1	The paradox of assessing greenhouse gases from soils for nature-
2	based solutions
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21 Abstract

22 Quantifying the role of soils in nature-based solutions requirer equires accurate estimates of 23 soil greenhouse gas (GHG) fluxes. Technological advances allow us to simultaneously 24 measure multiple GHGs simultaneously, and now it is possible to provide complete GHG 25 budgets from soils (i.e., CO₂, CH₄, and N₂O fluxes). We propose that there is a conflict 26 between the convenience of simultaneously measuring multiple soil GHG fluxes at fixed time 27 intervals (e.g., once, or twice per month) and the intrinsic temporal variability and patterns of 28 different GHG fluxes. Information derived from fixed time intervals -as is commonly done 29 during manual field campaigns- had limitations to reproduce reproducing statistical 30 properties, temporal dependence, annual budgets, and associated uncertainty, when compared 31 with information derived from continuous measurements (i.e., automated hourly 32 measurements) for all soil GHG fluxes. We present a novel approach (i.e., temporal 33 univariate Latin Hypercube sampling) that can be applied to provide insights and optimize 34 monitoring efforts of GHG fluxes across time. We suggest that multiple GHG fluxes should 35 not be simultaneously measured at a few fixed time intervals (especiallymainly when 36 measurements are limited to once aper month), but an optimized sampling approach can be 37 used to reduce bias and uncertainty. These results have implications for assessing GHG 38 fluxes from soils and consequently reduce uncertainty on the role of soils in nature-based 39 solutions.

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41 Keywords: Carbon dioxide, methane, nitrous oxide, representativeness, uncertainty

43 **1. Introduction**

44 Soils are importantessential for nature-based solutions for their role in climate mitigation 45 potential through the implementation of implementing different natural pathways (Griscom et 46 al., 2017; Bossio et al., 2020). The climate mitigation potential of soils is dependent on 47 multiple factors such as weather variability (Kim et al., 2012), ecosystem type (Oertel et al., 48 2016), soil structure (Ball, 2013), management practices (Shakoor et al., 2021), or 49 disturbances (Vargas, 2012), where soils can ultimately act as net sources or sinks of 50 greenhouse gases (GHGs). Therefore, accurate quantification of the magnitudes and patterns 51 of soil GHGs fluxes is needed to understand the potential of soils to mitigate or contribute to 52 global warming across ecosystems and different scenarios. 53 Most of our understanding of soil GHGs has come from manual measurements 54 performed throughout labor--intensive field campaigns and experiments (Oertel et al., 2016). 55 While most studies around the world have focused on soil CO₂ fluxes (Jian et al., 2020), 56 there are early examples reported coupled measurements of soil CO₂, CH₄, and N₂O fluxes across tropical forests (Keller et al., 1986) and savannas (Hao et al., 1988), 57 58 temperate forests (Bowden et al., 1993), and peatlands (Freeman et al., 1993). These pioneer 59 studies provided an early view of the importance of integrated measurements of multiple soil 60 GHG fluxes to understand the net global warming potential of soils, but also 61 demonstrated emonstrated the technical limitations and challenges associated with these 62 efforts. For example, it is known that manual measurements have the strength of providing 63 good spatial coverage during field surveys but provide limited information about the 64 temporal variability (Yao et al., 2009; Barba et al., 2021). 65 Technological advances have opened the opportunity to simultaneously measure 66 multiple soil GHG fluxes (i.e., CO₂, CH₄ and N₂O) at unprecedented temporal resolution 67 (e.g., hourly). These efforts have demonstrated differences in diel patterns and pulse events

68 (e.g., rewetting) due to wetting and drying cycles across tropical (Butterbach-Bahl et al., 69 2004; Werner et al., 2007), subtropical (Rowlings et al., 2012)(Rowlings et al., 2012), and 70 temperate (Savage et al., 2014; Petrakis et al., 2017)(Savage et al., 2014; Petrakis et al., 71 2017) ecosystems. These approaches provide more accurate information to calculate net 72 GHG budgets and the global warming potential of soils (Capooci et al., 2019). That said, 73 performing automated measurements of multiple GHGs is expensive, and this approach 74 usually has <u>a</u> lower representation of the spatial heterogeneity within ecosystems (Yao et al., 75 2009; Barba et al., 2021).

76 Ideally, we would like to measure everything, everywhere, and all the time, but this is 77 not possible impossible due to logistical, technological, physical, and economic constraints. 78 Light weightLightweight and low-powered laser-based spectrometers have reduced technical 79 barriers forto simultaneously measuring multiple GHGs fluxes from soils, and it. It is now 80 easier and faster to perform discrete manual surveys across time. This opportunity creates a 81 paradox concerning when to measure different GHG fluxes from soils when performing 82 manual measurements. In general, researchers Researchers generally tend to perform 83 simultaneous measurements of multiple GHGs during manual surveys, but this convenience 84 could result in biased information. We propose that there is a conflict between the convenience of measuring multiple GHGs at a few fixed time intervals and the intrinsic 85 86 temporal variability of magnitudes and patterns of different GHG fluxes. 87 Here, we present a proof-of-concept and test how a subset of measurements derived 88 from a fixed temporal stratification (FTS) for simultaneous measurements (i.e., stratified 89 sampling schedule) or using an optimized sampling (i.e., temporal univariate Latin 90 Hypercube sampling (tuLHs)), compared with automated measurements of soil CO₂ (F_ACO_2), 91 <u>CH₄ (F_A CH₄), and N₂O (F_A N₂O) fluxes from a temperate forest (Petrakis et al., 2018; Barba</u> et al., 2021, 2019). Here, we The underlying assumption supporting any FTS approach is that 92

93 <u>a few measurements in time can reproduce the statistical properties and temporal</u>

- 94 dependencies of soil CO₂, CH₄, and N₂O fluxes because these GHGs respond similarly to
- 95 biological and physical drivers. The *tuLHs* is a new statistical method for generating
- 96 subsamples of parameter values (i.e., soil GHG gas fluxes in this case study) to reproduce the
- 97 probability distribution and the temporal dependence of each original time series of GHG
- 98 fluxes. We reveal that reporting GHG fluxes using an FTS for simultaneous measurements
- 99 may result in biased information on temporal patterns and magnitudes. This study shows how
- 100 <u>a biased sampling schedule could influence our understanding of GHG fluxes and, ultimately,</u>
- 101 the climate mitigation potential of soils.
- 102 test how a subset of measurements derived from a fixed temporal stratification (FTS)
- 103 for simultaneous measurements (i.e., stratified sampling schedule) or using an optimized
- 104 sampling (i.e., temporal univariate Latin Hypercube sampling (*tuLHs*)), compared with
- 105 automated measurements of soil CO₂ (F₄CO₂), CH₄ (F₄CH₄), and N₂O (F₄N₂O) fluxes in a
- 106 temperate forest. We reveal that reporting measurements of GHG fluxes using a FTS for
- 107 simultaneous measurements, results in biased information of temporal patterns and

108 magnitudes. This study shows how a biased sampling schedule could influence our

109 understanding of GHG fluxes and ultimately the climate mitigation potential of soils.

