

# Extreme events driving year-to-year differences in gross primary productivity across the US

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1 **Abstract.** Solar-Induced chlorophyll Fluorescence (SIF) has previously been shown to strongly correlate with gross primary  
2 productivity (GPP), however this relationship has not yet been quantified for the recently launched TROPOspheric Monitoring  
3 Instrument (TROPOMI). Here we use a Gaussian mixture model to develop a parsimonious relationship between SIF from  
4 TROPOMI and GPP from flux towers across the conterminous United States (CONUS). The mixture model indicates the  
5 SIF-GPP relationship can be characterized by a linear model with two terms. We then estimate GPP across CONUS at 500-m  
6 spatial resolution over a 16-day moving window. We ~~find that CONUS GPP varies by less than 4% between 2018 and 2019.~~  
7 ~~However, we~~ observe four extreme precipitation events that induce regional GPP anomalies: drought in west Texas, flooding in  
8 the midwestern US, drought in South Dakota, and drought in California. Taken together, these events account for 28% of the  
9 year-to-year GPP differences across CONUS. Despite these large regional anomalies, we find that CONUS GPP varies by less  
10 than 4% between 2018 and 2019.

## 11 1 Introduction

12 Terrestrial gross primary productivity (GPP) is the total amount of carbon dioxide (CO<sub>2</sub>) assimilated by plants through pho-  
13 tosynthesis and represents one of the main drivers of interannual variability in the global carbon cycle Le Quéré et al. (2018).  
14 As such, quantifying the spatiotemporal patterns of terrestrial GPP is critical to understanding how the carbon cycle will both  
15 respond to and influence climate. Work over the past decade has shown satellite measurements of solar-induced chlorophyll  
16 fluorescence (SIF) to correlate strongly with tower-based estimates of GPP (e.g., Frankenberg et al., 2011a; Yang et al., 2015;  
17 Sun et al., 2017; Turner et al., 2020; Wang et al., 2020) and are often used as a remote-sensing proxy for GPP.

18 This relationship between SIF and GPP is typically expressed through a pair of light use efficiency models Monteith (1972)  
19 that relate GPP and SIF to the absorbed photosynthetically active radiation (APAR):

$$20 \text{ GPP} = \text{APAR} \times \Phi_{\text{CO}_2} \quad (1)$$

$$1 \text{ SIF} = \text{APAR} \times \beta \Phi_{\text{F}} \quad (2)$$

2 where  $\Phi_{\text{CO}_2}$  is the light use efficiency of  $\text{CO}_2$  assimilation,  $\Phi_{\text{F}}$  is the fluorescence yield, and  $\beta$  is the probability of fluoresced  
3 photons escaping the canopy. Solving for APAR and substituting, we can rewrite GPP as:

$$4 \text{ GPP} = \frac{\Phi_{\text{CO}_2}}{\beta \Phi_{\text{F}}} \text{SIF}. \quad (3)$$

5 The derivation follows from Lee et al. (2013), Guanter et al. (2014), Sun et al. (2017), and others.

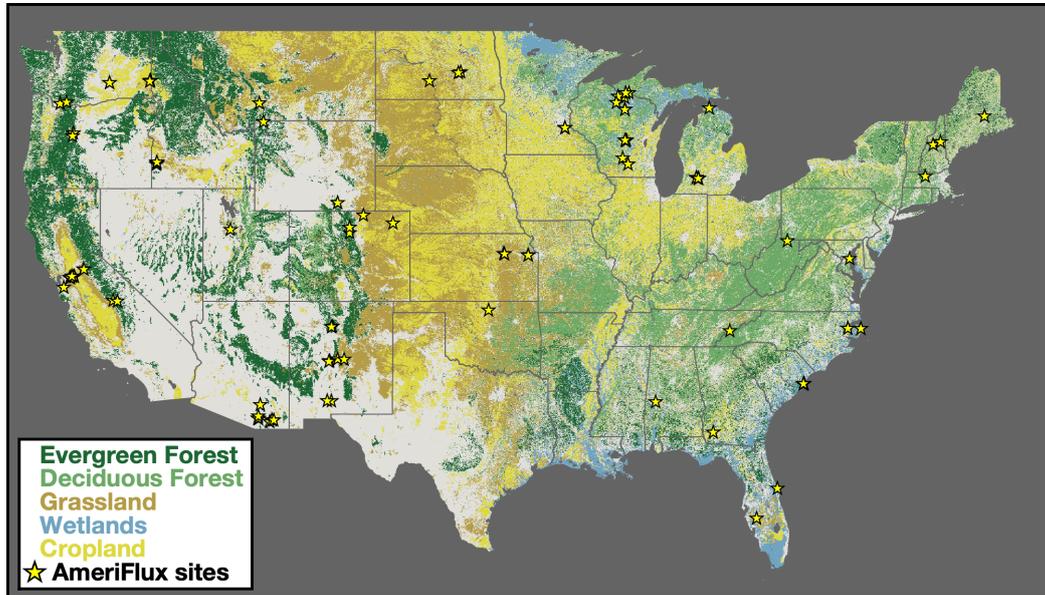
6 This seemingly straightforward relationship between SIF and GPP has been widely used to infer GPP from measurements of  
7 SIF (e.g., Frankenberg et al., 2011a; Parazoo et al., 2014; Yang et al., 2015, 2017; Sun et al., 2017, 2018; Magney et al., 2019;  
8 Turner et al., 2020) with some work showing that SIF captures more variability in GPP than APAR alone (e.g., Yang et al.,  
9 2015, 2017; Magney et al., 2019). However, there is much complexity encapsulated in the first term of Eq. 3 ( $\Phi_{\text{CO}_2}/\beta\Phi_{\text{F}}$ ).  
10 There is an ongoing debate about what *exactly* SIF is telling us about GPP (e.g., Joiner and Yoshida, 2020; Maguire et al., 2020;  
11 Dechant et al., 2020; He et al., 2020; Marrs et al., 2020) and the spatio-temporal scales at which SIF and GPP correlate well. A  
12 recent paper from Magney et al. (2020) presents a concise summary of the covariation between SIF and GPP at different spatio-  
13 temporal scales and how non-linear relationships at the leaf-scale often integrate to a linear response at the canopy-scale. This  
14 is due, in large part, to the fact that most of our satellite measurements occur near the middle of the day when the  $\Phi_{\text{CO}_2}$ - $\Phi_{\text{F}}$   
15 response is more-or-less linear and the observed signal is integrated over many leaves.

