance events (e.g., selective logging) on the dynamics of terra firme forests were studied by Hiltner et al. (2018). A follow-up study investigated the impacts of repeated logging events under continuously changing air temperatures and precipitation rates (Hiltner et al., 2021). If additional effects, such as climate change and forest management, will be added to the dynamic forest model's simulations of the present study, the reasons of biomass losses could be determined more accurately. Assessing this aspect would be interesting for future studies where the methodology presented here can be applied.

It was also found that the temporal patterns of establishing trees can change after disturbances such as modifications to the seed mortality of specific tree species, as such changes influence the competitive processes of trees within communities (Dantas de Paula et al., 2018). Here, we did not consider the influences of ~~tree mortality~~system mortality intensities rates on establishment processes, though this factor should be considered in future studies.

Regarding the mapping of the biomass loss rates in French Guiana, there are ~~three~~four important aspects. First, it is important to verify the quality of the forest model parameterization with field data as was done for ~~biomass loss~~biomass loss rates in this study and by Hiltner et al. (2018; 2021), who analysed biomass dynamics, tree size distribution, and functional species composition by comparing model results with data from forest inventories ~~of the Paracou study site in French Guiana. However, the~~Guiana. The amount of available 'ground truth data' is very small, and so a comprehensive validation of the simulation results and the biomass loss map was not possible. In future studies, the addition of more data sources could allowed for more extensive validation of the study results in future. Second, a multiple linear regression model predicting ~~biomass loss~~biomass loss rates can be valid only for a certain type of forest. In mapping ~~biomass loss~~biomass loss rates at the country level, we assumed the predominance of a similar type of forest, the terra firme forests in French Guiana (Hammond, 2005). For this forest type, Stach et al. (2009) calculated a forest cover of 95% of the country's land area. Third, site parameters across entire landscapes can be heterogeneous, affecting forest dynamics and structure. Various studies demonstrated that natural environmental factors, such as soil properties (Rödig et al., 2017; Soong et al., 2020), relief (Guitet et al., 2018), and climatic variations (Rödig et al., 2017; Wagner et al., 2012), as well as the ~~logging-silvicultural~~ history (Hiltner et al., 2018; Piconiot et al., 2016b, 2019), can affect the succession dynamics and states of tropical forests. In this study, such spatially heterogeneous environmental influences on forest dynamics in terra firme forests are indirectly considered in the forest model and the regression model via stochastic ~~tree mortality~~system mortality. Some of these environmental factors, which vary at the regional level, can be taken into account considered in future studies by adjusting some already existing model parameters including further processes (see Tab. S1, Tab. S2, Fischer et al., 2016). Examples are 1) effects of forest management and fire can be implemented, 2) the effects of weather variables such as temperature, rainfall variability and solar radiation can be taken into account, and 3) the relationships describing tree geometry can vary in space and time. Moreover, species diversity could play an important role, which was aggregated in this study using the plant functional type approach (Maréchaux and Chave, 2017). In further investigations, it is recommended that climatic and topographic effects or short-term disturbance events and forest management be implemented to further improve the approach developed here. Finally, changes in climatic conditions and tree

605 coverage have an impact on forest height, LAI, and subsequently, on biomass loss rates. For example, drought, uprooting due to storm or flooding, forest fire, insect calamity, and forest management can be possible drivers of variability in the LAI. Furthermore, forest height can vary, e.g., due to uprooting from storm and flooding, fire, and forest management. Those environmental drivers may also interact with each other. The effects of tree coverage and climate and their importance for driving the maps of forest height and LAI and subsequently the estimations of biomass loss rates should be explored in follow-up studies.

610 5 Conclusions

Here, we developed a framework for estimating ~~biomass loss~~biomass loss rates in tropical forests. We analysed the effects of ~~tree mortality~~system mortality intensity rates and its relation to forest productivity, forest structure, and biomass, based on the example of terra firme forests in French Guiana. By quantifying such effects through simulation experiments, it was possible to derive complex relationships between ~~biomass loss~~biomass loss rates and other forest attributes. Our approach revealed the
615 strong influences of the succession states and ~~tree mortality~~system mortality intensity rates on the ~~biomass loss~~biomass loss rates of forests.

We also linked individual-based forest modelling with remote sensing so that an estimation of ~~biomass loss~~biomass loss rates due to ~~tree mortality~~system mortality was feasible. The resulting sample map of biomass loss indicated that biomass is dying at a faster rate in the central, southern, and eastern regions than in the northern parts of French Guiana. The forest areas in the
620 north and northeast are used for timber production, agricultural activities, and housing (Bovolo et al., 2018; Stach et al., 2009), whereas the forest areas in the south are predominantly natural rainforests (Hammond, 2005).

The approach we developed here can be easily transferred to other forest biomes (e.g., boreal and temperate forests) using forest models that capture biome-specific forest dynamics and available remote sensing products. Estimating the spatiotemporal distribution of forest ~~biomass loss~~biomass loss rates has recently been identified as particularly relevant for the
625 monitoring of mortality hotspots (Hartmann et al., 2018). Moreover, improved estimations of the turnover times of carbon in forest stands have been made possible so that uncertainties in the global carbon cycle (Friend et al., 2014) can be reduced.

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Author's contributions

635 U.H., A.H., and R.F. conceived and designed the experiments; U.H. acquired and managed the data; U.H. performed the simulations; U.H., A.H., and R.F. contributed to analysis and discussion; U.H. wrote the manuscript; U.H., A.H., and R.F. reviewed the manuscript.

Data availability

640 The FORMIND parameterization (Hiltner et al., 2018) and the source code of FORMIND can be downloaded for free on the website www.formind.org. We included a comprehensive description of the model parameters used here in the supplemental material (Tab. S1 and Tab. S2). The [full simulation data set of the stem mortality scenarios and the biomass loss map of French Guiana](#) ~~is~~[are](#) freely available as online attachment. The input data of the MCD15A2H Version 6 MODIS for the LAI and the forest height map can be downloaded for free (Myneni et al., 2015; Simard et al., 2011)

Competing Interest

645 The authors declare no conflict of interest. The funders had no role in the design of the study, in the collection, analyses, or interpretation of data, in the writing of the manuscript, and in the decision to publish the results.

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