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111 2. Materials and Methods

112 2.1 Study site

113 The experiment was performed in a temperate forest located at the St Jones Estuarine

- 114 Reserve (a component of the Delaware National Estuarine Research Reserve [DNERR] in
- 115 Delaware, USA. The site has a mean annual temperature of 13.3 °C and <u>a</u> mean annual
- 116 precipitation of 1119 mm. Soils are classified as Othello silt loam with a texture of 40% sand,
- 117 48% silt, and 12% clay within the first 10 cm (Petrakis et al., 2018). The dominant plant

118 species are bitternut hickory (Carya cordiformis), eastern red cedar (Juniperus virginiana L.), 119 American holly (Ilex opaca), sweet gum (Liquidambar styraciflua L.), and black gum (Nyssa 120 sylvatica (Marshall)). The site has a mean tree density of 678 stems ha⁻¹ and a diameter at 121 breast height (DBH) of 25.7±13.9 cm (mean±SD) (Barba et al., 2021). 122 123 2.2 Automated measurements of soil GHG fluxes 124 We performed analyzed data from automated measurements (45 minutes 1hr time intervals) of 125 soil emissions of three GHGs (i.e., CO₂, CH₄, and N₂O) between September 2014-126 SeptemberJanuary and December 2015. This was a typical year with a mean annual 127 temperature of 13.4 °C and an annual precipitation of 1232 mm. Continuous measurements 128 of soil GHGs were taken by coupling a closed-path infrared gas analyzer (Li-COR LI-8100 129 A, Lincoln, Nebraska) and nine dynamic soil chambers (Li-COR 8100-104) controlled by a 130 multiplexer (Li-COR 8100-104) with a cavity ring-down spectrometer (Picarro G2508, Santa 131 Clara, California). Detailed description of experimental design, measurements protocol are 132 described in previous studies A detailed description of the experimental design and 133 measurements protocol is described in previous studies (Petrakis et al., 2018; Barba et al., 134 2021, 2019). Briefly, for each flux observation, we measured CO₂, CH₄. (Petrakis et al., 135 2018; Barba et al., 2021, 2019). Briefly, for each flux observation, we measured CO₂, CH₄ 136 and N₂O concentrations every second with the Picarro G2508 for 300 seconds and calculated 137 fluxes (at 45 minutes1 hr time intervals) from the mole dry fraction of each gas (i.e., 138 corrected for water vapor dilution) using the SoilFluxPro software (v4.0; Li-COR, Lincoln, 139 Nebraska, USA). Fluxes were estimated using both-linear and exponential fits, and we kept 140 the flux calculation with the highest R^2 . We applied quality assurance and quality control 141 protocols using information from all three GHGs as established in previous studies (Petrakis 142 et al., 2018; Barba et al., 2021, 2019; Capooci et al., 2019; Petrakis et al., 2017). Using these

time series, we extracted values to represent discrete temporal measurements based on FTS
and <u>using anused the</u> optimization approach as described below.

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146 2.3 Temporal subsampling of time series

147 Subsampling of time series was performed using FTS and a temporal optimization following 148 a univariate Latin Hypercube (tuLHs) approach. The difference between FTS and temporal 149 optimization is that the first approach is focused on a fixed schedule (e.g., sampling once per 150 month), and the second is focused on reproducing the statistical properties and temporal 151 dependence relationship of the original GHG time series with a subset of measurements. This 152 means that optimized subsamples may not be spaced systematically (e.g., every 15 days), 153 and selected dates may vary for each GHG flux due to their specific statistical properties and 154 temporal variabilitydependence.

155 FTS represents a traditional schedule for performing manual measurements of GHG 156 fluxes from soils. The FTS is usually performed with manual measurements because they 157 require extensive logistical coordination due to travel time and costs, availability of 158 instrumentation (e.g., gas analyzers) and), personnel to perform the measurements, and 159 weather conditions. During these scheduled visits, researchers usually collect fluxes from all 160 three GHGs and analyze them in a systematic mannersystematically to calculate magnitudes 161 and patterns throughout the length of the experiment. Usually, researchers perform manual 162 samples during the early hours of the day (between 9 am and 12 pm) to avoid confounding 163 effects due to large changes in temperature and moisture, as demonstrated by information 164 summarized by the soil respiration global database (Cueva et al., 2017; Jian et al., 2020). 165 Consequently, we selected subsamples from each original GHG time series (derived from 166 automated measurements) using flux measurements from 10 am at fixed intervals of once per 167 month (n=12), twice per month (n=24), or four times per month (n=48) starting on the first 168 week of available data from automated measurements.

We applied *tuLHs* as an alternative subsampling approach to obtain an optimized subsample with the same univariate statistical properties and temporal dependence relationship of the original GHG time series. Optimization was performed to select subsamples for each GHG flux using the same number of samples as for fix temporal stratification<u>FTS</u>: twelve (k=12), twenty-four (k=24), or forty-eight (k=48) measurements throughout the year of available data from automated measurements.

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176 2.4 Temporal Univariate Latin Hypercube Sampling (tuLHs)

177 Let $S = \{(x_1, y_1, z_1), (x_2, y_2, z_2), \dots, (x_n, y_n, z_n)\}$ be observations of the variables X, Y, and Z in a 178 time series, where X, Y₁ and Z are soil GHGs (i.e., CO₂, CH₄ and N₂O). Each measured 179 variable of the time series is characterized by two functions: the univariate probability 180 distribution function and the temporal dependency function. Once these two functions are 181 known, then the behaviors of the variable can be reproduced (Le et al., 2020; Chilès and Delfiner, 2009; Trangmar et al., 1986; Pyrcz and Deutsch, 2014)(Le et al., 2020; Chilès and 182 183 Delfiner, 2009; Trangmar et al., 1986; Pyrcz and Deutsch, 2014). The tuLHs consists of three steps: (1) modeling the univariate behavior of the variable using the empirical cumulative 184 185 univariate probability distribution function; (2) modeling the temporal dependence using the 186 empirical variogram function; and (3) optimizing a subsample applying a global optimization 187 method, differential evolution, using the previously obtained variogram function as an objective function. 188

First, to model the univariate behavior of the variables from the observations of S, the empirical univariate cumulative distribution function $F_n^*(x)$ of X is estimated by:

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$$F_n^*(x) = \frac{1}{n} \sum_{i=1}^n I\{x_i \le x\} \quad (1)$$

where *I* represents an indicator function equal to 1 when its argument is true, and 0 otherwise. Similarly, the empirical univariate distribution function of the variables *Y* and *Z* can be derived. Second, to model the temporal dependence of the variables from the observations of *S*, the empirical temporal correlation function (i.e., temporal variogram function) $\gamma^*(t)$ of X is estimated by:

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$$\gamma^*(t) = \frac{1}{2N(t)} \sum_{i=1}^{N(t)} [X(t_i + t) - X(t_i)]^2 \qquad (2)$$

198 where N(t) is the number of pairs $X(t_i + t)$ and $X(t_i)$ are separated by a time t. The variogram 199 functions of the variables Y and Z are analogous. Third, Toto optimize the subsample, it is required to choose the "optimal" data points with the selected sample size (i.e., k=12, 24, or 200 48; where $k \ll n$ that will have the same behavior of <u>as</u> the original observations of S (i.e., 201 202 GHG fluxes derived from automated measurements). To achieve this objective, we use the differential evolution, a global optimization method (Storn and Price, 1997)(Storn and Price, 203 204 <u>1997</u>), using the variogram function as an objective function. The procedure consists of dividing the univariate empirical probability distribution in Eq. (1) into k equiprobable strata, 205 206 which is equivalent to k ordered data subsets. From each subset, only one value must be chosen 207 to satisfy the condition of a univariate Latin hypercube. The differential evolution method is 208 applied to find the optimal points that minimize the difference between the subsample 209 variogram γ (t) and the data variogram γ * (t) in Eq. (3).