16 Here we focus on the ecosystem-scale relationship between SIF and GPP, as that is the relevant observable scale from  
17 space-borne instruments. We begin by characterizing the relationship between instantaneous SIF from TROPOMI and half-  
18 hourly GPP from flux towers. Following this, we use this ecosystem-scale relationship to infer GPP at a spatial resolution of  
19 500-m using TROPOMI SIF measurements and identify drivers of interannual variability in GPP. Previous work has identi-  
20 fied effects ~~such as seasonal redistribution Butterfield et al. (2020), and responses such as~~ drought (e.g., Sun et al., 2015), ~~and~~  
21 ~~flooding Yin et al. (2020) as important drivers of flooding (Yin et al., 2020), and seasonal redistribution (Butterfield et al., 2020) as~~  
22 ~~important factors controlling~~ interannual variability in GPP.

## 23 **2 Identifying distinct relationships between SIF and GPP**

24 We build on our previous work ~~Turner et al. (2020) (Turner et al., 2020)~~ downscaling measurements of SIF to 500-m spatial  
25 resolution. Briefly, the TROPOspheric Monitoring Instrument (TROPOMI; Veefkind et al., 2012) is a nadir-viewing imaging  
26 spectrometer. TROPOMI has a 2,600 km swath with a nadir spatial resolution of 5.6 km along track and 3.5 km across track.  
27 Köhler et al. (2018) presented the first retrievals of SIF from TROPOMI. As in Turner et al. (2020), we apply a *post hoc* bias  
28 correction to ensure positivity of monthly average values as systematically negative SIF values are non-physical. We downscale

29 individual TROPOMI scenes using the near-infrared reflectance of vegetation index ( $NIR_v$ ) that was proposed by Badgley et al.  
30 (2017, 2019). We use the MCD43A4.006 (v06) MODIS NBAR reflectances Schaaf et al. (2002) to compute  $NIR_v$ . Two notable  
31 differences from Turner et al. (2020) are: 1) the analysis is extended to cover all of CONUS and 2) we now use a 16-day moving  
32 window, thus including a full orbit cycle in each averaging window to minimize effects due to viewing-illumination geometry  
1 and noise. Supplemental Fig. S3 shows the improvement when averaging to longer temporal windows with an  $r$  of 0.66, 0.74,  
2 0.79, and 0.82 for instantaneous, 8-day, 16-day, and 32-day temporal windows, respectively.



**Figure 1. Dominant landcover over conterminous United States (CONUS).** Colors show the dominant landcover over CONUS. Classification is based on the 2019 USDA CropScape database USDA (2018). Forests are shown in green croplands in yellow, and wetlands in blue. Location of 82 AmeriFlux sites used in this study are shown as yellow stars.

3 The extension to CONUS facilitates comparison of TROPOMI SIF retrievals to flux tower data over a more representative set  
4 of ecosystems and robustly infer the SIF-GPP relationship. Specifically, there are 82 AmeriFlux sites Baldocchi et al. (2001)  
5 within CONUS that reported data in 2018, 2019, or 2020 whereas Turner et al. (2020) only included 11 sites and did not  
6 have data from forests. Figure 1 shows the location of these 82 AmeriFlux sites overlaid on the dominant landcover. These  
7 eddy covariance sites provide a direct measure of net ecosystem exchange (NEE;  $CO_2$  fluxes) Baldocchi et al. (1988). We  
8 ~~use GPP that has been partitioned by the group operating the site. If GPP is not provided we compute it~~ compute GPP at  
9 each site using nighttime measurements of NEE as a proxy for ecosystem respiration Reichstein et al. (2005) to partition the  
10 NEE into respiration and GPP. The AmeriFlux sites used here cover 10 ecosystems as defined by the International Geosphere-  
11 Biosphere Programme: evergreen needleleaf forest, deciduous broadleaf forest, mixed forest, grassland, cropland, wetland,  
12 woody savanna, savanna, open shrubland, and closed shrubland. These are the classifications reported with the AmeriFlux data  
13 as of July 2021 (<https://ameriflux.lbl.gov>).