210
$$OF_1 = \sum_{i=1}^{N(t)} [\gamma(t) - \gamma^*(t)]^2 \qquad (3)$$

211 where *OF* is the objective function, and the variograms γ (t) and γ * (t) are calculated using Eq. 212 (2).

214 2.5 Statistical analyses

215 The t-test was used to compare the means, and the Kolmogorov-Smirnov test was used to 216 compare the probability distribution of measurements derived from each different sampling 217 protocol. All tests were done with thea 95% confidence level. In addition, their statistical 218 properties, such as mean, median, standard deviation, and first and third quartile are, were 219 compared. The differences ofin the experimental semivariograms were calculated as a 220 comparison measure for the temporal dependence of the samples and the original time series 221 of GHG fluxes. For cumulative sums of GHG flux, their mean is calculated as the most likely 222 value and their quantile difference between 97.5 and 2.5 is used to quantify the range of 223 uncertainty., and their quantile difference between 97.5 and 2.5 is used to quantify the range 224 of uncertainty. All analyzes were performed using the R program (Team and Others, 2013). 225

226 **3. Results**

227 3.1 Relationships among GHG fluxes from soils

228 Justification in support of FTS for simultaneous measurements of GHG fluxes would require 229 evidence of strong linear correlations between magnitudes and temporal dependence among 230 soil GHG fluxes. First, we did not find strong linear relationships between any combination 231 of GHG fluxes from soils derived from automated measurements (Fig. A1S1). Therefore, our 232 data did not support the assumption that the magnitude of one GHG flux was associated with 233 a linear increase or decrease of another GHG flux. Second, semivariogram models 234 demonstrated differences in the temporal dependence for each GHG flux. Automated 235 measurements of soil CO₂ fluxes (F_ACO_2) showed a temporal dependence following a 236 Gaussian variogram model, with a nugget of 4, a sill plus nugget of 28, and a correlation 237 range of 80 days (Fig. A2aS2a). Automated measurements of soil CH₄ fluxes (F_A CH₄) also

showed a temporal dependence but followed a spherical variogram model, with a nugget of $7x10^{-8}$, a sill plus nugget of $1.5x10^{-7}$, and a correlation range of 110 days (Fig. A2bS2b). In contrast, automated measurements of soil N₂O fluxes (F_4 N₂O) did not show a temporal dependence, where a pure nugget effect was present, and with a correlation range of 0 days (Fig. A2eS2c). Consequently, the these GHG fluxes' magnitudes and temporal patterns of these GHG fluxes-were different and did not provide support in favor of FTS for simultaneous measurements of GHG from soils.

245

246 3.2 Optimization of GHG sampling protocols

We applied a *tuLHs* approach to identify subsamples that had with the same statistical properties and temporal dependence for each one of the original GHG time series from

automated measurements. Subsamples were identified for twelve (k=12), twenty-four

250 (k=24), or forty-eight (k=48) measurements throughout the year for each GHG time series.

251 All subsamples represent measurements collected at 10 am. Our results show that the

252 optimized measurement dates were different for each GHG flux (Fig. 1), and we provide

explicit examples for k=24 (Fig. 1) and k=12, 48 (Fig. A3, A4S3, S4).

254 The optimized CO₂ subsamples were well distributed throughout the year for all sampling scenarios (i.e., k from 12 to $48_{\frac{1}{2}}$) because F_ACO_2 had a strong temporal dependence 255 256 and a small nugget effect with respect to the sill (Fig. A2aS2a). The optimized CH₄ 257 subsamples were also relatively well distributed throughout the year, especially for scenarios 258 of k=24 and k=48, as F_4CH_4 also had a temporal dependence but with a higher nugget effect 259 with respect to the sill (Fig. A2bS2b). Finally, the optimized N₂O subsamples were more 260 difficult challenging to define, especially with a small sample size (i.e., k=12; Fig. A3eS3c) 261 because $F_A N_2 O$ did not have a temporal dependence (Fig. A2eS2c).

263 3.3 Differences in statistical properties and temporal dependency of subsamples

Overall, there were no statistically significant differences among between the mean values 264 265 derived from automated measurements and those from FTS or the *tuLHs* approach (Fig. 2 for k=24; Fig. A5S5 for k=12; Fig. A6S6 for k=48; Tables A1S1 and A2S2). Although this appears 266 267 to be a promising result, the, more than a simple comparison of the means is not enoughneeded to fully evaluate the information derived from different sampling scenarios.approaches. In 268 269 other words, it is possible to have a similar mean value without reproducing the probability 270 distribution nor the temporal dependence of the original time series (i.e., correct answer but for 271 the wrong reasons). Here, we present results based on comparing the means, standard deviation, probability distributions, and semivariograms derived from automated 272 273 measurements and the different sampling scenarios for all GHG fluxes.

The mean of $F_A CO_2$ was 5.9, $\mu mol CO_2 m^{-2} s^{-1}$, while the mean for FTS 5.5 $\mu mol CO_2$ 274 m⁻² s⁻¹, and 5.9 μ mol CO₂ m⁻² s⁻¹ for the *tuLHs* approach with *k*=24 (Fig. 3a-c). These results 275 were comparable with the means derived from FTS (5.4 and 5.4 μ mol CO₂ m⁻² s⁻¹), and from 276 the *tuLHs* approach (6.2 and 5.9 μ mol CO₂ m⁻² s⁻¹) using *k*=12 and *k*=48, respectively (Figs. 277 A5, A6S5, S6; Table A1S1). The standard deviation of F_ACO_2 was 3.9 and 3.2 μ mol CO₂ m⁻² 278 s⁻¹ for FTS, and 3.9 μ mol CO₂ m⁻² s⁻¹ for the *tuLHs* approach with *k*=24 (Figs. 3a-c). These 279 280 results were comparable with the standard deviations derived from FTS (3.1 and 3.3 µmol CO₂ $m^{-2} s^{-1}$ and from the *tuLHs* approach (4.1 and 3.9 µmol CO₂ $m^{-2} s^{-1}$) using k=12 and k=48, 281 respectively (Fig. A5, A6S5, S6; Table A1S1). Our results show that the semivariograms of 282 283 optimized samples using the *tuLHs* approach closely approximate the semivariograms of 284 automated measurements for k=24 (Fig. 4a) and k=12 and 48 (Figs. A7a, A8aS7a, S8a). These 285 results are consistent with the sums of absolute differences between the semivariograms of the 286 samples and the semivariogram of F_ACO_2 with differences of 69.31, 54.39, 49.42 for FTS, and 287 5.69, 1.99, 1.39 for the *tuLHs* approach for k=12, 24, 48, respectively (Table A2S2).

The mean of F_A CH₄ was -0.93 <u>nmol CH₄ m⁻² s⁻¹</u>, while -0.86 nmol CH₄ m⁻² s⁻¹ for FTS 288 and -0.94 nmol CH₄ m⁻² s⁻¹ for the *tuLHs* approach with k=24 (Fig. 3d-f). These results were 289 also comparable with the means derived from FTS (-0.83 and -0.88 nmol CH₄ m⁻² s⁻¹), and 290 from the *tuLHs* approach (-0.87 and -0.92 nmol CH₄ m⁻² s⁻¹) using k=12 and 48, respectively 291 (Figs. A5, A6S5, S6; Table A1S1). The standard deviation of F_A CH₄ was 0.36 and 0.26 nmol 292 CH₄ m⁻² s⁻¹ for FTS₅ and 0.34 nmol CH₄ m⁻² s⁻¹ for the *tuLHs* approach with k=24. These results 293 were comparable with the standard deviations derived from FTS (0.27 and 0.29 nmol CH₄ m⁻² 294 s^{-1} , and from the *tuLHs* approach (0.33 and 0.35 nmol CH₄ m⁻² s⁻¹) using k=12 and k=48, 295 296 respectively (Figs. A5, A6S5, S6; Table A1S1). The semivariograms of optimized samples using the *tuLHs* approach closely approximate the semivariogram of automated measurements 297 298 for k=24 (Fig. 4b) and k=12 and 48 (Figs. A7b, A8bS7b, S8b). Consequently, the sums of 299 absolute differences between the semivariograms of the samples and the semivariogram of 300 F_A CH₄ were 0.63, 0.48-,0.49 for FTS, and 0.06, 0.04, 0.02 for the *tuLHs* approach with k=12, 301 24, 48, respectively (Table A2S2).

Finally, the mean of F_4 N₂O was 0.45 and 0.61 nmol N₂O m⁻² s⁻¹ for FTS, and 0.51 nmol 302 N₂O m⁻² s⁻¹ for the *tuLHs* approach with k=24 (Fig. 3g-i). These results were also comparable 303 with the means derived from FTS (0.59 and 0.25 nmol N₂O m⁻² s⁻¹), and from the *tuLHs* 304 approach (0.58 and 0.49 nmol N₂O m⁻² s⁻¹) using k=12 and 48, respectively (Figs. A5, A6S5, 305 S6; Table A1S1). The standard deviation of F_AN_2O was 1.62 and 1.97 nmol N_2O m⁻² s⁻¹ for 306 FTS, and 1.54 nmol N₂O m⁻² s⁻¹ for the *tuLHs* approach with k=24. These results were 307 308 comparable with the standard deviations derived from FTS (1.38 and 0.91 nmol N₂O m⁻² s⁻¹), and from the *tuLHs* approach (1.58 and 1.54 nmol N₂O m⁻² s⁻¹) using k=12 and k=48, 309 310 respectively (Figs. A5, A6S5, S6; Table A1S1). Our results show that there is no temporal 311 dependence for N₂O fluxes, but the semivariograms of optimized samples using the *tuLHs* 312 approach closely approximate the semivariogram of automated measurements for k=24 (Fig.

313 4c) and k=12 and 48 (Figs. A7c, A8eS7c, S8c). Consistently, the sum of absolute differences 314 between the semivariograms of the samples and the semivariogram of F_AN_2O were 10.01, 315 12.25, 16.75 for FTS, and 0.82, 1.13, 3.57 for the *tuLHs* approach with k=12, 24, 48, 316 respectively (Table A2S2).

These results show that the *tuLHs* approach reproduced with greater precision the probability distribution and the temporal dependence of the time series derived from automated measurements with more precision than FTS for all GHGs. In the next section, we explore the implications of these differences for calculation of calculating cumulative GHG fluxes.

321

322 3.4 Calculation of cumulative GHG fluxes

323 We calculated the cumulative flux for all GHGs using available information from automated

324 measurements (Fig. 2; Table A3S3). The cumulative sum for available measurements of

325 *F*_ACO₂ was 5758.5 g CO₂ m⁻² [893.9, 13860.8; 95% CI]; for *F*_ACH₄ was -0.47 g CH₄ m⁻² [-

326 0.81, -0.19; 95% CI]; and 0.63 g N₂O m⁻² [-0.75, 5.19; 95% CI] for F_A N₂O.

We used the mean for each GHG flux derived from the *tuLHs* approach or the FTS to calculate the cumulative sum (Table <u>A3S3</u>). We found that the FTS underestimated the

329 cumulative flux (-8.4, -6.2, -7.1%) and the uncertainty (-32.6, -21.6, -19.3%) of F_ACO_2 for

330 k=12, 24, 48, respectively (Fig. 5a). In contrast, the *tuLHs* approach <u>slightly</u> overestimated

the cumulative flux (6.5, 1.1, 0.1%) and <u>slightly</u> underestimated the uncertainty (-9.1, -4.4, -

332 3.7%) for *k*=12, 24, 48, respectively (Fig. 5a).

The FTS underestimated the cumulative flux (-9.1, -6.1, -3.1%) and the uncertainty (-334 31.8, -27.3, -15.9%) of F_A CH₄ for k=12, 24, 48, respectively (Fig. 5b). In contrast, the *tuLHs* 335 approach underestimated the cumulative flux (-6.1%) only for k=12, but <u>slightly</u> 336 underestimated the uncertainty (-15.9, -6.8, -4.5%) for k=12, 24, 48, respectively (Fig. 5b).

337 The FTS substantially underestimated the cumulative flux (-168, -170, -173%) of 338 F_4N_2O for k=12, 24, 48, respectively. Uncertainty was overestimated for k=12 and 24 (3.6) 339 and 26%) and underestimated for k=48 (-31%; Fig. 5c). In contrast, the *tuLHs* approach 340 overestimated less the cumulative flux (29.5, 13.4, 9.1%) for k=12, 24, 48, respectively (Fig. 341 5c). This approach underestimated the uncertainty for k=12 (-11.2%) and k=24 by -11.2 and -342 (-13.8%) but overestimated the uncertainty by 2.9% for k=48 (Fig. 5c). These results show 343 that the *tuLHs* approach consistently provided closer estimates for cumulative sums and 344 uncertainty ranges than an FTS for all GHG fluxes.