14 We characterize the relationship between TROPOMI SIF and AmeriFlux GPP by plotting downscaled ~~instantaneous~~-SIF  
 15 observations against daily GPP from the nearest AmeriFlux GPP data in time site (see Supplemental Figs. S1-S3). The  
 1 TROPOMI overpass time varies over the orbit cycle. Frankenberg et al. (2011b) presented a simple approach to compute a  
 2 “daily corrected” SIF that accounts for variations in overpass time, length of day, and solar zenith angle:

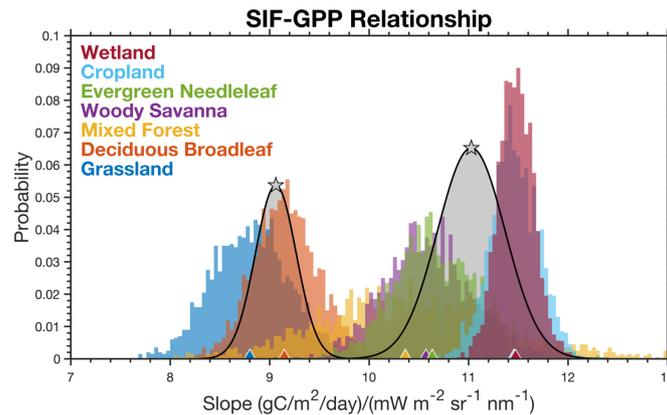
$$3 \quad \overline{\text{SIF}}(x, y, t) = \text{SIF}(x, y, \tau_s) \frac{\int_{\tau_0}^{\tau_f} \cos[\text{SZA}(x, y, \tau)] d\tau}{\cos[\text{SZA}(x, y, \tau_s)]} \quad (4)$$

4 where SIF( $x, y, \tau_s$ ) is the instantaneous SIF, SZA is the local solar zenith angle,  $\tau_0$  is sunrise,  $\tau_f$  is sunset, and  $\tau_s$  is the  
 5 hour corresponding to the TROPOMI overpass time. We compare this daily corrected SIF against the daily GPP for each  
 6 AmeriFlux site. Specifically, the 6-7 steps we take here are: 1) construct a timeseries of daily GPP from each AmeriFlux site,  
 7 2) apply the *post hoc* bias correction to the TROPOMI SIF data, 2) ~~find all TROPOMI scenes that cover an AmeriFlux site,~~  
 8 3) ~~compute the daily correction for TROPOMI SIF data,~~ 4) downscale TROPOMI scenes to 500-m spatial resolution using  
 9 MODIS NIR<sub>v</sub>, 4) 5) find all TROPOMI scenes that cover an AmeriFlux site, 6) construct a timeseries of SIF observations  
 10 from the 500-m grid cell that contains the AmeriFlux site, 5) ~~construct a timeseries of AmeriFlux GPP data that are coincident~~  
 11 ~~in time with the TROPOMI overpass, and 6) ~~regress SIF on and 7) regress coincident daily corrected TROPOMI SIF on~~~~  
 12 daily AmeriFlux GPP with a bisquare regression. The bisquare regression was chosen due to robustness against outliers.  
 13 Additionally, we force the regression through the origin based on the physical constraint that GPP should be zero if SIF is  
 14 zero. We observe a linear relationship between SIF and GPP when plotted against all ecosystems (Supplemental Figure S1)  
 15 and when separated by ecosystem (Supplemental Figure S2). Notable exceptions are closed shrubland, open shrubland, and  
 16 savanna ecosystems where SIF explains less than 10% of the variability in GPP for AmeriFlux sites in those ecosystems due,  
 17 in part, to a low signal-to-noise ratio. These ecosystems typically have a small SIF signal and the bright surfaces often result in  
 18 a higher retrieval uncertainty. This combination of a small signal and high retrieval uncertainty results in a low signal-to-noise  
 19 ratio, complicating efforts to derive a robust relationship between SIF and GPP for these ecosystems.

20 Many of the ecosystems exhibit a similar linear relationship between SIF and GPP, which begs the question: “*what ecosys-*  
 21 *tems have a distinct SIF-GPP relationship?*” To address this, we bootstrap the bisquare regression for each ecosystem 2000  
 22 times. The slopes from this bootstrap can be seen in Figure 2. The range of slopes vary from 13 to 18 ( $\mu\text{mol m}^{-2} \text{s}^{-1}$ ) / ( $\text{mW m}^{-2} \text{sr}^{-1} \text{s}^{-1}$ )  
 23 with grasslands at the low end and evergreen needleleaf forests at the high end. We then use a two component Gaussian mixture  
 24 model (see, for example, Bishop, 2007) to identify clusters of ecosystems with a similar SIF-GPP relationship. The implemen-  
 25 tation of our Gaussian mixture model is adapted from Turner and Jacob (2015). Parameters of the mixture model are obtained  
 26 via an iterative expectation-maximization algorithm. A drawback of these mixture models is they often find local minima. To  
 27 address this, we repeat the fitting of the mixture model with multiple initializations and use simulated annealing to search for  
 28 a global minimum. We tested a range of mixture model sizes and found a mixture of two Gaussians to be the most robust.  
 29 Adding additional terms in the model resulted in Gaussians that did not have the largest weighting factor for any ecosystem.  
 30 This is because ecosystems like Woody Savanna and Deciduous Broadleaf have a large spread in their slope. As such, there

31 [is a lot of uncertainty and the model does not find that they require a unique regression slope.](#) The resulting mixture model is  
32 overlaid on the histogram in Figure 2.