345

346 4. Discussion

347 Applied challenges, such as quantifying the role of soils in nature-based solutions, require 348 accurate estimates of GHG fluxes. To do this, two fundamental questionsproblems exist for 349 designing environmental monitoring protocols: where to measure and when to measure? 350 Ultimately a monitoring protocol aims to quantify the attributes of an ecosystem, so that it 351 can be compared in time within that ecosystem or with other ecosystems. Because we cannot 352 measure everything, everywhere, and all the time, we can argue that any monitoring protocol has assumptions that are based on physical, economic, social, and practical reasons to address 353 354 a specific scientific question. These assumptions for designing monitoring protocols could 355 result in misleading, biased, or wrong conclusions, and therefore is critical to assess the 356 consequences of different monitoring efforts. As Hutchinson described in "The Concept of Pattern in Ecology²², we do not always know if a given pattern is extraordinary or a simple 357 358 expression of something which we may learn to expect all the time (Hutchinson,

359 <u>1953)(Hutchinson, 1953)</u>.

360 Automated measurements of soil GHG fluxes have revolutionized our understanding
 361 of the temporal patterns and magnitudes of these fluxes in soils (Vargas et al., 2011; Savage

362 et al., 2014; Bond-Lamberty et al., 2020; Tang et al., 2006). That said, these types of 363 measurements have limitations to represent spatial variability and have higher equipment costs that limits their broad applicability across study sites Automated measurements have 364 365 revolutionized our understanding of the temporal patterns and magnitudes of soil GHG fluxes (Savage et al., 2014; Bond-Lamberty et al., 2020; Tang et al., 2006; Capooci and Vargas, 366 367 2022b). These measurements have limitations in representing spatial variability and have 368 higher equipment costs that limit their broad applicability across study sites (Vargas et al., 369 2011). Consequently, discrete manual measurements are a common approach to 370 simultaneously measure multiple GHG fluxes and report patterns, budgets, and information to parameterize empirical and process based models (Phillips et al., 2017; Wang and Chen, 371 372 2012). In this study, we argue that the convenience of simultaneously measuring multiple 373 GHGs using FTS may result in bias estimates; therefore, optimization of sampling protocols 374 is needed when there is a limited number of measurements in time (i.e., k=12, 24, 48). 375 We show that the magnitude of one GHG flux is not associated with a linear increase 376 or decrease of another GHG flux, and the temporal dependencies of each GHG flux are 377 different from each other (Fig. A1). Therefore, it is not possible to infer the dynamics of one 378 GHG flux based solely on information from another under the assumption that they share 379 similar (or autocorrelated) biophysical drivers. Multiple studies have shown that the 380 importance of different biophysical drivers (e.g., temperature, moisture, light) is different for soil CO₂, CH₄ or N₂O fluxes (Luo et al., 2013; Tang et al., 2006; Ojanen et al., 2010). Our 381 results show that soil CO₂ fluxes have a strong temporal dependence (Fig. A2a), likely as a 382 383 result of the strong relationship between these fluxes and soil temperature in temperate mesic 384 ecosystems (Hill et al., 2021; Bahn et al., 2010). The temporal dependence decreased for soil 385 CH₄ fluxes (Fig. A2b), where there is less evidence for such strong correlation with soil temperature (Bowden et al., 1998; Castro et al., 1995), and where multiple variables are 386

usually needed to explain the variability of these fluxes (Luo et al., 2013; Castro et al., 1994).
Soil N₂O fluxes had no temporal dependence (Fig. A2c), showing a strong decoupling from
soil CO₂- and CH₄ fluxes (Wu et al., 2010), likely as a result of independent biophysical
drivers regulating soil N₂O fluxes (Luo et al., 2013; Bowden et al., 1993; Ullah and Moore,
2011).

392 To address the limitations of a. Consequently, discrete manual measurements
393 are a common approach to simultaneously measure multiple GHG fluxes and report patterns,
394 budgets, and information to parameterize empirical and process-based models (Phillips et al.,
395 2017; Wang and Chen, 2012). In this study, we argue that the convenience of simultaneously

396 measuring multiple GHGs using FTS may result in biased estimates. Therefore, optimization

397 of sampling protocols is needed to provide insights to improve measurement protocols when

398 there is a limited number of measurements in time (i.e., k=12, 24, 48).

399 We show that the magnitude of one GHG flux is not associated with a linear increase

400 or decrease of another GHG flux, and the temporal dependencies of each GHG flux are

401 different (Fig. S1). Therefore, it is not possible to infer the dynamics of one GHG flux based

402 solely on information from another under the assumption that they share similar (or

403 autocorrelated) biophysical drivers. These results imply that the magnitudes and temporal

404 patterns of GHGs are different and therefore do not support an FTS approach for

405 simultaneous measurements of GHG fluxes from soils.

406 <u>Multiple studies have shown that the relevance of different biophysical drivers (e.g.,</u>

407 temperature, moisture, light) is different for soil CO₂, CH₄, or N₂O fluxes (Luo et al., 2013;

- 408 <u>Tang et al., 2006; Ojanen et al., 2010</u>). Our results show that soil CO₂ fluxes have a strong
- 409 temporal dependence (Fig. S2a), likely due to the strong relationship between these fluxes
- 410 and soil temperature in this and other temperate mesic ecosystems (Hill et al., 2021; Bahn et
- 411 al., 2010; Barba et al., 2019). The temporal dependence decreased for soil CH₄ fluxes (Fig.

412 S2b), where there is less evidence for such a strong correlation with soil temperature in this

413 and other temperate mesic ecosystems (Bowden et al., 1998; Castro et al., 1995; Warner et

414 al., 2019; Barba et al., 2019). It has been reported that multiple variables and complex

- 415 relationships are usually needed to explain the variability of soil CH₄ fluxes in forest soils, as
- 416 there is a delicate balance between methanogenesis and methanotrophy (Luo et al., 2013;

417 <u>Castro et al., 1994; Murguia-Flores et al., 2018). In contrast, soil N₂O fluxes had no temporal</u>

418 dependence (Fig. S2c), showing decoupling from the observed patterns of soil CO₂ and CH₄

419 fluxes (Wu et al., 2010), likely as a result of independent biophysical drivers regulating soil

420 <u>N₂O fluxes (Luo et al., 2013; Bowden et al., 1993; Ullah and Moore, 2011).</u>

421 <u>To address the limitations of an</u> FTS protocol, we propose a novel optimization

422 approach (i.e., *tuLHs*) to reproduce the probability distribution and the temporal dependence

423 of each original time series of GHG fluxes. Traditional approachesmethods usually optimize

424 subsamples by either individually focusing on reproducing the probability distribution of the

425 original information (Huntington and Lyrintzis, 1998), or by focusing on(Huntington and

426 <u>Lyrintzis, 1998) or</u> reproducing the temporal dependence of the original information

427 (Gunawardana et al., 2011).(Gunawardana et al., 2011). The tuLHs is a simple approach that

428 consists of usinguses the univariate probability distribution function and the temporal

429 correlation function (i.e., variogram) as objective functions for each GHG flux. Our results

430 show that optimized subsamples do not coincide in time for the three GHGs, suggesting that

431 information should be collected based on the<u>each GHG flux's</u> specific statistical and temporal

432 characteristics of each GHG flux (Fig. 1). This study provides a proof-of-concept for the

433 application of the *tuLHs*-and. It demonstrates how an optimization can be performed approach

434 provides insights to design monitoring protocols and improve soil GHG flux estimates of soil
435 GHG fluxes.