33 We observe a clustering of ecosystems with SIF-GPP relationships around  $16.4 (\mu\text{mol m}^{-2} \text{s}^{-1}) / (\text{mW m}^{-2} \text{sr}^{-1} \text{s}^{-1})$ .  
34 This grouping is the dominant weighting term for wetlands, evergreen needleleaf forest, deciduous broadleaf forest, mixed  
35 forest, cropland, and woody savanna. We refer to this cluster as the “Dominant Cluster” and assume that ecosystems not  
1 specifically mentioned elsewhere will have a response that is similar to this primary cluster. The other component of the  
2 mixture model corresponds to grasslands. Ecosystems not explicitly mentioned use the “Dominant Cluster” for scaling SIF  
3 to GPP. Table 1 lists the SIF-GPP relationships for these two clusters. [The uncertainty is the variance for the Gaussian for](#)  
4 [that particular cluster \(see Bishop, 2007; Turner and Jacob, 2015, for more on Gaussian mixture models\).](#) Previous work has  
5 also found unique SIF-GPP relationships between C3 and C4 plants using measurements from a tower including a non-linear  
6 response in C3 plants [He et al. \(2020\) \(He et al., 2020\)](#), we examined this here using two AmeriFlux sites in corn fields and  
7 two in potato fields. We do observe potential differences in the SIF-GPP relationship between these C3 and C4 systems (see  
8 Supplemental Figure S5). The difference in SIF-GPP relationship for C3 and C4 systems seen here is also similar to what  
9 was observed using  $\text{NIR}_v$  Badgley et al. (2019). These relationships can be used to reconstruct GPP from TROPOMI SIF as:  
10  $\text{GPP} = \text{SIF} \times (\sum_i f_i s_i)$  where  $s_i$  is the SIF-GPP relationship in Table 1 for the  $i^{\text{th}}$  cluster and  $f_i$  is the fraction of a grid cell  
11 represented by that cluster.



**Figure 2. Identifying distinct SIF-GPP relationships across ecosystems.** Histogram shows the distribution of slopes that map SIF to GPP using a bisquare regression and a 2000 member bootstrap. Colors denote the different ecosystems and triangles at the bottom show the mean for that ecosystem. Gray distributions are from a two-member Gaussian Mixture Model and the stars indicate the mean for that component.

1 TROPOMI is in low earth orbit and only observes a snapshot in time. The equatorial overpass time at nadir is 13:30 lo-  
2 cal time. ~~By assuming that GPP scales linearly with PAR (i.e., Eq. 1) we can compute a correction factor to estimate daily~~  
3 ~~integrated GPP.~~ [We compute a daily corrected SIF that accounts for variations in overpass time, length of day, and solar zenith](#)

**Table 1.** SIF-GPP relationships for different groupings.

Cluster	SIF-GPP relationship <sup>a</sup> ( $s_i$ )
Dominant Cluster	$16.4 \pm 0.2$
Grassland	$14.1 \pm 0.3$

<sup>a</sup>All SIF-GPP relationships have units of  $(\text{gC m}^{-2} \text{ day}^{-1}) / (\text{mW m}^{-2} \text{ sr}^{-1} \text{ nm}^{-1})$ .  
Uncertainty is the diagonal of the covariance matrix for the mixture model.

4 [angle \(Frankenberg et al., 2011b; Köhler et al., 2018\):](#)

$$5 \quad \overline{\text{SIF}}(x, y, t) = \text{SIF}(x, y, \tau_s) \frac{\int_{\tau_0}^{\tau_f} \cos[\text{SZA}(x, y, \tau)] d\tau}{\cos[\text{SZA}(x, y, \tau_s)]} \quad (5)$$

6 [where  \$\text{SIF}\(x, y, \tau\_s\)\$  is the instantaneous SIF, SZA is the local solar zenith angle,  \$\tau\_0\$  is sunrise,  \$\tau\_f\$  is sunset, and  \$\tau\_s\$  is the hour](#)  
7 [corresponding to the TROPOMI overpass time. We use this daily corrected SIF in conjunction with](#) More formally, we scale  
8 the instantaneous SIF by the ratio of the integral of the cosine of the solar zenith angle (SZA) over the day to  $\cos(\text{SZA})$  from  
9 the TROPOMI overpass time. Putting everything together, we estimate daily GPP from TROPOMI SIF observations as:

$$10 \quad \text{GPP}(x, y, t) = \overline{\text{SIF}}(x, y, t) \cdot \gamma \sum_i s_i f_i(x, y) \cdot \frac{\int_{\tau_0}^{\tau_f} \cos[\text{SZA}(x, y, \tau)] d\tau}{\cos[\text{SZA}(x, y, \tau_s)]} \quad (6)$$

11 where  $\text{SIF}(x, y, t)$  is the 500-m downscaled SIF using a 16-day moving window,  $\gamma$  is a unit conversion from  $\mu\text{mol}$  to  $\text{gC}$ ,  $s_i$  is  
12 the SIF-GPP relationship inferred from comparison with AmeriFlux GPP (see Table 1),  $f_i(x, y)$  is the fraction of the grid cell  
13 represented by the  $i^{\text{th}}$  cluster, SZA is the local solar zenith angle,  $\tau_0$  is sunrise,  $\tau_f$  is sunset, and  $\tau_s$  is the hour corresponding to  
14 the TROPOMI overpass time. We do not include information on cloud cover in our approach, this could potentially be included  
15 in the future to account for diurnal variations in PAR.

16 [Our estimate of GPP is proportional to SIF and the regression coefficients:  \$\text{GPP} \propto \overline{\text{SIF}} \cdot s\_i\$ . As such, we propagate our](#)  
17 [uncertainties in quadrature:](#)

$$18 \quad \sigma_{\text{GPP}} = \sqrt{\left( \frac{\partial \text{GPP}}{\partial \overline{\text{SIF}}} \sigma_{\overline{\text{SIF}}} \right)^2 + \sum_i \left( \frac{\partial \text{GPP}}{\partial s_i} \sigma_{s_i} \right)^2} \quad (7)$$

$$19 \quad = \sqrt{\left( \sigma_{\overline{\text{SIF}}} \gamma \sum_i s_i f_i(x, y) \right)^2 + \sum_i \left( \overline{\text{SIF}}(x, y, t) \cdot \sigma_{s_i} \gamma s_i f_i(x, y) \right)^2}$$

20 [where  \$\sigma\_{\overline{\text{SIF}}}\$  is the uncertainty in the daily corrected SIF and  \$\sigma\_{s\_i}\$  is the uncertainty in the SIF-GPP relationship.](#)

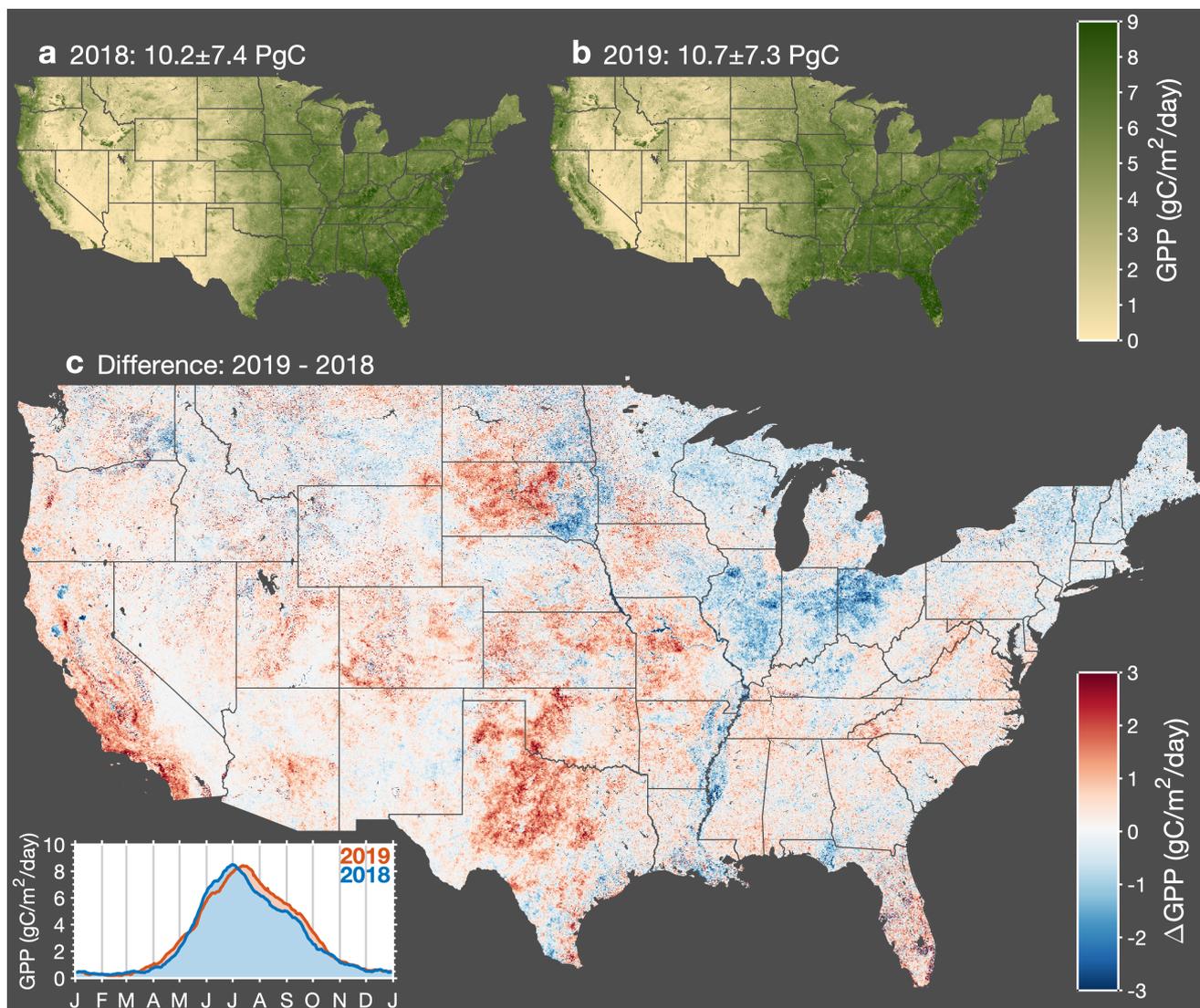
### 21 3 Drivers of interannual variations in US gross primary productivity