436 The more temporal data we can collect, the better, but in many cases, measurement protocols are limited to a few measurements per year (i.e., k=12 to 48). Our results 437 demonstrate that for a small sample size (i.e., k=12), the optimized measurements for soil 438 439 CO₂ fluxes are consistently spread across the year, and for soil CH₄ fluxes are centered 440 within the growing season, and for soil because of their strong temporal dependence. For the 441 case of soil N₂O fluxes, the variogram shows a constant temporal variability, meaning there 442 is no temporal dependence. Therefore, the optimized measurements are concentrated within 443 the fall season due to their distribution probability (Fig. 1a). Our optimization approach 444 shows how measurements can be distributed across time as more samples are available (i.e., 445 k=24 to 48; Fig. 1b-c) and demonstrates that optimization is critical when a limited number 446 of measurements are available. In other words, a few measurements properly distributed 447 across time provide better agreement with information derived from automated 448 measurements. We highlight that this optimization approach should be tested across different 449 ecosystems as it will result in site-specific recommendations. That said, aA similar 450 conclusion was proposed for the spatial distribution of environmental observatory networks, 451 where a network of few sites properly distributed (e.g., across a country) improves our 452 understanding of the target variable more than a spatially biased network (Villarreal et al., 453 2019). Thus, the need for representativeness assessment of information collected across time 454 and space is needed for accurate evaluation ofto evaluate environmental measurements and 455 quantification of quantify nature-based solutions accurately. 456 An initial approach suggested no statistical differences among the mean flux values 457 derived from different sampling protocols. Arguably, this simplistic approach is a false-458 negative due to biased information from the FTS that does not accurately represent the 459 probability distribution and the temporal variability of soil GHG fluxes (e.g., Figs. 3-4). In

460	contrast, the optimization approach resulted in closer probability distributions and temporal
461	variabilities for all GHGs, providing additional evidence against the FTS approach.
462	We highlight that this optimization approach should be implemented across different
463	ecosystems as it will result in site-specific recommendations. The <i>tuLHs</i> can be applied to
464	any time series length and with any time step (e.g., hours, days), but specific results will be
465	representative of the probability distribution and the temporal dependence of the selected
466	time series. What is essential is to question if a few measurements from an experiment
467	represent the <i>reality</i> of the physical world because if limited information is available, then the
468	actual probability distribution and temporal dependence of the phenomena could be an
469	unknown-unknown. In other words, with few measurements, we may not be aware, and we
470	will not be able to know which is the actual probability distribution and temporal dependence
471	of the studied phenomena. To address this challenge, we tested the tuLHs approach with
472	high-temporal frequency information representing the probability distribution of multiple soil
473	GHG fluxes at the daily time-step across a calendar year.
474	In this case study, the year chosen had typical climatological conditions and
475	demonstrates that the statistical properties of the GHG fluxes are different and do not support
476	an FTS approach. Therefore, longer time series (e.g., multi-year) may provide more robust
477	optimizations that can be applied to monitoring efforts in future years. Alternatively, forecast
478	scenarios can be predicted, or in some cases (e.g., soil CO2 efflux in temperate mesic
479	ecosystems), a proxy variable could be used (e.g., soil temperature) to inform the <i>tuLHs</i> and
480	provide insights for a sampling design. In those cases, the <i>tuLHs</i> can be used to suggest an
481	optimized sampling design under those specific assumptions. Ideally, automated
482	measurements should be co-located with manual efforts to adequately capture the temporal
483	and spatial variability of soil GHG fluxes at a specific site.
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484 There are several implications of biased monitoring protocols for the understanding of 485 soil GHG fluxes and nature-based solutions. First, temporal patterns and temporal 486 dependency may not be properly represented with the need to be revisited for studies using 487 an FTS approach. Soil GHG fluxes have complex temporal dynamics that vary from diurnal 488 to seasonal and annual scales that FTS is not able to(Vargas et al., 2010) that a few 489 measurements following an FTS approach cannot reproduce (Barba et al., 2019; Bréchet et 490 al., 2021).(Barba et al., 2019; Bréchet et al., 2021). Second, soil GHG fluxes could present 491 hot-moments, which are transient events with disproportionately high values that are often 492 missed with an FTS approach (Vargas et al., 2018; Butterbach-Bahl et al., 2004). Third, 493 cumulative sums and uncertainty ranges are biased or misleading when derived using an FTS approach (Capooci and Vargas, 2022; Tallec et al., 2019; Lucas-Moffat et al., 2018). For 494 495 this third point, our(Tallec et al., 2019; Lucas-Moffat et al., 2018; Capooci and Vargas, 496 2022b). Our study demonstrates that an optimized approach consistently provided closer 497 estimates for cumulative sums and uncertainty ranges when compared with automated 498 measurements (Fig. 5). We postulate that representing the variability of soil N_2O fluxes is 499 more sensitive to the FTS approach (>170% and >30% for cumulative sums and uncertainty 500 ranges, respectively) than for soil CH₄ and CO₂ fluxes. Fourth, it is possible that if the 501 information derived from thean FTS approach is biased, then functional relationships could 502 also be different from those derived from automated measurements (Capooci and Vargas, 503 2022). It has been discussed that hypothesis testing and our capability for(Capooci and 504 Vargas, 2022a). It has been argued that hypothesis testing and our capability of forecasting 505 responses of soil GHG fluxes to changing climate conditions is also biased with information 506 from the FTS approach (Vicca et al., 2014). Finally, because soils have a central role 507 for(Vicca et al., 2014). Finally, because soils have a central role in nature-based solutions 508 within countries and across the world (Griscom et al., 2017; Bossio et al., 2020), accurate

509 measurements are required to properly assess management practices, environmental 510 variability and the contribution of GHGs from soils (Anderegg, 2021). assess management 511 practices, environmental variability, and the contribution of GHGs from soils. 512 513 Conclusion 514 We highlight that we do not always know if a given pattern is extraordinary or a simple 515 expression of something which we may learn to expect all the time (Hutchinson, 1953). Furthermore, the "Knowledge Paradox" has been recognized for soil science, where 516 517 innovative knowledge has often not been accepted by or implemented in society (Bouma, 2010). Here, we postulate that with emergent technologies there is a convenience of 518 519 measuring multiple GHGs from soils; however, few measurements collected at fixed time 520 intervals results in biased estimates. 521 We We highlight that we only sometimes know if a given pattern is extraordinary or a simple expression of something which we may learn to expect all the time (Hutchinson, 1953). 522 523 Arguably, there is bias in our understanding of the probability distribution and temporal 524 dependency of soil GHG fluxes across the world because most results are based on a few 525 manual measurements (e.g., once a month) following an FTS approach. Currently, it is 526 unknown how large such bias could be across studies and ecosystems, but because most 527 studies lack high-temporal frequency information, the real probability distribution and 528 temporal dependency of soil GHG fluxes may remain unknown in most study sites. What is essential is to question if the observed patterns, derived from an FTS approach, are enough 529 530 for improving our understanding of soil processes or are results that we have learned to 531 expect.