22 Figure 3 shows annual mean GPP across CONUS inferred from TROPOMI SIF measurements using Eq. 6. A number of  
23 prominent features are visible such as the Central Valley of California, the Snake River Valley in Idaho, and the Adirondack  
24 Mountains in upstate New York. California's Central Valley and Idaho's Snake River Valley are both major agricultural regions  
25 in the western US (e.g., the Central Valley of California accounts for more than 15% of irrigated land in the US). The Adirondack  
26 Mountains are a roughly circular dome that rise above the surrounding lowlands, resulting in a shorter growing season and lower  
27 annual mean GPP. This shortened growing season can be seen in an animation of GPP over CONUS (Supplemental Movie S1).

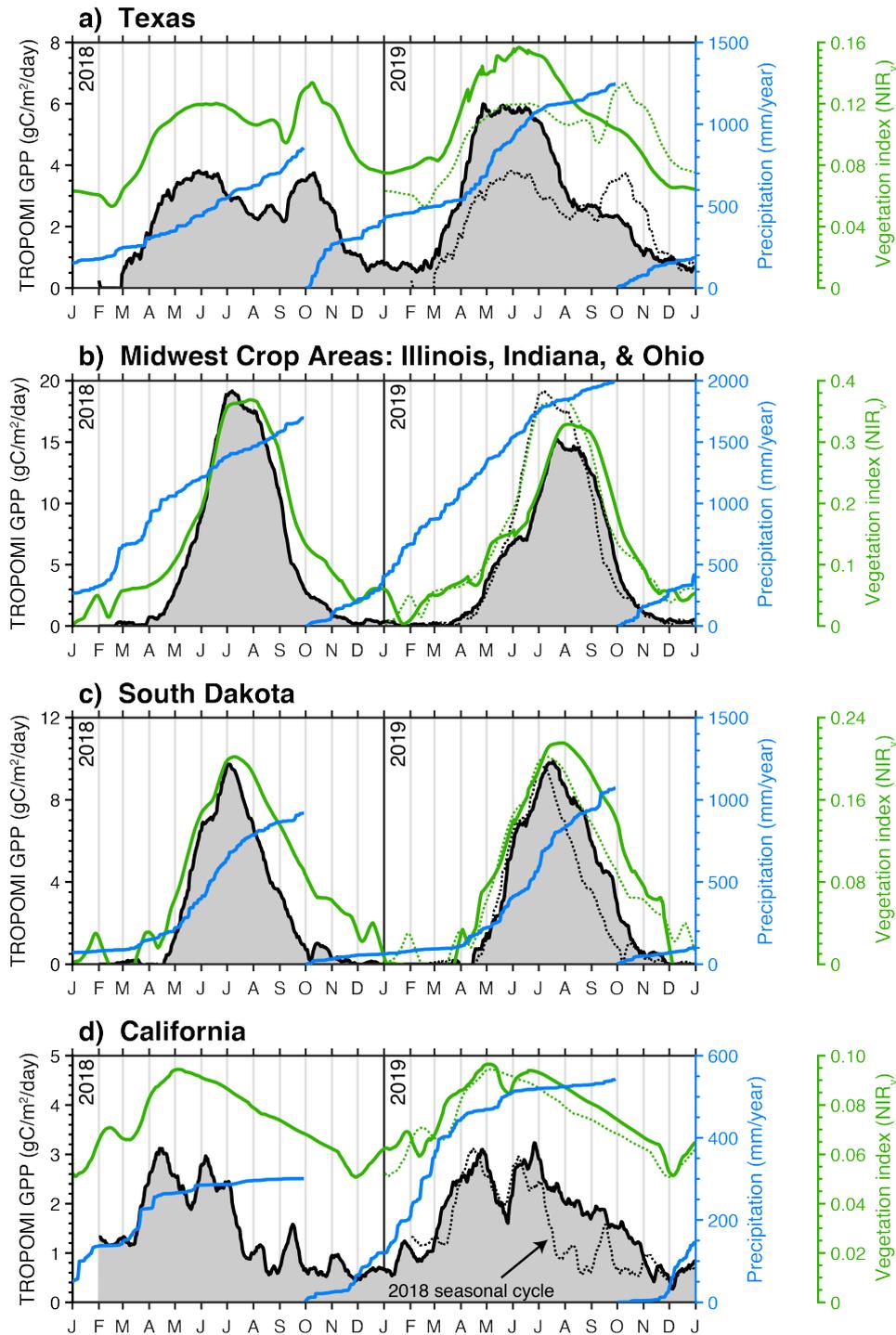
1 We observe substantial GPP across the eastern US (delineated here by  $98^{\circ}\text{W}$ ) with annual mean values generally in excess  
2 of  $5 \text{ gC/m}^2/\text{day}$ . This region accounts for less than half of the land but more than 70% of the annual GPP. This delineation in  
3 GPP roughly coincides with the location of drylands in CONUS that are more sensitive to changes in precipitation [as inferred  
4 by measurements from the Global Precipitation Measurement \(GPM\) mission \(specifically, we use the GPM\\_3IMERGDE.06  
5 product\)](#); drylands are also projected to expand in future climate Yao et al. (2020). Most of the large year-to-year differences  
6 occur in these western US drylands (see Fig. 3c), a notable exception being a negative GPP anomaly in 2019 relative to 2018  
7 that extended across Illinois, Indiana, and Ohio. Here we highlight four precipitation-driven GPP anomalies, which taken  
8 together, account for 28% of the interannual GPP variability across the United States: 1) 2018 drought in west Texas, 2) 2019  
9 midwestern crop flooding, 3) 2018 drought in South Dakota, and 4) 2018 drought in California. Figure 4 summarizes the  
10 interannual precipitation differences that we hypothesize are responsible for explaining these four GPP anomalies.