532 We postulate that with emergent technologies, there is a convenience of measuring

533 <u>multiple GHGs from soils; however, few measurements collected at fixed time intervals</u>

534 result in biased estimates. We recognize that potential measurement bias in measurements is 535 dependent depends on the each GHG flux's magnitudes and temporal patterns of each GHG 536 flux and could be site-specific. Nevertheless, evaluations are needed to quantify potential bias 537 in estimates of GHG budgets and information used for model parameterization and 538 environmental assessments. Furthermore, the underlying assumption that each GHG flux 539 responds similarly to biophysical drivers may need to be tested across multiple ecosystems to 540 quantify how few measurements influence our understanding of magnitudes and temporal 541 patterns of soil GHG fluxes. 542 In this study, we present a proof-of-concept and propose a novel-optimization 543 approach (i.e., temporal univariate Latin Hypercube sampling) that can be applied with site-544 specific information of different ecosystems to improve monitoring efforts and reduce the 545 bias of GHG flux measurements across time. We highlight that constant biased 546 environmental monitoring may provide confirmatory information, which we have learned to 547 expect, but modifications of monitoring protocols could shed light into extraordinary 548 patterns. These on new or unexpected patterns. These new patterns are the ones that will test 549 paradigms and push science frontiers. 550 551 552 Data Availability. All data used for this analysis is available at: 553 https://doi.org/10.6084/m9.figshare.19536004.v1https://doi.org/10.6084/m9.figshare.195360 554 04.v1. The R code used in this study is available at: https://github.com/vargaslab/temporal-555 univariate-Latin-Hypercube.git 556 557

558	Author Contributions. R.V. conceived this study, and V.H.L. designed and performed the
559	primary analysis with input from R.V in all phases. R.V. wrote the manuscript with
560	inputcontributions from V.H.L.
561	
562	Competing Interest Statement. None
563	
564	Acknowledgments. The authors thank the Delaware National Estuarine Research Reserve
565	(DNERR), the personnel from) and the St Jones Reserve personnel for their support
566	throughout this study. <u>Authors The authors</u> acknowledge the land on which they realized this
567	study as the traditional home of the Lenni-Lenape tribal nation (Delaware nation). This study
568	was funded by a grant from the National Science Foundation (#1652594).
569 570	
5/1	

572) and NASA Carbon Monitoring System (80NSSC21K0964).



580 Figure 1. Temporal distribution of fixed temporal stratification (i.e., stratified manual

581 sampling approach) and optimized sampling using a temporal univariate Latin Hypercube

- 582 (tuLHs) approach for: k=12 (a), k=24 (b), and k=48 (c). Fixed temporal stratification is in
- 583 black, soil CO₂ fluxes in red, soil CH₄ fluxes in blue, and soil N₂O fluxes in green.

584 <u>Time (x-axis) represents days from January 1 to December 31 of, 2015.</u>







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589 Figure 2. Time series of automated measurements (F_A) of soil greenhouse gas fluxes (black 590 circles) and optimized samples (k=24) using a temporal univariate Latin Hypercube sampling 591 (tuLHs) approach for soil CO₂ (a), soil CH₄ (b) and soil N₂O (c) fluxes. Horizontal The 592 horizontal red line represents the mean, and the horizontal blue line is the median of each 593 greenhouse gas flux derived from automated measurements. Selection The selection of 594 datapoints data points for k=12 and 48 are presented for each soil greenhouse gas time series 595 in Figs. A3S3 and A4S4, respectively. Time (x-axis) represents days from January 1 to 596 December 31 of, 2015. 597 598



Figure 3. Histograms for automated measurements of soil CO_2 ($F_A CO_2$; a), soil CH_4 (F_A CH₄; d+), and soil N₂O (F_A N₂O; g). Histograms for optimized samples (k=24) using a temporal univariate Latin Hypercube sampling (tuLHs) approach for soil CO₂ (b), soil CH₄ (e), and soil N₂O (h) fluxes. Histograms for fixed temporal stratification (i.e., stratified manual sampling schedule) (k=24) for soil CO₂ (c), soil CH₄ (f), and soil N₂O (i) fluxes. Appendix ASupplementary material includes results for measurements with k=12 (Fig. A5S5) and *k*=48 (Fig. A6S6).



- 617
- 618 **Figure 4**. Comparison of semivariograms between automated measurements (F_A) of soil
- 619 greenhouse gas fluxes (solid black line) and for optimized samples using a temporal
- 620 univariate Latin Hypercube sampling (*tuLHs*) approach (red circles) or fixed temporal
- 621 stratification (green circles) with k=24. Semivarograms are presented for soil CO₂ (a), CH₄
- 622 (d)), and N₂O (c) fluxes. Semivariograms for measurements with k=12 and k=48 are
- 623 presented<u>shown</u> in supplementarySupplementary Figs. A7S7 and A8S8, respectively.
- 624 Semivariogram fits were gaussian (Gau) or spherical (sph).
- 625 626



632

633 **Figure 5.** Comparison of percent differences from cumulative sums and associated

634 uncertainty (95% CI) between greenhouse gas fluxes derived from automated measurements

635 (F_A) and using an optimized sampling approach (*tuLHs*) or a fixed temporal stratification.

636 Differences are represented for of soil CO₂ (a), soil CH₄ (b), and soil N₂O (c) fluxes.

637 <u>Black The black</u> circle in the center (0,0) of a plot represents the values derived from 638 automated measurements (F_A) . Blue circles represent estimates from fixed temporal

639 stratification, and red circles represent estimates from an optimized sampling approach

640 (*tuLHs*). Estimates were calculated based on the 258 available automated measurements (Fig.

641 2)), and numeric estimates are in Table A3S3.