11 The largest positive GPP anomaly in 2019 relative to 2018 was observed across western Texas. This single event accounted  
12 for 11% of the year-to-year difference in GPP across CONUS with an annual GPP of  $0.65 \pm 0.47 \text{ PgC}$  in 2018 and  $0.76 \pm$   
13  $0.45 \text{ PgC}$  in 2019. From Figure 4a, we observe 50% higher GPP in spring 2019 compared to spring 2018. This increase in GPP  
14 was driven by a lack of precipitation in spring 2018. The cumulative precipitation from October 2017 through June 2018 was  
15 50% less than October 2018 through June 2019 (500 mm vs 1000 mm). The other notable difference between GPP in 2018 and  
16 2019 was a second peak during fall 2018 that was not present in 2019. This second peak coincided with a series of precipitation  
17 events beginning in early September. This tight coupling between GPP and precipitation is expected for dryland systems such  
18 as west Texas (e.g., Smith et al., 2019). The seasonal GPP dynamics inferred from TROPOMI SIF are also present in the  
19 MODIS vegetation index  $\text{NIR}_v$ , albeit with slight differences in magnitude, implying convergent responses in SIF and  $\text{NIR}_v$   
20 for this ecosystem.

21 The second largest anomaly is the reduction in 2019 GPP relative to 2018 across midwestern crop areas (specifically Illinois,  
22 Indiana, and Ohio) that accounted for 7% of the year-to-year difference in CONUS GPP. The 2018 annual GPP was  $0.70 \pm$   
23  $0.12 \text{ PgC}$  and  $0.63 \pm 0.14 \text{ PgC}$  in 2019. We observe a decrease in the maximum GPP between 2019 and 2018 as well as a two  
24 week delay in the timing of the maximum. This anomaly was highlighted in recent work from Yin et al. (2020) who attribute  
25 the anomaly to flooding in the midwestern US. The flooding delayed planting of crops by two weeks and resulted in decreased  
26 carbon uptake across the midwestern crop areas and Mississippi Alluvial Valley, where we also observe a negative anomaly in  
27 Figure 3c. Yin et al. (2020) provide a detailed discussion of these floods and their impacts on crop productivity.



**Figure 3. Interannual variations in gross primary productivity across CONUS.** Map of annual mean GPP for 2018 (panel a) and 2019 (panel b). (Panel c) Map of the difference in annual mean GPP between 2019 and 2018. Red indicates higher GPP in 2019 and red indicates higher GPP in 2018. Inset in bottom left corner shows a timeseries of the average GPP across CONUS for 2018 and 2019.



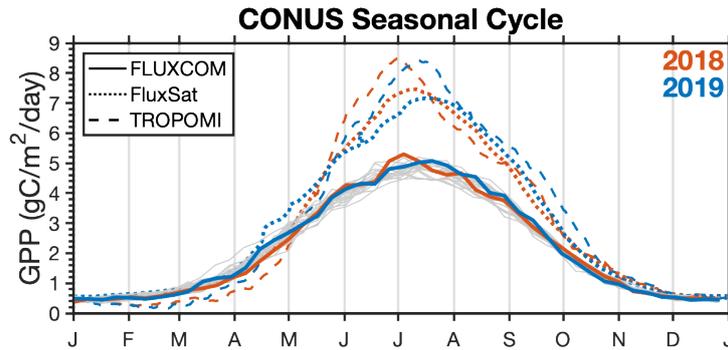
**Figure 4. Major drivers of interannual variability in CONUS GPP.** Black line shows the TROPOMI-derived GPP over Texas (a), the midwest crop region (b), South Dakota (c), and California (d). Blue line shows the cumulative precipitation over the water year as measured by the GPM satellite. Green line is  $\text{NIR}_v$  from MODIS. Black and Green dotted lines are 2018 GPP and  $\text{NIR}_v$  superimposed on the 2019 timeseries.

28 South Dakota exhibits a dipole with positive anomalies in 2019 in the west and negative anomalies in the east, again relative  
29 to 2018. The 2018 annual GPP was  $0.20 \pm 0.09$  PgC and  $0.63 \pm 0.08$  PgC in 2019. The negative anomalies in the east are  
30 driven by the flooding events discussed above and in Yin et al. (2020). However, the positive anomaly in western portion of  
31 the state is the dominant term. This positive anomaly is driven by a series of summer precipitation events that served to extend  
32 the growing season across the western plains. From Figure 4c, we can see three precipitation events throughout the mid-to-late  
33 summer that coincide with pauses in senescence: mid-July, early August, and mid-September. As with Texas, this highlights  
34 the tight coupling between GPP and precipitation for dryland systems. In toto, these precipitation events served to increase  
1 statewide GPP in 2019 relative to 2018.