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646 Appendix A – Supplementary Tables and Figures

647

- 648 Table A1. Statistical properties for automated measurements of soil CO₂ (F₄ CO₂), soil CH₄
- 649 $(F_{+}CH_{4})$ and soil N₂O $(F_{+}N_{2}O)$ fluxes, optimized samples (k=12, 28, 48) using a temporal
- 650 univariate Latin Hypercube sampling (*tuLHs*), and fixed temporal stratification (*k*=12, 28,
- 651 48). Units for soil CO₂-fluxes are in μmol m⁻² s⁻¹, and for soil CH₄ and N₂O fluxes in nmol m⁻
- 652 2-s-1-.
- 653

	Number of measurements (k)	1st. Quartile	Median	Mean	3rd. Quartile	Standard Deviation
F ₄ CO ₂	8259	2.81	5.03	5.87	8.65	3.85
	12	3.19	5.30	6.25	8.88	4 .06
<i>tuLHs</i> approach (€€92)	24	3.00	5.13	5.93	8.44	3.90
	48	2.84	4.97	5.88	8.54	3.87
Fixed temporal	12	2.68	<u>5.82</u>	5.37	7.10	3.15
stratification	24	2.69	5.66	5.50	7.07	3.2 4
(02)	48	2.69	5.53	5.45	8.05	3.29
F ₄ CH4	8259	-1.14	-0.92	-0.93	-0.67	0.36
	12	-1.11	-0.89	-0.87	- 0.66	0.33
tuLHs approach (CH 4)	24	-1.14	-0.92	-0.94	-0.66	0.34
	48	-1.13	-0.91	-0.92	-0.66	0.35
	12	-1.01	-0.83	-0.83	- 0.67	0.27
Fixed temporal stratification	24	-1.01	-0.89	-0.86	-0.68	0.26
(CH 4)	48	-1.10	-0.86	-0.88	-0.66	0.29
<i>F</i> _A N₂O	8259	-0.18	0.01	0.45	0.49	1.62
	12	-0.18	-0.01	0.58	0.50	1.58
<i>tuLHs</i> approach (N ₂O)	24	-0.18	0.03	0.51	0.45	1.54
	4 8	-0.17	0.02	0.49	0.45	1.54
	12	-0.35	0.51	0.59	0.83	1.38
Fixed temporal stratification	24	-0.21	-0.08	0.61	0.36	1.97
(<u>N2O)</u>	4 8	-0.31	0.00	0.25	0.53	0.91

- **Table A2.** Comparison of errors between experimental variogram for automated measurements of soil greenhouse gases (F_A ; k=8259) and experimental variograms for data using temporal univariate Latin Hypercube sampling (*tuLHs*) and fixed temporal
- stratification.
- 659

	Approach	Number of measurements (k)	Error (Sum of absolute differences)
		12	69.31
		24	54.39
Soil CO2	Fixed	4 8	<u>49.42</u>
fluxes		12	5.69
		2 4	1.99
	tuLHs	4 8	1.39
		12	0.63
		24	0.68
Soil CH4	Fixed	48	0.49
fluxes		12	0.06
		24	0.04
	tuLHs	4 8	0.02
		12	10.01
		24	12.25
Soil N₂O	Fixed	4 8	16.75
fluxes		12	-0.82
		24	1.13
	tuLHs	4 8	3.57

664 Table A3. Cumulative sum and associated uncertainty of greenhouse gas (GHG) fluxes 665 derived from automated measurements $(F_{\mathcal{A}})$ and using an optimized sampling approach (tuLHs) or a fixed temporal stratification. Cumulative sum represents the total flux from 666

667 available measurements derived from automated measurements for all GHG fluxes.

- 668 669

	Number of measurements (k)	Cumulative Sum	Uncertainty 95% CI		Uncertainty Range
<i>F</i> .4 CO 2 (g CO2 m ²)	8259	5758	893	13860	12966
	12	6130	1423	13218	11794
tuLHs approach (g CO ₂ m ²)	2 4	5818	1046	13438	12391
(č -)	48	5766	946	13429	12482
Fixed	12	5273	1376	10117	8740
temporal stratification	24	5402	1196	11356	10160
(g CO₂ m[≠])	48	5351	1162	11621	10458
F ₄ CH ₄ (g CH ₄ m ²)	8259	-0.33	-0.58	-0.14	0.44
	12	-0.31	-0.49	-0.12	0.37
tuLHs approach (2 CH 4 m ²)	24	-0.33	-0.57	-0.16	0.41
	48	-0.33	-0.56	-0.14	0.42
	12	-0.3	-0.45	-0.15	0.3
Fixed temporal stratification	2 4	-0.31	-0.46	-0.14	0.32
(g CH ₄ m ²)	48	-0.32	-0.51	-0.14	0.37
<i>F_AN₂O</i> (g N ₂ O m ²)	8259	0.44	-0.53	3.67	4 .2
	12	0.57	-0.48	4.19	4.67
tuLHs approach (c.N2Om ²)	24	0.5	-0.43	4 .35	4 .78
(81.20 m)	48	0.48	-0.5	3.58	4.08
	12	-0.3	0.83	3.52	4.35
	2 4	-0.31	-0.43	4.86	5.29

	Fixed temporal stratification (g N ₂ O m ²)	4 8	-0.32	-0.7	2.21	<u>2.91</u>
670 671	·					
672						





- 680 N₂O) fluxes. None of these relationships were significant at α =0.05. These relationships were
- 681 derived using all available data from automated measurements (F_A) of soil greenhouse gas
- 682 fluxes.





Figure A3. Time series of automated measurements (FA) of soil greenhouse gas fluxes (black circles) and optimized samples (k=12) using a temporal univariate Latin Hypercube sampling (tuLHs) approach for soil CO2 (a), soil CH4 (b) and soil N2O (c) fluxes. Horizontal red line represents the mean and horizontal blue line the median of each greenhouse gas flux derived from automated measurements.



Figure A4. Time series of automated measurements (FA) of soil greenhouse gas fluxes
(black circles) and optimized samples (*k*=48) using a temporal univariate Latin Hypercube
sampling (*tuLHs*) approach for soil CO₂ (a), soil CH₄ (b) and soil N₂O (c) fluxes. Horizontal
red line represents the mean and horizontal blue line the median of each greenhouse gas flux
derived from automated measurements.



725 726

Figure A5. Histograms for automated measurements of soil CO₂ (F_{4} CO₂; a), soil CH₄ (F_{4} CH₄; d) and soil N₂O (F_{4} -N₂O; g) fluxes. Histograms for optimized samples (k=12) using a temporal univariate Latin Hypercube sampling (*tuLHs*) approach for soil CO₂ (b), soil CH₄ (e) and soil N₂O (h) fluxes. Histograms for fixed temporal stratification (i.e., stratified manual sampling schedule; k=12) for soil CO₂ (c), soil CH₄ (f) and soil N₂O (i) fluxes.



Figure A6. Histograms for automated measurements of soil CO₂ (F₄CO₂; a), soil CH₄ (F₄ CH4; d) and soil N2O (FA N2O; g) fluxes. Histograms for optimized samples (k=48) using a temporal univariate Latin Hypercube sampling (tuLHs) approach for soil CO2 (b), soil CH4 (e) and soil N₂O (h) fluxes. Histograms for fixed temporal stratification (i.e., stratified manual sampling schedule; k=48) for soil CO₂ (c), soil CH₄ (f) and soil N₂O (i) fluxes.









761Figure A8. Comparison of semivariograms between automated measurements (F_4) of soil762greenhouse gas fluxes (solid black line) and for optimized (red circles) or fixed temporal763stratification (green circles) with k=48. Semivarograms are presented for soil CO₂ (a), CH₄764(d) and N₂O (c) fluxes. Semivariogram fits were gaussian (Gau) or spherical (sph).



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