2 The final notable anomaly is California's positive GPP anomaly in 2019. The 2018 annual GPP was  $0.27 \pm 0.24$  PgC and  
3  $0.33 \pm 0.26$  PgC in 2019. 2018 was a mild drought in California with  $\sim 80\%$  of the state being classified as abnormally dry;  
4 2019 had 50% more precipitation during the water year than 2018 (Figure 4c). Two consequences of this drought in 2018 were:  
5 a delayed onset of photosynthesis and a mid-summer senescence. The onset of photosynthesis in 2018 coincided with a series  
6 of atmospheric rivers that delivered about a third of the total precipitation that year, indicating a water limitation up to that  
7 point. In contrast, 2019 had ample precipitation through the winter and we observe both an earlier onset of photosynthesis and  
8 an extension of the growing season into the fall. Evergreen forests are the main contributor to the SIF signal during the summer  
9 and fall Turner et al. (2020) and, as such, will be more sensitive to the accumulated precipitation. The spatial pattern of the  
10 differences in August-November GPP (Fig. S4) strongly correlate with evergreen forests.

11 In contrast to the anomalies presented earlier, the SIF-derived GPP and MODIS-based vegetation index ( $NIR_v$ ) show di-  
12 vergent seasonal dynamics for California.  $NIR_v$  shows small differences between 2018 and 2019 with a strong similarity to  
13 the 2019 SIF-derived GPP. The seasonality of  $NIR_v$  is similar to that of the leaf area index (LAI) derived from MODIS (see  
14 Supplemental Figure 6), implying a biophysical signal. Vegetation indices [derived from the red and NIR part of the spectrum](#)  
15 estimate *photosynthetic capacity* provided optimal soil moisture, temperature, and PAR are known Sellers (1985). As such, this  
16 suggests that we observed a down-regulation of photosynthesis from evergreen forests in response to a water limitation during  
17 fall 2018, whereas these forests were close to photosynthetic capacity in fall 2019 resulting in a similar seasonality to 2018 and  
18 2019  $NIR_v$ . Sims et al. (2014) also report a low sensitivity of MODIS vegetation indices to drought stress in forests.

19 [We additionally compare our GPP estimated from TROPOMI SIF to previous work developing gridded GPP products using](#)  
20 [machine learning. Specifically, the FLUXCOM initiative \(<http://www.fluxcom.org/>; Jung et al., 2020\) and FluxSat \(Joiner and Yoshida, 20](#)  
21 [trained machine learning models to predict gridded-GPP from eddy covariance sites using remote sensing data \(including](#)  
1 [MODIS\). Figure 5 shows the CONUS seasonal cycle from both FLUXCOM, FluxSat, and TROPOMI. The seasonal cycles of](#)  
2 [GPP inferred from TROPOMI and FluxSat are generally in agreement with a similar magnitude while FLUXCOM predicts](#)  
3 [35% less GPP. Additionally, the interannual variability in GPP over CONUS inferred from TROPOMI SIF is larger than what](#)  
4 [is predicted by either FLUXCOM or FluxSat, both of which show little interannual variability. The low interannual variability](#)  
5 [is particularly evident in FLUXCOM where we can see small spread in the variability from 2001-2017 \(gray lines\).](#)



**Figure 5.** Comparison of the seasonal cycle inferred from TROPOMI SIF to FLUXCOM and FluxSat. Red lines indicate the 2018 seasonal cycle and blue lines indicate the 2019 seasonal cycle for TROPOMI (dashed lines), FluxSat (dotted lines), and FLUXCOM (solid lines). Thin gray lines are years 2001–2017 for FLUXCOM.

## 6 4 Conclusions

7 We have developed a parsimonious relationship between measurements of SIF from TROPOMI and GPP inferred from flux  
 8 towers. This relationship allows for estimation of GPP directly from TROPOMI SIF measurements. We combine this SIF-  
 9 GPP relationship with work downscaling TROPOMI data to 500-m spatial resolution to construct estimates of GPP across  
 10 the conterminous United States in 2018 and 2019. ~~Our estimate of US GPP varies by less than 4% between 2018 and 2019.~~  
 11 ~~We do, however, We~~ observe large regional anomalies that are driven by extreme precipitation events. Namely, west Texas,  
 12 South Dakota, and California experienced droughts in 2018 while midwestern US crop areas (Illinois, Indiana, and Ohio)  
 13 experienced flooding in 2019. Taken together, these four events account for 28% of the year-to-year variability in GPP across  
 14 the conterminous United States. Despite these large regional anomalies, our estimate of US GPP varies by less than 4% between  
 15 2018 and 2019.

16 The impact of the west Texas drought, South Dakota drought, and midwestern flooding are observed in other remote-  
 17 sensing measures of photosynthetic capacity such as  $NIR_v$  while the California drought shows a divergent result using SIF; the  
 18 divergent responses are driven by specific ecosystems such as evergreen forests. Our work suggests that SIF provides a measure  
 19 of *photosynthetic activity* as opposed to *photosynthetic capacity*, and converge with other remote-sensing measures under non-  
 20 stressed conditions. Future work investigating the response to extreme events ~~across ecosystems in evergreen systems~~ may  
 21 provide additional insight into these divergent responses in remote-sensing measurements related to photosynthesis.

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